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# A Case–Control Study on the Origin of Atypical Scrapie in Sheep, France

**Alexandre Fediaevsky, Eric Morignat, Christian Ducrot, and Didier Calavas**

A matched case–control study (95 cases and 220 controls) was designed to study risk factors for atypical scrapie in sheep in France. We analyzed contacts with animals from other flocks, lambing and feeding practices, and exposure to toxic substances. Data on the *prnp* genotype were collected for some case and control animals and included in a complementary analysis. Sheep dairy farms had a higher risk for scrapie (odds ratio [OR] 15.1, 95% confidence interval [CI] 3.3–69.7). Lower risk was associated with organic farms (OR 0.15, 95% CI 0.02–1.26), feeding corn silage (OR 0.16, 95% CI 0.05–0.53), and feeding vitamin and mineral supplements (OR 0.6, 95% CI 0.32–1.14). Genetic effects were quantitatively important but only marginally changed estimates of other variables. We did not find any risk factor associated with an infectious origin of scrapie. Atypical scrapie could be a spontaneous disease influenced by genetic and metabolic factors.

Atypical scrapie is a transmissible spongiform enceph-alopathy (TSE) of small ruminants; it was recently defined by the European Food Safety Authority according to phenotypic features  $(1)$ . This disease was identified in 1998 (*2*), and little is known about its etiology and epidemiology (*3*), which contrasts with the etiology and epidemiology of classical scrapie (*4*). Diagnosis of atypical scrapie is impaired by discrepant clinical diagnostic results (*5*) of rapid diagnostic tests and variable accumulation of scrapie prion protein ( $PrP^{Sc}$ ) in the brainstem (3). In France, the average apparent prevalence of atypical scrapie detected by active surveillance with tests recommended by the Euro-

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pean Food Safety Authority (*6*,*7*) for the detection of this disease in brainstem samples was 6 cases/10,000 tested animals during 2002–2006. This prevalence was comparable elsewhere in Europe (*8*).

The origin of atypical scrapie is still unclear, and whether the disease has an infectious origin remains a major question. This disease has been transmitted experimentally to Tg-mice (*9*) and sheep (*10*), but histopathologic features of atypical scrapie have suggested similarities with human spontaneous TSE (Gerstmann-Sträussler-Scheinker syndrome)  $(2,11)$ . The few reports on  $>1$  case of atypical scrapie in the same flock provide insufficient information to draw conclusions on natural transmission of this disease (*3*). If atypical scrapie had an infectious origin, it could be influenced by risk factors associated with a pattern of infectious disease transmission as described for classical scrapie (*12*–*14*). In 2006, a case–control study of atypical scrapie in Norway did not detect such factors, but it showed that the removal of the placenta at lambing could have a protective effect (*15*). Feeding of vitamin and mineral supplements showed an adverse effect, which was interpreted as interaction of some minerals with cellular prion protein (*16*,*17*), rather than a feed contamination. Such an effect warrants confirmation.

Genetic factors should be considered when investigating risk factors for atypical scrapie because some mutations of the *prnp* gene, which codes for prion protein (PrP), modify the risk for this disease (*2*,*18*,*19*). Because all described genotypes of the *prnp* gene confer susceptibility to sheep, a purely genetic origin is unlikely but a confounding effect could occur.

Other possible origins for atypical scrapie could involve exposure to toxic substances, particularly pesticides, which were shown to be involved with other neurodegenerative diseases involving protein disorders such as Par-

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#### **Materials and Methods**

#### **Study Design**

The epidemiologic unit was the animal, and most of the data collected concerned its birth cohort, assuming that in each flock all animals born during the same birth campaign ( $C_0$ , defined from July 1 of year n – 1 to June 30 of year n) shared the same exposure. Cases and controls were matched by frequency matching on their birth cohort  $(C_0)$  so that their distributions were similar over the birth campaigns.

