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1 **Spatial vs. temporal effects on demographic and genetic structures: the roles of**
2 **dispersal, masting and differential mortality on patterns of recruitment in *Fagus***
3 ***sylvatica*.**

4
5 Sylvie ODDOU-MURATORIO*,

6 Etienne K. KLEIN §, Giovanni G. VENDRAMIN#, Bruno FADY*,

7 * *INRA, UR 629, Ecologie des Forêts Méditerranéennes, F-84914 Avignon France*

8 § *INRA, UR 546, Biostatistique et Processus Spatiaux, F- 84914 Avignon France*

9 # *CNR, Plant Genetics Institute, Via Madonna del Piano 10, 50019 Sesto Fiorentino*

10 *(Firenze), Italy*

11 *Corresponding author: Sylvie ODDOU-MURATORIO*

12 *Phone: +33-490 135 914, Fax: +33-490 135 959.*

13 *e-mail: oddou@avignon.inra.fr.*

14
15 Short title: Demo-genetics of beech recruitment.

16 Keywords: contemporary gene flow, spatial genetic structure, microsatellite, tree, spatially
17 explicit mating model.

18

19

20 ***Abstract***

21 Trees' long life span, long-distance dispersal abilities and high year-to-year variability in
22 fecundity, are thought to have pervasive consequences for the demographic and genetic
23 structure of recruited seedlings. However, we still lack experimental studies quantifying the
24 respective roles of spatial processes such as restricted seed and pollen dispersal and temporal
25 processes such as mast seeding, on patterns of regeneration. Dynamics of European beech
26 (*Fagus sylvatica*) seedling recruitment was monitored in three plots from 2004 to 2006. Six
27 polymorphic microsatellite genetic markers were used to characterize seedlings and their
28 potential parents in a 7.2 ha stand. These seedlings were shown to result from 12 years of
29 recruitment, with one predominant year of seedling recruitment in 2002 and several years
30 without significant recruitment. Using a spatially explicit mating model based on parentage
31 assignment, short average dispersal distances for seed ($\delta_s = 10.9$ m) and pollen (43.7 m $<$
32 $\delta_p < 57.3$ m) were found, but there was also a non-negligible immigration rate from outside the
33 plot ($m_s = 20.5\%$; $71.6\% < m_p < 77.9\%$). Hierarchical analyses of seedling genetic structure
34 showed that (1) most of the genetic variation was within plots; (2) the genetic differentiation
35 among seedling plots was significant ($F_{ST} = 2.6\%$) while (3) there was no effect of year-to-
36 year seed rain variation on genetic structure. In addition, no significant effect of genetic
37 structure on mortality was detected. The consequences of these results for the prediction of
38 population dynamics at ecological time-scales are discussed.

39 ***Introduction***

40 Understanding the effects of demographic and genetic processes on the amount and spatial
41 distribution of genetic variation in natural populations is one of the main objectives of modern
42 evolutionary ecology. These issues have gained a renewed interest in the context of increased
43 habitat fragmentation and rapid environmental change, which stress the need to understand
44 and predict evolutionary trajectories of populations at ecological time scale. At short time and
45 spatial scales, demographic (survival, growth, competition) and evolutionary forces
46 (selection, gene flow and genetic drift) tightly interact to shape spatial patterns of allelic
47 frequencies across life stages or generations. In return, genetic variation affects population
48 dynamics by determining individual capacities of survival, growth, and reproduction (Lande,
49 1982). The study of this interplay between demographic and short-term evolutionary
50 processes is sometimes referred to as demo-genetics and builds on the theoretical framework
51 mainly developed by Lande (1982) and recently adapted by Coulson *et al.* (2006).
52 In trees, the period spanning from seed dispersal to early seedling recruitment is thought to be
53 a major transition step where important demo-genetic interactions take place and have
54 pervasive consequences on the structure and dynamics of tree populations (Petit & Hampe,
55 2006). This step is characterised in particular by a massive mortality of seeds/seedlings
56 produced during an individual life time, with typically only one seed in a million surviving as
57 a reproductive adult (Petit & Hampe, 2006). Moreover, recruitment studies highlight the role
58 of various demographic and genetic processes and of their variation on patterns of
59 regeneration; major processes are the spatial distribution and density of reproductive plants,
60 their seed outputs, the shape and form of their dispersal kernel and the spatial patterns of
61 microsites favorable for seedling establishment (Clark *et al.*, 2007; Clark *et al.*, 1999; Nathan
62 & Muller-Landau, 2000). In the following section, we will focus in particular on: (1) dispersal
63 limitation, which is a major factor shaping variation of recruitment through space; (2) mast

64 seeding (i.e. synchronous intermittent production of large seed crops) which is an important
65 factor shaping the variation of recruitment through time particularly in temperate forests
66 (Kelly & Sork, 2002; Piovesan & Adams, 2001); (3) mortality at early stages of the tree life
67 cycle that causes major variation in the demographic and genetic structure across life stages.

68 The combined role of propagule (seed and pollen) production and dispersal on patterns
69 of recruitment is widely acknowledged and studied in population dynamics and genetics. Seed
70 dispersal and individual seed production shape both the initial spatial pattern of seedling
71 abundance (Clark *et al.*, 2007; Clark *et al.*, 1999; Nathan & Muller-Landau, 2000) and genetic
72 relatedness among established individuals (Wright, 1943). Unless pollen production is
73 strongly limited or population density is very low (Sagnard *et al.*, 2011), pollen dispersal is
74 usually considered as driving mainly patterns of genetic relatedness. By contrast, the temporal
75 component of the regeneration pattern due to inter-annual variation in seed production has
76 received less attention, in particular from a population genetics perspective. It has been
77 suggested that the genetic consequences of high variation in seed production among
78 individuals are reduced over time due to the fact that seed production ranking among trees
79 varies strongly between years (Krouchi *et al.*, 2004). This phenomenon tends to increase the
80 effective population size and therefore decrease the spatial variation in genetic relatedness
81 over the whole regeneration phase. From a population dynamics perspective, the few studies
82 focusing on recruitment patterns across space and time show contrasting results. Some studies
83 have found across-year consistency in seed-fall and seedling distribution with a strong site
84 effect (Wright *et al.*, 2005), while in others, across-year variation was higher than variation
85 across sampling sites (Beckage *et al.*, 2005). It thus remains largely unknown how spatial and
86 temporal recruitment dynamics interact across heterogeneous landscapes, which of these
87 components has a greater effect on patterns of regeneration in long-lived plants, and under

88 which conditions (Alvarez-Buylla *et al.*, 1996; Clark *et al.*, 1999; Hampe *et al.*, 2008; Jones
89 & Hubbel, 2006).

