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# 1 TITLE

2 **Similar patterns of background mortality across Europe are mostly driven by drought in**  
3 **European beech and a combination of drought and competition in Scots pine.**

4

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21

## 22 1 INTRODUCTION

23 Tree mortality plays a major role in shaping forest dynamics, structure and composition  
24 (Franklin *et al.*, 1987; Ruiz-Benito *et al.*, 2017a), species range shifts (Benito Garzón *et al.*, 2013),  
25 ecosystem functioning and services (Millar & Stephenson, 2015), carbon fluxes and feedback to the

26 global climate system (Sitch *et al.*, 2008). Therefore, understanding and predicting tree mortality is  
27 a key challenge in ecology, particularly in a changing climate.

28         Global change is exacerbating drought-induced tree mortality (Allen *et al.*, 2015). Recent  
29 forest die-off events have occurred in all major biomes and on every wooded continent (Allen *et al.*,  
30 2015) and background tree mortality also appears to have increased in North America (Mantgem *et*  
31 *al.*, 2009; Hember *et al.*, 2017) and in Spain (Carnicer *et al.*, 2011). Less conspicuous than die-offs  
32 events, minor large-scale changes in tree background mortality can have a huge impact on forest  
33 ecosystems and dynamic, including changes in productivity rates, functional composition and  
34 species turnover (Stephenson & Mantgen, 2005; Ruiz-Benito *et al.*, 2017b). Yet, an empirical  
35 quantification of background tree mortality at continental scale is missing and whether or not forest  
36 mortality follows an increasing global trend that will keep rising under global change remains  
37 unclear (Hartmann *et al.*, 2018). Moreover, mortality is a major process which delimits species  
38 range (Gaston, 2009), notably at the driest edge of their distribution (Benito Garzón *et al.*, 2013).  
39 Therefore, large scale studies that capture the entire species distribution are essential to determine  
40 how climate change induced mortality might affect species distribution.

41         Understanding and predicting background tree mortality patterns at large scales remains  
42 challenging for several reasons (but see Das *et al.* (2016) and Neuman *et al.* (2017) for examples of  
43 large-scale studies). First of all, mortality is a stochastic phenomenon (Franklin *et al.*, 1987), which  
44 is therefore difficult to predict. Secondly, it is often the result of a complex and gradual process with  
45 multiple interacting drivers (Manion, 1981), that act at different spatial and temporal scales (Dietze  
46 & Moorcroft, 2011). Thirdly, there may be a lag time between episodic stressful conditions and tree  
47 mortality responses (Cailleret *et al.*, 2016; Jump *et al.*, 2017). Lastly, background tree mortality  
48 rates are difficult to estimate due to the small sample size of dying trees in local studies, while large  
49 samples are needed to understand mortality patterns.

50 In European forests, background tree mortality is strongly driven by climate variability  
51 (Neumann *et al.*, 2017). Among the climatic factors affecting tree mortality, drought plays a major  
52 role (McDowell *et al.*, 2008; Benito Garzón *et al.*, 2013; Ruiz-Benito *et al.*, 2013; Allen *et al.*,  
53 2015) and particularly affects populations at the driest edge of species distributions (Benito Garzón  
54 *et al.*, 2018). Among the biotic factors, competition for limited resources may be an important cause  
55 of tree mortality and may also interact with climate, notably through a higher increase in mortality  
56 rates in areas that are both dry and dense (Ruiz-Benito *et al.*, 2013; Vilà-Cabrera *et al.*, 2013; Young  
57 *et al.*, 2017). Moreover, tree mortality responses can differ widely depending on whether we  
58 consider intra- or inter-specific competition (Condés & del Río, 2015). However, how intra- and  
59 interspecific competition interact with climatic drought to shape range-wide mortality patterns  
60 remains unknown.

61 Tree mortality sensitivity to biotic and abiotic factors vary along species' ecological  
62 strategies, from stress-tolerators to competitors and from angiosperms to gymnosperms (Choat *et*  
63 *al.*, 2012; Ruiz-Benito *et al.*, 2017a). European beech (*Fagus sylvatica* L.) and Scots pine (*Pinus*  
64 *sylvestris* L.) are two widely distributed European tree species with different life history strategies.  
65 Beech is a highly competitive, shade-tolerant and late-successional species while Scots pine is a  
66 weakly competitive and light demanding pioneer tree (San-Miguel-Ayanz *et al.*, 2016). Scots pine is  
67 considered as a drought-avoiding species, which has narrow xylem vessels (tracheids) and  
68 withstands droughts by closing its stomata early and maintaining a reduced metabolism (isohydric  
69 response; McDowell *et al.*, 2008). By contrast, beech is known to be sensitive to drought (van der  
70 Maaten, 2012; Chen *et al.*, 2015), has wide vessels more prone to cavitation and maintains a more  
71 constant metabolism but a narrower margin of hydraulic safety under droughts (anisohydric  
72 response; McDowell *et al.*, 2008). Regional scale studies suggested that both species are being  
73 progressively replaced by other species in the southern part of their distribution (Vilà-Cabrera *et al.*,

74 2013; Galiano *et al.*, 2010) and in some inner Alpine valleys in the case of *P. sylvestris* (Rigling *et*  
75 *al.*, 2013).

76 Our main objective was to understand and predict range-wide patterns of background  
77 mortalities in Scots pine and European beech. To that end, we parameterised individual-level  
78 logistic regression models, as a function of climatic drought and basal area of heterospecific and  
79 conspecific trees (used as a proxy of inter and intra-competition or facilitation, at the plot level),  
80 using records from five National Forest Inventories covering the entire European latitudinal  
81 gradient, from Spain to Finland. We hypothesised that (i) mortality in both species is influenced by  
82 climatic drought, basal area and their interaction but with a higher influence of basal area in the case  
83 of Scots pine; and (ii) that despite these differences in their sensitivity to drought and basal area,  
84 both species display similar spatial patterns of mortality across their ranges: high mortality in the  
85 south resulting from increasingly dry climates, especially in the Mediterranean biome.