Cases were recruited among cases detected by the active surveillance program during January 2006–March 2007. The index case had to be a female that was born and reared all its life in the same flock, with a known  $C_0$ . A total of 137 cases met these criteria.

Two controls per case were selected. The control animal was an animal born in  $C_0$ , kept until birth campaign 2006 ( $C_{2006}$ ) in the same flock, and originated from a control flock randomly selected among the list of flocks from which >1 sheep had been tested in 2006 with a test recommended for the detection of atypical scrapie. All results from TSE rapid tests on sheep from control flocks had to be negative for atypical scrapie and classical scrapie. Each control had 9 replacement animals randomly selected from sheep flocks from the same county (French département).

Flocks of case and control animals were required to have no history of scrapie and  $>20$  ewes kept for reproduction. Males were not included in the study because they have a low incidence of atypical scrapie and because farming practices used with rams are different from those used with ewes.

#### **Data Collection**

Four persons interviewed farmers during the summer of 2007. The questionnaire, which was available in French on request, was divided in 5 parts: 1) 13 questions on structure and economic context of the flock,  $2$ ) 7 questions on purchase of sheep and contacts with other flocks, 3) 3 questions on lambing management, 4) 16 questions on feeding practices including list of feed, and 5) 8 questions on exposure to toxic products, including the list of products used. Questions related to structure of the flock were asked for  $C_0$ and  $C_{2006}$  to check if changes had occurred. Questions related to exposure during the first months of life were asked only for  $C_0$ , questions related to general feed exposure were asked for the period between  $C_0$  and the 2 subsequent reproduction campaigns ( $C_0-C_{0+2}$ ) and questions related to exposure to toxic products and mineral feeding were asked for  $C_0 - C_{2006}$ . For each flock, the number of animals tested with a recommended test for atypical scrapie during active surveillance programs during 2002–2006 was extracted from the Base Nationale des Encéphalopathies Spongiformes Transmissibles Animales.

*Prnp* genotypes at codons 136, 141, 154, and 171 were determined by Labogena (Jouy en Josas; France). For cases, material examined consisted of a sample of soft tissue (muscle or ear) or brainstem. For controls, the matching constraint was relaxed to enable interviewers to sample some hair from any ewe born during  $C_0 - C_{0+2}$ .

#### **Data Management**

Data were entered into a Microsoft (Redmond, WA, USA) Access 2000 database. All statistical analyses were performed by using R 2.6.1 for Windows (*22*).

Three types of toxic exposure were assessed: pesticides on crops, insecticides on premises, and antiparasitic treatments. For each category, active components of products reported were identified from databases (23–26). The Direction des Végétaux et de l'Environnement from the Agence Française de Sécurité Sanitaire des Aliments specified the known or suspected neurotoxic components and their mechanism of action. Three categorical variables were created and were assigned a value of 1 for a case of exposure to any neurotoxic product of the category of concern during  $C_0 - C_{2006}$  and a value of 0 otherwise. Missing values from the questionnaire were imputed by using available covariates as predictors after verifying that the missingness pattern was compatible with random missing (*27*).

Genotypes were linearly classified into 5 levels of risk on a log scale (Table 1) according to the odds ratio (OR) estimated for the sheep population in France (*19*). Many genotypes were missing, mainly for controls because of



\*Groups showed homogeneous odds ratios for atypical scrapie. The level of risk is the value of the corresponding log linear variable introduced into the multivariate model. *prnp*, prion protein.

difficulties in extracting DNA from hair samples ( $n = 117$ ) and for a few cases because of unsuitable samples  $(n = 13)$ . Missing values for controls were randomly imputed by using distribution of genotypes per breed. From the distribution of all available genotypes of cases of atypical scrapie in France, 20 datasets were imputed. To account for geographic distribution of flocks, France was divided into 9 sheep production areas according to sheep farming density and production patterns.