90 Massive seed and seedling mortality during recruitment has been shown to strongly
91 affect tree population structure. At early life-history stages, high density-dependent mortality
92 due to seedling competition, predation and/or sensitivity to pathogens (Howe & Smallwood,
93 1982; Janzen, 1970; Nathan & Casagrandi, 2004) can result in higher average distances
94 between mothers and successfully established offspring than those expected from seed
95 dispersal alone. The existence of Janzen-Connell effects in forest trees is supported by
96 different studies across different forest systems including temperate deciduous forest (Hille
97 Ris Lambers & Clark, 2003). Janzen-Connell effects are expected to result in a decrease in
98 structure and relatedness from the initial seed rain to recruited seedlings (Trapnell *et al.*,
99 2008). Alternatively, an increase in genetic structure and relatedness from the initial seed rain
100 to recruited seedlings can be expected when mortality is driven by microsite heterogeneity
101 and genotype–microsite interactions, especially in spatially variable environments (Sagnard *et*
102 *al.* 2010). Moreover, because the genetic load of trees is high (Petit & Hampe, 2006), the
103 purging of inbred individuals may contribute to seedling mortality (Ferriol *et al.*, 2011).
104 However, such processes are notoriously difficult to demonstrate in natural populations, and
105 require monitoring the demo-genetics of naturally established seedlings through time (Kalisz
106 *et al.*, 2001). Most frequently, genetic structure studies taking a life-stage approach compare
107 very distant cohorts, typically seeds, seedlings and adults (Alvarez-Buylla *et al.*, 1996; Jones
108 & Hubbel, 2006).

109 Here we used a demo-genetic approach to investigate the consequences of the spatio-
110 temporal patterns of regeneration on the demo-genetic structure of European beech (*Fagus*
111 *sylvatica* L.). This wind-pollinated species is both gravity- and animal-dispersed (Jensen,
112 1985) and produces beech nuts in irregular mast years. Previous demographic estimates of

113 seed dispersal obtained through seed trapping and inverse modeling methods showed
114 restricted dispersal abilities, with a median distance of seed dispersal of ~6.50 m (Sagnard *et*
115 *al.*, 2007). The European beech is described as a shade tolerant species able to survive under
116 1-2% of full above-canopy light, but showing optimal growth potential at 30-40% of above-
117 canopy light (Kunstler *et al.*, 2007).

118 The originality of this study lies in the fact that we investigated the role of various
119 genetic and demographic processes (inter-annual variation in seed production, dispersal, and
120 mortality) on the distribution pattern of genetic diversity in three regeneration plots with
121 different canopy closure in a beech stand where seedling recruitment was monitored from
122 2004 to 2006, and extrapolated back to 1993 by estimating seedling age. Because it is difficult
123 to disentangle the different ecological factors shaping genetic structure, we also directly
124 investigated the seed and pollen dispersal processes using parentage/paternity analyses. This
125 approach allowed us to better test the consequences of a process (dispersal) and its variation
126 in space (with stand density) or time (among year).

127 Based on this data set, the following specific questions were addressed: (1) how do
128 patterns of seedling density and survival correlate in space and time? (2) How spatially
129 restricted are contemporary pollen and seed dispersal? (3) Which spatial or temporal
130 processes are the major factors shaping seedling genetic structure? (4) Does mortality affect
131 the genetic structure of the seedlings?

132

133

134 **Material and Methods**

135 **Study species and sampling design:** The European beech, *Fagus sylvatica* L. (Fagaceae), is
136 a monoecious diploid ($2n = 24$) late-successional forest tree. It is a highly outcrossing wind-
137 pollinated species (selfing <10% ; Merzeau *et al.*, 1994). Reproductive trees (typically older
138 than 60-80 years in dense stands) produce beech nuts in irregular mast years, with an inter-
139 mast interval of at least 2 years (Nilsson & Wastljung, 1987; Teissier du Cros, 1981). Seeds
140 are primarily dispersed by gravity, and secondarily dispersed by rodents (*Apodemus*
141 *flavicollis*, *Clethrionomys glareolus*) and birds (*Glandus glandularius*) that scatter hoard them
142 (Jensen, 1985).

143 The study site is a mixed beech-oak stand located within a large forest dominated by beech
144 (Haye forest, North-Eastern France, Longitude: 06° 06' 36" E; Latitude: 48° 38' 23"N) .

145 Within the site (~7.8 ha), all 342 adult beech trees were mapped using a Rangemaster 900
146 Scan telemeter and a compass, with a precision of more than 5 meters (Figure 1). All were
147 sampled for genetic analyses. Adult density varied across the site from <30 stem/ha to >90
148 stem/ha as a result of the 1999 Lothar storm.

149 [Figure 1 around here]

150 Young seedlings were sub-sampled in 3 plots (A: 9 m², B: 5 m² and C: 8 m² in area)
151 corresponding to a gradient in the levels of photosynthetically active radiation (PAR)
152 measured in micromol/m²/s with a SunScan ® system. Within each plot 5-6 measurements
153 were averaged and converted into percentage of transmitted PAR ($\frac{\text{Transmitted PAR}}{\text{Incident PAR}}$) by
154 reference to the average measure of 157.9 micromol/m²/s obtained in the open (=incident
155 PAR). Light availability under the canopy strongly decreased from 55% of transmitted light
156 (plot A) to 31% of incident light (plot B) and finally to 6% of incident light (plot C), in
157 relation to the variation in adult tree density in a 50 meter radius around the plots (A: 56 trees
158 ha⁻¹; B: 72 trees ha⁻¹; C: 92 trees ha⁻¹).

159 In total, 462 young seedlings were exhaustively mapped, among which 254 were sampled
160 for genetic analyses (Table 1). Among the 462 seedlings, 371 were found in the first year of
161 the survey (2004) while 91 new seedlings germinated in spring 2005 and were mapped in
162 summer 2005. Survival of all seedlings was recorded from 2004 to 2006. Germination year
163 was estimated retrospectively by counting node scars and ranged from 1993 to 2005. Finally,
164 372 seeds were collected in the autumn of 2004 in the crown of 29 fruiting trees (5-16
165 seeds/tree, mean = 12.8, S.D. = 2.98).