86

## 87 **2 MATERIAL AND METHODS**

### 88 **2.1 Forest inventory data**

89 We used mortality data from five national forest inventories (NFIs) covering the entire  
90 European latitudinal gradient, from the Mediterranean to boreal biome. Data from four of the NFIs  
91 had been previously harmonised as part of FunDivEUROPE project (Spain, Germany, Sweden and  
92 Finland) and the French NFI was added to this study. In each NFI, trees were recorded in temporary  
93 or permanent plots depending on the country. Plots in the German, Finnish and Swedish inventories  
94 are gathered within clusters (see Appendix S1 for details of the survey design and sampling  
95 methods for each NFI). We selected plots in which at least one of our two target species (i.e. *F.*  
96 *sylvatica* or *P. sylvestris*) was recorded. These plots were classified into Mediterranean, cool  
97 temperate and boreal biomes (see the map of biome boundaries in Fig. S1.1) and were unevenly  
98 distributed along the latitudinal gradient (Fig. S1.2). The final datasets contained 57,191 beech trees

99 and 161,720 Scots pine trees in 10,150 plots and 16,669 plots, respectively. From those trees, 1,490  
100 (2.6%) and 7,649 (4.7 %) were recorded as dead for beech and Scots pine, respectively.

101 As explanatory variables of tree mortality, we selected tree *DBH* (diameter at breast  
102 height) as *DBH* is known to influence individual tree mortality (Ruiz-Benito *et al.*, 2013). We  
103 additionally calculated three proxies of indirect competition between trees (or facilitation) (Fig.  
104 S2.1): basal area of neighbouring trees considering all tree species (i.e. *BAall*, m<sup>2</sup> ha<sup>-1</sup>), basal area of  
105 neighbouring conspecifics (i.e. *BAintra*, m<sup>2</sup> ha<sup>-1</sup>) and basal area of neighbouring heterospecifics (i.e.  
106 *BAinter*, m<sup>2</sup> ha<sup>-1</sup>).

## 107 **2.2 Drought-related variables**

108 Climatic drought intensity over the study period (Fig. S2.2) was characterised by a water  
109 availability index:  $WAI = (MAP - PET) / PET$ , where *MAP* is the mean annual precipitation (mm) and  
110 *PET* the mean potential evapotranspiration (mm). For each plot, *PET* was extracted from the CRU  
111 v3.24.01 monthly gridded dataset at 0.5-degree resolution (Harris *et al.*, 2013) and *MAP* was  
112 calculated from a downscaled version of E-OBS at 1 km resolution (Moreno & Hasenauer, 2016).  
113 For each plot, *WAI* was averaged over the period between two years before the first survey date and  
114 the second survey date to include delayed effects of drought on mortality (Greenwood *et al.*, 2017).

115 Changes in climatic drought intensity over the study period (i.e. temporal variability of  
116 drought intensity) were described by the Standardized Precipitation Evapotranspiration Index  
117 (*SPEI*, Fig. S2.2; Vicente-Serrano *et al.*, 2009), obtained from a gridded dataset at 0.5-degree  
118 resolution (Beguería & Vicente Serrano, 2017). *SPEI* is a multi-scalar drought index whose  
119 variations have been shown to be highly correlated with tree response to climate (Greenwood *et al.*,  
120 2017). Its calculation considers both *PET* and *MAP*, with *PET* derived from the Penman-Monteith  
121 equation. *SPEI* compares drought intensity during a long-term reference period (i.e. from 1901 to  
122 2015) to that of a given period from 3 to 48 months. In our study, we selected a 12-month period to  
123 consider both current and previous year water shortage. *SPEI* is expressed as a standardised index

124 relative to each site, with a standard deviation of 1, where negative values indicate more intense  
125 drought over the timescale considered compared to reference conditions. For each plot, we  
126 calculated mean *SPEI* (hereafter *SPEI*) over the period from two years before the first survey date to  
127 the second survey date.

## 128 **2.3 Model description**

129 We parameterised two species-specific models, where  $P_i$  is the annual probability of  
130 mortality for each individual tree  $i$ . We used a logistic regression model with a link *cloglog* to allow  
131 the sigmoidal curve of the mortality probability to be asymmetrical and deal with zero inflated  
132 distributions (Zuur *et al.*, 2009):

$$133 P_i = 1 - \exp(-\exp(\alpha_0 + \alpha_{country} + k_{i,sp} + \log(t_i))) \quad (1)$$

134 where  $\alpha_0$  is an intercept term (set to zero);  $\log(t_i)$  is an offset variable that takes into account the  
135 survey interval length  $t_i$  (years) for each tree  $i$ ;  $\alpha_{country}$  is the random country intercept to include the  
136 sampling differences between each NFI and  $k_{i,sp}$  is a species-specific linear function that includes  
137 the relationship between the mortality of tree  $i$  of the species  $sp$  (i.e. *F. sylvatica* or *P. sylvestris*) and  
138 the explanatory fixed-effect variables. Although clusters and plots could be considered as a source  
139 of variation for each tree, we did not consider cluster and plot as random terms because most of the  
140 clusters contained only one plot and in many plots no trees died between the two survey dates. We  
141 used the function “glmer” of the “lme4” package to run the model described in equation 1 in R 3.3.3  
142 (R Core Team 2017).

143 For both species, we explained mortality patterns using four fixed-effect predictors with  
144 low collinearity (i.e. Spearman correlation coefficient:  $r < 0.59$ , and Variance Inflation Factor: VIF  
145  $< 2$ ; Dormann *et al.*, 2013), namely: *BAintra*, *BAinter*, *WAI* and *SPEI*. Conspecific and  
146 heterospecific basal area (i.e. *BAintra* and *BAinter*) were both included in the model as they can  
147 have different effects on tree mortality (Condés & del Río, 2015). To ensure a linear relationship

148 between each explanatory variable and tree mortality, *BAinter*, *WAI* and *SPEI* were log transformed  
149 (see Appendix S3 for details).

