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**Chilling and forcing temperatures interact to predict the onset of wood formation in Northern Hemisphere conifers.**

**Running head:** Temperature dependence of spring xylem onset

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56

## 57 **Abstract**

58 The phenology of wood formation is a critical process to consider for predicting how trees  
59 from the temperate and boreal zones may react to climate change. Compared to leaf  
60 phenology, however, the determinism of wood phenology is still poorly known. Here, we  
61 compared for the first time three alternative ecophysiological model classes (*threshold*  
62 *models*, *heat-sum* models and *chilling-influenced heat-sum* models) and an empirical model in  
63 their ability to predict the starting date of xylem cell enlargement in spring, for four major  
64 Northern Hemisphere conifers (*Larix decidua*, *Pinus sylvestris*, *Picea abies* and *Picea*  
65 *mariana*). We fitted models with Bayesian inference to wood phenological data collected for  
66 220 site-years over Europe and Canada. The chilling-influenced heat-sum model received  
67 most support for all the four studied species, predicting validation data with a 7.7-day error,  
68 which is within one-day of the observed data resolution. We conclude that both chilling and  
69 forcing temperatures determine the onset of wood formation in Northern Hemisphere  
70 conifers. Importantly, the chilling-influenced heat-sum model showed virtually no spatial bias  
71 whichever the species, despite the large environmental gradients considered. This suggests  
72 that the spring onset of wood formation is far less affected by local adaptation than by  
73 environmentally-driven plasticity. In a context of climate change, we therefore expect rising  
74 winter-spring temperature to exert ambivalent effects on the spring onset of wood formation,  
75 tending to hasten it through the accumulation of forcing temperature, but imposing a higher  
76 forcing-temperature requirement through the lower accumulation of chilling.

77

78 **Keywords:** wood phenology, cambium, phenological models, chilling temperatures, forcing  
79 temperatures, conifers.

## 80 **Introduction**

81 The seasonality of physiological processes is an essential component of terrestrial ecosystem  
82 models (TEMs; Delpierre et al., 2012; Kramer, 1995), but is usually poorly represented being  
83 mostly confined to the simulation of leaf onset and leaf loss (Delpierre, Vitasse, et al., 2016).  
84 In such models, the phenology of non-leaf organs or tissues (e.g. wood) is simulated (i)

85 simultaneous or relative to leaf phenology or (ii) using generic, non-organ-specific  
86 temperature functions for modulating the allocation of carbon (Delpierre, Vitasse, et al., 2016  
87 ; but see Schiestl-Aalto, Kulmala, Mäkinen, Nikinmaa, & Mäkelä, 2015). This reflects the  
88 state of our knowledge on the phenology of trees, which is far more developed for leaves as  
89 compared with other organs or tissues (Delpierre, Vitasse, et al., 2016; Ford, Harrington,  
90 Bansal, Gould, & St. Clair, 2016). It is difficult to quantify how strongly this knowledge gap  
91 affects the predictive ability of TEMs, but it certainly jeopardizes their biological realism  
92 (Guillemot et al., 2017). For example, it has been demonstrated in evergreen conifers that the  
93 spring resumption of cambium activity generally occurs before budburst (Cuny, Rathgeber,  
94 Lebourgeois, Fortin, & Fournier, 2012; Gruber, Strobl, Veit, & Oberhuber, 2010; Huang,  
95 Deslauriers, & Rossi, 2014; Michelot, Simard, Rathgeber, Dufrêne, & Damesin, 2012; Rossi  
96 et al., 2009). Moreover, several studies have shown that, independent from leaf phenology,  
97 the duration of the wood growing season *per se* is a major determinant of wood production  
98 (Delpierre, Berveiller, Granda, & Dufrêne, 2016; Lempereur et al., 2015), so that an earlier  
99 onset of cambium activity, or a later cessation may result in a higher cell production (Lupi,  
100 Morin, Deslauriers, & Rossi, 2010; Mäkinen, Jyske, & Nöjd, 2018). Consequently, there is a  
101 clear need for the development of wood phenology modules for inclusion into TEMs.

102 In order to develop wood phenology modules for TEMs, we first have to understand the  
103 causal climatic drivers of wood phenology. In the temperate and boreal regions of the  
104 Northern Hemisphere, the formation of wood is seasonal and occurs from late spring to early  
105 autumn (Rossi et al., 2016, 2008). In spring, cambial mother cells start dividing, producing  
106 new derivatives of phloem outward and xylem inward (Larson, 1994; Vaganov, Hughes, &  
107 Shashkin, 2006). As a base model for this cycle, several authors have proposed that, just as  
108 for buds, the spring resumption of cambium activity is the outcome of a two-phase dormancy  
109 period (Begum et al., 2018; Begum, Nakaba, Yamagishi, Oribe, & Funada, 2013; Ford et al.,  
110 2016; Little & Bonga, 1974; Rensing & Samuels, 2004). According to this model, cambium  
111 activity is prevented by tree's *internal* factors (e.g. physiological state, signals) during the  
112 *endo*-dormancy phase; while it resumes during the *eco*-dormancy phase when the *external*  
113 conditions are favourable.

114 The main candidate for external conditions driving the resumption of cambium activity in  
115 temperate and boreal ecosystems is the spring temperature (as reviewed in Begum et al.,  
116 2018; Delpierre, Vitasse, et al., 2016; Larson, 1994). Field observation have shown that spring  
117 cambium resumption is usually delayed at high altitudes and latitudes as compared to low

118 altitudes and latitudes (Jyske, Mäkinen, Kalliokoski, & Nöjd, 2014; Moser et al., 2010; Rossi  
119 et al., 2016; Rossi, Deslauriers, Anfodillo, & Carraro, 2007; Rossi et al., 2008). Furthermore,  
120 local stem heating activated the formation of wood (Gričar et al., 2007), with a gradually  
121 increased response to heat applied from winter to spring (Oribe & Kubo, 1997).

122 Based on these evidences, previous studies have developed different model formulations  
123 based on spring temperature to predict the timing of cambial resumption. A first model class  
124 uses a *temperature threshold* for predicting the onset of cambial activity in conifers from cold  
125 biomes (Deslauriers, Rossi, Anfodillo, & Saracino, 2008; Rossi et al., 2007, 2008). However,  
126 although this model is able to identify likely periods of cambial activity, its accuracy for  
127 predicting the onset of cambial activity from temperature time series is probably low (Fig.  
128 S1). Another model class is that of *heat sums* (Giagli, Gricar, Vavrcik, & Gryc, 2016;  
129 Schmitt, Jalkanen, & Eckstein, 2004; Seo, Eckstein, Jalkanen, Rickebusch, & Schmitt, 2008;  
130 Swidrak, Gruber, Kofler, & Oberhuber, 2011). Their underlying hypothesis is that the  
131 cambium resumes its activity (cell division followed by cell differentiation) after sufficient  
132 exposure to temperatures above a threshold (so-called *forcing* temperatures). Thus, *heat sum*  
133 models mimic the progress of cambium through the eco-dormancy phase, making the implicit  
134 hypothesis that the endo- and eco-dormancy phases are sequential, and that endo-dormancy  
135 stops at the date when heat accumulation starts (Delpierre, Vitasse, et al., 2016). In practice, a  
136 degree-days accumulation is calculated by summing temperatures above a threshold ('base  
137 temperature') of typically +5°C (or more rarely lower values e.g. 0-1°C, see Antonucci et al.,  
138 2015; Li et al., 2017) from a given day, fixed a priori, before the onset date of cambial  
139 reactivation. However, there is no consensus concerning the day or period of year from which  
140 the cambium becomes sensitive to *forcing* temperatures. Some studies choose January 1 or  
141 spring equinox (Giagli et al., 2016; Schmitt et al., 2004), whereas others (Seo et al., 2008)  
142 consider the starting date occurring when trees have experienced a daily mean temperature  
143 above +5°C for at least five consecutive days. Moreover, *heat sum* models usually fail in  
144 identifying a species-specific heat sum threshold above which cambium would systematically  
145 be active (Giagli et al., 2016; Moser et al., 2010), which is indicative of their low structural  
146 realism and thus low predictive ability. More recently, *chilling-influenced heat sum* models  
147 have been shown able to predict spring cambial reactivation in Douglas fir (Ford et al., 2016).  
148 Similar to *heat sum* models, those models were originally designed for describing the progress  
149 of primary meristems (i.e. leaf or flower buds) from dormancy to budburst. Their basic  
150 hypothesis is that the cambium requires a lower accumulation of *forcing* temperatures during

151 the eco-dormancy phase when exposed to increasing levels of cold temperatures (so-called  
152 *chilling* temperatures, (Cannell & Smith, 1983; Little & Bonga, 1974) during the endo-  
153 dormancy phase, which may precede or be concomitant to the eco-dormancy phase (Chuine,  
154 Garcia de Cortazar-Atauri, Kramer, & Hänninen, 2013). The underlying physiological basis  
155 of such *chilling-influenced heat sum* models is not fully understood (Rinne et al., 2001; Singh,  
156 Svystun, AlDahmash, Jönsson, & Bhalerao, 2017). Last, a recent study made use of empirical  
157 models (linear regression of spring-averaged temperature) to predict the timing of cambial  
158 resumption (Rossi et al., 2016).

