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William Budnick, Thibault Leboucher, Jérôme Belliard, Janne Soininen, Isabelle Lavoie, et al.. Local and regional drivers of taxonomic homogenization in stream communities along a land use gradient. Global Ecology and Biogeography, 2019, 28 (11), pp.1597-1609. 10.1111/geb.12976. hal-02609712

HAL Id: hal-02609712 https://hal.inrae.fr/hal-02609712v1

Submitted on 9 Jan 2024

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Running title: Land use effects on stream diversity

LOCAL AND REGIONAL DRIVERS OF TAXONOMIC HOMOGENIZATION IN STREAM COMMUNITIES ALONG A LAND USE GRADIENT

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ACKNOWLEDGEMENTS

We thank W. Xu for helpful comments on the null model. Support by the National Science

Foundation (grant NSF DEB-1745348) to SIP is gratefully acknowledged.

Author Biosketch: William R. Budnick is a Ph.D. candidate studying macroecology at the University of Texas at Arlington. His primary research interests lie with the application of cutting-edge numerical and spatial techniques to study and predict how global change forces will affect macroecological patterns particular to aquatic ecosystems. His interests also include integrating basic ecological theory with fisheries science to assist with the global conservation of threatened fish and crayfish fauna.

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2	COMMUNITIES ALONG A LAND USE GRADIENT
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8	ABSTRACT
9	Aim: The interaction of land use with local vs. regional processes driving biological
10	homogenization (β -diversity loss) is poorly understood. We explored: i) stream β -diversity
11	responses to land cover (forest vs. agriculture) in terms of physicochemistry and
12	physicochemical heterogeneity, ii) whether these responses were constrained by the regional
13	species pool, i.e. γ -diversity, or local assembly processes through local (α) diversity, iii) if local
14	assembly operated through the regional species abundance distribution (SAD) or intraspecific
15	spatial aggregation, and iv) the dependency on body size, dispersal capacity, and trophic level
16	(producer vs. consumer).
17	
18	Location: United States of America, Canada, and France
19	
20	Time Period: 1993-2012
21	
22	Major Taxa Studied: Stream diatoms, insects, and fish
23	

24	Methods: We analyzed six datasets totaling 1,225 stream samples. We compared diversity
25	responses to eutrophication and physicochemical heterogeneity in forested vs. agricultural
26	streams with regression methods. Null models quantified contribution of local assembly to β -
27	diversity (β -deviance, β_{DEV}) for both land covers and partitioned it into fractions explained by the
28	regional SAD (β_{SAD}) vs. aggregation (β_{AGG}).
29	
30	Results: Eutrophication explained homogenization and more uneven regional SADs across
31	groups, but local and regional biodiversity responses differed across taxa. β_{DEV} was insensitive to
32	land use. β_{SAD} largely exceeded β_{AGG} and was higher in agriculture.
33	
34	Main Conclusion: Eutrophication but not physicochemical heterogeneity of agricultural streams
35	underlay β -diversity loss in diatoms, insects and fish. Agriculture did not constrain the
36	magnitude of local vs. regional effects on β -diversity, but controlled the local assembly
37	mechanisms. While the SAD fraction dominated in both land covers, it further increased in
38	agriculture at the expense of aggregation. Notably, the regional SADs were more uneven in
39	agriculture, exhibiting excess common species or stronger dominance. Diatoms and insects
40	diverged from fish in terms of biodiversity, SAD shape, and β_{DEV} patterns, suggesting an
41	overriding role of body size and/or dispersal capacity compared to trophic position.
42	
43	Key words: β-diversity, biodiversity loss, taxonomic homogenization, diatoms, fish, insects,
44	land use, local assembly, spatial aggregation, species abundance distribution

45 **INTRODUCTION**

46 Landscape transformations from continuous undeveloped expanses to agricultural fields and 47 urban sprawls have accelerated the global biodiversity decline (Newbold, Hudson, Hill, Contu, 48 Lysenko et al., 2015). Human land use (hereafter land use) underlies declines in both regional 49 richness, i.e. γ -diversity (Barlow, Lennox, Ferreira, Berenguer, Lees et al., 2016), and 50 dissimilarity among biological communities, i.e. β -diversity, resulting in taxonomic 51 homogenization across space and time (Petsch, 2016). Biodiversity losses from land use stem 52 from habitat loss, fragmentation, eutrophication, and physicochemical stress, altogether 53 considered among the primary threats facing global biodiversity (Sala, Stuart Chapin, Iii, 54 Armesto, Berlow et al., 2000; Devictor, Julliard, Clavel, Jiguet, Lee et al., 2008). Preventing 55 biodiversity losses and mitigating subsequent homogenization remain a top priority because both 56 can translate into decreased biological integrity and ecosystem resilience (de Juan, Thrush & 57 Hewitt, 2013; Socolar, Gilroy, Kunin & Edwards, 2016). Therefore, it is critical from a 58 conservation planning standpoint to continue investigating how land use affects ecological 59 processes underlying global diversity in order to mitigate the ongoing biodiversity crisis. 60 Land use effects on biodiversity occur across scales, operating either in a top-down or 61 bottom-up fashion or both (Flohre, Fischer, Aavik, Bengtsson, Berendse et al., 2011). Top-down 62 mechanisms function through the regional species pool (γ -diversity), which is a product of speciation and extinction, large-scale dispersal, climate, and evolutionary, geological, and land 63 64 use history (Zobel, 2016). Bottom-up mechanisms include local-level assembly processes, e.g.