150 Tree size (*DBH*) was included as a covariate in our model, as we were not directly  
151 interested in the importance of tree size on mortality. As we required a single parameter per  
152 predictor to estimate the relative importance of each predictor (see section 2.5), we calculated a  
153 non-linear variable from *DBH*:  $DBHnl_{sp} = DBH + r_{sp} \times \log(DBH)$  (see Appendix S3 for details).

154 To understand how tree mortality was affected by basal area and climatic drought, we  
155 included the main effect of each variable and first-order interaction terms between abiotic and biotic  
156 variables. Herewith, the function *k* from equation 1 took the form:

$$\begin{aligned} 157 \quad k_{i,sp} = & \beta_{1,sp} \times \log(WAI_i) + \beta_{2,sp} \times \log(SPEI_i) \\ 158 \quad & + DBHnl_{i,sp} \times (\beta_{3,sp} + \gamma_{1,sp} \times \log(WAI_i) + \gamma_{2,sp} \times \log(SPEI_i)) \\ 159 \quad & + BAintra_i \times (\beta_{4,sp} + \gamma_{3,sp} \times \log(WAI_i) + \gamma_{4,sp} \times \log(SPEI_i)) \\ 160 \quad & + \log(BAinter_i) \times (\beta_{5,sp} + \gamma_{5,sp} \times \log(WAI_i) + \gamma_{6,sp} \times \log(SPEI_i)) \end{aligned} \quad (2)$$

161 where  $\beta_x$  and  $\gamma_x$  are the estimated coefficients of the main and interaction effects, respectively (Table  
162 S3.1).

## 163 **2.4 Model performance and evaluation**

164 Binned residuals plots were used to ensure our final species-specific models were well-  
165 calibrated (Fig. S3.3-4). To evaluate the discrimination accuracy of our models, we computed the  
166 mean area under the curve (AUC) on 100 bootstrap samples among the predicted and observed  
167 values. AUC values of 0.6-0.7 show a fair discrimination accuracy, between 0.7 and 0.8 good and  
168 above 0.8 excellent (Hurst *et al.*, 2011). We used independent cross-validation to measure the  
169 generalisation power of the model, for which we used 75% of the data to fit the model and the  
170 remaining 25% to independently validate our predictions.

171



## 172 **2.5 Relative importance of climatic drought and basal area on mortality**

173           Following Ratcliffe *et al.*, (2016), we explored the relative importance of each predictor on  
174 individual tree mortality in relation to the other predictors by considering the predictors' main  
175 effects and their interactions. For doing so, we first computed the absolute importance of each  
176 predictor using our model coefficients. For instance, to compute  $A_{BAintra,i}$  the absolute importance of  
177  $BAintra$  on the probability of mortality of the tree  $i$ , we applied the following equation separately  
178 for each species:

$$179 \quad A_{BAintra,i,sp} = \beta_{4,sp} + \gamma_{3,sp} \times \log(WAI_i) + \gamma_{4,sp} \times \log(SPEI_i) \quad (3)$$

180 where  $\beta_x$  and  $\gamma_x$  are the estimated coefficients of the single predictors and their interaction effects  
181 respectively;  $WAI_i$  and  $SPEI_i$ , are the plot values corresponding to these variables.

182           Secondly, the relative importance of each predictor was computed for each tree by dividing  
183 the absolute importance of the focal predictor by the maximum absolute importance between all  
184 predictors of the target tree. For instance, to estimate the relative importance of  $BAintra$  for the tree  
185  $i$ , we calculated for each species:  $|A_{BAintra,i,sp}| / \max(|A_{BAintra,i,sp}|, |A_{SPEI,i,sp}|, |A_{WAI,i,sp}|, |A_{BAinter,i,sp}|)$ ;  
186 where  $A_{SPEI,i}$ ,  $A_{WAI,i}$  and  $A_{BAinter,i}$  are the absolute importance of  $SPEI$ ,  $WAI$  and  $BAinter$  for tree  $i$ ,  
187 respectively. For each tree  $i$ , the predictor that had the greatest influence on individual tree mortality  
188 probability had a relative importance of one.

189

## 190 **3 RESULTS**

### 191 **3.1 Model performance and validation**

192           Scots pine and beech models showed good agreement between observed and predicted  
193 values (AUC = 0.73 and 0.71, respectively). The Scots pine model performed well in predicting  
194 annual tree mortality probability across the European latitudinal gradient as predicted and observed  
195 values exhibited similar patterns (Fig. 1a). Nevertheless, caution is needed to interpret the results at  
196 the southern part of the latitudinal gradient where Scots pine mortality probability was slightly

197 underestimated. In the beech model, mortality probability was accurately predicted in the southern  
198 half of the latitudinal gradient but was overestimated between 48° and 54° latitude, which  
199 corresponds mainly to northern Germany (Fig. 1b). Model and partial residual plots for each  
200 predictor showed no strong spatial patterns, thus supporting the validity of the models (Fig. S3.3-4).

### 201 **3.2 Relative importance of climatic drought and basal area across latitude**

202 In the case of Scots pine, basal area variables (i.e. *BA<sub>intra</sub>* and *BA<sub>inter</sub>*) were more  
203 important than drought-related variables (i.e. *WAI* and *SPEI*) in explaining the probability of  
204 mortality across the latitudinal gradient (Fig. 2a and Table 1). The conspecific basal area was the  
205 most important driver from south to north with a mean relative importance of 0.96 (Table 1). The  
206 order of importance of the four predictors was stable across latitude, except from 43° to 45° latitude  
207 (corresponding to the French part of the Mediterranean biome) where drought-related variables  
208 (mainly *SPEI*) were nearly as important as basal area variables (Fig. 2a). From south to north, high  
209 levels of both conspecific and heterospecific basal area and increases in drought intensity (i.e. low  
210 *SPEI*) were correlated with higher probability of mortality (Fig. 2a). In contrast, low *WAI* was  
211 associated with high mortality probabilities in the Mediterranean biome and with low mortality  
212 probabilities in the boreal biome (see changes from negative to positive influence in Fig. 2a).