159 Though previous studies evaluated the ability of the three abovementioned model classes  
160 separately in simulating the date of the resumption of cambium activity in spring (threshold-  
161 type, Rossi, Morin, Deslauriers, & Plourde, 2011; *heat sums*, Seo et al., 2008; Swidrak et al.,  
162 2011; *chilling-influenced heat sums*, Ford et al., 2016; empirical regression, Rossi et al.,  
163 2016), there has been no comparison of those models merits on the same dataset. Here, we  
164 make use of a large number of field observation data collected over Europe and Canada  
165 (GLOBOXYLO database) to conduct for the first time a systematic evaluation of the causal  
166 factors affecting the breaking of cambial dormancy, and to propose an improved model of  
167 cambial spring resumption. Specifically, by identifying which model structure receives most  
168 support from observed data, we aim to evaluate: (1) if the resumption of cambium activity of  
169 Northern Hemisphere conifers in spring is more likely caused by the crossing of a given  
170 temperature threshold or by an accumulation of heat (“do *threshold* models outperform *heat*  
171 *sum* models?”) and; (2) if observation data support the existence of a separate endo-dormancy  
172 phase that can be broken by chilling exposure (“do *chilling-influenced heat sum* models fit the  
173 data best?”). Our hypotheses are (1) *threshold* models are fine for identifying a thermal  
174 probability of cambium activity but have low predictive ability since the daily variability of  
175 temperature superimposed to seasonal variations cannot serve as a reliable cue for trees; (2)  
176 that over large geographical gradients, models incorporating both the effects of chilling and  
177 forcing temperature are better able to describe the variability in the beginning of wood  
178 formation (since over large climate zones, multiple climate limitations interact). Having  
179 identified the model structure best supported by the data, we then evaluate the biological  
180 reliability of its inferred parameters, for future use in Terrestrial Ecosystem Models.

## 181 **Material and methods**

### 182 **Study sites**

183 The selected study sites were extracted from the GLOBOXYLO database<sup>a</sup>, a dataset gathering  
184 wood formation and meteorological information collected over the past 15 years from several  
185 research teams all over the world. The selected data concern the four most observed  
186 coniferous species (*Larix decidua* Mill. (LADE), *Pinus sylvestris* L. (PISY), *Picea abies* L.  
187 Karst. (PCAB) and *Picea mariana* (Mill.) BSP (PCMA)), covering a wide range of  
188 temperature and photoperiod conditions in the Northern Hemisphere (from 40.0°N to 67.5°N  
189 latitude, 79.2°W to 29.4°E longitude, and from 30 m to 2150 m altitudes) (Fig. 1, Table S1).  
190 Specifically, the dataset includes wood formation critical dates from 2001 to 2013 over 46  
191 study sites for a total of 220 site-years, representing 1105 tree-site-year observations. All  
192 sampled trees were dominant individuals. The average ( $\pm$ SD) tree age was  $124 \pm 70$  years,  
193 with a diameter at breast height (DBH) of  $44 \pm 30$  cm, and a tree height of  $21 \pm 8$  m (Table  
194 S1).

## 195 **Wood formation data**

### 196 *Microcore sampling and preparation*

197 At each study site, on average  $5 \pm 2$  trees were chosen and sampled weekly from March-April,  
198 depending on local climate conditions, to monitor wood formation. The collection,  
199 preparation, and analysis of wood samples followed a common protocol across sites. Wood  
200 microcores of 2 mm in diameter and 15-20 mm in length were collected weekly at breast  
201 height ( $1.3 \pm 0.3$  m) over the growing season, using a Trephor® tool (Rossi, Anfodillo, &  
202 Menardi, 2006) or surgical bone sampling needles (Deslauriers, Morin, & Begin, 2003).  
203 Microcores were then cut with rotary or sledge microtomes in transverse sections of 10-30  $\mu$ m  
204 thick, stained with safranin and astra blue or cresyl violet acetate and observed under bright-  
205 field and polarized light after coloration (Rossi, Deslauriers, & Anfodillo, 2006).  
206

### 207 *Determination of the spring resumption of xylem formation*

208 We focus on the beginning of xylem cell enlargement (bE) as a critical, well-defined marker  
209 corresponding to the spring start-up of wood formation. Ultrastructural changes in cambial  
210 cells are the very first stage of growth reactivation. The bE occurs somewhat later than the  
211 onset of ultrastructural changes in cambial cells; but the latter is very difficult to observe

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<sup>a</sup> <https://www6.nancy.inra.fr/foret-bois-lerfob/Projets/Projets-en-cours/GLOBOXYLO>



212 accurately and involves both xylem and phloem cells (Prislan, Čufar, Koch, Schmitt, &  
213 Gričar, 2013; Prislan, Schmitt, Koch, Gričar, & Čufar, 2011). It is therefore not often reported  
214 in wood formation monitoring studies. To quantify bE, the number of cells in each  
215 differentiation zone (cambial, enlargement, thickening, and mature) was counted along at  
216 least three radial files on the anatomical sections. Enlarging tracheids were characterized by  
217 radial diameter at least twice that of a cambial cell. We defined, at the tree level, the  
218 beginning of the enlargement phase (bE) as the date (day of year, DoY) when more than 50%  
219 of the observed radial files present at least one first enlarging tracheid (Rathgeber,  
220 Longuetaud, Mothe, Cuny, & Le Moguédec, 2011).

### 221 **Temperature and photoperiod data**

222 Mean daily temperatures have been collected at the study sites (Fig. 1). However, local  
223 weather stations were usually not installed before the start of the wood formation monitoring.  
224 To be able to consider in our models weather conditions also before the monitoring period, we  
225 used, for European sites, the WATCH gridded meteorological dataset (grid-resolution = 0.5°,  
226 Weedon et al., 2014) to extrapolate those missing data, after establishing linear regression  
227 between the local and corresponding WATCH temperature data (correlation between  
228 overlapping local and WATCH temperature time series was  $0.95 < r < 0.99$ ), and removing  
229 the (low) biases of WATCH data. For Canadian sites, i.e. for *Picea mariana*, we did not  
230 extrapolate the temperature time series. Day length (the daily duration of the photoperiod)  
231 was calculated daily as a function of latitude, using astronomical formulae<sup>b</sup>.

### 232 **Models description**

233 We compared three classes of ecophysiological models and one empirical model (Table 1) in  
234 their ability to predict the date of onset of xylem cell enlargement phase (bE) in the four tree  
235 species of interest. The three model classes are: (i) *threshold* models, (ii) *heat sum* models,  
236 (iii) *chilling-influenced heat sum* models. Since the patterns of xylem formation have been  
237 strongly related to mean temperatures over large geographical gradients (Rossi et al., 2016),  
238 we used an empirical model relating bE to early season (January-June) average temperature as  
239 a benchmark for ecophysiological models.

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<sup>b</sup> See for example Pr Dennis Baldocchi's biometeorology course, lecture number 7  
(<https://nature.berkeley.edu/biometlab/index.php?scrm=espm129>)

240 For all ecophysiological models, we used photoperiod thresholds to delineate the start and end  
241 of the endo- and eco-dormancy periods, different to most earlier phenological modelling  
242 studies, which usually considered temperature accumulation to start at a given day of year  
243 (e.g. usually January 1 in most phenological studies considering *heat sum* models; Linkosalo,  
244 Carter, Hakkinen, & Hari, 2000; Seo et al., 2008). This choice was motivated by the fact that  
245 our study covers a large latitudinal gradient over which a given calendar day (not perceptible  
246 by trees *per se*) may correspond to a large variations in photoperiod (a signal which is  
247 perceptible by trees).

### 248 ***Temperature- and photoperiod-threshold models***

249 In this class of models, we assumed that bE occurs when a given temperature and/or  
250 photoperiod threshold has/have been crossed. A first formulation of this model (henceforth  
251 referred to as *Tt* model) is:

$$252 \quad bE = \min(d) \text{ such that } T(d) \geq T^* \text{ and } d > -10 \quad (1)$$

253 where *bE* is the beginning of the xylem enlargement period (DoY), *d* is a day of year (DoY),  
254 *T* is the daily average temperature, and *T\** is a temperature threshold (°C). We assume that the  
255 passing of the temperature threshold necessarily occurs after winter solstice of the previous  
256 year (i.e. DoY 355 of the previous year, or DoY -10 of current year).

257 In case *bE* occurs when the thresholds of both temperature and photoperiod have been  
258 exceeded, the model (henceforth *TDLt* model) writes:

$$259 \quad \begin{cases} bE = \min(d) \text{ such that } T(d) \geq T^* \text{ and } d > j \\ \text{with } j = \min(d) \text{ such that } DL(d) \geq DL^* \text{ and } d > -10 \end{cases} \quad (2)$$

260 where *DL* is the daily photoperiod (hours) and *DL\** is a photoperiod-threshold (hours).

### 261 ***Heat sum model***

262 In the *heat sum* model, we assumed that bE occurs when a given accumulation of heat (above  
263 a temperature threshold, i.e. forcing temperatures) has been reached. The model (henceforth  
264 *HS* model) takes the form:

$$265 \quad bE = \min(d) \text{ such that } F(d) \geq F^* \quad (3)$$

266 with 
$$\begin{cases} F(d) = \sum_{F_{start}}^d T_{diff}(d) \\ with T_{diff}(d) = \begin{cases} T(d) - T_f, & \text{if } T \geq T_f \\ 0, & \text{if } T \leq T_f \end{cases} \end{cases} \quad (4)$$

267 where  $T_f$  is a temperature-threshold above which forcing temperatures are accumulated,  $F(d)$   
 268 is the heat sum at day  $d$  (degree-days) and  $F^*$  is the forcing units requirement at which bE  
 269 occurs (degree-days). In this model, the accumulation of forcing temperature starts at a given  
 270 photoperiod threshold  $DL_{Fstart}$  (hours), occurring after the winter solstice of the previous year  
 271 such that:

272 
$$F_{start} = \min(d) \text{ such that } DL(d) \geq DL_{Fstart} \quad (5)$$
  
 273 with  $d > -10$

274 This model simulates the progress of cambium through the eco-dormancy phase and makes  
 275 the implicit hypothesis that the preceding endo-dormancy phase ends on day  $F_{start}$ .