65 environmental filtering, interspecific interactions, and small-scale dispersal (Márquez & Kolasa,

66 2013), which constrain local (α) diversity and subsequently affect site-to-site community

67 dissimilarity. Studies across terrestrial and freshwater systems have reported a general decline in

68 γ -diversity because of land use, but divergent patterns of α -diversity, including decreased α -69 diversity, owing to losses of sensitive and endemic species, and stable, or even increased α -70 diversity, owing to greater rates of species invasion and colonization (Vellend, Baeten, Myers-71 Smith, Elmendorf, Beauséjour et al., 2013; Newbold et al., 2015; Gonzalez, Cardinale, 72 Allington, Byrnes, Arthur Endsley et al., 2016). Thus, land use likely exerts differential impact 73 on the species pool and local assembly processes that may cause γ - and α -diversity, respectively, 74 to vary at different rates, which in turn influences β -diversity response (Kraft, Comita, Chase, 75 Sanders, Swenson et al., 2011).

76 β -diversity is usually treated as a scalar linking average α -diversity with γ -diversity, thus 77 reflecting spatial or temporal differences among localities. One can then measure the influence of 78 α - and γ -diversity as proxies of local and regional drivers of β -diversity, respectively. 79 Specifically, null models that constrain the observed species pool variation (i.e., γ -diversity) can 80 assess the role of local assembly by calculating a β -diversity measure (β_{DEV}) corresponding 81 solely to α -diversity variation (e.g., Kraft *et al.*, 2011) (Fig. 1a). β_{DEV} can be further decomposed 82 into fractions reflecting roles of intraspecific spatial aggregation (i.e., the spatial clumping 83 pattern of individuals within species) and the regional species abundance distribution (SAD, 84 vector of species abundances) (Xu, Chen, Liu & Ma, 2015) (Fig. 1b). Intraspecific spatial 85 aggregation results from dispersal, competitive, and environmental mechanisms that cluster 86 individuals of species across fewer sites, thus bolstering β -diversity (Veech, 2005). However, 87 regional SADs influence β -diversity because rare species are less likely to be locally sampled 88 due to low regional abundance (He & Legendre, 2002). Although examined across latitudes, the 89 two fractions of local assembly have not been studied in other contexts and it is unknown 90 whether these components are responsive to strong ecological influences (e.g., land use).

91 Studying how local assembly and regional species pool processes interplay is an ongoing 92 area of research in terrestrial systems because it may explain how β -diversity varies with land 93 use (Socolar et al., 2016). Surprisingly, little attention is focused on freshwater systems, even 94 though freshwater biodiversity is more vulnerable to land use relative to terrestrial systems, 95 particularly through habitat modification (Sala et al., 2000; Wiens, 2016) and eutrophication 96 from agriculture (Withers, Neal, Jarvie & Doody, 2014). Although primary productivity in 97 agricultural streams could increase with eutrophication, forest streams, which are usually low in 98 nutrients and have more shading, tend to harbor higher biodiversity stemming from greater 99 physical and environmental heterogeneity that translates into greater ecosystem complexity 100 (Penaluna, Olson, Flitcroft, Weber, Bellmore et al., 2017). Agriculture probably causes changes 101 in physicochemical heterogeneity as well, but this subject is poorly explored. Thus, the scarcity 102 of data, especially for aquatic taxa, has inhibited general understanding of how land use 103 influences local and regional processes driving β -diversity. 104 Impacts of agricultural eutrophication on β_{DEV} are not understood, although null models 105 have been used to assess environmental disturbance (e.g., Myers, Chase, Crandall & Jiménez, 106 2015). We hypothesize β -diversity response to eutrophication, including variation in β_{DEV} , 107 depends on trophic level, body size, and dispersal capacity. For example, many unicellular 108 producers, like diatoms, have high nutrient demands and may benefit from increased nutrients 109 (Passy, 2008; Soininen, Jamoneau, Rosebery & Passy, 2016). Diatom microscopic size, high 110 local abundance, and broad geographic distributions, allowing both in-stream and overland 111 passive dispersal (Finlay, 2002), may result in weak β -diversity and β_{DEV} response to agriculture. 112 Smaller bodied macroscopic organisms, such as aquatic insects, may be more constrained in

113 active dispersal capacity during larval stages but exhibit greater overland mobility during winged

114	adult life stages, which could offset some harmful agricultural effects. In contrast, larger
115	consumers with more limited geographic dispersal capacity, such as fish, may be negatively
116	affected by eutrophication due to ammonia toxicity, loss of suitable habitat, and lower quality
117	food sources (Allan, 2004).
118	In this study, we compared spatial patterns of biodiversity and abundance in streams with
119	watersheds dominated by agriculture vs. forest. Our objectives were to determine: i) how β -
120	diversity and related biodiversity properties respond to agriculture (through nutrient enrichment
121	or physicochemical heterogeneity), ii) if agriculture alters the relative contribution of local
122	assembly effects to β -diversity, iii) whether agriculture differentially constrains the fractions of
123	local assembly explained by spatial aggregation vs. the SAD, and iv) if the relationships outlined
124	in i) to iii) vary across organismal groups (Table 1).
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137 United States

138 US community data, spanning 19 latitudinal degrees and 55 longitudinal degrees, were obtained 139 from the National Water-Quality Assessment (NAWQA) Program of the United States 140 Geological Survey and the National Rivers and Streams Assessment (NRSA) of the United 141 States Environmental Protection Agency. Communities were collected in the warm months 142 during low flow conditions (July through September) to constrain seasonal succession and 143 variation in temperature and flow. NAWQA communities (diatoms, insects, and fish) were 144 sampled between 1993-2010, whereas NRSA communities (fish), between 2011 and 2012. 145 Diatoms were collected from the richest-targeted habitats, encompassing hard substrates or 146 macrophytes. Depending on available substrate, a defined area of 25 cobbles, 5 woody snags or 5 147 macrophyte beds was sampled within a stream reach and the samples were composited. Diatoms 148 were identified generally to species in counts of 400-800 cells. Benthic insects (class Insecta) 149 were composed of combined sieved samples taken from the richest-targeted habitats (i.e., riffles, 150 main-channel, and natural-bed instream habitats). Insects were identified to the lowest possible 151 category (order to species) in counts of 400-800 individuals. Both NAWQA and NRSA fish were 152 generally identified to species in counts of 400-950 individuals taken from riffle, pool, and run 153 habitats using electrofishing equipment with seines.