213 For beech trees, drought-related variables were more important than basal area variables in  
214 explaining mortality probability across the major part of the latitudinal gradient (except in the  
215 south) with a mean relative importance of 0.74 and 0.70 for *WAI* and *SPEI*, respectively (Fig. 3a  
216 and Table 1). Low *WAI* and *SPEI* were associated with higher mortality rates (see negative  
217 influence in Fig. 3a). The relative importance of conspecific basal area remained stable across  
218 latitude whereas that of heterospecifics varied from being the most important variable explaining  
219 beech mortality in the Mediterranean biome to being the least important one in the cool temperate  
220 biome (Fig. 3a and Table 1). Beech mortality probability increased with conspecific basal area and  
221 decreased with heterospecific basal area (Fig. 3A and Table 1).

### 222 **3.3 Interactions between climatic drought and basal area**

223 In the Scots pine model, all interactions between drought-related variables (i.e. *WAI* and  
224 *SPEI*) and basal area variables (i.e. *BA<sub>intra</sub>* and *BA<sub>inter</sub>*) were significant (Table S3.1). The  
225 strongest interaction was between climatic drought intensity and conspecific basal area (i.e. *WAI*  
226 and *BA<sub>intra</sub>*; Fig 3b & Table S3.1): regardless of drought intensity, the probability of mortality  
227 remained weak when the conspecific basal area was low or intermediate, whereas it strongly  
228 increased in dry areas where the conspecific basal area was high (Fig. 2b; see Fig. S4 for the other  
229 interactions that affected mortality weakly, albeit significantly).

230 In the beech model, the only significant interaction was that between climatic drought and  
231 heterospecific basal area (*WAI* and *BA<sub>inter</sub>*; Table S3.1): the probability of mortality increased in  
232 dry areas where heterospecific basal area was low or intermediate, while the probability of mortality  
233 remained stable (and always low) in dry areas where heterospecific basal area was high (Fig. 3b).

### 234 **3.4 Spatial patterns of predicted tree mortality across Europe**

235 Across their range, the predicted annual probability of Scots pine mortality was on average  
236 higher than that of beech (0.0061 and 0.0038, respectively; Table 1) but followed the same trend  
237 across the latitudinal gradient (Fig. 4). The highest predicted mortality rates for both species were in  
238 south-eastern France, at the ecotone between the Mediterranean and cool temperate biomes (Fig. 4).

239 The predicted rates of Scots pine mortality were highest in the Mediterranean biome (mean  
240 value of 0.0077 for 62,165 trees), intermediate in the cool temperate biome (mean value of 0.0063  
241 for 62,914 trees) and lowest in the boreal biome (mean value of 0.0033 for 36,641 trees) (Table 1).  
242 Similarly, the predicted individual probability of beech mortality was higher in the Mediterranean  
243 biome (mean value of 0.0052 for 9,315 trees) than in the cool temperate biome (mean 0.0035 for  
244 47,876 trees) (Table 1). However, the gap between mortality rates in the Mediterranean biome and  
245 the more northern biomes is likely to be higher than predicted as the Scots pine model slightly

246 underestimated mortality probability in the Mediterranean biome (Fig. 1a) and the beech model  
247 overestimated mortality probability in the cool temperate biome (Fig. 1b).

248

## 249 **4 DISCUSSION**

250 Exploring the drivers of background tree mortality at a continental scale opens a new  
251 perspective for understanding tree mortality patterns across species' ranges, including some  
252 demographic events observed at a smaller scale (Carnicer *et al.*, 2011). Although considerable  
253 attention has been paid to the effects of drought and basal area on tree mortality (Mantgem *et al.*,  
254 2009; Greenwood *et al.*, 2017; Hember *et al.*, 2017; Ruiz-Benito *et al.*, 2013), our results  
255 demonstrate that the combination of the two, through direct and indirect effects that vary along  
256 geographical gradients and between the two species (Fig. 2 and 3), is shaping background mortality  
257 across species' ranges (see also Ruiz-Benito *et al.*, 2013; Jump *et al.*, 2017; Young *et al.*, 2017).  
258 Interestingly, both species had similar patterns of predicted mortality, with the highest mortality  
259 rates in the southern French part of the Mediterranean biome (Fig. 4).

### 260 **4.1 Increase in climatic drought intensity associated with higher mortality rates**

261 Drought-related variables were more important for beech mortality than Scots pine (Fig. 2  
262 and 3), probably reflecting functional differences in species responses to drought (Choat *et al.*,  
263 2018). Scots pine is a drought-avoiding species (e.g. a species which rapidly closes its stomata to  
264 maintain high water status; McDowell *et al.*, 2008), that can survive from wet to dry environments  
265 (San-Miguel-Ayanz *et al.*, 2016), whereas beech is a drought-sensitive species (van der Maaten,  
266 2012; Chen *et al.*, 2015) with an anisohydric response to drought (e.g. a species that keeps its  
267 stomata open until late during droughts to maintain carbon uptake; McDowell *et al.*, 2008).  
268 Nevertheless, both beech (i.e. an angiosperm and broad-leaved species) and Scots pine (i.e. a  
269 gymnosperm and evergreen species) exhibited higher mortality rates in areas that were subject to  
270 increasing droughts during the study period (negative *SPEI*; Fig. 2 and 3). This result suggests that

271 major phylogenetic and functional groups could display a similar mortality response to increasing  
272 drought (Greenwood *et al.*, 2017) and is consistent with the results of a multi-species study  
273 suggesting that climatic extremes (like extreme droughts) are affecting tree mortality in Europe  
274 (Neumann *et al.*, 2017).