166 **Genotyping:** DNA was isolated from buds (adult trees and seedlings) and embryos (seeds)
167 using the Qiagen DNeasy Plant kit. Individuals were genotyped using 4 nuclear microsatellite
168 markers (FS1-03, FS1-25, FS3-04, FS4-46) developed for *Fagus sylvatica* (Pastorelli *et al.*,
169 2003) and 2 nuclear microsatellite markers (FCM5 and SFC-0161) developed for *Fagus*
170 *crenata* (Asuka *et al.*, 2004; Tanaka *et al.*, 1999) (Table 1), following PCR conditions given
171 by the authors. Adult and seed PCR products were separated using an automated 96-capillary
172 MegaBACE™ 1000 sequencer (GE Healthcare). Genotypes were sized using the internal size
173 standards ET400 and the MegaBACE™ Fragment Profiler ver. 1.2 software (GE Healthcare).
174 Seedling PCR products were separated using a LICOR automated gel-sequencer (some adults
175 were also genotyped on LICOR sequencer for homogenous sizing).

176 **Null alleles and quality of the marker set:** Null allele frequencies (NAF) were first
177 estimated by direct counting in maternal progeny arrays (for details see Oddou-Muratorio *et*
178 *al.*, 2009). Additionally, we estimated NAF in adult and seedling cohorts using the maximum-
179 likelihood method implemented in ML-NullFreq software and accounting for genotyping
180 error (Kalinowski & Taper, 2006).

181 The non-negligible prevalence of null alleles in European beech was confirmed, with four loci
182 out of six affected (Table A1). NAF were > 10% in 2 locus/cohort combinations (adults at
183 locus FS1-25 and FS4-46) and >5% in 9 locus/cohort combinations out of 18. Adults tended

184 to have higher NAF than seedlings at loci FS1-25 and FS4-46, but the opposite trend was
185 observed at locus FS1-03. The difference in NAF between adult and seedling cohorts was
186 highest (~0.13) at locus FS 4-46, despite low genetic differentiation among adult and seedling
187 cohorts (F_{ST} averaged across loci = 0.54%).

188 When high frequencies of null alleles were found ($p > 0.05$), we evaluated their impact on
189 consanguinity and differentiation estimates using the ENA method proposed by Chapuis &
190 Estoup (2007). This method consists (1) in correcting the original data set by statistical
191 adjustment of genotype frequencies based on estimated null allele frequencies and then (2) in
192 re-estimating F_{IS} or F_{ST} based on visible alleles only. It is designed to yield unbiased F_{ST}
193 values, but not to yield unbiased F_{IS} -values, which are likely to be underestimated.

194 Significantly positive F_{IS} values were estimated at the four loci affected by null alleles (Table
195 A1). However after correcting the data set for null alleles using the ENA method proposed by
196 Chapuis & Estoup (2007), no significant heterozygote deficiency could be detected in either
197 adults or seedlings (results not shown). By contrast, F_{ST} values were of the same order of
198 magnitude in the raw and corrected datasets, showing that genetic differentiation between
199 groups of seedlings was not affected by null alleles.

200 According to these results, we later considered in our analyses that: (1) null alleles occurred at
201 non-negligible frequencies in the data set, frequencies that were not estimated with precision
202 as depicted by their variation across cohorts at the same locus; (2) some of the observed
203 positive F_{IS} -values were probably partly affected by null alleles, but also by the mating
204 system; F_{IS} -values were thus assumed to be relevant in a comparative context; (3) genetic
205 differentiation as estimated by F_{ST} was not affected by null alleles.

206 **Genetic diversity:** Genotyping problems occurred at locus FCM5, and particularly for
207 seedlings (81% missing data). Given that missing data affect the estimation of genetic
208 differentiation, all analyses were run with only 5 loci, excluding FCM5.

209 Expected heterozygosity (H_e), allelic richness (A) and heterozygote deficiency (as measured
 210 by F_{IS}) were estimated using the Fstat software (Goudet, 2000). Allelic richness was rarefied
 211 to a minimum sample size of 10 individuals. Significance of F_{IS} - values was assessed at the
 212 5% confidence level after Bonferonni correction.

213 **Spatio-temporal genetic structure:** To investigate the spatial and temporal components of
 214 genetic structure in the seedlings, we used a hierarchical AMOVA design (Excoffier *et al.*,
 215 1992). We analyzed the respective effects of spatial (plot) versus temporal (year of
 216 germination) processes, by testing successively two two-level nested models (year within plot
 217 and plot within year) for the genotypic frequency vector G_{ijk} of individual k germinated in
 218 year j and located in plot i :

$$219 G_{ijk} = \mu + f_i + t_{j(i)} + w_{ijk} \quad (1)$$

$$220 \text{ and } G_{ijk} = \mu + f'_j + s_{i(j)} + w_{ijk} \quad (2)$$

221 where f_i is the average effect of plot i , f'_j the average effect of year j , $t_{j(i)}$ is the average effect
 222 of year j nested within plot i , $s_{i(j)}$ is the average effect of plot i nested within year j and w_{ijk} is
 223 the replication error associated with the k_{th} individual from the i_{th} plot germinated in the j_{th}
 224 year. For microsatellites assumed to follow a stepwise mutation model, Slatkin (1995)
 225 recommends measuring variation in allelic frequencies and genetic differentiation with alleles
 226 ordered according to their size rather than with unordered alleles (identity in state).

227 Accordingly, we used both F_{st} and R_{st} as estimators of genetic differentiation. Significance of
 228 F-statistics was assessed by means of 5000 permutations. All computations were done using
 229 the Arlequin software (Schneider *et al.*, 2000). Detailed AMOVA design is described in
 230 online appendix A1.

231 **Fine-scale spatial genetic structure (SGS) within plots / cohorts:** The classical analysis of
 232 SGS consists in plotting the variation in average genetic relatedness among individuals
 233 against distance (or logarithm of distance in a two-dimensional space). Under isolation by

234 distance, this relationship is expected to be linear in a part of the distance range, and shows a
235 decay rate proportional to $1/d_e \sigma_e^2$, d_e being the effective population density, and σ_e^2 the axial
236 variance in gene dispersal distance (Rousset, 2000). Here, our sampling design with only
237 three seedling plots was not conceived to investigate the variation of genetic relatedness
238 among pairs of seedlings over a large range of distances. Instead, we focused on the
239 “between-generation” component of SGS, by computing coefficients of genetic relatedness
240 (F_{ij}) among all pairs of individuals that involved one seedling ($i = 1$ to N_S , where N_S = total
241 number of seedlings) and one adult ($j = 1$ to N_A where N_A = total number of adults). Computed
242 in this way, F_{ij} coefficients reflect the parent-offspring component of genetic structure, with
243 expected values equal to 0.25 when i and j are related (parent-offspring) or 0 when i and j are
244 unrelated.