276 ***Chilling-influenced heat sum model***

277 In the chilling-influenced heat sum model (*CiHS* model), the progress of cambium through  
 278 the endo- and eco-dormancy phases is explicit, and bE occurs at the end of the eco-dormancy  
 279 phase. During endo-dormancy, cambium division is inhibited by tree internal factors, the  
 280 effects of which are counteracted by low temperatures. Following the approach proposed by  
 281 (Cannell & Smith, 1983) for bud meristems, this hypothesis translates into an accumulation of  
 282 chilling temperatures, quantified as a number of chilling units ( $C_{tot}$ , in chill units C.U.).  $C_{tot}$  is  
 283 calculated on a daily basis from  $C_{start}$  (DoY), up to the  $C_{end}$  date as follows:

284 
$$C_{tot}(d) = \sum_{C_{start}}^{C_{end}} R_c(T(d)) \quad (6)$$

285 where the daily rate of chilling ( $R_c$ ) can be calculated as a linear function of temperature:

286 
$$R_c(T(d)) = \begin{cases} 1 & \text{if } T(d) < T_c \\ 0 & \text{if } T(d) \geq T_c \end{cases} \quad (7)$$

287 where  $T_c$  is the temperature threshold ( $^{\circ}\text{C}$ ) below which chilling accumulation occurs.

288 Besides the accumulation of chilling, the model assumes that the progression of the cambium  
 289 towards bE during eco-dormancy is favoured by the accumulation of forcing temperatures  
 290  $F(d)$ , as described in eq. 3-4. The *CiHS* model postulates that, as the accumulation of chilling

291 proceeds, the requirement for forcing temperatures decreases, such that the critical sum of  
292 forcing  $F^*$  is defined daily, and linearly depends on  $C_{tot}$ :

$$293 \quad F^*(d) = g \times C_{tot}(d) + h \quad (8)$$

294 where  $g$  is the slope of the relation between required forcing units and chilling-accumulation  
295 (degree-days per C.U.), and  $h$  is the forcing units requirement in the absence of chilling  
296 (degree-days).

297 In this model, both the period of cambium sensitivity to chilling temperatures (delimited by  
298 days of year  $C_{start}$  and  $C_{end}$ , eq. 6) and the start of forcing temperature accumulation (on day  
299 of year  $F_{start}$ , eq. 4) are parameterized as photoperiods (through parameters  $DLC_{start}$ ,  $DLC_{end}$   
300 and  $DLF_{start}$ , respectively; see eq. 5 for the correspondence of e.g. day of year  $F_{start}$  with  
301 photoperiod  $DL_{F_{start}}$ ). We set the parameter bounds such that  $DLC_{start}$  ( $DLC_{end}$ ) cannot occur  
302 earlier than the autumn equinox (winter solstice) of previous year. Letting the model inference  
303 procedure free to find the most likely photoperiod limits for chilling and forcing accumulation  
304 within a large range (from autumn equinox of the previous year up to summer solstice of the  
305 current year), our model may represent several temporal combinations of the chilling and  
306 forcing temperature accumulation functions, corresponding to different hypotheses of the  
307 interplay between the endo- and eco-dormancy phases (i.e. sequential and parallel; see  
308 (Chuine et al., 2013).

### 309 *Empirical relation with spring average temperature*

310 This empirical model (analogous to Rossi et al., 2016) assumes that bE can be related to  
311 spring temperature via a linear regression, such that:

$$312 \quad bE = m_{T_{spg}} \times T_{spg} + p_{T_{spg}}, \quad (9)$$

313 where  $T_{spg}$  is the average January-June temperature ( $^{\circ}\text{C}$ ) calculated for each site-year, and  
314  $m_{T_{spg}}$  and  $p_{T_{spg}}$  are parameters of the regression line.

315

### 316 **Parameter estimation and model comparison through Bayesian inference**

317 To assess the models' abilities to simulate bE dates, we randomly split the bE data observed at  
318 the tree scale into calibration vs. validation subsets, with 70% of the data for calibration, and  
319 30% for validation. We checked that the distribution of the random calibration and validation

320 bE subsets did not differ (Wilcoxon rank sum test,  $p > 0.50$ ). Since the model fitting ability and  
321 inferred parameters may depend on the calibration subset used, we repeated the calibration  
322 procedure 30 times, using different calibration vs. validation subset combinations. The model  
323 evaluation results we report concern validation data, unless indicated.

324 Model parameters were fitted via Bayesian inference (see, e.g. Gelman, Carlin, Stern, &  
325 Rubin, 2004, and Fu, Campioli, Van Oijen, Deckmyn, & Janssens, 2012, for application in  
326 phenological modelling). The Bayesian framework calculates a posterior estimate and  
327 uncertainty for the model parameters, based on a prior distribution and the likelihood, defined  
328 as the probability of obtaining the observed data, given the model assumptions with their  
329 respective parameters. We use a Gaussian likelihood for all models:

$$330 \quad L(\theta) = \prod_{i=1,n} \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{P(\theta)_i - O_i}{\sigma}\right)^2\right], \quad (\text{eq. 10})$$

331 where  $O_i$  is the observed bE date (DoY) for site-year-tree  $i$ ;  $P(\theta)_i$  is the bE date (DoY)  
332 predicted by the model at point  $\theta$  in the parameter space, and  $\sigma$  is the standard deviation of  
333 the Gaussian distribution.

334 For all models considered, the fitted parameters included temperature and day length  
335 thresholds, for which natural extremes are given by the temperature and day length observed  
336 across the dataset. We therefore used uniform priors with these values as boundaries.

337 Posterior distributions were estimated with a differential evolution MCMC (DEzs,  
338 implemented in the ‘BayesianTools’ R package, (Hartig, Minunno, & Paul, 2017)). For each  
339 model and species, we ran 200,000 MCMC iterations and confirmed convergence of the chain  
340 after burn-in using the Gelman-Rubin criterion (Gelman, Meng, & Stern, 1996), requiring the  
341 *psrf* value for all parameters to be smaller than 1.05.

342 As a criterion to compare the models, we used posterior model weights, related to the Bayes  
343 factor (BF; Kass & Raftery, 1995, based on the model fit on the validation data. Assuming an  
344 equal prior weight on all models, the posterior weight for each model (PMW) is given by:

$$345 \quad PMW_i = \frac{ML_i}{\sum_j ML_j} \quad (\text{eq. 11}),$$

346 where ML is the marginal likelihood of model  $i$  or  $j$ . The marginal likelihood is the likelihood  
347 of the model for a given dataset, averaged over the parameter uncertainty. In our case, we  
348 calculated the ML for the validation data, with parameter uncertainties derived from the

349 posterior estimated with the calibration data. This approach of calculating the ML on a model  
350 calibrated by a subset of the data circumvents the known problem of the BF to be highly  
351 dependent on parameter priors (see, e.g., O’Hagan, 1995; van Oijen et al., 2013). The PMW  
352 can be intuitively interpreted as the probability that the respective model is ‘true’. In order to  
353 get a representative evaluation of the model abilities, we averaged PMW calculations across  
354 the 30 model-validation procedures.

355 Beside PMWs, we calculated for illustration the models’ root mean square error of prediction  
356 (RMSE) and Akaike Information Criteria (AIC), at the mode of their posterior parameter  
357 distributions (MAP).

### 358 **Quantifying bias in the model predictions**

359 We quantified the bias in model predictions of validation data at the scales of the tree, the  
360 site-year, the site (“is the model able to represent the inter-site variability of bE?”) and the  
361 year (“is the model able to represent the local annual anomaly of bE after removing the local  
362 bE average?”). Since there is no consensus in the statistical literature on how to evaluate  
363 model bias, we used two different methods. *Method 1*: We plotted and computed the  
364 coefficients of the linear ordinary least-squares regression of observed (y-axis) versus  
365 predicted (x-axis) data, as recommended by (Piñeiro, Perelman, Guerschman, & Paruelo,  
366 2008), and tested the null hypothesis: “the slope of the linear regression equals one and the  
367 intercept equals zero” (Wald test) using the LinearHypothesis function from the ‘car’ R  
368 package (Fox & Weisberg, 2011). *Method 2*: we performed a major axis (type II) linear  
369 regression of predicted (y-axis) versus observed (x-axis) data, and checked if the 95%-  
370 confidence intervals of the slope and intercept included one and zero, respectively (Mesplé,  
371 Troussellier, Casellas, & Legendre, 1996).

372

## 373 **Results**

### 374 **Wood phenological observations**

375 The observed bE dates spanned 90 days, ranging from March 16 (DoY 75) for a PISY tree at  
376 the southernmost site from the database (‘Moncayo’ site, Spain) to July 2 (DoY 183) for a  
377 LADE tree located at 1900-m on an altitudinal gradient (‘Lötschental site’, Switzerland; Table  
378 2). In this dataset, PISY was the earliest species to resume xylem cell enlargement in spring,  
379 showing ca. three-week earlier average bE than PCAB and 7-week earlier than PCMA and

380 LADE. The amplitude of bE dates spanned by each species varied from 49 days in PCMA to  
381 101 days in PISY, consistent with the size of the climate space occupied by each species in  
382 the dataset (Fig. 2).

383

### 384 **Performance of the models**

385 Whatever the tree species, the chilling-influenced heat sum model (*CiHS*) was identified as  
386 the best-supported (most likely) model for predicting bE, displaying the highest posterior  
387 model weights over validation data with  $PMW_{\text{valid}}$  from 0.67 to 1.00 (average 0.90; Table 3).  
388 The *CiHS* model largely outperformed models belonging to the *threshold* (i.e. *Tt* and *TDLt*  
389 models) or the *heat sum* (*HS*) classes, which both showed nil  $PMW_{\text{valid}}$  (Table 3). The  
390 prediction error of *CiHS* was substantially lower than that of other models structures (e.g.  
391 validation RMSE of *CiHS* was on average 1.3 days lower as compared to the heat-sum model  
392 *HS*, 3.6 days lower as compared to the temperature-and-photoperiod threshold model (*TDLt*),  
393 9.8 days lower as compare to the temperature-threshold model (*Tt*), Table 3). In PCMA, the  
394 empirical model predicting bE as a linear function of spring temperature (*MST*) received some  
395 support ( $PMW_{\text{valid}}=0.33$ ), but substantially less than *CiHS* ( $PMW_{\text{valid}}=0.67$ ). Beside its  
396 performance at the tree scale (Table 3), the *CiHS* model was also good at representing the  
397 variability of bE across site-years (Fig. 3), across sites (Suppl. Fig. S2), and across years  
398 (Suppl. Fig. S3). The *CiHS* model yielded unbiased predictions of the observations at all  
399 aggregation scales according to Method 1 for model bias testing (Table 4). Method 2 pointed  
400 more contrasted results: it confirmed the absence of bias at the scales of the site and of the  
401 site-year (except for LADE in the latter case; Table 4). However, it pointed biased results at  
402 the tree scale, and as regards annual anomalies (except for PCMA). In those cases, Method 2  
403 returned that the *CiHS* overestimated early bE and underestimated late bE dates (i.e. slopes of  
404 the major axis regression of predicted versus observed dates were less than one).