275         The increase in drought intensity that occurred at about 45° latitude during the study period  
276 (see the lowest *SPEI* values in Fig. S2.1 and S2.3b) could be responsible for the higher tree  
277 mortality rates in the Mediterranean biome (Fig. 4), which is also supported by the high relative  
278 importance of the increase in drought intensity at this latitude (see the highest values of *SPEI* in Fig.  
279 2 and 3). Moreover, we observed higher mortality rates in the driest areas (i.e. low *WAI*), as already  
280 reported for Scots pine in some inner Alpine valleys (Rigling *et al.*, 2013) and in the Iberian  
281 Peninsula (Vilà-Cabrera *et al.*, 2011; Galiano *et al.*, 2011). Nevertheless, the stronger effect of  
282 increasing droughts over the study period (i.e. *SPEI*) than that of drought intensity (i.e. *WAI*) on  
283 Scots pine mortality could mean that mortality events tend to occur when drought conditions exceed  
284 the average in a given area, suggesting a certain degree of Scots pine adaptation to local conditions  
285 (Savolainen *et al.*, 2007).

286         Drought-related variables were key drivers of beech mortality and were comparatively  
287 more important than heterospecific and conspecific basal area. A regional study of tree mortality  
288 suggested that competition between trees is more important than climate (Ruiz-Benito *et al.*, 2013),  
289 but that study did not cover a climatic gradient as large as our study. Drought-induced mortality  
290 could also be studied under the assumption that mortality events follow a period of reduced growth  
291 (Caillleret *et al.*, 2016). This assumption needs to be used with caution for beech, which can survive  
292 long periods of reduced growth before death (Hülsmann *et al.*, 2018). In addition, beech growth-  
293 based studies produced contradictory results, showing both drought-induced reduction in growth  
294 (Jump *et al.*, 2006) and drought-associated increase in growth (Tegel *et al.*, 2014). Contrarily, Scots  
295 pine growth variations (i.e. and those of gymnosperms in general; Caillleret *et al.*, 2016) can be used

296 to predict upcoming drought-induced mortality events because, although more drought-resistant,  
297 Scots pine does not support long periods of reduced growth (Hülsmann *et al.*, 2018).

## 298 **4.2 Conspecific and heterospecific neighbours can affect individual tree** 299 **mortality differently**

300 Competition is a critical driver of forest structure (Kunstler *et al.*, 2016), which strongly  
301 influences tree mortality and is comparatively more important for shade-intolerant than shade-  
302 tolerant species (Ruiz-Benito *et al.*, 2013). High mortality rates were associated with high  
303 conspecific basal area in both species and high heterospecific basal area in Scots pine. However,  
304 high heterospecific basal area was correlated with low mortality rates in beech (Fig. 2 and 3). Scots  
305 pine is a shade-intolerant tree which is highly sensitive to competition for light (Ruiz-Benito *et al.*,  
306 2013), which might explain why both intra and inter-specific competition strongly and positively  
307 influenced its mortality rate (Condés & del Río, 2015). In contrast, beech is a late successional and  
308 shade-tolerant species (Hülsmann *et al.*, 2018) that outcompetes other species in fertile sites  
309 (Condés & del Río, 2015). This is consistent with our observation of high mortality rates with high  
310 conspecific basal area but also with low heterospecific basal area: beech mainly suffers from the  
311 presence of conspecific neighbours, but not from heterospecific neighbours, which are necessarily  
312 less competitive species. This result is supported by growth studies showing that beech benefits  
313 from admixture with other species but is highly sensitive to intra-specific competition (Pretzsch *et*  
314 *al.*, 2013a; Ratcliffe *et al.*, 2015).

315 The heterospecific basal area affected the mortality rates of both species less than the  
316 conspecific basal area (Table S3.1, Fig. 2 and 3). The dominant nature of both Scots pine and beech  
317 in European forests may partly explain this difference as the basal area of heterospecifics was much  
318 lower than that of conspecifics all along the latitudinal gradient (Fig. S2.1). Nevertheless, the  
319 overdominance of intra-specific competition, a key process for stabilising ecosystems, is a globally-  
320 observed pattern (Kunstler *et al.*, 2016), which could be linked to how interspecific differences

321 determine complementarity mechanisms and, consequently, individual resource-use and coexistence  
322 mechanisms (Ruiz-Benito *et al.*, 2017b).

### 323 **4.3 The effects of climatic drought and basal area should be considered jointly in** 324 **mortality studies**

325 Competition with neighbours can be expressed as asymmetric competition for light on  
326 small suppressed trees (Ruiz-Benito *et al.*, 2013) but also as symmetric competition for limited  
327 resources, like water or nutrients (Franklin *et al.*, 1987; Gessler *et al.*, 2017). Drought-induced  
328 mortality may be strong in areas with high levels of competition, because plants are more stressed  
329 and small changes in water availability could result in massive mortality events (Bradford and Bell,  
330 2017; Ruiz-Benito *et al.*, 2013; Young *et al.*, 2017). In the case of Scots pine, the strong interaction  
331 between drought intensity and conspecific basal area reinforces this assumption (Table S3.1).  
332 Indeed, mortality rates were high in dry areas with high conspecific basal area whereas in areas with  
333 lower conspecific basal area, trees had still sufficient resources to survive despite reduced water  
334 availability (Fig. 2b). This result suggests that Scots pine suffers from the presence of neighbouring  
335 trees only when resources are scarce (Young *et al.*, 2017).