245 As proposed by Hampe *et al.* (2010), we analyzed between-generation SGS for different
246 group of seedlings, i.e. seedlings grouped by spatial plots (A, B and C) or by germination
247 year. This allowed us to investigate the variation in SGS with spatial adult tree density (which
248 decreased from plot C to A) and temporal adult tree density (high in the masting year 2002
249 and low in the other years).

250 All kinship analyses were performed using Spagedi 1.2 (Hardy, Vekemans, 2002), which
251 makes it possible to specify adult-seedling pairs to be compared. To measure genetic
252 relatedness, we used the kinship coefficient (F_{ij}) of Loiselle *et al.* (1995). F_{ij} -values were
253 estimated using 5 loci (excluding FCM5), and assumed not to be affected by null alleles,
254 similarly to F_{ST} (Rousset 2000). The allele frequencies from the whole population (i.e.
255 grouping adults and seedlings) were used as a reference sample. To visualize SGS, F_{ij} -values
256 were averaged over a set of distance classes (d) (with a minimum number of 80 pairs of
257 individuals per distance class) and plotted against distance. To test SGS, F_{ij} values were
258 regressed on $\ln(d_{ij})$, where d_{ij} is the spatial distance between individuals i and j , to provide the

259 regression slope b_{log} . Then, the spatial positions of all individuals were permuted 5,000 times
260 in order to get the frequency distribution of b_{log} under the null hypothesis that F_{ij} and d_{ij} were
261 uncorrelated. Approximate standard errors for the multi-locus estimates of F_{ij} within each
262 distance class were obtained through a jackknife procedure that consisted of deleting each
263 locus one at a time. This assumes that the different loci provide independent replicates of the
264 genetic structure process.

265 **Impact of mortality on genetic structure:** The genetic differentiation between dead and
266 alive seedlings in year 2006 was first investigated using the hierarchical AMOVA design
267 described above by equations (1) and (2), replacing the average effect of year by that of status
268 (dead/alive) in f_j , $t_{j(i)}$ and $s_{i(j)}$ and considering that w_{ijk} was the replication error associated
269 with the k_{th} individual from the i_{th} plot in state j .

270 Then, fine-scale patterns of between-generation SGS were investigated as detailed above to
271 test whether gene dispersal patterns from adults to seedlings differed among dead and alive
272 seedlings.

273 **Estimation of the seed and pollen dispersal kernel based on established seedlings:** the
274 spatially explicit mating model (SEMM) developed by Burczyk *et al.* (2006) and Oddou-
275 Muratorio & Klein (2008) was used to estimate the shape and range of seed and pollen
276 dispersal kernels from genotypes and positions of established seedlings and their potential
277 parents. The model considers that each seedling i can be mothered either (1) by a mother tree
278 located outside the study site due to seed immigration (with probability m_s) or (2) by a local
279 mother tree located within the study site (with probability $(1 - m_s)$). In the latter case, the
280 model considers that offspring i may be the result either of self-pollination (with probability
281 s), pollen flow from outside the neighbourhood (with probability m_p), or pollen from a
282 sampled male (with probability $1 - s - m_p$). The genotypes of seedlings and candidate
283 mothers/fathers are used to define the compatible offspring-parent triplet, and to compute

284 transition probabilities in a fractional parentage analyses design. The contribution of a
 285 sampled and genetically compatible mother tree j to the seedling rain at the location of
 286 seedling i is modelled as the product of the probability of a seed to disperse from j to i (the
 287 seed dispersal kernel), and of mother intrinsic fecundity. Similarly, the contribution of each
 288 sampled father tree k to the pollen cloud above mother tree j at is modelled as the product of
 289 the probability of a pollen grain to disperse from k to j (the seed dispersal kernel), and of the
 290 father intrinsic fecundity. Here, we used the exponential power function to model the seed and
 291 pollen dispersal kernels:

$$p(a,b;d) = \frac{b}{2\pi a^2 \Gamma(2/b)} \exp\left(-\left(\frac{d}{a}\right)^b\right) \quad (3)$$

293 where d is the distance of interest (mother-seedling or father-mother) and Γ is the classically
 294 defined gamma function. The parameter b is the shape parameter affecting the tail of the
 295 dispersal function and a is a scale parameter homogeneous to distance. When $b=1$ this model
 296 simplify to an exponential; when $b<1$ the kernel is fat-tailed and when $b>1$ the kernel is thin-
 297 tailed. Through parameter b , this model thus allows to estimate whether long-distance
 298 dispersal events are respectively more or less important as compared to the exponential
 299 kernel.

300 The model allows for a simultaneous estimation of seed and pollen immigration levels (m_s
 301 and m_p), selfing rate (s) along seed and pollen dispersal kernels parameters (a_s, b_s, a_p, b_p), as
 302 detailed in Oddou-Muratorio & Klein (2008) and in online Appendix A2 (supplementary
 303 material).

304 SEMM requires as input the mapped locations of all sampled seedlings and all potential
 305 reproductive adult males and females within a local population, the multilocus genotypes of
 306 seedlings and adults, and allele frequencies of the same species in surrounding (background)
 307 populations. The genotypes and spatial positions of all the 342 adult trees found within the

308 site and of the 221 seedlings were used for these analyses (33 seedlings were eliminated
309 because they could not be genotyped at more than 2 loci and/or because of missing spatial
310 position data). Note that locus FCM5 was included in these analyses, as missing data do not
311 bias this estimation procedure. Background allele frequencies were assumed to be similar to
312 that of the local population.

313 **Estimation of the pollen dispersal kernel based on maternal progenies:** SEMM was also
314 used to estimate the shape and the range of the pollen dispersal kernel from genotypes and
315 positions of maternal trees, progeny arrays, and potential fathers (online Appendix A3
316 supplementary material). The model is simpler than the seedling model above (Oddou-
317 Muratorio *et al.*, 2005) and considers that a given seed i collected on mother-tree j may be the
318 result of self-pollination (with probability s), pollen flow from outside the neighbourhood N
319 (with probability m_p), or pollen from a sampled male (with probability $1-s-m_p$).

320 The genotypes and spatial positions of all 342 adult trees as well as the genotypes of the 372
321 seeds collected on 29 fruiting trees were used for these analyses.