Land use and cover data were generated by the NAWQA and NRSA using National Land Cover Datasets 1992 and 2006, 30 m resolution. We selected 400 streams for diatoms and 126 streams for insects split equally between both land cover categories. Since fish communities and environmental data in both the NAWQA and NRSA data were sampled with similar methods, we combined both fish datasets into a single dataset comprising 231 streams (116 agricultural and 115 forested streams).

160 France

161	French diatom data were sourced from a national dataset including field collections of 200
162	streams from 2011. Algae were collected from stones during the low flow period in June
163	through September with a standardized sampling method (Afnor, 2007). Diatoms were identified
164	generally to species in counts of about 320-475 cells. The French fish dataset was collected by
165	the French National Agency for Water and Aquatic Environments (ONEMA) during low flow
166	periods between May and October 2011. The dataset comprised 200 streams with fish identified
167	to species in counts of 10-3300 individuals sampled with electrofishers. For both French
168	datasets, we used 100 agricultural and 100 forest streams, spanning 8 latitudinal and 14
169	longitudinal degrees. Land use cover data were obtained from the CORINE land cover database
170	(European Environment Agency, 2013)
171	
172	Canada
173	Canadian diatom data included 46 stream samples (23 streams in both land cover categories)
174	collected in August to September during the low flow period between 2002 and 2009 (Lavoie,

175 Campeau, Zugic-Drakulic, Winter & Fortin, 2014) spanning 3 latitudinal and 6 longitudinal

176 degrees. Samples were composites of rock scrapes (5-10 rocks) per stream reach, targeting riffles

and runs. Diatoms were mainly identified to species in counts of at least 400 valves. Land use

178 cover data were compiled from government GIS databases, including the Ecoforestry

179 Information System, Annual Crop Inventory, and the Insured Crop Database.

180

181 Environmental data

182 All datasets had associated physicochemical and coordinate data (i.e., GCS coordinates re-

183 projected with Lambert Conformal Conic). Environmental variables in our analyses included

184 water temperature, air temperature, nitrite + nitrate (or total nitrogen when absent), ammonia, 185 orthophosphate, total phosphorus, specific conductance, and pH (Appendix 1, Table S1.1 in 186 Supplemental Information). Environmental data for the US datasets consisted of the average for 187 the month of sample collection. Environmental data for French diatoms included the median of 188 measurements obtained 30 days before and 15 days after the diatom sample date. The French fish 189 environmental data represented the average of 12 monthly measurements prior to fish sampling. 190 Air temperature for French diatom data were not recorded at the time of sampling and were 191 obtained from the WorldClim database (Hijmans, Cameron, Parra, Jones & Jarvis, 2005), 192 whereas air temperatures for French fish streams were measured at the stream. Canadian 193 environmental data were seasonal averages calculated from water samples collected from July to 194 September.

195

196 Diversity, spatial aggregation, and species abundance distribution

197 We calculated $\overline{\alpha}$ -diversity (average richness across samples), γ -diversity (total richness per land 198 use), and β -diversity of stream samples for both land cover categories for each dataset. We used 199 equation (1) to calculate the observed β -diversity (β_{OBS}),

200 201

$$\beta_{\rm OBS} = 1 - \frac{\overline{\alpha}}{\gamma} \tag{1}$$

which indicated the average proportion of the species pool absent from a stream.

We used the null model framework developed by Xu *et al.* (2015) to quantify i) the magnitude of the local assembly effect on β -diversity after controlling for γ -diversity and ii) the contributions of the SAD vs. intraspecific spatial aggregation to local assembly (Fig. 1b). First, the difference (i.e., β -deviance, β_{DEV}) between β_{OBS} and the expected β -diversity (β_{EXP} , i.e., β diversity expected assuming completely random sampling, see Appendix S2) was taken to

208 quantify local assembly absent the effect of γ -diversity. β_{DEV} is bounded between 0 and 1, with 209 larger β_{DEV} corresponding to greater local control. Secondly, we calculated β -diversity predicted 210 when intraspecific spatial aggregation is constant across all species (β_{PRED}). Then, the difference 211 between β_{PRED} and β_{EXP} reveals what fraction of β_{DEV} is contributed by the SAD (β_{SAD}), while the 212 remaining fraction of β_{DEV} is attributed to spatial aggregation (β_{AGG}). In this model, β_{SAD} can 213 exceed β_{DEV} if β_{PRED} exceeds β_{OBS} . The corresponding aggregation fraction will in turn be 214 negative because the sum of the two fractions, β_{SAD} and β_{AGG} , must equal 1, thus meaning that 215 the pattern is less aggregated than expected by the null model. To test whether the two land 216 covers differ in their magnitude of intraspecific spatial aggregation, we used maximum 217 likelihood methods and calculated the aggregation parameter, k, across samples within each land 218 cover (Appendix S3). Because smaller k corresponded to greater aggregation, we analyzed the 219 reciprocal of the parameter for easier interpretation. In summary, the procedure yielded six 220 measurements: β_{EXP} , β_{PRED} , β_{DEV} , β_{SAD} , β_{AGG} , and 1/k. 221 Regional SADs for both land cover categories was analyzed by summing abundances of 222 each species across all stream samples and calculating the standard deviation (parameter σ) of 223 the Poisson-lognormal distribution fit of the abundance data using the `sads' R package (Prado, 224 Mirands & Chalom, 2017). Parameter σ reflects SAD evenness with greater σ values 225 corresponding to increased unevenness. To determine if changes in σ were associated with 226 prevalence of rare vs. common species, we also examined the relationship of σ with the skewness 227 ('skewness' function from R package `moments', Komsta & Novomestky, 2015) of the log-228 transformed regional species abundances for each land cover category. Skewness was significant if skewness divided by the standard error of the skewness (i.e., $(6/n)^{0.5}$, where n = 229 230 number of species) was greater than 2. Significant positive skew indicates greater prevalence of

abundant species, while significant negative skew reveals higher number of rare speciescompared to the lognormal distribution.