#### **Univariate and Multivariate Analyses**

Analyses were conditional to the matching variable and based on univariate and multivariate generalized linear mixed models with the logit link function for the outcome and  $C_0$  as a random coefficient (28,29). ORs and their 95% confidence intervals were derived from the coefficient estimates and variance parameters. When variables could not be introduced simultaneously in the multivariate analysis because they were collinear, the most biologically sound variable was selected.

Variables for the multivariate model were selected according to the recommendations of Hosmer and Lemeshow (*28*). Candidate variables for the multivariate model were backward selected according to the log-likelihood ratio test. Candidate variables with a p value <20% in univariate analyses were tested before other variables were tested. The effect of variables with a p value  $>20\%$  on the coefficient parameter of the selected variables was then verified 1 at a time. Best parameterization of continuous variables and statistical significance of interactions terms were then checked. A false discovery rate (FDR) (*30*) was calculated by using p values of the log-likelihood ratio tests for tested variables and interaction terms.

A complementary model was used to assess if genetics influenced stability of the final model. For each of the datasets imputed, level of genetic risk was introduced in the final model as an ordinal covariate: coefficients, standard errors, and Wald test p values of different variables were inferred according to the method of Little and Rubin (*27*).

#### **Sensitivity Analysis**

The national database used to sample controls did not enable us to take into account the size of the flocks. Therefore, counties with a large percentage of small flocks  $\left( \leq 20 \right)$ ewes) may have been overrepresented. To assess the influence of geographic selection bias, we conducted a sensitivity analysis by using 2 methods: 1) weighting of controls in the final model with weights being defined for each county as the ratio of the percentage of flocks  $>20$  ewes in the county divided by the percentage of flocks  $>20$  ewes at the national scale, and 2) introduction of sheep production areas as random coefficients in the final model.

#### **Results**

#### **Study Population**

Among 137 selected farms containing cases, 11 did not satisfy the selection criteria. In addition, 11 farmers refused to participate and 20 could not be reached. A total of 95 cases were included in the study. For controls, 1,131 farmers were contacted to participate in the study; 621 controls did not satisfy the selection criteria  $(374$  because flocks had  $\leq 20$ ewes, 41 because matching criteria could not be satisfied, 20 because flocks had <20 ewes and matching criteria were not satisfied, and 186 because of other reasons). In addition, 124 farmers refused to participate and 161 could not be reached. A total of 225 controls were included in the study.

The 95 cases and 225 controls were located throughout France (Figure 1). Case animals were born during 1994–2005. Because of exclusion of cases independent of the selection of their matched controls, the average ratio of controls per case was 2.4 instead of 2 and varied according to  $C_0$  (Figure 2). There were a few missing values for questionnaire variables (0.4%).

#### **Univariate Analyses**

Flocks containing cases were larger than flocks containing controls, had more animals tested for TSEs, and were present more often on sheep dairy farms (Table 2). Moreover, these 3 variables showed a significant correlation (Pearson coefficient of correlation between size of flock and number of animals tested  $\rho = 0.67$ , Spearman coefficients of correlation between size of flocks and dairy production  $\rho = 0.23$  and between number of animals tested



Figure 1. Distribution of cases of atypical scrapie and controls (no. cases/no. controls) in sheep, France, 2007. Sheep production areas are outlined in black, and counties are outlined in gray.



Figure 2. Distribution of  $C_{0}$  for cases of atypical scrapie and controls in sheep, France, 1994–2005.  $C_{0}$ , birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year  $n - 1$  to June 30 of year n) shared the same exposure.

and dairy production  $\rho = 0.20$ , p<0.001). None of the variables associated with a hypothesis of infectious origin was associated with atypical scrapie (Table 3).

Variables associated with feeding practices were not associated with increased risk for atypical scrapie (Table 4). Feeding milk replacers, which was negatively correlated with dairy production (Spearman coefficient of correlation  $\rho = -0.41$ , p<0.001), and feeding corn silage showed an inverse association with atypical scrapie.