322 **Results**

323 **Demographic dynamics in seedlings plots**

324 Effective recruitment patterns were highly variable among years, with on average 41.3% of
325 seedlings germinated in 2002, and several years without significant recruitment (Table 1). For
326 seedlings germinated before 2004, these variations can result either from low seed production
327 and/or germination, or from high seedling mortality between the germination year and 2004.

328 These effective recruitment patterns are consistent with the expected effect of the Lothar
329 storm in 1999, as most of the seedlings observed alive between 2004 and 2005 germinated
330 after the stand canopy was significantly opened by the storm. Overall, seedling density
331 observed in 2004 was lower in plot A (12.7 m⁻²) than in plot B (31 m⁻²) or plot C (24 m⁻²)
332 (Figure 2).

333 [Table 1 around here]

334 [Figure 2 around here]

335 The average mortality rate was 22.7% between 2004 and 2005, and 8.2% between 2005 and
336 2006. From year 2004 to 2005, there was a trend of higher mortality under low light
337 conditions (high canopy closure), with mortality rate increasing from 15% in plot A to 21 %
338 in plot B and finally to 27% in plot C (χ^2 test: p-value = 0.10). By contrast, from year 2005 to
339 year 2006, the mortality rate was lower in plot C (0.3%) than in plots A (9%) and B (12%) (χ^2
340 test: p-value = 0.01). Overall, mortality from 2004 to 2006 (average mortality rate = 29%)
341 tended to reduce variation in seedling density among plots (as measured by the coefficient of
342 variation (CV) of seedling density, $CV_{2004}= 0.36$ versus $CV_{2006}=0.31$), with final seedling
343 density in 2006 ranging from 9.8 m⁻² (plot A) to 16.9 m⁻² on plot C and to 21.2 m⁻² on plot B.

344 **Genetic diversity within seedling and adult cohorts**

345 Levels of diversity did not differ among adult and seedling cohorts, among seedling plots, or
346 among dead and alive seedlings (Table 2). Nei's genetic diversity was high both in seedlings
347 and adults ($H_e = 0.71$ and 0.72 , respectively), and allelic richness was also comparable in
348 seedlings and adults ($A = 5.71$ and 5.94 , respectively). By contrast, the within-individual
349 structure of genetic diversity differed between adult and seedling cohorts, with a higher
350 heterozygote deficiency in the adults ($F_{IS} = 0.131^{**}$) than in the seedlings ($F_{IS} = 0.069^{**}$).
351 However, when the three loci affected by null alleles (FS1-25, FS4-46, FS1-03) were
352 removed, F_{IS} -values did not differ significantly from 0.

353 [Table 2 around here]

354 **Spatio-temporal genetic structure of seedlings**

355 Spatial (among plots) versus temporal (among year classes) components of seedling genetic
356 structure were first investigated using two different two-level AMOVA models: (Model1)
357 years nested within plots (Table 3) and (Model2) plots nested within years. The main effect

358 for years was not significant (Model2 see Table A2), while the main effect for plots was
359 significant (Model1). Detailed analyses of Model1 (Table 3) showed that the “among-plot”
360 component of genetic variation was not negligible ($F_{RT} = 2.6\%$ of the total variation)
361 considering the small spatial scale investigated (average distance among plots ~100 m). By
362 contrast, year-to-year variation in a given plot was not significant ($F_{SR} = 0.5\%$) but
363 contributed to overall differentiation ($F_{ST} = 3.1\%$). Differentiation estimates using R -statistics
364 were similar to these values.

365 [Table 3 around here]

366 To investigate whether the lack of a significant among year/plot effect could be due to an
367 over-representation of seedlings germinated in year 2002 in the data set (68% of the 234
368 seedlings), we rarefied the sample to balance sample size within year (so that year 2002
369 represented 47% of 139 seedlings) and ran the AMOVA analyses again (Table 3). Results
370 were consistent with those obtained with the complete data set, with a significant main spatial
371 effect ($F_{RT} = 3.2\%$).

372 Pairwise F_{ST} among seedling plots (Table A3) showed that only plot C was significantly
373 differentiated from plots B and A ($F_{ST} = 1.3\%$ in both cases). All seedling plot genetic
374 frequencies significantly differed from that of the adult populations (F_{ST} ranging from 0.9%
375 for B plot to 2.3% for A plot).

376 **Fine-scale spatial genetic structure (SGS) within plots and within cohorts**

377 The fine-scale variations of ‘between-generation’ SGS were analyzed by plotting genetic
378 relatedness among all seedling-adult pairs against distance (Fig 3A). Then, similar plots were
379 obtained by computing genetic relatedness among seedling-adult pairs for seedlings belonging
380 to the same plot (Fig 3B) or the same year class (Fig 3C). Patterns of SGS were always
381 markedly significant, with regression slopes of F_{ij} on $\log(\text{distance})$ different from zero in all
382 cases ($p < 0.001$). Overall, the ‘between-generation’ SGS was strong and decreased rapidly

383 with distance, with a marked peak of SGS in the first distance interval ($F_{10m} = 0.039$)
384 followed by a notable decrease between 20 and 40 m and no more significant trend after 40 m
385 (Fig 3A). The adult SGS tended to be even stronger ($F_{10m} = 0.065$; Fig 3A).

386 [Figure 3 around here]

387 ‘Between-generation’ SGS varied among plots (Fig 3B). However, trends were not significant
388 due to the high standard errors of F_{ij} and b estimates (Table A4). There was also no clear trend
389 of increasing ‘between-generation’ SGS during years of low recruitment, as could have been
390 expected with a reduced contribution of adults to reproduction in non-masting years. Year
391 2001 and 2004 (low recruitment) showed higher and lower ‘between-generation’ SGS,
392 respectively, whereas year 2002 showed intermediate SGS (Fig 3C).

393 Direct gene flow estimates

394 Maximum-likelihood estimates of seed and pollen dispersal as well as mating system
395 parameters were obtained using SEMM (Table 4). Both for pollen and seed dispersal, the
396 exponential kernel (*i.e.* fixing $b_p = 1$ and $b_s = 1$ and estimating solely a_p and a_s) provided a
397 better fit than the Gaussian kernel (*i.e.* fixing $b_p = 2$ and $b_s = 2$). The exponential power kernel
398 with joint estimates of b and a parameters did not improve the fit (results not shown). We
399 estimated a larger mean distance for pollen dispersal ($\hat{\delta}_p \approx 57$ m; CI = 30.6 – 123.4 m) than for
400 seed dispersal ($\hat{\delta}_s \approx 11$ m; CI = 9.4 -12.9 m). Also, the pollen immigration rate ($\hat{m}_p = 71.6\%$;
401 CI = 60.2 – 85%) was significantly higher than the seed immigration rate ($\hat{m}_s = 20.5\%$; CI =
402 13.5 -27.1%). The selfing rate was significantly positive ($s = 3.5\%$).