### 405 **Posterior parameter estimates for the *CiHS* model**

406 Since the *CiHS* model predicted unknown data best, we looked at its posterior parameter  
407 estimates to evaluate their biological reliability. We first note that most parameters of the  
408 *CiHS* model could be estimated well (meaning that prior uncertainty was considerably  
409 reduced), and that the estimates were similar across the 30 calibration-validation splittings of

410 the data (Fig. 4, see Table S2 for parameter values at the mode of the merged 30 posterior  
411 distributions).

412 In all species, chilling accumulation ( $DL_{Cstart}$ ) started earlier than or close to vernal equinox  
413 (corresponding to 12-hour photoperiod, Fig. 4, occurring on DoY 81, Fig. 5) and generally  
414 lasted up to late dates (defined by  $DL_{Cend}$ ), potentially up to the summer solstice when  
415 applicable. Notable exceptions were high-latitude PISY and PCAB. In PCAB, the duration of  
416 chilling accumulation was very short at high latitudes, virtually non-existent at low latitudes  
417 (Fig. 5) and presented a maximum duration of ca. 15 days at intermediate latitudes (ca. 54°N)  
418 due to latitudinal variations of the photoperiod course in spring.

419 Chilling accumulation resulted in an actual reduction of the forcing requirement for bE (all  $g$   
420 parameters were negative, Fig. 4), with a strong sensitivity to chilling exposure in LADE (-  
421 14.9 degree-days / chill unit) and PCAB (-27.9 degree-days / chill unit). The upper  
422 temperature threshold for chilling accumulation ( $T_c$ ) ranged from -5.6°C in PCAB to +6.1°C  
423 in PCMA (Fig. 4) with a median across species of +1.6°C. The lower temperature threshold  
424 for forcing accumulation ( $T_f$ ) ranged from -2.9°C in PCAB to +3.4°C in LADE (Fig. 4) with a  
425 median across species of +0.15°C. The start of forcing accumulation (defined by  $DL_{Fstart}$ )  
426 looked bounded by vernal equinox (Fig. 5). It occurred later than the start of chilling  
427 accumulation in both spruce species (PCAB and PCMA), but earlier than the start of chilling  
428 accumulation in PISY and LADE (Fig. 5).

429

## 430 Discussion

431 The purpose of this study was to improve our understanding of the phenology of wood  
432 formation, and in particular to unravel the causal triggers for the spring onset of xylem growth  
433 in coniferous species. To this end, we evaluated the ability of three families of  
434 ecophysiological models and one empirical model to predict the start of the enlargement  
435 period of the xylem cells. Our results demonstrate that models based on temperature sums  
436 perform better than those based on temperature- and photoperiod-thresholds do (Table 3).  
437 Moreover, our results clearly support the chilling-influenced heat sum model (*CiHS*),  
438 explicitly considering the processes of chilling and forcing temperature accumulation, for the  
439 prediction of the spring onset of wood formation. Beside its high posterior probability  
440 compared to the other models, the *CiHS* model also predicted the spring onset of xylem



441 formation with good accuracy. Its RMSE on the validation data, averaging 7.7 days (Table 3),  
442 is close to the temporal resolution of micro-core sampling from the trees (i.e. 7 days), and  
443 similar to the typical prediction accuracy of budburst (i.e. primary meristems), when deployed  
444 over continental gradients (e.g. Basler, 2016). The clear support for a chilling-influenced heat  
445 sum for the modelling of spring xylem phenology is different from what is reported in  
446 budburst model comparisons. For the latter, heat sums and chilling-influenced heat sums do  
447 not usually differ in their fit (Basler, 2016; Vitasse et al., 2011).

448 The identification of the *CiHS* model as receiving most support from the inference procedure  
449 suggests that both forcing and chilling temperatures play a role in determining the spring  
450 resumption date of xylem formation. To our knowledge, there is no *direct* evidence in the  
451 literature of a modulation of the date of onset of xylem cell formation in trees exposed to  
452 various chilling temperatures during winter and/or spring. Stem heating experiments showed  
453 that an artificial resumption of cambial activity can be triggered during late winter, but not in  
454 early winter (Begum, Nakaba, Oribe, Kubo, & Funada, 2010). This observation supports the  
455 existence of an endo-dormancy phase, during which the cambium activity is repressed by  
456 unknown tree internal factors (Delpierre, Vitasse, et al., 2016, but see Singh et al., 2017, for a  
457 review of dormancy processes in primary meristems). However, it does not prove, nor does it  
458 quantify the role of chilling temperatures in hastening the reactivation of xylem formation in  
459 spring. Thus, there is a clear need for quantifying the actual role of chilling temperatures in  
460 modulating the spring resumption of xylem formation, in line with pioneer works regarding  
461 buds and seeds (see Sarvas, 1974, reviewed in Hänninen, 2016), which have recently been  
462 actualized (e.g. Flynn & Wolkovich, 2018).

463 We delineated the time periods for the accumulation of chilling or forcing temperatures with  
464 photoperiod limits, instead of day of year (DoY) as usually done in phenological modelling  
465 (see e.g. Olsson and Jönsson, 2014; Basler 2016) for examples over large latitudinal  
466 gradients). If the use of DoY is perfectly sound in local studies (i.e. for which the relation  
467 between DoY and photoperiod is unequivocal), it is questionable in studies spanning  
468 continental scales since plants sense time from variations in the photoperiodic signal. Across a  
469 latitudinal gradient, a given photoperiod is reached at different DoYs (except the 12-hour  
470 photoperiod occurring at spring equinox (March 20) across the entire gradient). This resulted  
471 in large differences in our southern vs. northern study sites as regards the timing of the  
472 chilling accumulation for PCAB and the duration of both chilling and forcing accumulation in  
473 PISY (Fig. 5), two species spanning large latitudinal gradients in our dataset. Whether such

474 variations of the actual dates of cambium sensitivity to temperatures are realistic remains to  
475 be determined. This could experimentally be done by comparing the sensitivity of cambium to  
476 chilling in genetically identical plants placed in various photoperiod conditions, either in  
477 climate chambers (as done for budburst, e.g. Basler and Körner, 2014) or in natural conditions  
478 (e.g. International Phenological Gardens, Chmielewski & Rötzer, 2001).

479 In the *CiHS* model, the threshold temperatures promoting the progress of the eco-dormancy  
480 phase (forcing temperature threshold,  $T_f$ , ranging from  $-2.9^{\circ}\text{C}$  to  $+3.4^{\circ}\text{C}$ , Fig. 4) were  
481 comparable to values generally used in the modelling of budburst (typically  $0^{\circ}\text{C}$  or  $+5^{\circ}\text{C}$ ,  
482 Hänninen, 2016), based on experimental results (from  $-5^{\circ}\text{C}$  to  $+1^{\circ}\text{C}$  in, Heide, 1993). On the  
483 other hand, the values of threshold chilling temperatures ( $T_c$ ) determined by the parameter  
484 inference procedure span a larger range (from  $-5.6^{\circ}\text{C}$  to  $+6.1^{\circ}\text{C}$ ) and appear quite low in the  
485 cases of PCAB ( $-5.6^{\circ}\text{C}$ ) and LADE ( $-1.1^{\circ}\text{C}$ ) as compared to the values either determined  
486 experimentally in buds and seeds (for which Sarvas, 1974, reports  $-3^{\circ}\text{C}$  as a lower limit for  
487 chilling effectiveness) or considered by expert judgment ( $0^{\circ}\text{C}$  to  $+4.5^{\circ}\text{C}$  in Coville, 1920;  
488  $+2^{\circ}\text{C}$  to  $+4^{\circ}\text{C}$  for cambium in Little & Bonga, 1974) as effective for chilling.

489 From a larger perspective, the questions about the plausibility of parameter values we inferred  
490 are further linked with the range of environmental conditions in which the bE data were  
491 obtained. Indeed, inferring model parameters from data acquired from trees growing under  
492 natural conditions, inevitably exposed to multiple interacting environmental factors (think e.g.  
493 of the strong concurrent latitudinal temperature and photoperiod gradient), is not equivalent to  
494 inferring them from a controlled experiment where the environmental conditions can be at  
495 least partially be disentangled (Verdier et al., 2014), and their biological interpretability is  
496 necessarily less generic. However, we noticed that our species-specific parameterizations of  
497 the *CiHS* model were able to reproduce the locally observed between-species difference in bE  
498 at those sites where two species of interest co-occur (Fig. 6), giving credit to the overall  
499 plausibility of the inferred parameters.

500 In this study we used model formulations initially developed for simulating the occurrence of  
501 budburst, assuming similar environmental controls of the phenology of primary and  
502 secondary meristems (Delpierre, Vitasse, et al., 2016). Even for budburst, those models lack  
503 an indisputable biological support (Clark, Salk, Melillo, & Mohan, 2014; Delpierre, Vitasse,  
504 et al., 2016). New model formulations for the phenology of budburst appear in the literature  
505 from time to time, considering more complex interactions of chilling and forcing temperatures

506 in interaction with photoperiod (e.g. Blümel & Chmielewski, 2012; Caffarra, Donnelly, &  
507 Chuine, 2011). Similar to the necessary effort to calibrate and compare those continuous-state  
508 budburst models to continuous data (for instance by measuring the release of plasmodesmata  
509 closure by callose, which is an indicator of bud endo-dormancy, Singh et al., 2017), a  
510 biologically-undisputable modelling of spring cambial activity will require the evaluation of  
511 those models with continuous seasonal markers of cambial cells activity (i.e. cytoplasmic  
512 changes in cambial cells such as presence and form of microtubules, vacuoles, lipid droplets,  
513 plastids and other cell organelles; Begum et al., 2012; Chaffey & Barlow, 2002; Prislán et al.,  
514 2013; Rensing & Samuels, 2004), or metabolite content.