233

234 Statistical analyses

235 *Resampling scheme*

236 Generally, the described procedures in our study typically produced a single value without any 237 estimate of error, which inhibits statistical comparisons between datasets. Therefore, to test for 238 abiotic and biotic differences between land covers, we conducted a resampling procedure where 239 we randomly selected 50% of the streams within each land cover category for each dataset 240 without replacement 999 times. Each loop calculated the median of each physicochemical 241 variable, an estimate of physicochemical heterogeneity, biodiversity ($\overline{\alpha}$ -, β -, and γ -diversity), 242 SAD, and null model measures including the null model β -diversity values, and the within group 243 intraspecific aggregation (1/k). This procedure generated six new datasets that contained 244 resampled physiochemistry data and biotic measures, which were used further statistical 245 analyses. R scripts are available as supplementary material for online publication only (see 246 Appendices S3 and S4).

247

248 Eutrophication and physicochemical heterogeneity

249 We employed principal components analysis with all resampled, standardized median

250 physicochemical variables (mean = 0, standard deviation = 1) to create a synthetic variable

- 251 corresponding to the major physicochemical trend. The first PCA axis represented a
- eutrophication gradient and explained between 53.1% (French diatom samples) and 94.3%
- 253 (Canadian diatom samples) of the variation among samples (Appendix 1, Fig. S1.1).

To estimate physicochemical heterogeneity within each land cover, we used permutational analysis of multivariate dispersion on standardized physiochemical data with the 'betadisper' function from R package `vegan' (Anderson, Ellingsen & McArdle, 2006; Oksanen, Blanchet, Friendly, Kindt, Legendre et al., 2017). In this procedure, physicochemical heterogeneity is calculated as the average distance from a multivariate group median (group = land cover) with larger distances corresponding to greater within-group heterogeneity.

260

261 Environmental effects

262 We determined how land use-driven eutrophication and physicochemical heterogeneity affected 263 diversity components using a combination of univariate and multivariate techniques and variance 264 partitioning. For each dataset, we used permutational MANOVA function `adonis' from package 265 vegan' to test for differences in the multivariate mean of the $\overline{\alpha}$, β , and γ -diversity between 266 land covers. If the permutational MANOVA was significant, we followed with permutational 267 ANOVA using the `perm.anova' function provided in `RVAideMemore' (999 permutations; 268 Herve, 2018) for each dependent variable. We then used RDA-based variance partitioning 269 models ('vegan' function 'varpart') on each dataset to identify major explanatory factors 270 underlying diversity patterns, with eutrophication, physicochemical heterogeneity, and land 271 cover (coded as dummy variables) as predictors and the diversity measures ($\overline{\alpha}$ -, β -, and γ -272 diversity) as response variables.

273 We employed permutational MANOVA and permutational ANOVA to determine if the 274 resampled β_{DEV} , β_{SAD} , β_{AGG} , 1/k, σ , and skewness differed between land covers. Because total 275 abundance and γ -diversity influence the shape of the regional SAD, we controlled their 276 influences by regressing parameter σ against total abundance and γ -diversity of the resample and

obtaining the residuals, which were then used in subsequent analyses. To further explore if β_{DEV} was sensitive to variation in SAD unevenness (residual σ) and intraspecific spatial aggregation (1/*k*), we calculated Pearson correlations within both land cover categories for all datasets. Pearson correlations were also used to assess whether residual σ correlated with skewness and 1/*k*. We then implemented variance partitioning to determine if eutrophication, physicochemical heterogeneity, land cover, or their covariance explained the variation in β_{DEV} .

283

284 **Results**

285 Eutrophication and Environmental Heterogeneity Effects on Diversity and the SAD

286 Permutational MANOVA and permutational ANOVAs of environmental data showed that all 287 physiochemistry levels were significantly elevated (P < 0.05) in agricultural land use across all 288 datasets. Permutational ANOVAs also indicated greater physicochemical heterogeneity among 289 agricultural streams in all but the Canadian diatom dataset (higher in forest land cover) and the 290 US Fish dataset (no differences, Fig. 3). MANOVA of $\overline{\alpha}$ -, β -, and γ -diversity against land use 291 revealed that land use significantly affected the diversity measures across all datasets. Following 292 our first objective, we demonstrated that β -diversity declined with agriculture across all datasets. 293 Gamma diversity usually decreased, whereas $\overline{\alpha}$ -diversity often increased with agriculture (Table 294 2). Except for French fish, SADs were generally significantly more uneven for agricultural land 295 use than forest (higher residual σ), although the differences were mainly small (Fig. 4, columns 296 1-3). Intraspecific aggregation (1/k) was always greater in forest than in agriculture and was 297 negatively correlated with residual σ , meaning more even SADs were always associated with 298 higher aggregation (Appendix 1, Table S1.2). Skewness was significantly positive in the insect 299 and all three diatom datasets, but non-significant in the two fish datasets. When positive,

300 skewness correlated positively with residual σ regardless of land cover (although weakly for
301 diatoms), indicating that SAD unevenness was generally characterized by greater abundances of
302 more common species.

303 Our first objective was to determine how biodiversity explained by land use, 304 eutrophication, and physicochemical heterogeneity. Variation in all diversity measures was 305 primarily explained by covariance effects, while pure land cover, pure eutrophication, and pure 306 physicochemical heterogeneity contributed minorly (Fig. 5). In general, covariance of 307 eutrophication with land cover explained most of the variation, indicating that land use 308 constrained biotic variability mainly through eutrophication rather than physicochemical 309 heterogeneity. However, the insect dataset differed from the rest in that the covariance fraction of 310 land cover, eutrophication, and physicochemical heterogeneity captured most of the variation. 311

312 Eutrophication-associated shifts in local assembly across organismal groups

For our second objective, we found local assembly weakly drove diatom and insect β -diversity (β_{DEV} generally less than 0.26 across land covers) but had a relatively greater influence on fish β diversity (β_{DEV} between 0.38-0.45). β_{DEV} differed significantly between forest and agriculture (permutational ANOVA) in all datasets except insects (no difference). However, the magnitude of the difference in β_{DEV} was usually small (3.49 to 16.04%), with the direction of the difference depending on organismal group and biogeographic region (Fig. 4, column 4).