403 [Table 4 around here]

404 Independent estimates of pollen dispersal and selfing rate based on maternal progenies fell
405 within the same range of that of estimates based on established seedlings. The best fit for

406 pollen dispersal was also obtained with the exponential kernel (fixing b_p) with an estimated
407 $\hat{\delta}_p$ value ≈ 44 m. The selfing rate was significantly positive ($s = 2.1\%$).

408 **Impact of mortality on genetic structure**

409 Because there was strong among-plot genetic structure, the relationship between survival and
410 genetic differentiation was tested using a two-level AMOVA design with status (alive/dead)
411 nested within plot. Neither the main effect nor the interaction effect of status was significant
412 (result not shown), showing that dead and alive seedling were not genetically differentiated
413 within plot.

414 The fine-scale variations of 'between-generation' SGS were analyzed by plotting genetic
415 relatedness among seedling-adult pairs against distance for dead and alive seedlings (Fig 3D).

416 The 'between-generation' SGS of dead and alive seedlings was very similar, with a marked
417 peak in the first distance interval (>10 m), followed by a notable decrease between 20 and 40
418 m.

419

420 ***Discussion***

421 Our results highlight some major spatial and temporal characteristics in the development of a
422 demographic and genetic structure within beech populations. These results were based on a
423 more than 10 year survey, with seedling establishment monitored from 2004 to 2006, and
424 extrapolated back to 1993 by estimating seedling age. This 10 year time scale is relevant
425 considering that in most forests across Europe, the management strategy of beech is high
426 forest (even-aged stands resulting from natural regeneration). In this type of silviculture,
427 stands are regenerated over a period spanning 10 to 20 years. At the beginning of the
428 regeneration phase, adult trees are selectively logged to leave approx. 100- 150 mature adult
429 trees/ha. Because of masting and strong seedling vigor, all beech seedlings that will
430 effectively contribute to the new reproductive stands are often recruited in less than 20 years.

431 **Demographic and genetic structure show opposite spatio-temporal effects**

432 Our results first indicate strong temporal and low spatial heterogeneity in recruited seedling
433 density, contrasting with the strong spatial and low temporal heterogeneity in their genetic
434 structure. Beginning with seedling density, we observed a strong temporal heterogeneity on
435 early-recruitment patterns in European beech, with 41.3% of seedlings germinated in 2002,
436 and several years out of the 13 under study (from 1993 to 2005) without significant
437 recruitment. These variations can result from low seed production, low germination, high
438 mortality, or a combination of these factors. Many sites favorable for seedling establishment
439 were opened by the 1999 Lothar storm, explaining the lack of recruitment before 1999.
440 However, recruitment patterns between years 1999 and 2004 are consistent with a massive
441 seedling germination event in 2002 and with the assumption that seed production is a limiting
442 factor for recruitment in beech (Piovesan & Adams, 2001).
443 By contrast, spatial heterogeneity in seedling density was weak, with a trend of lower initial
444 density in the plot with open canopy (PAR= 55%, 12.7 seedlings.m² in 2004), compared to

445 plots with intermediate canopy closure (PAR= 31%, 31 seedlings.m⁻²) or high canopy closure
446 (PAR=6%, 24 seedlings.m⁻²). However, our experimental design with only three plots was not
447 conceived to address the impact of canopy closure on initial seedling density or to separate
448 this effect from that of seed-tree density. Still, our results show that despite variable seed-tree
449 density across the plot (from <30 to >90 trees per ha⁻¹), there was a high density of seedlings
450 even under unusually high canopy openness (>9,700 ha⁻¹ in 2006, which is several times
451 higher than recommended for afforestation rates). This is consistent with other studies, which
452 have found that seedling germination is almost independent of light availability (Szwagrzyk *et*
453 *al.*, 2001), contrary to subsequent growth and long-term survival (Kunstler *et al.*, 2005;
454 Szwagrzyk *et al.*, 2001).

455 In contrast to their density, the genetic structure of recruited seedlings was significantly
456 shaped by spatial processes and poorly affected by temporal heterogeneity of the seed rain.
457 Here, the stand-level spatio-temporal genetic structure was investigated by testing
458 successively the main and nested effects of spatial location and year of germination on genetic
459 differentiation among seedlings as measured by F_{ST} . The main effect of spatial location was
460 strongly significant, and translated into a significant level of differentiation between plots of
461 $F_{ST} = 2.6\%$. By contrast, genetic differentiation among temporal cohorts within plots was not
462 significant. Moreover, genetic frequencies significantly differed between plot C and the two
463 other plots ($F_{ST\ C-A} = F_{ST\ C-B} = 1.3\%$) and between each seedling plot and adults ($0.9 < F_{ST\ Adult-}$
464 $seedling < 2.3\%$; Table A2). These F_{ST} -values may look weak, but by comparison, genetic
465 differentiation at allozyme loci measured over 389 populations across Europe were not larger
466 than 5.9% (Comps *et al.*, 2001). Using SSR markers to measure genetic differentiation among
467 10 populations across Europe, Buiteveld *et al.* (2007) reported pairwise F_{ST} -values ranging
468 between 0.8% and 5.3%, with an overall F_{ST} of 5.3%.

469 **Genetic structure revealed low levels of genetic drift despite restricted dispersal**

470 The major role of spatial versus temporal processes in shaping plant genetic structure has
471 been acknowledged in previous studies, usually by comparing genetic structure across
472 different life-stages (Alvarez-Buylla *et al.*, 1996; Chung *et al.*, 2003; Jones & Hubbel, 2006).
473 In their pioneer demo-genetic study in the tropical tree *Cecropia obtusifolia*, Alvarez-Buylla
474 *et al.* (1996) showed that patchy recruitment in gaps markedly affect the genetic composition
475 of the seed rain, with higher differentiation among gaps than among-life-stages. More
476 recently, Jacquemyn *et al.* (2009) used multi-stage spatial genetic structure analyses
477 combined with parentage analyses in the perennial *Orchis mascula* to show that patterns of
478 SGS were mostly shaped by pollen and seed-mediated gene dispersal, and were consistent
479 across life-stage.

