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► To cite this version:

Bérenger Janzac, Frederic Fabre, Alain Palloix, Benoît Moury. Constraints on evolution of virus avirulence factors predict the durability of corresponding plant resistances. *Molecular Plant Pathology*, 2009, 10 (5), pp.599-610. 10.1111/j.1364-3703.2009.00554.x . hal-02666562

HAL Id: hal-02666562

<https://hal.inrae.fr/hal-02666562>

Submitted on 31 May 2020

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A new predictor of plant resistance durability

Constraints on evolution of virus avirulence factors predict the durability of corresponding plant resistances.

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Summary

Understanding the factors driving pathogen emergence and re-emergence is a major challenge, particularly in agriculture where the use of resistant plant cultivars imposes strong selective pressures on plant pathogen populations and leads frequently to “resistance breakdown”. Presently, durable resistances are only identified after a long period of large-scale cultivation of resistant cultivars. We propose a new predictor of the durability of plant resistance. Because resistance breakdown involves modifications in the avirulence factors of pathogens, we tested for correlations between the evolutionary constraint acting on avirulence factors or their diversity, and the durability of the corresponding resistance genes in the case of plant-virus interactions. An analysis performed on 20 virus species-resistance gene combinations revealed that the selective constraint applied on amino acid substitutions in virus avirulence factors correlates with the observed durability of the corresponding resistance genes. Based on this result, a model predicting the potential durability of resistance genes as a function of the selective constraint applied on the corresponding avirulence factors is proposed to help breeders to select the most durable resistance genes.

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Code de champ modifié

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Introduction

Understanding the mechanisms linked to the emergence of pathogenic micro-organisms is a major challenge considering their impact on human health and activities. Most attention has been paid to the role of ecological factors (including migration, climate, agricultural practices and spread of biological disease vectors) in the emergence of plant diseases (Anderson *et al.*, 2004). However, the role of the ecological factors is often difficult to disentangle from that of the genetic changes operating within the pathogen populations (Fargette *et al.*, 2006; Desprez-Loustau *et al.*, 2007; Holmes and Drummond, 2007). The emergence of pathogens requires both their genetic adaptation to new hosts and their subsequent spread and maintenance within host populations. Diverse strategies have been developed to control plant pathogens, the use of resistant varieties by growers being the most efficient and simplest one. However, the widespread use of genetic resistances imposes strong selective pressures on the targeted pathogens which are able to adapt, resulting in the breakdown of the resistance. Resistance breakdown is a case of pathogen emergence which involves the jump of a host barrier at the intraspecific level and whose genetic bases have been best documented (Flor 1971; Gabriel 1999; Skamnioti and Ridout, 2005). Combining surveys of resistance breakdown on the long term with genetic analyses of pathogen populations can therefore help understanding the respective roles of intrinsic evolutionary processes and external ecological factors in pathogen emergence.

As defined by Johnson (1979), a resistance is durable if it remains effective in a cultivar for a long period of time during its widespread cultivation under environments favourable to the disease development. According to this definition, the durability of resistances can be measured only *a posteriori* after their large-scale deployment. Ability to predict the durability of resistance represents a major economic issue in agriculture. However, the criteria currently retained by breeders to evaluate resistance sources do not include their potential durability

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because of the lack of predictors. According to several authors, the mechanism of action and the genetic determinism of plant resistance to pathogens explain only a small part of their durability (Fraser 1990; McDonald and Linde, 2002; García-Arenal and McDonald, 2003). The first predictor of resistance durability to pathogens was proposed by McDonald and Linde (2002) and further applied to virus resistances by García-Arenal and McDonald (2003). This predictor, called ‘evolutionary potential’ of pathogens focused on processes that govern pathogen population evolution (mainly effective population size, migration and reproduction system). However, this approach ignores the molecular events involved in resistance breakdown or the fitness cost associated to virulence, i.e. ability of a pathogen genotype to overcome a resistance in the case of gene-for-gene interactions (Sacristán and García-Arenal, 2008), and therefore cannot explain why different resistance genes directed toward the same pathogen can display different durabilities.