Pesticides and insecticides on the premises correlated with an increased risk for atypical scrapie (Table 5). These 2 variables correlated with dairy production (Spearman coefficients of correlation  $\rho = 0.33$  and  $\rho = 0.39$ , respectively,  $p \leq 0.001$ ).

#### **Multivariate Analyses**

The set of candidate variables included 36 categorical variables and 1 continuous variable (Tables  $6, 7$ ). The final



\*C0, birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year n – 1 to June 30 of year n) shared the same exposure; SE, standard error; C<sub>2006</sub>, birth campaign 2006; C<sub>0</sub>-C<sub>0+2</sub>, period between C<sub>0</sub> and the 2 subsequent reproduction campaigns.



Table 3. Univariate analyses of contact with sheep from other flocks and afterbirth exposure variables conditional to  $C_0$  for atypical scrapie in sheep, France\*

 $C_0$ , birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year n - 1 to June 30 of year n) shared the same exposure;  $C_{2006}$ , birth campaign 2006;  $C_{2005}$ , birth campaign 2005.

model included 5 variables and 1 interaction term (Table 6). The random coefficient had a null variance, and the scale parameter was close to 1, which indicated an absence of shrinkage.

No variable associated with a hypothesis of infectious origin was present in the final model. The number of animals tested and sheep dairy farming were associated with disease (Table 7). Organic farms, feeding corn silage, and use of vitamin and mineral supplements showed an inverse association with disease. Use of these supplements, which was not significant by univariate analysis, was significant after adjustment for sheep dairy farming, and the 2 variables had a significant interaction term. The uncontrolled FDR for our analysis was 33%, which indicated that one third of the variables in the final model were spurious.

After we introduced the genetic effect, estimates of other variables did not vary by >25% of their initial values (Table 8). In addition, genetics showed a strong effect; OR for atypical scrapie ranged from 2.6 for genotypes in group 2 to 48.4 for genotypes in group 5 (Table 1).

#### **Sensitivity Analysis**

Sensitivity analysis to check possible geographic selection bias led to the same results as analysis without taking into account geographic selection bias when either weighting of samples (Figure 3) or adjusting for sheep production areas (Figure 1) was used. These results indicate that putative bias had no statistical effect on the results.



 $^{\star}C_{0}$ , birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year n – 1 to June 30 of year n) shared the same exposure;  $C_{2006}$ , birth campaign 2006;  $C_0-C_{0+2}$ , period between  $C_0$  and the 2 subsequent reproduction campaigns.

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Variable	Controls	Cases	Odds ratio	p value
Use of mineral drugs during $C_0 - C_{2006}$				
<b>No</b>	99	38	1.2	0.50
Yes	126	57		
Pesticides containing neurotoxic components used on crops during $C_0 - C_{2006}$				
<b>No</b>	155	51	1.9	0.009
Yes	70	44		
Insecticides containing neurotoxic components used on premises during $C_0 - C_{2006}$				
<b>No</b>	169	55	2.2	0.002
Yes	56	40		
Antiparasitic treatments containing neurotoxic components during $C_0 - C_{2006}$				
<b>No</b>	100	47	0.8	0.42
Yes	125	48		
*C <sub>0</sub> , birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year n – 1 to June 30 of year n) shared the same exposure; C <sub>2006</sub> , birth campaign 2006.				

Table 5. Univariate analyses of exposure to toxic product variables during the specified period and conditional to  $C_0$  for atypical scrapie in sheep, France\*

#### **Discussion**

There was no evidence of a relationship between risk for atypical scrapie and factors related to an infectious origin of the disease in France. Our results are consistent with those of Hopp et al (*15*). The only difference in the study by Hopp et al. was that removal of the placenta was associated with a decreased risk for disease. Other studies of atypical scrapie also suggested that the disease could have a noninfectious origin (*3*). The variables in our dataset were not associated with disease, and their corresponding ORs in univariate and multivariate analyses, regardless of the parameterization of the variables, were  $\approx$ 1, which suggests that they did not tend to be risk factors.