480 By contrast however, we focussed here on a more fine-scale temporal structure. We
481 investigated the genetic structure among seedlings recruited in consecutive years, and thus
482 belonging to a single life-stage from the point of view of most previous studies. Because
483 beech trees produce nuts in irregular mast years, we expected significant genetic
484 differentiation among year-classes due to inter-annual variance in reproductive success. Such
485 fine-scale temporal differentiation has been reported in marine perennial organisms (e.g.
486 (Planes & Lenfant, 2002)). In our case, the across-year genetic homogeneity of the seed rain
487 suggest low levels of genetic drift, i.e. high effective population size and/or relatively even
488 contributions of all adult trees to reproduction either as male or female, even in non-masting
489 years. Fine-scale patterns of between-generation SGS confirmed (1) the across-year genetic
490 homogeneity of the seed rain (Fig 3C) and (2) the major role of gene dispersal and effective
491 population size on spatial patterns of allelic frequencies.

492 **Seed and pollen dispersal direct estimates**

493 Contemporary estimates of seed and pollen dispersal based on parentage/paternity analyses
494 shed light on dispersal processes and their ecological determinants in European beech.

495 Considering first seed dispersal patterns as estimated from established seedlings, our results
496 reflected preferential short distance dispersal as depicted by the low average distance of seed
497 dispersal ($\hat{\delta}_s \approx 11$ m). However, events of long distance dispersal appeared not negligible and
498 may contribute to high effective population size in beech, as depicted by the seed immigration
499 rate ($\bar{m}_s = 20.5\%$) and the exponential-shaped dispersal kernel. These estimates convert into a
500 median dispersal distance of 7.6 m and are consistent with previous demographic estimates of
501 seed dispersal in beech (median dispersal distance of 6.49 m in Sagnard *et al.*, 2007), or with
502 recent genetic estimates in beech species across different sites (Oddou-Muratorio *et al.* 2010).
503 They are also consistent with life history traits of the beech dispersers. The rodents involved
504 in secondary dispersal of beech seeds have been shown to remove seeds a few meters away
505 from the source tree (4.1 m on average, Jensen, 1985), whereas frequent 1 km- dispersal
506 events have been reported for jays (Nilsson, 1985).

507 Our results suggest greater dispersal abilities for pollen than for seeds, with both a higher
508 mean dispersal distance ($\delta_p = 43.7 - 57.2$ m versus $\delta_s = 10.9$ m), and a higher immigration rate
509 ($m_p = 71.6\%$ to 77.9% versus $m_s = 20.5\%$). The results based on established seedling or open-
510 pollinated progeny were highly consistent (Table 4). This supports the accuracy of the
511 SEMM. Moreover, it shows that early selective processes acting between seed release and
512 seedling establishment may not be driven by the genetic origin of the pollen grain (no
513 outbreeding or inbreeding depression).

514 From direct estimates of seed and pollen dispersal, we can estimate real-time, total gene flow
515 estimates (σ_{rt}), as detailed in Oddou-Muratorio & Klein (2008). In a two dimensional space,
516 for hermaphrodite, outcrossing species: $\sigma_{rt}^2 = \frac{1}{2}\sigma_{p-rt}^2 + \sigma_{s-rt}^2$, where σ_{p-rt}^2 and σ_{s-rt}^2 are the
517 respective second moments of the pollen and seed dispersal kernels. In our case, $\sigma_{rt} = 51.35$ m
518 (CI = 29-108m); this is consistent with the shape of the between-generation auto-correlograms
519 (Figure 3) which show significant SGS up to ~50m.

520 We reasonably assumed that differentiation (F_{ST}) and SGS estimates were not biased by the
521 frequencies of null alleles estimated in *F. sylvatica* (between 0 and 14% depending on cohort
522 and method of estimation, with all but two estimates <10%; see Table A1). However,
523 attention should be put in evaluating the potential impact of null alleles on direct estimates of
524 seed and pollen dispersal. Somehow reassuringly, Dakin & Avise (2004) showed using
525 simulations that the range of null allele frequency observed in this study (NAF<10%) equates
526 to a less than 5% risk of falsely excluding an actual parent of a heterozygous offspring in
527 parentage/paternity analyses. Additionally, we also estimated NAF using our mating model
528 parameters, using a modified version of the SEMM and adult and seedlings genetic and
529 spatial data as input (Chybicki & Burczyk, 2010). Interestingly, “direct” NAF estimates were
530 lower (2% on average, see Table A1) than those obtained by traditional methods (between 4.3
531 and 6.7%). The reason for this discrepancy may be that “direct” NAF estimates account for
532 the SGS present in the population. By contrast, in the case of significant SGS and preferential
533 local mating, biparental inbreeding can result in a deficit of heterozygotes (similar to a
534 Wahlund effect) that traditional methods could misinterpret as a signature of null alleles

535 **Evolutionary and ecological drivers of mortality**

536 The overall mortality rate over the two years of the survey (from 2004 to 2006) was low
537 (28.8% for all seedlings, and 20% for those germinated in 2002). In a 10-year mortality
538 survey, Szwagrzyk *et al.* (2001) reported mortality rates close to 100% after 4 years (but
539 under lower light availability, with PAR ranging from <3% to 15%). During the year of high
540 mortality (2004-2005, mortality rate 22.3%), low light availability tended to induce higher
541 mortality, in agreement with previous results in European beech (Szwagrzyk *et al.*, 2001).
542 To investigate evolutionary drivers of mortality, we first estimated genetic differentiation
543 between dead and alive seedlings (as measured by hierarchical *F-statistics* within plots). We
544 did not observe any significant differences between the two groups. Moreover, levels of

545 heterozygote deficiency were consistent among groups ($F_{IS} = 0.073^*$ in alive seedlings versus
546 0.058^{NS} in dead seedlings, Table 2). Although the level of inbreeding estimated by F_{IS} may be
547 affected by null alleles, F_{IS} -values can still be used to compare dead and alive seedlings
548 because of the absence of genetic differentiation between these groups. Thus, there was no
549 evidence that mortality is driven by the purging of selfed individuals in the studied beech
550 stand. Finally, between-generation patterns of spatial genetic structure (SGS) were also
551 consistent for dead and alive seedlings (Fig. 3), indicating that levels of genetic relatedness
552 within the stand did not significantly contribute to mortality. Overall, we did not find any
553 evidence that mortality is driven by inbreeding or lack of local adaptation.

554 Interestingly, patterns of inbreeding and relatedness coefficients were actually stronger for
555 adults compared to seedlings ($F_{IS} = 0.131$ in adults versus 0.069 in seedlings). This indicates
556 that massive recruitment during a single mast year does not reduce effective population size
557 as could have been expected.