319

320 Contribution of the SAD vs. intraspecific spatial aggregation to β_{DEV}

321 For objective three, the partitioning of β_{DEV} revealed that β_{SAD} generally exceeded 100% and

 β_{AGG} was negative, regardless of land cover except for the US and French diatom datasets, which

323 showed $\beta_{SAD} < 100\%$ and positive β_{AGG} for forest land use (Fig. 4, columns 5-6). As changes in 324 β_{SAD} correspond to equal and opposite changes in β_{AGG} , we focus on β_{SAD} for brevity. β_{SAD} 325 represented nearly all of β_{DEV} regardless of dataset and land cover type (~ 90-110% of total 326 deviance) and was significantly (although marginally) larger in agricultural land use than in 327 forest cover. Further, β_{DEV} was generally negatively correlated with residual σ , regardless of land 328 cover or organismal group, implying that increased SAD unevenness was usually associated with 329 greater contribution of the regional species pool (Appendix 1, Table S2). Variance partitioning of 330 β_{DEV} across datasets showed mixed patterns among and within organismal groups over what 331 effects best explained β_{DEV} (Fig. 6).

332

333 Variability across organismal groups

Consistent with our fourth objective, we demonstrated that smaller organisms (diatoms and insects) with greater dispersal capacity were more similar in terms of SAD and β_{DEV} patterns, but diverged from fish. However, we also observed divergence in some ecological patterns between datasets within organismal groups (i.e., diatoms and fish) in that $\overline{\alpha}$ -diversity, γ -diversity, SAD skewness, and β_{DEV} responses varied between country of origin, which indicated context dependency of our results.

340

341 **DISCUSSION**

342 In this comprehensive study of stream organisms from two continents, agriculture and

343 subsequent eutrophication were generally associated with reduced β - and γ -diversity and

344 increased $\overline{\alpha}$ -diversity. First, covariance of land use with physicochemical gradients, rather than

345 with physicochemical heterogeneity, characterized regional biodiversity loss with land use.

Second, all datasets showed significant shifts in magnitude of β_{DEV} with eutrophication but the direction (i.e., stronger or weaker local assembly effects) depended on organismal group and potentially biogeographical factors. Third, the regional SAD overrode intraspecific spatial aggregation in explaining β_{DEV} and its influence and unevenness increased with agriculture.

351 Eutrophication and Environmental Heterogeneity Effects on Diversity and the SAD

352 With respect to objective one, regional biodiversity loss, local diversity gains, and increased 353 community similarity in aquatic taxa were correlated with agricultural land use, consistent with 354 patterns expected for taxonomic homogenization (Petsch, 2016). Recently, Ribiero et al. (2015) 355 explored the generality of floral homogenization consequential of agricultural land use and noted 356 that too many studies focus on a single spatial scale or a single taxon. For aquatic taxa, 357 agriculturally-associated changes in β -diversity have been reported, however, we have only 358 begun to examine these changes at broader spatial scales. For example, Winegardner et al. 359 (2017) attributed greater temporal β -diversity of diatoms across modified US landscapes to 360 richness gains and losses stemming from disproportionate influence of contemporary vs. past 361 land use, yet observed no changes in spatial β -diversity. In contrast, diatom spatial β -diversity 362 declined with eutrophication in French streams (Jamoneau, Passy, Soininen, Leboucher & Tison-363 Rosebery, 2018). Our investigation, exploring diatoms, insects, and fish across regional to 364 subcontinental scales, demonstrates that the detrimental effects of agriculture on the regional 365 biodiversity in stream ecosystems are independent of species biology or scale.

We further revealed that biodiversity variation between forest and agriculture was mainly driven by land use differences in physicochemistries rather than physicochemical heterogeneity, a result contrary to conventional wisdom that higher environmental heterogeneity brings greater

369 turnover. While agriculture may homogenize the landscape, we show that it tended to lead to 370 greater stream physicochemical heterogeneity, possibly due to variability in fertilization and 371 landscape management regimes. Heterogeneity is an important mechanism of co-existence 372 because it offsets competitive exclusion (Tilman & Pacala, 1993). However, we observed that 373 physicochemical heterogeneity poorly explained β -diversity, because eutrophication in 374 agricultural streams may have exceeded the physiological thresholds of sensitive species and 375 decoupled compositional and environmental variability (Bini, Landeiro, Padial, Siqueira & 376 Heino, 2014). The lack of a relationship may also be due to our measure of heterogeneity, which 377 did not incorporate other aspects of heterogeneity, such as variability in substrate size, known to 378 diminish with agriculture (Allan, 2004).

379 Increased prevalence of common species over spatial and temporal scales is a hallmark of 380 taxonomic homogenization (Olden & Rooney, 2006), but our findings are restricted to the spatial 381 dimension. Notably, while across datasets SADs were generally more uneven in agriculture, 382 they were more positively skewed compared to forest only in two datasets, i.e. US insects and 383 French diatoms. In these datasets, homogenization in agriculture was characterized by greater 384 prevalence of common relative to rare species, which has also been observed in terrestrial 385 arthropods (Simons, Gossner, Lewinsohn, Lange, Türke et al., 2015; Komonen & Elo, 2017). 386 However, SADs were more positively skewed in forest cover than in agriculture for two datasets 387 (US and Canadian diatoms), and not skewed for both fish datasets. This suggested that stronger 388 SAD unevenness in agriculture resulted from either buildup of common species or greater 389 regional dominance by a relatively few species. Like recent terrestrial and tropical studies 390 (Vázquez & Gaston, 2004; Lohbeck, Bongers, Martinez-Ramos & Poorter, 2016), we showed 391 that SAD unevenness was associated with agriculturally-driven homogenization. Future research

on homogenization should incorporate novel methods and procedures, like we employed, to
elucidate how habitat modification and trait distribution contribute to the two forms of
unevenness, i.e. asymmetry vs. dominance.