515 Even if the *CiHS* model has no clear mechanistic foundation, we remind that the exposure to  
516 chilling temperature promotes soluble sugars accumulation from starch conversion, especially  
517 sucrose (along with raffinose, stachyose and other metabolites; Sakai & Larcher, 1987;  
518 Strimbeck, Schaberg, Fossdal, Schröder, & Kjellsen, 2015) that remain high until spring de-  
519 hardening. Since cell production is limited by local sucrose availability (Deslauriers, Huang,  
520 Balducci, Beaulieu, & Rossi, 2016), we posit that exposure to chilling temperatures may  
521 constitute a local pool of sucrose readily available for cell production when temperatures  
522 become favourable for mitosis and/or cell expansion. In case of low chilling, this local  
523 sucrose pool would be low, and carbon-fueling for cell formation would rely more on the  
524 resumption of photosynthesis, which responds to forcing temperature accumulation (Mäkelä,  
525 Hari, Berninger, Hänninen, & Nikinmaa, 2004; Pelkonen & Hari, 1980). This mechanistic  
526 hypothesis is coherent with the general behaviour of the *CiHS* model (the required forcing  
527 accumulation decreases with increasing chilling exposure), and would explain why we infer in  
528 some species low temperature thresholds for chilling accumulation (-1.1°C in LADE, -5.6°C  
529 in PCAB). Indeed, the rate of starch to sugar conversion has been shown to be maximum at  
530 temperatures from -3°C to -5°C, and continued down to -15°C (in *Salix sachalinensis* twigs,  
531 Sakai, 1966).

532 Moreover, the successful use of model structures designed and used to predict budburst to  
533 simulate the resumption of cambial spring activity raises the question of the coordination and  
534 interaction of the phenologies of tree organs (Delpierre, Vitasse, et al., 2016). Phytohormones  
535 can play a significant role; with e.g. auxins produced in expanding buds influencing the rate  
536 of stem cambial divisions (see review of Sorce, Giovannelli, Sebastiani, & Anfodillo, 2013).  
537 Yet, the important role of auxin is also interconnected with cytokinin in the vascular  
538 cambium. Although auxin peak in the middle of cambium and cytokinin in the middle of

539 phloem, the latter acts as a positive regulator of cell division in the vascular cambium by  
540 increasing the number of cambial cell (Immanen et al., 2016) because of its crucial role on the  
541 cell division cycle (Schaller, Street, & Kieber, 2014). Thus, the resumption of xylem  
542 formation in spring is at least partially independent from auxin-producing buds, as clearly  
543 demonstrated in stem heating experiments (where xylem formation resumes artificially  
544 whereas buds remain dormant, Begum et al., 2010; Gricar et al., 2006; Oribe, Funada, &  
545 Kubo, 2003), and from the observed earlier timing of enlargement of new xylem cells, as  
546 compared to bud elongation in the evergreen coniferous trees studied here (*Picea abies*, *Pinus*  
547 *sylvestris*, *Picea mariana*; Antonucci et al., 2015; Cuny et al., 2012; Huang et al., 2014;  
548 Michelot et al., 2012). The presence of auxins in overwintering tissues (Egierszdorff, 1981),  
549 and of a local pool of sucrose (see above) may decouple the onset of cambium division and  
550 xylem enlargement from the timing of bud elongation, as observed from stem heating  
551 experiments; along with the presence of signal-transduction chains involving phytochromes  
552 (proteins acting as photoreceptors, i.e. able to sense modifications of the photoperiod) in the  
553 cambium (Petterle, Karlberg, & Bhalerao, 2013), this suggests that the cambium may well  
554 respond to variations of environmental conditions independently from buds. This hypothesis  
555 is supported by inter-annual variability in the delays between the spring phenophases of wood  
556 and leaves in both gymnosperms (Cuny et al., 2012) and angiosperms (Takahashi, Okada, &  
557 Nobuchi, 2013).

558  
559 The chilling-influenced heat sum model produced mostly unbiased results when the data were  
560 aggregated at the site-year or at the site scale (Table 4), pointing to its overall accurate  
561 capacity of to simulate the spring resumption of xylem formation in coniferous species. Yet,  
562 one of our bias-detection methods (method 2) suggested that the model underestimated the  
563 range of tree individual bE (in all species, Table 4) and the annual bE anomalies (in 3 out of 4  
564 species, Table 4, Fig. S3). Though our models rely on environmental (temperature and  
565 photoperiod) data collected at the tree population scale, we conducted the parameter inference  
566 with the most basic level of information available (i.e. at the individual tree level, see section  
567 2.5). It is clear that part of the model bias that is detected at the individual scale is related to  
568 the model structural incapacity to simulate the variety of individual tree responses to the same  
569 environment that is observed in a tree population (Delpierre, Guillemot, Dufrêne, Cecchini, &  
570 Nicolas, 2017) and can actually be quite large (e.g. the within-population SD of observed bE  
571 dates for a given year is 5 days on average, Table 2). Bias in the predictions of annual bE  
572 anomalies may further originate from the simplicity of the model structure, which probably

573 does not represent the whole range of environmental interactions resulting in the spring onset  
574 of xylem formation.

575 A study aiming at simulating the date of budburst of *Betula pendula* and *Picea abies*  
576 individuals from central to Northern Europe (i.e. a bioclimatic scale comparable to the one  
577 considered in our work) reported a lower performance over validation data as compared to our  
578 results for bE (with prediction RMSE of 8.9 and 9.1 days, respectively for their best heat sum  
579 model), along with a non-homogeneous bias over the continent, suggesting a role for the local  
580 adaptation of trees phenological traits (Olsson & Jönsson, 2014). It is not clear whether the  
581 latitudinal bias observed in Olsson & Jönsson (2014) originates from local adaptation (that  
582 has been evidenced several times for budburst, see e.g. Chuine, Mignot, & Belmonte, 2000;  
583 Osada et al., 2018; Vitasse, Delzon, Bresson, Michalet, & Kremer, 2009; von Wuehlich,  
584 Krusche, & Muhs, 1995) or is related to the uncertainty of budburst observations recorded  
585 through local phenological protocols. The data we use in our work are less prone to such  
586 problems since the observations were collected and processed according to a common  
587 protocol across the entire study zone (Rathgeber et al., 2011; Rossi et al., 2016). To this  
588 respect, we conclude from the absence of bias in the prediction of site average dates of bE  
589 (Table 4, Fig. S2) that local adaptation is, if any, of marginal influence in determining bE  
590 (Perrin, Rossi, & Isabel, 2017) as compared to the plasticity of bE driven by varying  
591 temperature and photoperiod conditions.

592 This study is the first comparative assessment of ecophysiological models aiming at  
593 simulating the spring resumption of xylem formation in trees. We demonstrated that chilling-  
594 influenced heat sum models are best supported by the data for the four coniferous species  
595 studied. Thus, analogous to what is commonly observed for buds, we state that winter-spring  
596 temperatures exert ambivalent effects on the spring onset of wood formation (bE) (i.e. on the  
597 one hand, warmer temperatures tend to hasten the occurrence of bE through the accumulation  
598 of *forcing* temperature, but on the other hand they are associated to less *chilling*, imposing a  
599 higher forcing-temperature sum to trigger wood formation). Previous results from (Rossi et  
600 al., 2011) suggested that spring warming would result in a continuous trend to earlier bE in  
601 the next decades. Our results question these predictions, since warming reduces the number of  
602 chilling days. This is probably the cause of the recently evidenced reduced sensitivity of  
603 spring leaf phenology to warm temperatures (Fu et al., 2015), which we also forecast to  
604 happen for wood formation (note that the length of wood phenology time series is much  
605 shorter than for bud phenology, so that this hypothesis remains to be tested).

606 Our work paves the way for the development of ecophysiological models simulating the  
607 whole phenological sequence of wood formation. We expect the *CiHS* model to be included  
608 as a component of schemes representing the whole seasonal cycle of wood formation, into  
609 which subsequent wood formation phases would partially depend on the occurrence of bE  
610 (Hänninen & Kramer, 2007; Lupi et al., 2010). Such a model is also urgently needed in  
611 ecosystem models of the carbon cycle (Delpierre, Vitasse, et al., 2016) which are undergoing  
612 core changes in their representation of wood growth (Guillemot et al., 2017; Schiestl-Aalto et  
613 al., 2015).

614

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635

### 636 **Authors' contributions**

637 N.D. and C.B.K.R initiated the project. N.D. designed the study, performed the research and  
638 wrote the manuscript. F.H. advised on the Bayesian inference framework. N.D., S.L.,  
639 C.B.K.R., F.H. and A.D. analysed results. J.J.C., H.C., K.C., A.D., P.F., J.G., J.-G.H., C.K.,  
640 M.L., H.M., E.M.C., P.N., W.O., P.P., S.R., V.T., H.V. and C.B.K.R. collected xylem micro-  
641 cores and produced phenological data. C.B.K.R. compiled the phenological database. E.M.C.  
642 made Figure 1. All authors commented on the manuscript.

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911

## 912 Tables

913 **Table 1. Overview of the tested models and their parameters.**  $T_a$ = daily average air  
 914 temperature (°C);  $DL$ = photoperiod (hours). See text for definition of the model parameters.