558 **Perspectives**

559 This study highlights different magnitudes of temporal versus spatial effects on demographic
560 and genetic patterns of early recruitment. The high heterogeneity among year classes in
561 recruited seedling density revealed a major effect of mast seeding on demographic patterns of
562 recruitment. By contrast, the low genetic differentiation among seedlings recruited in different
563 years indicates balanced contribution of adult trees to reproduction within year. The
564 significant spatial genetic structure was consistent with the strong spatial limitation of both
565 seed and pollen dispersal detected using parentage analyses and neighborhood mating models.
566 As a direct consequence for forest managers, our results highlight that genetic diversity within
567 beech stands is mostly shaped by gene dispersal and adult tree density. Consequently, high
568 levels of genetic diversity can be maintained within stand even if young seedlings are
569 recruited in a reduced number of mast years.

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580 ***Author information box***

581 Sylvie Oddou-Muratorio is interested in estimating tree species abilities of pollen and seed
582 dispersal and their role in adaptation in the context of climate change. Etienne Klein is
583 interested in modelling and estimating long-distance dispersal through pollen and seed.
584 Giovanni G. Vendramin is a population geneticist involved in the application of molecular
585 markers to population genetics and genomics of forest tree species. Bruno Fady is a
586 researcher studying patterns of genetic diversity in forest trees in Mediterranean regions.

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715 **Figure legend**

716 Figure 1. Haye forest study plot, with exhaustive mapping of adult trees (plotted as circles: ○
717 /●) and 3 seedling plots (plotted as squares : ■). Among all the 342 adult trees, 29 were
718 chosen for collecting seed maternal progeny (black circles)

719 Figure 2. Patterns of mortality across year and plots. Density of alive seedlings were
720 monitored during 3 years (2004 in black, 2005 in grey, 2006 in white) on 3 plots (A, B, C)
721 differing for light availability under the canopy as measured by percentage of
722 photosynthetically active radiation (PAR). Hatched bars correspond to result based on
723 seedlings germinated in year 2002 (masting year), whereas filled bars correspond to all
724 seedlings cohorts.

725 Figure 3. Between-generation patterns of spatial genetic structure : A. for all seedling-adult
726 pairs (bold line), with the kinship values among all adult pairs plotted as a reference (broken
727 line); B. among the different spatial plots; C. among the different temporal cohorts; D. for
728 dead and live seedlings.

729

730 **Table 1.** Effective recruitment patterns of beech seedlings in the three studied regeneration
 731 plots. Total count of seedlings in year 2004 (NS) and the number of seedlings used for
 732 genetic analyses (NG) per plot and per year of germination.

Plot		Year of germination										Total
		1993	1995	1997	1999	2000	2001	2002	2003	2004	2005*	
Plot A	NS	0	0	0	0	4	0	71	18	13	9	115
	PAR=55% NG							57	15	10		82
Plot B	NS	1		1	1	12	1	63	7	28	41	155
	PAR=31% NG						1	56	5	19		81
Plot C	NS	0	1	2	3	51	9	57	23	5	41	192
	PAR=6% NG						7	46	17	2		72
Total	NS	1	1	3	4	67	10	191	48	46	91	462
	NG						8	159	37	31		235

733
 734 *Seedlings germinated in year 2005 were counted in summer 2005, while all the other
 735 were counted in summer 2004 .

736 **Table 2:** Stratified genetic diversity indexes averaged over 5 loci. N = number of
 737 genotyped seedlings; N_{NA} =sample size corrected for missing data, H_e = Nei's expected
 738 heterozygosity, N_a = No. of alleles, A = Allelic richness computed using rarefied sample
 739 of 10 individuals, F_{IS} = Fixation index.

Cohort	Group	N	N_{NA}	H_e	N_a	A	F_{IS}
Seedlings-	Live	66	55.00	0.68	8.60	5.1306	0.034
Plot A	Dead	16	15.00	0.75	6.80	6.0794	0.007
	<i>All</i>	82	70.00	0.70	9.80	5.3164	0.027
Seedling-	Live	62	60.25	0.69	9.20	5.2766	0.053
Plot B	Dead	19	18.00	0.68	6.00	4.9594	0.054
	<i>All</i>	81	78.25	0.68	9.40	5.18	0.052
Seedling-	Live	56	48.25	0.73	8.60	5.7908	0.086
Plot C	Dead	16	14.00	0.75	6.60	5.816	0.085
	<i>All</i>	72	62.25	0.73	9.20	5.7784	0.085
All seedlings	Live	184	163.50	0.71	12.40	5.7318	0.073*
	Dead	51	47.00	0.73	8.60	5.7064	0.058
	<i>All</i>	235	210.20	0.71	13.20	5.714	0.069*
Adult trees		342	327.80	0.72	15.80	5.9412	0.131*

740 * significant at 5% confidence level
 741

742

743 **Table 3.** Nested analysis of molecular variation for genetic variation among 3 beech seedlings
 744 plots, each sampled in 3 successive years. 234 seedlings were used for these analyses.
 745

234 individuals					135 individuals			
Source of variation	df *	Sum of squares	Est. Var.	% total variance	df *	Sum of squares	Est. Var.	% total variance
Among subplots	2	17.07	0.05	2.59	2	13.77	0.06	3.21
Among Years/subplots	7	14.03	0.01	0.50	7	11.78	0.01	0.44
Within years	352- 452	709.60	1.75	96.91	225 - 269	433.56	1.75	96.36
Total		740.69	1.80			459.11	1.82	

746
 747 * df = degree of freedom for the within year component. df varied among loci due to missing
 748 data.
 749
 750

751 **Table 4:** Direct estimates of selfing (s), pollen and seed migration rates (resp. m_p and m_s) and
 752 pollen and seed dispersal distance (δ_p and δ_s) through spatially explicit mating models
 753 (SEMM) applied either on seedlings (e.g. Oddou-Muratorio and Klein 2008) or seeds (e.g.
 754 Oddou-Muratorio et al 2005).

Parameter	Seedlings SEMM		Seeds SEMM	
	Estimate	Confidence	Estimate	Confidence
		Interval		Interval
s	3.5%	4.2-23.8%	2.1%	0.4-3.8%
m_p	71.6%	60.2-85.0 %	78.0%	72.1-83.0 %
δ_p (m)	57.25	30.6-123.4	43.5	25.4-61.7
m_s	20.5%	13.5-27.1 %	-	-
δ_s (m)	10.92	9.4-12.9	-	-

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