Flor (1955, 1971) was the first to show that resistance or susceptibility of plants to pathogens result from an intimate molecular relationship governed by a gene-for-gene interaction model. In this model, the interaction between two partners, the resistance gene of the plant (under at least two allelic forms: “resistant” and “susceptible”) and the avirulence gene of the pathogen (under at least two allelic forms: “avirulent” and “virulent”), determine the resistance or the susceptibility of a plant to a pathogen (Flor 1971). Breakdown of resistance occurs when the two following steps can be fulfilled by the pathogen: (i) Mutation and/or recombination events should appear in the avirulence gene of a pathogen to generate virulent variants and (ii) the fitness of these virulent variants should permit them to increase in frequency in the plant population and spread in the agro-ecosystem (Leach *et al.*, 2001). Since resistance durability is linked to the mechanisms and dynamics that control these two steps, we propose a new predictor of plant resistance durability based on the properties of pathogen avirulence factors. We focussed our study on plant viruses, for which a large number of avirulence factors

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corresponding to diverse plant resistance genes have been determined (Kang *et al.*, 2005b) and for which estimates of resistance durability are available from field data. We checked whether the evolutionary constraint acting on amino acid substitutions in avirulence factors determines the durability of the corresponding resistance. Indeed, it is likely that the higher the constraint on the avirulence factor, the higher the fitness penalty conferred by mutations and/or recombination in this avirulence factor and consequently, the higher the resistance durability. In addition, we tested the hypotheses that the nucleotide or amino acid diversity of avirulence factors determines resistance durability. Indeed, the higher the diversity in avirulence factors, the higher the probability for the virus to evolve a virulent form of the avirulence factor with a minimum number of nucleotide or amino acid changes and hence, the lower the resistance durability. It should be noted that, in theory, there is no expected correlation between the selective constraint applied on genes and their diversity (Kimura 1985; Wyckoff *et al.*, 2005), justifying to consider these hypotheses separately.

Results

Dataset construction and evaluation

The first step to test the correlation between the resistance durability and its potential predictors (evolutionary constraint and diversities of avirulence factors) was to build a dataset allowing to estimate precisely all these variables. Altogether, 20 virus species-plant resistance gene combinations satisfying this condition were retained (see Experimental procedures). These combinations correspond to 17 different avirulence factors, 13 different virus species and seven virus genera (Table 1). All viruses are RNA viruses, with a majority of potyviruses and tobamoviruses, and they infect a large diversity of host species. Seven viruses are transmitted by insects, two by soilborne fungi and four by contacts between plants. Six recessive and 14 dominant resistance genes are included.

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The 20 virus species-plant resistance gene combinations composing the dataset were classified among four durability classes of increasing durability according to the time scale of protection efficiency and the proportion of virulent variants described in the agroecosystem (Table 1 and see Experimental procedures). The diversity of avirulence factors and the evolutionary constraint acting on these factors were estimated from sequences available in GenBank (see supplementary material). The p-distance was used to estimate the diversity of avirulence factors both at the nucleotide and amino acid levels and the ω ratio, which is the ratio between the non-synonymous and the synonymous substitution rates (Kimura 1983), was used to estimate the evolutionary constraint acting on amino acid substitutions in the avirulence factors (Yang *et al.*, 2000; Yang and Swanson, 2002) (see Experimental procedures). We observed that the nucleotide diversity of the avirulence factors retained in the dataset was usually low and varied from 0.005 (movement protein of ToMV) to 0.160 (VPg of BYMV). Similarly, the amino acid diversity varied from 0.005 to 0.109 (Table 1). The ω ratio was generally low (≤ 0.3) indicating that avirulence factors were globally submitted to negative selection and that most amino acid changes were deleterious in these proteins (Kimura 1983). However, for the P25 protein of BNYYV a higher value was estimated ($\omega = 0.97$). For all the avirulence factors analyzed, no amino acid sites were found under significant positive selection ($\omega > 1$) with the SLAC method.

To get rid of possible biases in the estimation of the putative predictors of resistance durability, we investigated the minimum number of sequences required for correct estimation of the diversities and ω ratio. We noticed that eight to ten sequences were usually enough to obtain satisfactory estimates (see Experimental procedures and Fig.1) and that the lower the variable, the smaller the number of sequences required. Consequently, for the analyses we have retained only datasets with a minimum of ten sequences of the avirulence factor except in three datasets (P3-HcPro of SMV, 183 kDa protein of ToMV and 25 kDa protein of PVX)

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for which only eight or nine sequences of the avirulence factor were available, but which presented relatively low ω ratios and diversities. We also validated that the number of geographic origins of isolates or the presence of virulent isolates did not introduce biases in our estimates of the diversity or ω ratio of avirulence factors (see experimental procedures for details).