Model name	Type	Environmental variables	Fitted parameters (number)	Equation reference
$T_t$	temperature threshold	$T_a$	$T^*$ (1)	1
$TDL_t$	temperature and photoperiod thresholds	$T_a$ , $DL$	$DL^*$ , $T^*$ (2)	2
$HS$	Heat sum	$T_a$ , $DL$	$DLF_{start}$ , $T_f$ , $F^*$ (3)	3-5
$CiHS$	Chilling-influenced heat sum	$T_a$ , $DL$	$DLC_{start}$ , $DLC_{end}$ , $DLF_{start}$ , $T_c$ , $T_f$ , $g$ , $h$ (7)	6-8
$MST$	Regression line	January-June average temperature	$mT_{spg}$ , $pT_{spg}$ (2)	9

915

916 **Table 2. Overview of the wood phenology data.** bE= date of the beginning of xylem cell  
 917 enlargement (DoY),  $\Delta$ bE= amplitude of bE dates (days). The ‘within site-year SD’ metric is  
 918 the average standard deviation of bE among trees sampled on a given site-year.

Tree species	Number of site-years	Number of observations	Mean bE (DoY)	SD of bE (days)	Min. bE (DoY)	Max. bE (DoY)	$\Delta$ bE within-species (days)	Within-site-year SD (days)
LADE	62	300	150	12	118	183	65	5.4
PISY	37	175	112	20	75	176	101	5.4
PCAB	77	336	136	16	101	177	76	4.2
PCMA	42	294	152	9	128	177	49	4.8



919 **Table 3. Model performance comparison.** PMW= posterior model weight (eq. 11); RMSE= root mean square error (days);  $\Delta AICc$ =  
 920 differential Akaike Information Criterion, corrected for small sample biases (calculated as the difference from minimum AICc across all models;  
 921 according to this metric, the best model at maximum likelihood has a score of 0). PMWs are established over the whole posterior distribution.  
 922 RMSE and AICc were calculated at the point of maximum likelihood (MAP). We report here the medians of those metrics, established across the  
 923 30 calibration re-samplings. The *CiHS* model results appear in bold characters, as displaying the highest PMW over validation data in all species.  
 924

	Model class	Model name	PMW <sub>calib</sub>	PMW <sub>valid</sub>	RMSE <sub>calib</sub>	RMSE <sub>valid</sub>	$\Delta AICc_{calib}$	$\Delta AICc_{valid}$
LADE (n calib= 210, n valid=90)	threshold	Tt	0.00	0.00	14.7	15.0	226	89
	threshold	TDLt	0.00	0.00	9.9	10.3	95	30
	heat sum	HS	0.00	0.01	8.2	8.7	36	8
	chilling- influenced	<b>CiHS</b>	<b>1.00</b>	<b>0.94</b>	<b>7.5</b>	<b>8.1</b>	<b>0</b>	<b>0</b>
	heat sum regression	MST	0.00	0.00	8.7	8.9	53	13
PISY (n calib = 123, n valid = 52)	threshold	Tt	0.00	0.00	21.5	24.6	208	98
	threshold	TDLt	0.00	0.00	14.6	15.2	119	46
	heat sum	HS	0.00	0.00	11.4	11.2	63	18
	chilling- influenced	<b>CiHS</b>	<b>1.00</b>	<b>1.00</b>	<b>8.4</b>	<b>9.3</b>	<b>0</b>	<b>0</b>
	heat sum							

	regression	MST	0.00	0.00	15.6	15.6	133	49
	threshold	Tt	0.00	0.00	16.8	17.2	378	151
	threshold	TDLt	0.00	0.00	12.4	12.6	221	92
PCAB (n calib = 236, n valid = 100)	heat sum	HS	0.00	0.00	9.8	10.1	119	35
	chilling- influenced	<b>CiHS</b>	<b>1.00</b>	<b>1.00</b>	<b>7.5</b>	<b>7.9</b>	<b>0</b>	<b>0</b>
	heat sum							
	regression	MST	0.00	0.00	11.5	11.6	154	62
	threshold	Tt	0.00	0.00	13.1	13.2	334	139
	threshold	TDLt	0.00	0.00	7.3	7.3	116	45
PCMA (n calib = 206, n valid = 88)	heat sum	HS	0.00	0.00	5.8	6.1	38	15
	chilling- influenced	<b>CiHS</b>	<b>1.00</b>	<b>0.67</b>	<b>5.2</b>	<b>5.6</b>	<b>0</b>	<b>0</b>
	heat sum							
	regression	MST	0.00	0.33	6.7	6.9	32	1

926 **Table 4. Assessment of model bias on validation data.** We tested the model ability to produced unbiased predictions of bE from the validation  
 927 subsets at different scales, with two different methods (see Material and Methods 2.6 for details). The slopes and intercepts estimates are reported  
 928 with their 95%-confidence intervals between parentheses. Unbiased predictions are characterized by both slope= 1 and intercept=0. In Method 1,  
 929 we report the p-value of the Wald test (testing for unit slope and zero intercept as the null hypothesis). In Method 2, we identify biased  
 930 predictions when either the slope or intercept confidence intervals do not include one or zero, respectively. ‘yes’ / ‘no’ mark biased / unbiased  
 931 predictions.

		Method 1					Method 2		
bE data aggregation scale	Species	slope	intercept	F	P(>F)	Bias ?	slope	intercept	Bias ?
tree	LADE	1.04 (0.93, 1.15)	-6.2 (-23.0, 10.6)	0.41	0.66	<b>no</b>	0.67 (0.60, 0.74)	49.6 (38.4, 60.0)	yes
	PISY	1.02 (0.91, 1.12)	-4.2 (-16.5, 8.1)	2.57	0.08	<b>no</b>	0.79 (0.71, 0.87)	25.9 (16.3, 34.8)	yes
	PCAB	0.98 (0.92, 1.04)	2.5 (-5.4, 10.4)	0.34	0.71	<b>no</b>	0.88 (0.83, 0.93)	16.1 (8.9, 22.9)	yes
	PCMA	1.04 (0.94, 1.15)	-7.2 (-23.2, 8.7)	1.31	0.27	<b>no</b>	0.68 (0.61, 0.75)	49.2 (38.5, 59.2)	yes
site-year*	LADE	1.06 (0.90, 1.22)	-9.4 (-33.5, 14.6)	0.34	0.71	<b>no</b>	0.80 (0.68, 0.92)	30.6 (11.52, 47.6)	yes
	PISY	1.02 (0.85, 1.19)	-4.1 (-24.4, 16.2)	0.84	0.44	<b>no</b>	0.88 (0.74, 1.04)	15.4 (-3.1, 31.4)	<b>no</b>
	PCAB	0.99 (0.88, 1.10)	2.2 (-12.7, 17.1)	0.19	0.83	<b>no</b>	0.91 (0.81, 1.02)	11.8 (-2.6, 24.9)	<b>no</b>
	PCMA	1.01 (0.87, 1.16)	-3.1 (-25.1, 19.0)	1.83	0.18	<b>no</b>	0.91 (0.79, 1.05)	14.5 (-6.2, 33.3)	<b>no</b>
site*	LADE	1.13 (0.87, 1.40)	-19.3 (-58.8, 20.2)	0.72	0.51	<b>no</b>	0.82 (0.65, 1.04)	25.9 (-6.1, 52.3)	<b>no</b>
	PISY	1.03 (0.82, 1.23)	-6.0 (-31.3, 19.2)	0.97	0.4	<b>no</b>	0.92 (0.75, 1.13)	11.7 (-12.5, 31.8)	<b>no</b>
	PCAB	1.01 (0.83, 1.20)	-1.8 (-26.6, 23.1)	0.02	0.98	<b>no</b>	0.91 (0.76, 1.09)	11.47 (-13.0,	<b>no</b>

								32.2)	
	PCMA	1.27 (0.86, 1.68)	-42.0 (-102.1, 18.2)	3.55	0.11	<b>no</b>	0.75 (0.53, 1.03)	38.4 (-1.8, 70.2)	<b>no</b>
year anomaly**	LADE	1.02 (0.81, 1.23)	0 (-1.1, 1.1)	0.02	0.98	<b>no</b>	0.73 (0.59, 0.89)	0 (0, 0)	yes
	PISY	1.20 (0.67, 1.74)	0 (-2.0, 2.0)	0.3	0.74	<b>no</b>	0.43 (0.25, 0.64)	0 (0, 0)	yes
	PCAB	1.17 (1.00, 1.34)	0 (-1.1, 1.1)	2.02	0.14	<b>no</b>	0.69 (0.60, 0.80)	0 (0, 0)	yes
	PCMA	0.89 (0.74, 1.04)	0 (-0.8, 0.8)	1.06	0.36	<b>no</b>	1.02 (0.86, 1.21)	0 (0, 0)	<b>no</b>

932 \* bE dates were simulated at the tree individual scale, and subsequently averaged at the site-year or site scale; \*\* bE dates were simulated at the tree individual  
933 scale. For calculating annual anomalies, we subtracted the average bE date, established along the observation period, to bE data averaged at the site-year scale.

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935 **Figure captions**

936 **Figure 1. Location of the study sites.**

937 **Figure 2. Climate space at the observed date of bE.** Each bE datum is placed in a climate  
938 space defined by the day length at bE (x-axis) and the average temperature over the 15-day  
939 interval preceding bE (y-axis).

940 **Figure 3. Chilling-influenced heat sum (CiHS) model evaluation over validation data.**  
941 Predictions are reported at the tree scale (grey dots) and aggregated site-year scale (points,  
942 colours according to the average January-June temperature of the site-year). The thick black  
943 line is the least square regression line of predicted versus observed data. The one-to-one  
944 relation appears as the thin grey line. *NSE*= Nash-Sutcliffe model efficiency; *slope*= slope of  
945 the linear regression; *int*= intercept of the linear regression. The displayed statistics are  
946 calculated for site-year aggregated data. See Table 3 for statistics on tree-scale data.

947 **Figure 4. Posterior parameter distributions.** Parameters are shown for the *CiHS* model,  
948 which performed best over the validation data for each species. Grey lines represent each of  
949 the 30 inference procedures, with the overall distribution appearing as coloured line. For each  
950 parameter, the limits on the x-axis mark the bounds set to the uniform prior density. The mode  
951 of the overall distribution appears for each parameter on the upper left-hand corner (e.g.  
952  $DL_{Cstart} = 12.7$  hours for LADE). See Material and Methods for parameters description, and  
953 Table S2 for parameter values at the mode of the merged 30 posterior distributions.