336 In the case of beech, the influence of conspecific basal area on mortality was not  
337 modulated by drought (Table S3.1), suggesting that resource depletion does not exacerbate  
338 competitive pressure among beech trees. However, the probability of beech mortality in the driest  
339 areas was considerably higher when heterospecific basal area, the most important predictor in the  
340 Mediterranean biome (Fig. 3a & Table 2), was low (Fig. 3b). These findings suggest that beech  
341 survival in the driest part of its range is positively influenced by its neighbours (facilitation), which  
342 are mainly *Q. pyrenaica*, *P. sylvestris* and *C. sativa* (Table S5.1). Our results can only be compared  
343 to those of growth studies because the effect of mixing species has been more investigated in  
344 growth than mortality studies. Beech trees were shown to be more resilient and resistant to drought  
345 in mixed stands with oaks (Pretzsch *et al.*, 2013b). By contrast, Bosela *et al.* (2018) found that the

346 growth of beeches mixed with fir trees or in pure stands was equally negatively affected by long-  
347 term droughts but they didn't explore the south-western part of beech distribution. Overall, these  
348 results suggest that beech growth and mortality are influenced by interspecific interactions that vary  
349 along the European drought gradient: from neutral interactions in wet areas where beech co-occurs  
350 mainly with *Abies alba* and *Picea abies*, to facilitation in dry areas where beech co-occurs with  
351 more Mediterranean species (Fig 3b & Table S5.1). However, the mechanisms behind these  
352 interspecific interactions, particularly in dry areas, are still largely unknown. Identifying  
353 associations of species that can survive droughts could help to better understand drought-related  
354 mortality patterns in the coming years.

355 In the case of Scots pine, previous regional studies reported contradictory interaction  
356 effects between competition and drought: higher rate of decline in dry areas but only at low  
357 competition levels (Vilà-Cabrera *et al.*, 2013), low mortality rates related to high heterospecific  
358 basal area in wet areas (Condés & del Río, 2015) and only additive effects of competition and  
359 drought on mortality with no interaction effects (Galiano *et al.*, 2010). Our study is the first to  
360 describe interaction patterns between drought and basal area at the scale of the distribution of each  
361 species (Fig. 2, 3 and S4). As we found four significant interactions (albeit three of which only  
362 slightly affect mortality) influencing Scots pine mortality and only one in the case of beech (Table  
363 S3.1), we can assume that Scots pine mortality is affected directly and indirectly by drought through  
364 interactions with basal area while beech mortality was more directly affected by drought.

#### 365 **4.2 Tree mortality patterns along latitude and potential associated range shifts**

366 Predicted probability of mortality in both beech and Scots pine was higher in the southern  
367 part of their distribution, mainly corresponding to the French part of the Mediterranean biome and  
368 the Pyrenees in the case of beech (Fig. 4). In these areas, beech and Scots pine mortality rates were  
369 accurately predicted (Fig. 1), except at the southern end of Scots pine range where mortality rates  
370 are likely to be slightly higher than predicted (Fig. 1a). These accurate predictions in the southern



371 part of species ranges were expected as we chose climatic variables related to droughts, generally  
372 more important in the Mediterranean biome (Fig. S2.3). Surprisingly, the association of drought and  
373 competition-related variables alone explained Scots pine mortality patterns in the northern part of  
374 its distribution (Fig. 1a) but overestimated the probability of beech mortality in northern Germany  
375 (Fig. 1b) suggesting that other factors come into play in these areas to explain beech mortality  
376 patterns (see 4.3 Limitations).

377         An unexpected result was that French Mediterranean Scots pines and beech trees suffered  
378 even more from climatic drought than those in Spain, where several studies reported high mortality  
379 or defoliation rates in the Iberian Peninsula in both species (Carnicer *et al.*, 2011; Vilà-Cabrera *et*  
380 *al.*, 2011, 2013; Benito Garzón *et al.*, 2013). Nevertheless, this pattern may be explained by the high  
381 altitudes at which both species occur in Spain, and the calcareous soils of southeastern France,  
382 which do not retain water and are consequently very dry. In the case of Scots pine, we also  
383 hypothesise that local adaptation to temperature explains our underestimated mortality predictions  
384 in the southernmost part of the gradient (Savolainen *et al.*, 2007): populations in these areas may be  
385 highly locally-adapted to drought conditions and therefore less resistant to changing climate (Benito  
386 Garzón *et al.*, 2011).

387         The high mortality rates predicted in the French part of the Mediterranean biome could be  
388 explained by the increase in drought intensity during the study period in that region (Fig. S2.3b),  
389 suggesting that mortality plays a critical role in delimiting the driest part of the species ranges  
390 (Gaston, 2009; Benito Garzón *et al.*, 2013; Ruiz-Benito *et al.*, 2017a), in particular in the  
391 Mediterranean biome, which is expected to face drier conditions in the coming decades. In addition  
392 to direct effects of climate change, Scots pine and beech are exposed to more intense fires in the  
393 driest parts of their range (Fréjaville *et al.*, 2018) and these should increase the likelihood of range  
394 contraction at the ecotone between Mediterranean and cool temperate biomes.

### 395 **4.3 Limitations**

396           Until recently, European forests have been extensively exploited and forest management is  
397 still widespread, particularly in the Scandinavian countries (Schelhaas *et al.*, 2018). Although we  
398 removed the direct effects of management in our study (i.e. by removing plots in which trees were  
399 noted as harvested), management may still result in both an overestimation (e.g. by reducing  
400 competition pressure in thinned plots) and an underestimation of natural mortality rates through  
401 salvage loggings (i.e. the harvest of dead trees after a natural disaster) or sanitation fellings (i.e. the  
402 harvest of diseased trees).

403           Other factors also affect tree mortality, either directly, indirectly or through interactions,  
404 such as: changes in disturbance regimes (Seidl *et al.*, 2017), insect outbreaks (Anderegg *et al.*,  
405 2015), mistletoe (Dobbertin & Rigling, 2006), atmospheric pollutants (Dietze & Moorcroft, 2011),  
406 populations genetic differentiation and plasticity (Benito Garzón *et al.*, 2011), soil characteristics  
407 (Dietze & Moorcroft, 2011). However, given our concern to limit the model complexity and the  
408 lack of large-scale data, we decided not to include them in our study and to focus on comparing the  
409 effects of drought and competition on mortality.