Statistical analyses of the correlation between resistance durability and its potential predictors

The distribution of the virus species-resistance gene combinations in the durability classes was slightly unbalanced, with class 4 of “very high resistance durability” underrepresented compared to the other classes (Table 1). Consequently, classes 3 and 4 were grouped into a single durability class named 3+4. Correlations between the three variables related to the avirulence factors or the evolutionary potential of viruses as defined by García-Arenal and McDonald (2003) and the durability of the corresponding resistances were tested with two statistical methods: An ordinal multinomial model and a Kruskal-Wallis (KW) test (see Experimental procedures and Table 2). In agreement with our hypothesis, a clear relationship was observed between the ω ratio of the avirulence factors and the durability of the corresponding resistance: The lower the ω ratio, the higher the resistance durability (Fig. 2A). This relationship was highly significant with the ordinal multinomial model (P-value = 0.008) and not determined solely by the particularly high ω value of the P25 protein of BNYVV. Indeed, a nonparametric KW test, which has the advantage to give the same weight to each datum whatever its ω value, was also significant (P-value = 0.031) (Table 2). Because no noticeable differences were observed between the ω ratio of classes 2 and 3+4 (Fig. 2A), an analysis of the correlation between resistance durability and ω ratio was carried out using only two durability classes: Class 1 and class 2+3+4. Considering only two durability classes, the relationship observed between resistance durability and ω ratio was slightly better (P-value =

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0.003; ordinal multinomial model) than with three durability classes (P-value = 0.008) (Table 2). The relationship between the diversity of avirulence factors and resistance durability was, on the whole, opposite to that expected: The higher the diversity, the higher the resistance durability (Fig. 2 B and C). For the nucleotide diversity, a marginally significant relationship was detected using the KW test (P-value = 0.052, Table 2). Note that the results obtained with the ordinal multinomial model cannot be exploited because the proportional odds assumption was not satisfied (score-test = 0.039). No significant relationship was found between the resistance durability and either (i) the amino acid diversity, (ii) the evolutionary potential of viruses defined by García-Arenal and McDonald (2003) or (iii) the type of resistance (dominant or recessive) (Fig. 2 C and D, Table 2). Similarly, the separate analysis of the relationships between resistance durability and the three components of the evolutionary potential of viruses did not reveal any significance (data not shown). The smaller size of the dataset used here, as well as differences in the analytical methods could explain that difference with García-Arenal and McDonald's analyses. A likelihood ratio test indicates that a model including the ω ratio estimated on avirulence factors and the evolutionary potential of viruses or the type of resistance does not explain better the resistance durability than a model involving only the ω ratio (Table 2).

Discussion

McDonald and Linde (2002) first proposed a predictor of the durability of resistances based on the biological properties of the targeted pathogen. Later García-Arenal and McDonald (2003) applied the same approach to virus resistances and defined a risk index for virus species which correlates to resistance durability. With the increasing knowledge on viral avirulence factors, it has been postulated that some relationships might exist between the resistance durability and (i) the number of mutations involved in virulence (Harrison 2002,

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Ayme *et al.*, 2007), or (ii) the fitness cost associated to these mutations (Leach *et al.*, 2001; Jenner *et al.*, 2002). Fitness cost due to mutations involved in virulence property seemed important to predict the resistance durability in several plant-virus pathosystems (Boyd 2006; Lecoq *et al.*, 2004). We have consequently explored the possible relationships between the observed resistance durability and properties of virus avirulence factors.

Causality relationships between properties of viral avirulence factors and resistance durability

We observed a marginal relationship between resistance durability and the nucleotide diversity of the avirulence factor (Table 2). This relationship was opposite to that expected since the higher the diversity, the higher the resistance durability (Fig. 2 B). This results mainly from the lower diversity observed in durability class 1. A possible explanation is that three (TSWV, BNYVV and PMMoV) out of four viruses in the durability class 1 emerged relatively recently on a large geographical scale (independently of the resistance breakdown phenomenon), while the emergence of viruses in other classes was older (at least 20 years ago given our definition of durability classes 2, 3 and 4). As a consequence, viruses in the durability class 1 may have had less time to get diversified than viruses corresponding to classes 2, 3 and 4. The marginal correlation between resistance durability and nucleotide diversity of avirulence factors would then be both indirect and due to a bias in our dataset. More precise data about resistance breakdowns would be necessary to validate this assumption. It is important to note that this putative bias has, in theory, no consequences on our analyses involving the ω ratio (Kimura 1985; Wyckoff *et al.*, 2005).