954 **Figure 5. Variations of chilling and forcing accumulation time intervals along latitudinal  
955 gradients.** This figure displays the temporal interval of chilling accumulation (with the  
956 starting date plotted as '\*' and the ending date as 'o', linked by a straight line) and the starting  
957 date of forcing (plotted as '△'). The colour of the symbols indicates the northernmost (blue)  
958 or southernmost (red) latitude by species. For PCAB, we also illustrate an intermediate  
959 situation (latitude = 54°N, grey symbols). Dashed black line represents vernal equinox;  
960 continuous black line represents summer solstice.

961 **Figure 6. Comparing observed and simulated interspecific differences in the date of bE.**  
962 For those site-years where two species of interest have been sampled simultaneously, we  
963 plotted the observed and predicted between-species differences in bE dates ( $\Delta bE$ , days). Each

964 single point represents one site-year. (a, b): compare the distribution of differences; (c, d):  
965 compare observed and predicted differences for each site-year.

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966 **Supporting information**

967 Additional supporting information may be found in the online version of this article.

968 **Figure S1.** Logistic models are precise in determining temperature thresholds for the  
969 beginning of xylem growth, but are not predictive.

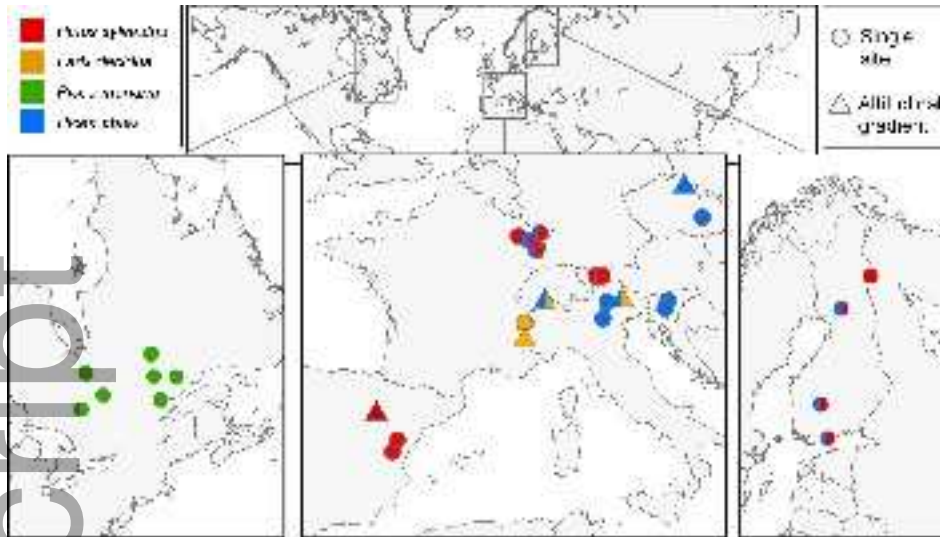
970 **Figure S2.** Model evaluation performance over validation data, aggregated per site.

971 **Figure S3.** Model evaluation performance over validation data, for annual anomalies.

972 **Table S1.** Study sites.

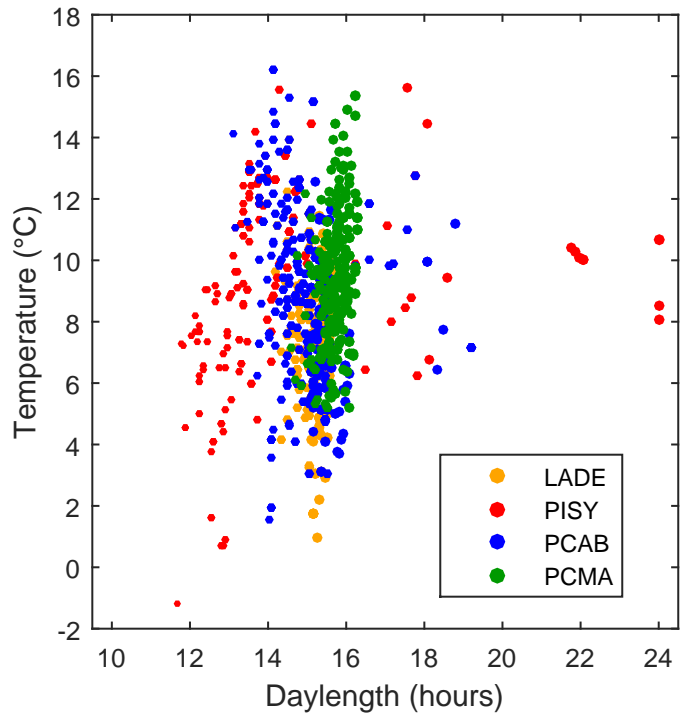
973 **Table S2.** Parameter values for the chilling-influenced heat sum (*CiHS*) model.

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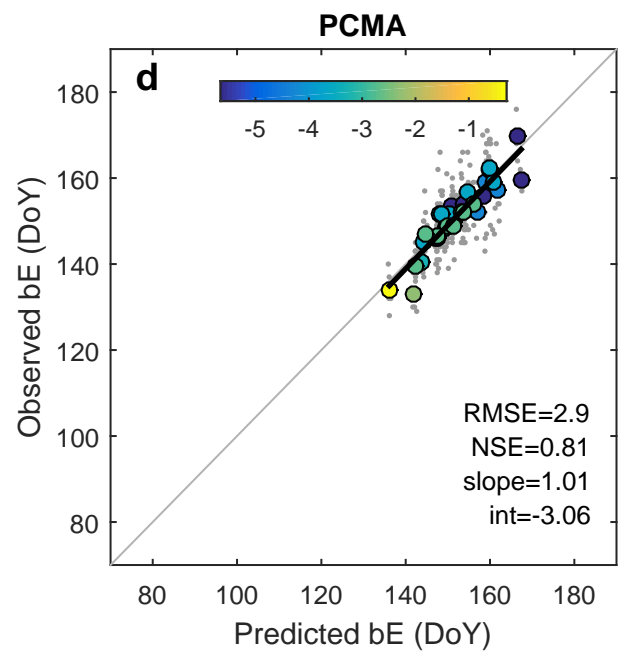
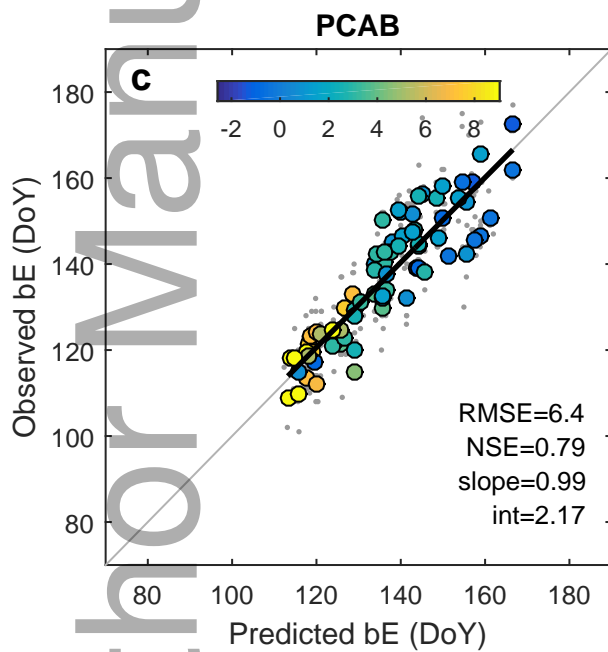
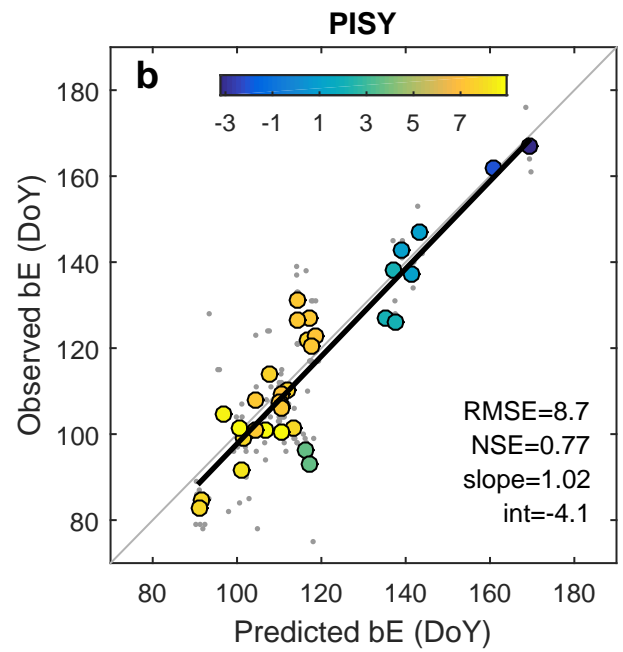
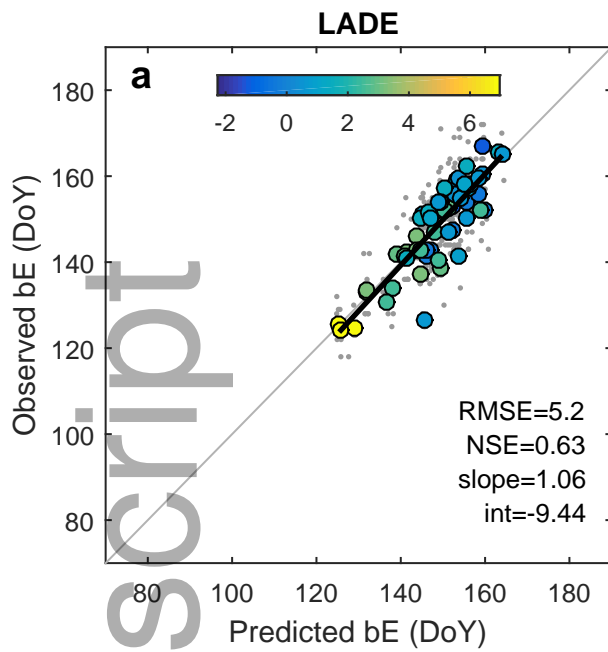