These results contrasts with those of case–control studies on classical scrapies, which found relationships between risk for disease and introduction of ewes (*12*–*14*), grazing on the same pasture as other flocks  $(14)$ , exposure to placenta (*13*), and feeding concentrates (*12*). Green et al. reported that flocks in which cases of scrapie (atypical or classical) were detected were larger and had more exchange of animals between flocks than control flocks (31). When they studied movement of sheep between flocks of equivalent status, flocks positive for classical scrapie were interconnected but flocks positive for atypical scrapie were

not. These authors indicated that their results were compatible with atypical scrapie not being infectious.

Our results showed the influence of nutritional and metabolic factors. Although sheep dairy farming covers a broad category of farms with many factors, sheep dairy farms often use more sophisticated technology, and dairy ewes are more exposed to metabolic disorders because of high levels of exported nutrients, including minerals, during milk production. Thus, some feed components such as vitamin and mineral supplements or corn silage could alleviate the risk for disease. This lessening of risk may also occur with less harsh farming conditions found on organic farms.

There is evidence suggesting that minerals, especially copper, manganese, and zinc, could play a role in the physiopathology of prion diseases (*16*,*17*). In contrast with our results, Hopp et al. (*15*) found that feeding of vitamin and mineral supplements was associated with atypical scrapie. The difference in the association between atypical scrapie and vitamin and mineral supplements in the 2 studies could be explained by differences in local conditions or roles of some minerals (*17*). Proper balance of minerals is complex because many interactions occur between mineral intake, diet of the animals, and physiologic conditions. Results



\*For categorical variables, the reference value was no. C<sub>0</sub>, birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year  $n - 1$  to June 30 of year n) shared the same exposure.

†By log likelihood ratio test.

Table 7. Adjusted odds ratios of atypical scrapie associated with variables computed from the final model in sheep, France\*

Variable	Adjusted odds ratio	95% CI
No. animals tested increased by 5	1.22	$1.11 - 1.35$
Sheep dairy farm when vitamin and mineral supplements not given	15.06	3.25-69.73
Sheep dairy farm when vitamin and mineral supplements given	2.77	$1.21 - 6.37$
Organic farm	0.15	$0.02 - 1.26$
Corn silage	0.16	$0.05 - 0.53$
Vitamin and mineral supplements used on sheep dairy farms	0.18	$0.03 - 1.04$
Vitamin and mineral supplements not used on sheep dairy farms	0.6	$0.32 - 1.14$
*CI, confidence interval.		

of the few epidemiologic studies conducted on this topic (*32*–*35*) were rarely conclusive but such associations are difficult to demonstrate with observational data.

Among different mechanisms that should be explored to understand occurrence of atypical scrapie, we believe that toxic exposure should not be overlooked. In our study, organic farms were less at risk, and univariate analysis showed that pesticides on crops and insecticides on the premises were risk factors for disease. Multivariate analysis showed a confounding effect of dairy farming but this finding should not rule out a possible effect of pesticides. Organophosphates and pyrethroids were frequently identified among reported products; these 2 groups of active components, which may be associated with Parkinson disease (*21*,*36*), interfere with the mitochondrial chain, which is a possible pathway for prion diseases (*37*). Another study on toxic exposures and gene–environment interaction could be more conclusive regarding this issue (*38*). In that regard, flocks recently exposed to insecticide treatment for control of bluetongue vectors could constitute interesting cohorts to follow up for atypical scrapie.

Genetic factors had no confounding effect on variables in the final model, and associated ORs were high  $(\leq 48$  for animals with the highest genetic risk). This high magnitude of risk, compared with other factors, suggests that genetic factors play a role in the epidemiology of atypical scrapie. However, exploring the role of genetic factors in the origin of atypical scrapie would require estimation of penetrance of different mutations, but this estimate is not currently

available. In addition, genetic factors are known to be a major risk factor for classical scrapie despite a contagious origin (*4*).