A significant relationship was detected between resistance durability and the constraint exerted on amino acid changes (ω ratio) in avirulence factors (Table 2). Predicting resistance durability from the value of the ω ratio implies a causality relationship. The observed association between resistance durability and the ω ratio could be due to the following

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relationships: (i) The ω ratio determines the resistance durability; (ii) the resistance durability determines the ω ratio; (iii) both the resistance durability and the ω ratio are determined by at least one third factor. Since we could not find any evident third factor that would determine a negative correlation between resistance durability and ω ratio of avirulence factors, the first two scenarios only will be examined. Could resistance durability determine variations in the ω ratio of avirulence factors? Datasets corresponding to durability classes 1 and (to a lower extent) 2 contain a proportion of sequences of virulent isolates which have undergone amino acid substitutions in their avirulence factors. Since these particular amino acid positions could be subjected to positive selection (Moury 2004; Schirmer *et al.*, 2005), this could increase the ω ratio for durability classes 1 and 2 while affecting less (or not at all) the ω ratio for the durability classes 3 and 4. However, the facts that (i) in durability classes 1 and 2 a limited proportion of virulent isolates (0% to 22 %) are characterized (except for BNYVV-*RzI*, BaYMV-*rym4*, BYMV-*wlv*, PMMoV-*L*³, SMV-*RsvI* and PVY-*pvr2*¹) (Table 2 in supplementary materials), (ii) that the change in virulence properties generally depends only on a small number of amino acid substitutions, usually one or two (Harrison 2002), and (iii) that the presence of virulent isolates in datasets did not introduce significant biases into the ω ratio estimates for the BNYVV-*RzI*, BaYMV-*rym4* and PMMoV-*L*³ combinations (the BYMV-*wlv*, SMV-*RsvI* and PVY-*pvr2*¹ datasets did not include enough virulent and avirulent isolates for such analyses), suggest that the ω ratio in viral avirulence factors indeed determines the level of resistance durability. In addition, and most convincingly, when all virulent isolates were withdrawn from the sequence dataset, the negative correlation between the ω ratio and the resistance durability was still significant (P-value = 0.019; ordinal multinomial model) (Table 2). Note, however, that this relationship was not significant (P-value = 0.11) with the KW test, probably because four pathosystems were excluded from that analysis, due to insufficient number of sequences, and because the KW test has a low

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statistical power (see Experimental procedures). The mechanism behind this causal link resides in the fact that the ω ratio measures the propensity of viruses to change amino acids in their avirulence factors without high fitness costs under the assumption that synonymous nucleotide changes are neutral.

For viruses and other plant pathogens, gains of virulence toward resistant genotypes are frequently associated to fitness costs in susceptible ones (Meshi *et al.*, 1989; Leach *et al.*, 2001; Jenner *et al.*, 2002; Desbiez *et al.*, 2003; Ayme *et al.*, 2007). Even though limited numbers of amino acid positions have been shown to be involved in virulence properties (Harrison 2002), measure of the evolutionary constraint on the whole avirulence factor seemed adequate to our study for two reasons. First, it is plausible that much more virulence mutations than presently known exist for a given virus species-resistance gene interaction because reverse genetic analyses of avirulence mutations have frequently been performed with a small number of virulent and avirulent isolates. Second, when virulence mutations in the virus impose a fitness cost, compensatory mutations may appear somewhere else in the avirulence factor, making worthwhile the measure of the ω ratio also in these parts of the protein.

Predicting resistance durability as a function of the ω ratio

Since the evolutionary constraint acting on amino acid substitutions (ω ratio) in avirulence factors is correlated to the durability of resistance, it is possible to assess, for a given resistance gene, the probability of being in a given durability class as a function of a measured ω ratio using Eq.1 (see Experimental procedures) (Fig. 3). According to this model, resistance genes corresponding to avirulence factors with ω values higher than 0.50 are associated to the class of lower durability (class 1) with a fairly high probability ($P > 0.80$). On the opposite, resistance genes for which avirulence factors have ω values lower than 0.15 are assigned either to the classes of medium or high durability (classes 2 or 3+4) with a probability $P \geq 0.80$.