395

396 Land use-associated shifts in local assembly across organismal groups

397 Following objective two, we examined how local assembly (β_{DEV}) varied between forested and 398 agricultural streams. In general, β_{DEV} marginally differed between land covers, suggesting that 399 the strength of local vs. regional mechanisms was relatively unaffected by physicochemical 400 stressors, consistent with prior work, reporting that fire disturbance altered β -diversity but not its 401 causes (Myers *et al.*, 2015). Community comparisons revealed that the magnitude of β_{DEV} 402 usually increased with body size, which here was linked with dispersal capacity. Smaller β_{DEV} 403 values in diatoms and insects indicated that the observed species pool exerted greater influence 404 on β -diversity relative to local assembly. These results are corroborated by earlier research 405 showing that diatom and insect communities are unsaturated, whereby local richness is limited 406 by the size of the regional pool as opposed to local interactions (Passy, 2009; Al-Shami, Heino, 407 Che Salmah, Abu Hassan, Suhaila et al., 2013; but see Thornhill, Batty, Death, Friberg & 408 Ledger, 2017). Therefore, it is possible that regional effects play a greater role in structuring 409 local richness and β -diversity of smaller and more dispersive organisms than of larger and less 410 dispersive organisms, and these relationships are not consistently affected by eutrophication. 411 In contrast, β_{DEV} in both fish datasets approaching 0.50 suggested relatively similar local 412 and regional control of β -diversity, in agreement with prior observations of comparable 413 contributions of regional and local factors to fish richness (Angermeier & Winston, 1998). 414 Taxonomic homogenization is a particularly prevalent phenomenon among freshwater fish

415 (Petsch, 2016) and our study elucidated that the possible causes include both local and regional416 processes.

417 Other nearly uniform patterns, independent of land use, were the negative correlation of 418 residual σ of the regional SAD and the positive correlation of intraspecific aggregation (1/k) with 419 β_{DEV} . These correlations indicated that more even regional SADs and increased intraspecific 420 spatial aggregation were associated with stronger local constraints on β -diversity. Recent work 421 has only begun to explore the relationship of SAD evenness with taxonomic homogenization, 422 showing clear links between the two with implications for conservation (e.g., Simons et al., 423 2015; Komonen & Elo, 2017). Our study is novel in that it demonstrates that local and regional 424 processes controlling β -diversity are dependent on SAD evenness—a finding that could guide 425 future stream conservation and management decisions, which need to be scale-explicit. For 426 example, if preserving β -diversity, then adopting practices promoting abundance of less common 427 species may be beneficial, given that SAD evenness is positively correlated with β -diversity. 428

429 The contribution of the SAD vs. intraspecific spatial aggregation to β_{DEV}

430 To our knowledge, we are the first to explore how land use affects partitioning of β_{DEV} into SAD 431 vs. spatial aggregation fractions, i.e. β_{SAD} vs. β_{AGG} (objective three). β_{SAD} accounted for most of 432 β_{DEV} , similar to observations for global tree communities (Xu *et al.*, 2015), but opposite to 433 findings, with a different null model, for Czech forests (Sabatini, Jiménez-Alfaro, Burrascano, 434 Lora & Chytrý, 2017). We further discovered that β_{SAD} largely exceeded β_{AGG} across organismal 435 groups, datasets, and land cover types. However, β_{SAD} was significantly higher in agriculture 436 compared to forest in all datasets. The two land covers also diverged in β_{AGG} —less spatial 437 aggregation than predicted by the null model ($\beta_{AGG} < 0$) was detected in agriculture across all

438 datasets, while some aggregation ($\beta_{AGG} > 0$) was observed in forest streams in four out of six 439 datasets. These results suggest that although land use did not constrain the magnitude of local 440 assembly effects (β_{DEV}), it did control the mechanisms of local assembly, i.e. land use increased 441 the role of the SAD, but diminished the influence of aggregation.

442

443 Organismal and geographic dependencies in biodiversity response to homogenization

444 In pursuit of our fourth objective, we found that organismal groups responded differently to land 445 use, as reported by other studies (e.g., Angermeier & Winston, 1998; Thornhill et al., 2017). Insects 446 resembled diatoms in biodiversity, SAD shape, and β_{DEV} patterns, which suggested that body 447 size and dispersal capacity may be more important than trophic position (autotroph vs. 448 heterotroph) in predicting ecological responses to agricultural eutrophication. We generally 449 expected consistent responses of these metrics to agriculture, regardless of country of origin (i.e., 450 diatoms and fish). We reasoned that agriculture, being a major habitat alteration, will override all 451 other influences, yet within both groups, there was divergence depending on region. We ensured 452 that variation in individual counts and mean counts among samples and differences in 453 geographic spread across datasets did not contribute to their dissimilarity (data not shown). Thus, 454 our findings of within-taxon variability with respect to biodiversity and the SAD highlighted the 455 importance of considering context dependency. Histories of land use disturbance among 456 geographic regions can set biodiversity and relative abundance patterns on different trajectories 457 by affecting processes underlying β -diversity (Cramer, Hobbs & Standish, 2008). For example, 458 European fish diversity has been historically depauperate relative to North American fauna 459 owing particularly to differences in glacial influence (Oberdorff, Hugueny & Guégan, 1997).

460 Furthermore, French aquatic communities have been impacted by agricultural activities far461 longer than their North American counterparts (Hahn & Orrock, 2015).