In an exploratory study such as ours, interpretation of variables in the final model is not straightforward. Particular caution should be given to the risk for purely statistical associations and to selection, classification, or confusion biases. The risk for purely statistical associations increases with the number of variables tested. An FDR estimates the rate of such spurious associations (*30*), which was high in our final model (FDR 33%). To limit the FDR to a low value  $(5\%)$ , one would only consider as confident results variables with a p value  $\leq 6 \times 10^{-4}$ , which are the number of animals tested, dairy farming, and use of corn silage. However, a high FDR should not prevent discussion of findings. Rather, this value provides a safeguard against overinterpretation of results when testing many hypotheses.

We identified a geographic selection bias in selection of controls. However, sensitivity analysis indicated that this bias did not influence the results. Misclassification problems for some variables could not be excluded, especially those regarding recall bias and memory failure. To minimize these problems, farming documents and account books were used when available, and some data were directly obtained from national databases.

Year of birth could have been a confounding factor because cases of atypical scrapie are usually found in old animals, and exposure to risk factors could be time-dependant. We matched controls on the birth cohort and accounted for



\*Genetic risk from multiple imputation parameters was estimated by the method of Little and Rubin (27). C<sub>0</sub>, birth cohort assuming that in each flock all animals born during the same birth campaign (defined from July 1 of year n – 1 to June 30 of year n) shared the same exposure. †By Wald test.

‡Because the variable is ordinal, the odds ratio (OR) for a given level of genetic risk is the exponential of the coded level (see Table 1) multiplied by the estimated coefficient  $\beta$ . The ORs for groups 2–5 are 2.6, 7.0, 18.4, and 48.4, respectively.



Figure 3. Distribution of control weightings calculated as the ratio of the percentage of flocks with >20 ewes in the county over the average percentage of flocks with >20 ewes for atypical scrapie in sheep, in France, 2007. Ranges represent classes of weightings.

year of birth as a random coefficient in a generalized linear mixed model that is recommended with this design (*28*,*29*). Null variance of the random coefficient indicated that there was no variation of risk between different birth cohorts.

In several studies on scrapie (atypical or classical), size of the fl ock was a risk factor for disease (*12*–*14*,*31*,*39*,*40*). There are many possible relationships between size and status of the animals. In the hypothesis of infectious origin, larger flocks are more exposed because of increased contacts, whereas in the hypothesis of spontaneous disease, expectancy of the number of cases increases with the size of the flock. Farming practices could also affect risk for developing disease and could be linked to the size of the flock. Moreover, larger flocks have higher number of animals tested as part of active surveillance, which increases the probability of detecting an animal with disease. The number of animals tested is determined by a combination of many structural factors that involve the size of the flock, local conditions of implementation of the surveillance program, mortality and culling rates, and use of a TSE qualification program. We could not simultaneously adjust for size of the flock and number of animal tested, and we prioritized control of surveillance bias. However, the 2 options were numerically equivalent.

Genetic analysis suggested no confounding effect but a strong association with the disease. However, results must be interpreted with caution because sensitivity analysis was conducted after imputing missing data for 53% of the controls and 13% of the cases.

Our final model suggested that atypical scrapie in sheep could be a spontaneous disease with a genetic determinant and possible influence of environmental and metabolic factors. On the basis of our results, there was no risk factor linked to an infectious origin. In particular, atypical scrapie is unlikely to originate from purchase of sheep. Other epidemiologic approaches such as spatial analyses or surveys on occurrence of secondary cases could help substantiate these findings. If infectious origin is confirmed, this finding would indicate that movement limitations of animals from flocks positive for atypical scrapie would not be a key measure in controlling the disease.

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