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On the whole, cross validation tests (see Experimental procedures) allowed assignments of resistance genes to the correct durability class in 44% of the cases (33% of correct assignments would be expected by chance). Even if our model still lacks power, 65% of resistance genes were correctly assigned in the durability class 3+4 corresponding to the class of genes of highest interest for resistance durability. It could also be noted that our model cannot account for the fact that different resistance alleles or genes interacting with the same viral avirulence gene show different durabilities as, for example, *pvr2*¹ and *pvr2*² for PVY (Table 1). Additional virus species-resistance gene combinations or additional predictive variables would be needed to improve the precision of those predictors. In the same way, using ω ratio estimates on the specific domain of avirulence protein which interact with resistance factors could improve the power of this model.

A novel strategy for resistance screening: Targeted search for durable resistance

Resistance durability is presently evaluated downstream of the breeding programs, when the resistance genes have been deployed in commercial varieties at a large scale and during several years. Incorporating resistance durability earlier in the breeding scheme would be highly desirable (Boyd 2006). Beyond their predictive value for the durability of resistances already incorporated in plant cultivars, ω ratios estimated from viral avirulence genes could also be integrated upstream into resistance breeding programs in two ways. First, comparing estimates of the ω ratios from avirulence factors corresponding to several candidate resistance sources could allow breeders to evaluate and select the potentially more durable resistances. Second, new resistance sources could be identified via the directed search for resistance factors elicited by candidate avirulence factors. Thus, it was shown that transient expression of avirulence proteins in plant tissues allowed to detect resistance genes in plants through the elicitation of a necrotic response characteristic of dominant resistances (Laugé *et al.*, 1998; Bendahmane *et al.*, 2000; Mestre *et al.*, 2000). Even if the efficiency of the detected

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resistance should be further confirmed by inoculation tests since necrosis could also be induced independently of the specific resistance phenomenon, using this method with the most constrained viral proteins could allow the *a priori* identification of the most durable resistances targeting them.

Conclusion

As proposed by McDonald and Linde (2002) and García-Arenal and McDonald (2003), we analyzed the ability of some pathogen traits to predict the capacity of emergence of adapted pathogens in crop plants carrying specific resistance genes. We focused on plant-virus interactions, but our analytical approach could however have a broader interest. The validity of the method depends on two functional assumptions: (i) The control mechanism (here plant resistance genes) interacts either directly or indirectly with a specific protein target and (ii) adaptation of the targeted organism to circumvent the control mechanism occurs via mutations in that protein. The observed correlation between the ω ratio of viral avirulence factors and the durability of corresponding plant resistance genes strongly suggests that these assumptions are widely satisfied in the case of plant-virus interactions. Notably, the fact that the molecular interactions between both partners operate at the protein but not the RNA level, from the virus side, is not trivial since it has been shown that RNA segments can act as viral avirulence factors independently of their possible coding capacity (Szittyá and Burgyan 2001; Diaz *et al.*, 2004). Since the two above assumptions are commonplace among plant-pathogen interactions, it would be worth evaluating the generalization of our approach to other kinds of plant pathogens.