462 In summary, we determined eutrophication is a major driver of β -diversity losses among 463 stream taxa, although the importance of geographic context was shown through the varied 464 biodiversity responses within taxonomic groups. Local assembly generally was weakly affected 465 by agriculture. However, in agriculture the regional SAD became significantly more uneven and 466 its effect on local assembly significantly increased compared to forest, which may be the 467 underlying causes of taxonomic homogenization. Biodiversity, SAD shape, and β_{DEV} depended 468 more strongly on body size and/or dispersal than trophic position. Future research should explore 469 how local and regional processes operate in tandem with the SAD to uncover whether 470 homogenization drivers are specific to organismal groups and the regions from which they were 471 sampled. Although we examined β -diversity loss from a taxonomic perspective, we recommend 472 future investigations on whether agriculture leads to phylogenetic and functional homogenization 473 across space and time. Then, taxonomic, phylogenetic, and functional diversity responses to 474 agriculture could be compared to generate more holistic understandings of the causes and 475 patterns of biotic homogenization.

476

477

478 ACKNOWLEDGEMENTS

We thank W. Xu for helpful comments on the null model. Support by the National Science
Foundation (grant NSF DEB-1745348) to SIP is gratefully acknowledged.

481

482

483 DATA AVAILABILITY STATEMENT

- 484 Community and environmental datasets for the US are available for download from the USGS
- 485 NAWQA Program via the Water Quality Portal (https://www.waterqualitydata.us/contact_us/)
- 486 and EPA NRSA (https://www.epa.gov/national-aquatic-resource-surveys/data-national-aquatic-
- 487 resource-surveys) databases. DOI with associated URLs for all community and environmental
- 488 datasets analyzed for this project will be made freely available for download from Dryad Digital
- 489 Repository upon publication.
- 490

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- 622
- 623

Table 1. Summary of procedures and analyses performed with corresponding expectations and observations.

Analyses Exp

Expectations

Observations

1) Determine the differences in	PCA,	Land cover would be characterized	1) All analyses clearly separated streams
physicochemistry and	PERMDISP,	effectively by physicochemical	into two groups, corresponding to forest and
physicochemical heterogeneity	MANOVA	parameters and potentially by	agriculture;
between land covers.		physicochemical heterogeneity.	2) Agricultural streams had elevated

nutrient levels, suggestive of eutrophication;

greater among agricultural streams except in

3) Physicochemical heterogeneity was

the Canadian diatom dataset.

2) Reveal the responses of $\overline{\alpha}$ -, γ -, and MANOVA, β -diversity, SAD evenness, and SAD Variance skewness to physicochemistry and partitioning physicochemical heterogeneity. The responses of biodiversity components to physicochemistry and physicochemical heterogeneity may differ depending on body size, dispersal capacity, and trophic level (autotroph vs. heterotroph).

1) In general, β - and γ -diversity were negatively related to eutrophication, whereas $\overline{\alpha}$ -diversity increased. SADs tended to be more uneven in agricultural streams due to buildup of common species and/or increased dominance; 2) Covariance of land use with physiochemistry explained most of the diversity variation across datasets, whereas environmental heterogeneity poorly explained diversity; 3) Pure land cover and pure physicochemistry generally explained some additional variation in the diversity components.

3) Determine if land use influences	Null models,	The contribution of local assembly	1) The role of local assembly was generally
the relative roles of local assembly	Permutational	should be responsive to agricultural	weakly affected by land use, and not in a
and the regional species pool in	ANOVA,	land use, however, the magnitude	consistent way across datasets, suggesting a
driving β -diversity.	Variance	and direction of the response may	potential influence of organismal type and
	partitioning	vary across organismal groups.	biogeography.

4) Determine if β -deviation (β_{DEV}) is Null models, explained by the species abundance Permutationa distribution (SAD) or intraspecific ANOVA spatial aggregation.

Null models, It is unknown how land use may Permutational influence the fractions of β_{DEV} ANOVA explained by the SAD and intraspecific spatial aggregation. The SAD was the dominant fraction of β_{DEV} and this pattern was independent of
 land use and organismal group. However,
 the SAD fraction was significantly higher in agriculture across all datasets, which may be
 the underlying factor of taxonomic
 homogenization;
 Intraspecific spatial aggregation fraction
 was negative for agricultural streams and
 positive for forest streams, indicating that

intraspecific aggregation was lower than

expected across disturbed streams;

Table 2. Summary of the impact of agricultural land use on resampled diversity measures as positive or negative percent change relative to forest cover. Significant differences between land covers were detected in all comparisons (permutational MANOVA and ANOVA, P < 0.05).

		% Change from agriculture			
Taxonomic group	Country	α	γ	¹ β _{OBS}	
Diatoms	US	+20.71	-3.54	-1.09	
	France	+13.33	-7.46	-2.22	
	Canada	-12.15	-23.14	-4.64	
Insects	US	-20.42	-22.98	-0.59	
Fish	US	+9.55	+12.97	-2.29	
	France	+54.99	+26.91	-6.41	

 ${}^{1}\beta$ -diversity = $1 - \frac{\overline{\alpha}}{\gamma}$

FIGURE LEGENDS

Figure 1. a. Conceptual model depicting the land use effect on the species abundance distribution (SAD) and intraspecific spatial aggregation, which in turn interact with local (α) and regional (γ) diversity. β-diversity is calculated as a function of average α-diversity and γdiversity. Interactions that were controlled for by the null models of Kraft *et al.* (2011) and Xu *et al.* (2015) are marked with a thick dotted line. **b.** Diagram summarizing the Xu *et al.* (2015) partition of β_{DEV} into fractions explained by the SAD and intraspecific spatial aggregation using an occupancy-abundance based null model procedure. The null model β_{DEV} is taken as the raw difference between expected β-diversity (β_{EXP}) and observed β-diversity (β_{OBS}). The fraction of β_{DEV} explained by the SAD, β_{SAD}, is the difference between predicted β-diversity (β_{PRED}) and expected β-diversity (β_{EXP}), whereas the fraction of β_{DEV} explained by intraspecific aggregation (β_{AGG}) represents the difference between β_{OBS} and β_{PRED}.