410

### 411 **CONCLUSIONS**

412           Mortality of Scots pine and beech was affected by climatic drought intensity and indirect  
413 competition from neighbouring trees, but in different ways. Drought directly affected beech  
414 mortality rates and beech trees benefited from mixing with other species, particularly in the  
415 Mediterranean biome. Scots pine mortality suffered mostly from competition and was indirectly  
416 affected by drought through interactions with competitors, especially in southeastern France. In this  
417 area, which experienced a marked increase in drought intensity during the study period, high  
418 mortality rates were predicted for both species, as expected for temperate trees for which the  
419 Mediterranean biome corresponds to the southernmost part of the distribution. In a warming

420 climate, our study is a step further in understanding geographical patterns of tree mortality in  
421 Europe and shed light on the high mortality risks faced by European tree species, regardless of their  
422 different life-history strategies, especially at the ecotone between the Mediterranean and cool  
423 temperate biomes. In this priority area, beech could benefit from mixing with other species and pine  
424 from reduced competition.

425

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435

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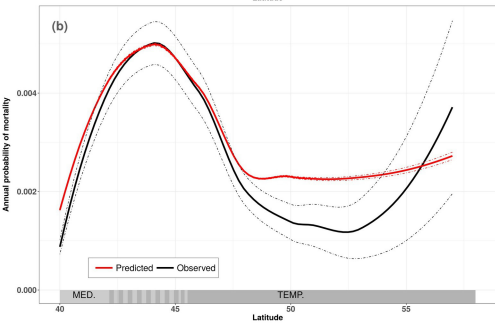
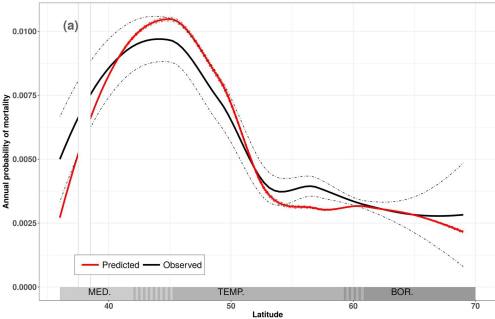
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#### 444 **DATA ACCESSIBILITY**

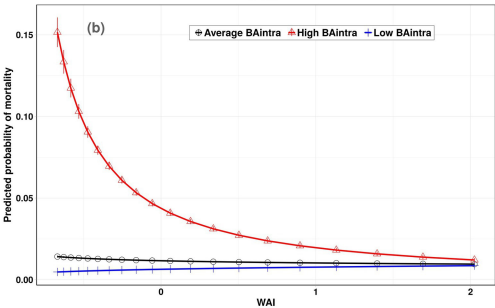
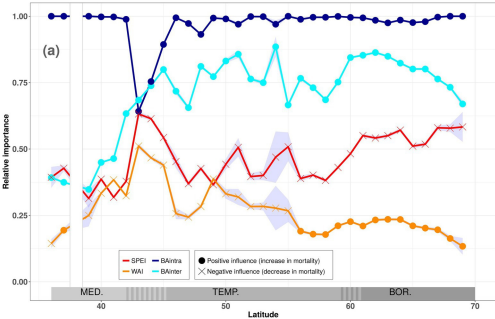
445           The data are available upon request to the co-authors.



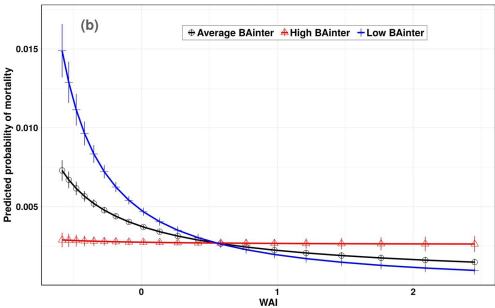
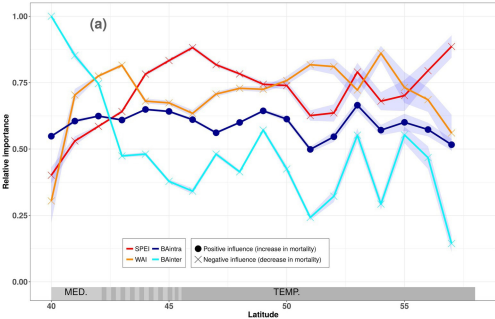
**Figure 1. Predicted and observed annual probability of mortality along the latitudinal gradient covered by the NFIs plots a) for *P. sylvestris* and b) for *F. sylvatica*.** Predicted and observed values were estimated at the individual-level and were clustered at 1° latitude resolution. A locally weighted regression was used to obtain the smooth solid lines (“loess” method of the `geom_smooth` function in “ggplot2” R package). Dotted lines indicate 95% confidence intervals. The acronyms MED., TEMP. and BOR. in grey bars refer to the Mediterranean, cool temperate and boreal biome, respectively. The white section for *P. sylvestris* in the Mediterranean biome represents missing data (due to its distribution in Spain).