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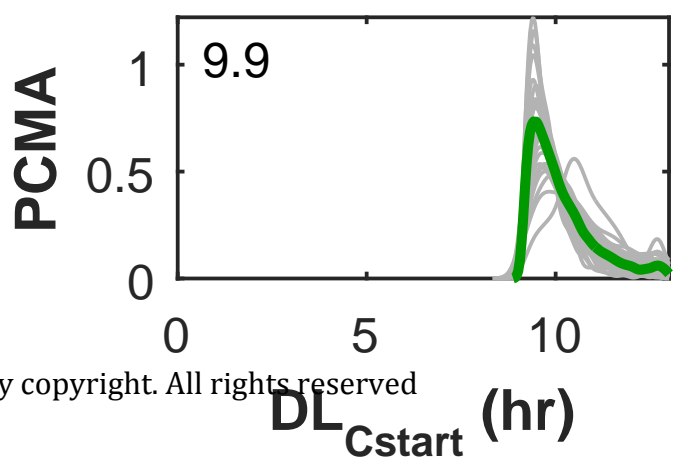
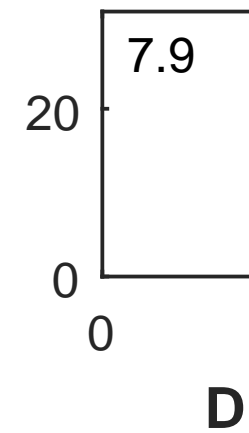
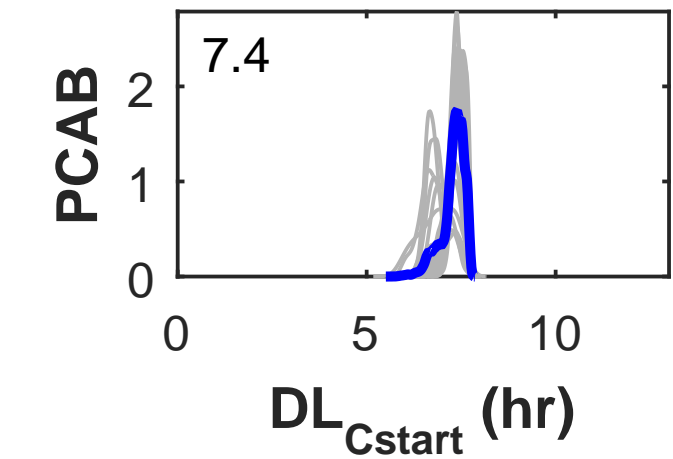
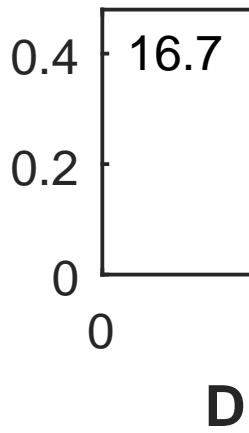
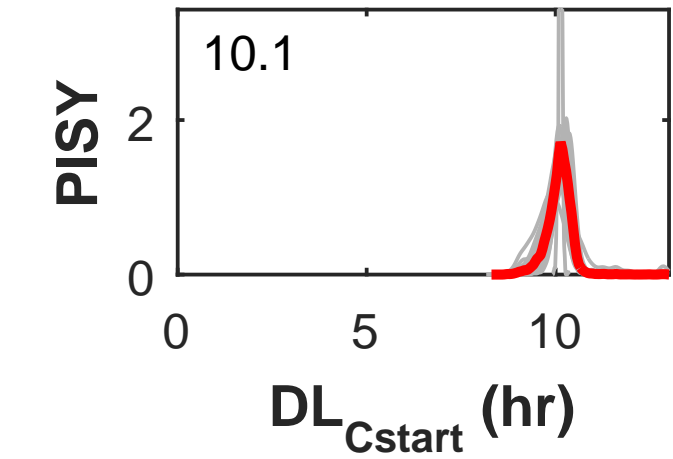
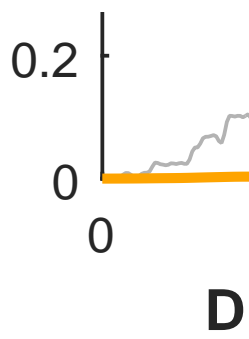
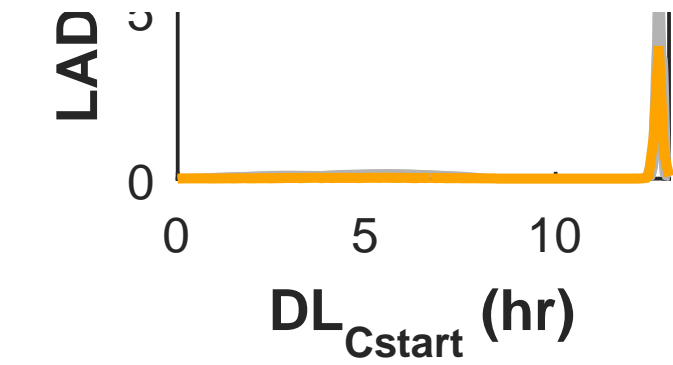


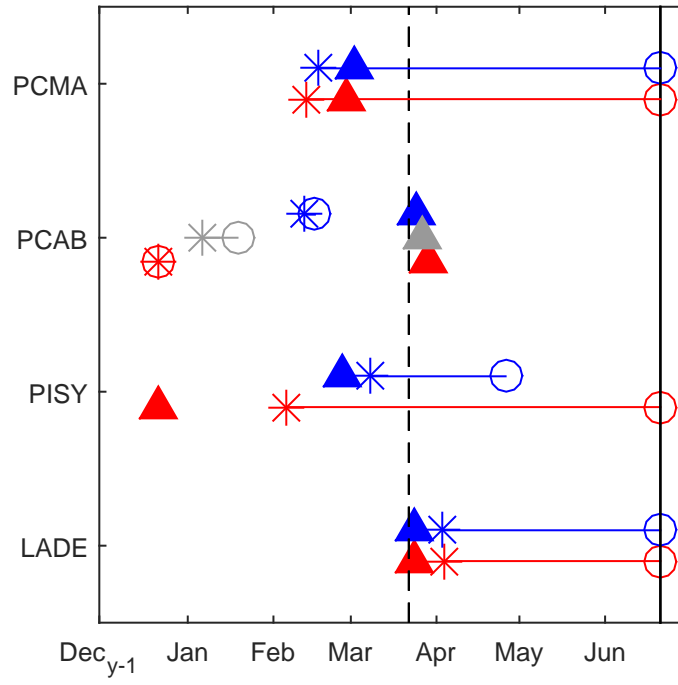


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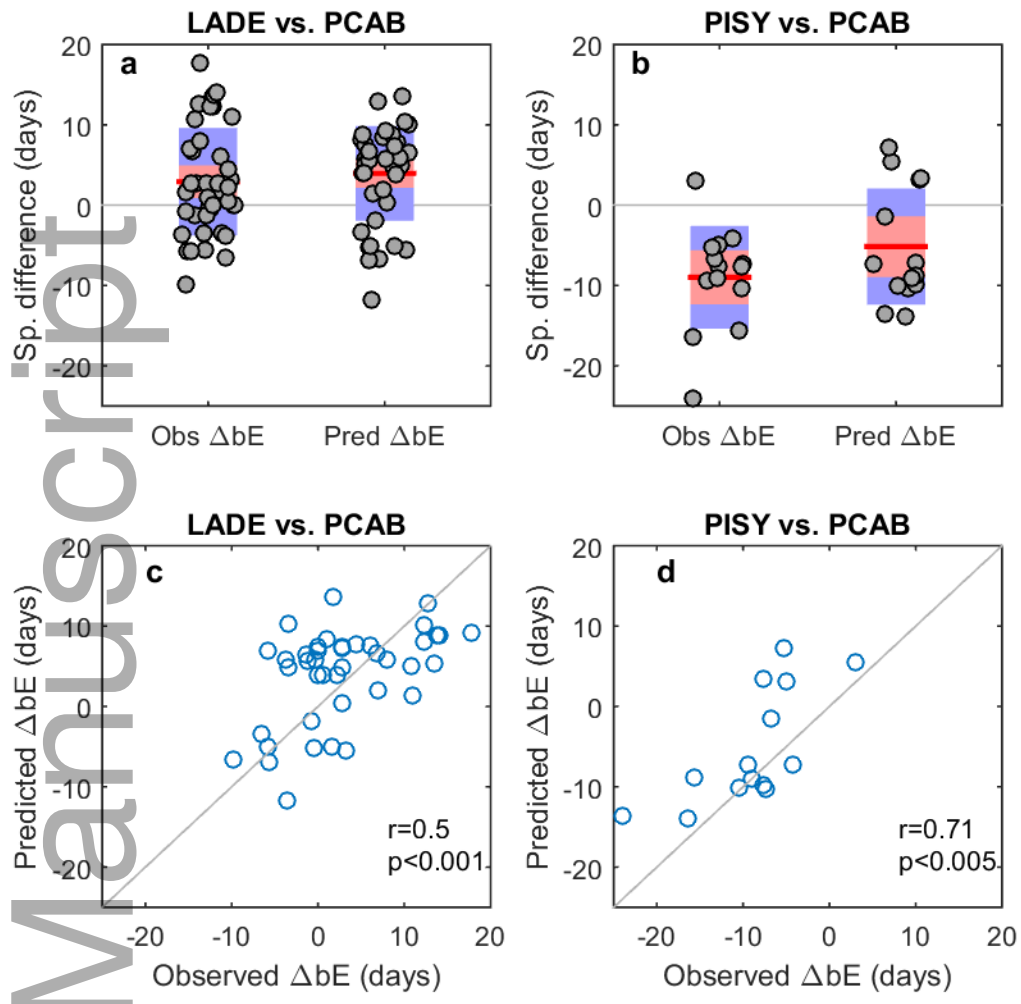


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