Figure 2, a-f. Maps of diatom, macroinvertebrate, and fish sampling localities in the US, France, and Canada. Grey triangles represent agriculture samples, whereas black circles represent forest samples. a = US diatoms, b = US insects, c = US fish, d = French diatoms, e =French fish, f = Canadian diatoms.

Figure 3, a-f. Boxplots showing differences in resampled physicochemical heterogeneity between land covers for each dataset. a = US diatoms, b = French diatoms, c = Canadian diatoms, d = US insects, e = US fish, f = French fish. Points indicate resamples that fall outside

the interquartile range. Different letters denote significant differences in mean heterogeneity (permutational ANOVA, P < 0.05).

Figure 4, a-f. Boxplots of resampled SAD and null model metrics showing the differences between land covers for each dataset. a = US diatoms, b = French diatoms, c = Canadian diatoms, d = US insects, e = US fish, f = French fish. Significant differences were observed in all comparisons (permutational ANOVA, P < 0.05) except β_{DEV} for US insects (panel D3, denoted by asterisk).

Figure 5, a-f. Venn diagrams showing output of redundancy analysis-based variance partitioning of diversity measures ($\overline{\alpha}$ -, β -, and γ -diversity). Values represent model adjusted R² values. Values in intersections represent covariance fractions, whereas values in circles represent pure fractions.

Figure 6, a-f. Venn diagrams showing output of regression-based variance partitioning of β_{DEV} . Values represent model adjusted R² values. Values in intersections represent covariance fractions whereas values in circles represent pure fractions.

Figures

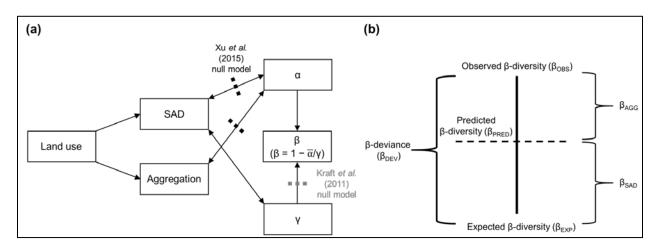
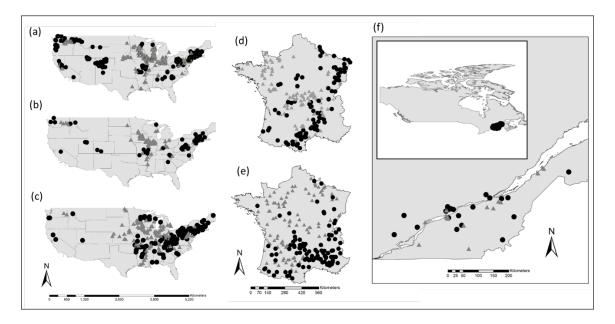


Figure 1





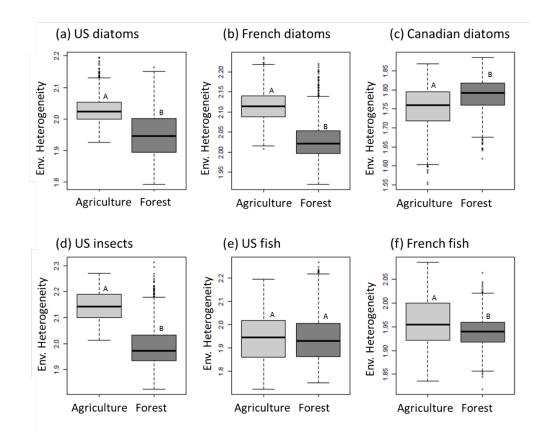


Figure 3

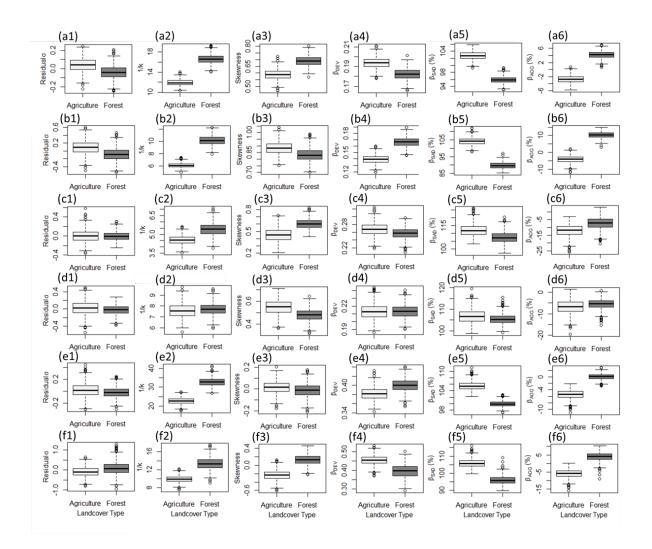


Figure 4

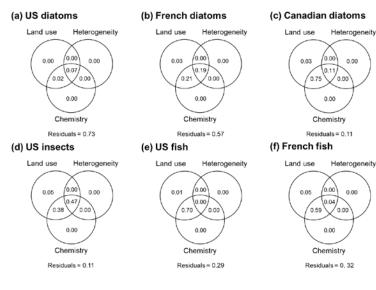


Figure 5

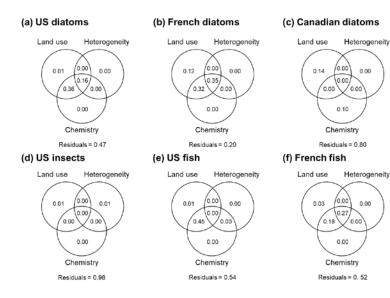


Figure 6

Supporting Information Appendix Short Titles

- Appendix 1: Expanded description of environmental data and null model correlation results
- Appendix 2: Description of Null Model Machinery
- Appendix 3: R-code script for looping procedures
- Appendix 4: R-code script for analyses of loop output