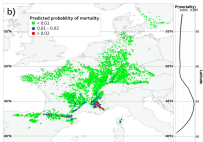
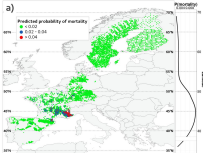
**Figure 2. Effects of drought-related variables and basal area on Scots pine mortality. a) Relative importance of the changes in climatic drought intensity over the study period (i.e. *SPEI*), climatic drought intensity (i.e. *WAI*), conspecific basal area (i.e. *BA<sub>intra</sub>*) and heterospecific basal area (i.e. *BA<sub>inter</sub>*) on Scots pine predicted probability of mortality.** The relative importance of each variable was computed for each tree from the logistic regression model (see section 2.5), by giving a value of one to the most influencing variable and scaling the remaining variables accordingly. For each variable, the relative importance values were aggregated by 1° latitude resolution and the points of the graph correspond to the average values. The grey areas around the curves correspond to the 95% confidence intervals. The acronyms MED., TEMP. and BOR. in grey bars refer to the Mediterranean, cool temperate and boreal biome, respectively. The white section corresponds to missing data at that latitude (due to Scots pine distribution in Spain). **b) Interactions between conspecific basal area (i.e. *BA<sub>intra</sub>*) and climatic drought intensity (i.e. *WAI*) on Scots pine probability of mortality.** This interaction was considered significant if its z value was lower than -2 or higher than 2 and was the most important interaction influencing Scots pine mortality (Table S3.1). Scots pine mortality was predicted at three different levels of conspecific basal area (mean value, 99.5<sup>th</sup> percentile and 0.005<sup>th</sup> percentile; proxies of average, high and low competition, respectively) along a drought gradient while the other predictors were fixed at their mean value.



**Figure 3. Effects of drought-related variables and basal area on beech mortality. a) Relative importance of the changes in climatic drought intensity over the study period (i.e. *SPEI*), climatic drought intensity (i.e. *WAI*), conspecific basal area (i.e. *BA<sub>intra</sub>*) and heterospecific basal area (i.e. *BA<sub>inter</sub>*) on beech predicted probability of mortality.** The relative importance of each variable was computed for each tree from the logistic regression model (see section 2.5), by giving a value of one to the most influencing variable and scaling the remaining variables accordingly. For each variable, the relative importance values were aggregated by 1° latitude resolution and the points of the graph correspond to the average values. The grey areas around the curve correspond to the 95% confidence intervals. The acronyms MED. and TEMP. in grey bars refer to the Mediterranean and cool temperate biome, respectively. **b) Interaction between heterospecific basal area (i.e. *BA<sub>inter</sub>*) and climatic drought intensity (i.e. *WAI*) on beech probability of mortality.** This interaction was considered significant as its z value was higher than 2 (see Table S3.1). Beech mortality was predicted at three different levels of heterospecific basal area (mean value, 99.5<sup>th</sup> percentile and 0.005<sup>th</sup> percentile; proxies of average, high and low competition, respectively) along a drought gradient while the other predictors were fixed at their mean value.



**Figure 4. Spatial projection of the annual predicted probability of mortality at the individual-level across Europe for a) *P. sylvestris* and b) *F. sylvatica*.** Graphs in the right panels display predictions (noted as  $P(\text{mortality})$ ) across latitude. For both species, predictions were calculated for all trees from the logistic regression model and were clustered at 1° latitude resolution. A locally weighted regression was used to obtain the smooth solid lines (“loess” method of the `geom_smooth` function in “ggplot2” R package). Grey areas indicate 95% confidence intervals (almost confused with the curves). The white section for *P. sylvestris* in the Mediterranean biome represents missing data at that latitude (due to its distribution in Spain).





		<i>P. sylvestris</i>				<i>F. sylvatica</i>		
		All biomes	Mediterranean biome	Cool temperate biome	Boreal biome	All biomes	Mediterranean biome	Cool temperate biome
Relative importance	<i>BAintra</i>	0.96 (0.9616; 0.9630)	0.95 (0.9455; 0.9482)	0.96 (0.9625; 0.9644)	0.99 (0.9860; 0.9870)	0.61 (0.6134; 0.6162)	0.64 (0.6411; 0.6465)	0.61 (0.6076; 0.6107)
	<i>BAinter</i>	0.67 (0.6755; 0.6737)	0.52 (0.5235; 0.5258)	0.73 (0.7308; 0.7331)	0.83 (0.8293; 0.8314)	0.54 (0.5377; 0.5426)	0.80 (0.7915; 0.8010)	0.49 (0.4878; 0.4928)
	<i>WAI</i>	0.31 (0.3089; 0.3114)	0.35 (0.3509; 0.3552)	0.32 (0.3191; 0.3233)	0.22 (0.2169; 0.2201)	0.74 (0.7390; 0.7442)	0.71 (0.6997; 0.7130)	0.75 (0.7456; 0.7513)
	<i>SPEI</i>	0.44 (0.4370; 0.4397)	0.40 (0.3976; 0.4017)	0.42 (0.4218; 0.4258)	0.53 (0.5265; 0.5311)	0.70 (0.7018; 0.7066)	0.61 (0.6055; 0.6168)	0.72 (0.7196; 0.7249)
Annual predicted mortality		0.0061 (0.00611; 0.00618)	0.0077 (0.00763; 0.00775)	0.0063 (0.00619; 0.00631)	0.0033 (0.00332; 0.00337)	0.0038 (0.00374; 0.00382)	0.0052 (0.00506; 0.00530)	0.0035 (0.00347; 0.00354)

**Table 1. Mean relative importance of each predictor and mean annual predicted probability of both species mortality per biome.** See Fig. S1.1 for biome boundaries. The relative importance of each variable was first computed for each tree from the logistic regression model (see section 2.5), by giving a value of one to the most influencing variable and scaling the remaining variables accordingly. Secondly, the relative importance and the annual predicted probabilities of mortality ( $P_i$  in the equation 1) were average for each biome. Numbers in brackets refer to 95% confidence intervals. *BAintra*: conspecific basal area ( $\text{m}^2 \text{ha}^{-1}$ ), *BAinter*: heterospecific basal area ( $\text{m}^2 \text{ha}^{-1}$ ), *WAI*: water availability index (adimensional), *SPEI*: Standardised Precipitation-Evapotranspiration Index (adimensional).

*Pinus sylvestris* L.

*Fagus sylvatica* L.