Experimental procedures

Dataset

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To analyze the durability of plant resistance to viruses, we retained the virus species-resistance gene (or allele) combinations which satisfied the three following conditions: (i) The virus avirulence factor corresponding to the resistance gene was identified, (ii) nucleotide sequences encoding the avirulence factor could be obtained from a sufficient number of virus isolates to estimate their diversity and evolutionary constraint and (iii) plant cultivars carrying the resistance gene have been largely exploited and field data about the durability of this resistance are available. All resistance genes composing the dataset control qualitative resistance (Table 1). Nucleotide sequences of avirulence factors were retrieved from Genbank, aligned with the ClustalW program (Thompson *et al.*, 1994) and analyzed with the RDP version 3 software (Martin *et al.*, 2005) to remove putative recombinant sequences which could create biases in the estimation of the evolutionary constraints (Schierup and Hein, 2000). For *Potato virus Y* (PVY) and *Tobacco etch virus* (TEV), five and twelve additional sequences of the VPg-coding region, respectively, were obtained (accession numbers EU334778 to EU334794). For the *pvr2*¹ and *pvr2*² resistance alleles in pepper, only VPg sequences from pepper isolates of PVY belonging to the C phylogenetic group were analyzed. In case of overlapping open reading frames in the cistron coding for the avirulence factor (183 or 184 kDa proteins and movement protein of tobamoviruses and 25 kDa protein of potato virus X), the overlapping region was removed to avoid artifacts in the estimation of their diversity and evolutionary constraint. GenBank accession numbers of sequences retained are available in supplementary materials (Table 2). To avoid artifactual polymorphism, we discarded the sequences of virus isolates that have been maintained on the long term in laboratory through repeated inoculations and sequences of virulent isolates selected artificially.

Estimation of avirulence factors diversity, constraint exerted on amino acid substitutions and evolutionary potential of viruses

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For each virus species-resistance gene combination, four following putative predictors of resistance durability were determined: (i) The diversity of the avirulence factors at the nucleotide level, (ii) the diversity of the avirulence factors at the amino acid level, (iii) the evolutionary constraint on amino acid substitutions in the avirulence factors and (iv) the evolutionary potential of viruses as defined by García-Arenal and McDonald (2003). Sequence alignments of avirulence factors were used to estimate the first three predictors. The average pairwise p-distance was used to estimate the diversity of the avirulence factors, both at the nucleotide and amino acid levels, using MEGA version 3.1 (Kumar *et al.*, 1994). The ω ratio, which is the ratio between the non-synonymous (amino acid altering) and the synonymous (silent at the amino acid level) substitutions, was used to estimate the evolutionary constraint acting on amino acid substitutions in the avirulence factors (Kimura 1983; Yang *et al.*, 2000; Yang and Swanson, 2002). This ω ratio was estimated on whole avirulence factor sequences with the SLAC (Single-Likelihood Ancestor Counting) method in the HyPhy software (Kosakovsky Pond and Frost, 2005), a modification of the site-by-site analytical method of Suzuki and Gojobori (1999). The probability threshold to detect particular amino acid sites undergoing positive selection ($\omega > 1$) was set to $P = 0.10$.

The evolutionary potential of viruses (García-Arenal and McDonald, 2003) corresponds to a compound risk index based on the biology of the viruses and takes into account: (i) The population size of the virus, (ii) the amount of gene and genotype flow in the virus population, and (iii) its frequency of recombination. Evolutionary potential of virus species were obtained from García-Arenal and McDonald (2003), or estimated according to the criteria defined by these authors (Table 1 and supplementary materials).

Precision of the estimates of diversity and evolutionary constraints of avirulence factors

The precision of estimates for the diversity and evolutionary constraint of avirulence genes depends on representative samples of sequences of these genes. A method to evaluate the

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representativeness of samples is to measure the variance of the studied variables with random subsamples of increasing sizes (Hughes et al. 2001). Here, subsamples of sequences of increasing size were randomly chosen among the five largest sequence sets (P25 of BNYVV, CP of ToMV, VPg of BaYMV, CP of PVX and NSs of TSWV) presenting contrasted diversity and evolutionary constraint (Table 1). In each subsample, ten random sets of sequences were analyzed allowing to estimate the variance of the three variables (nucleotide and amino acid diversities and ω ratio) and to perform mean comparisons. The minimal number of sequences that allowed to obtain satisfactory estimates of the ω ratio or diversities of avirulence factors (*i.e.* minimum sample delivering estimates that are not significantly different from those obtained with the whole dataset) was comprised between eight (CP of PVX and NSs of TSWV) and ten (CP of ToMV and VPg of BaYMV) (Fig. 1). The only exception concerned the P25 protein of BNYVV that has a higher ω ratio and, consequently, required a minimum of 20 sequences to be estimated properly (Fig. 1).

The same approach was used to determine the effect of increasing the number of geographical localities (countries) on the precision of variable estimates. Subsamples of ten sequences with increasing numbers of geographic origins chosen among the four sequence sets showing the highest diversity of geographical origins (TSWV-NSs, TuMV-CI, PMMoV-CP, PVY-NiaPro) were used to estimate the variance of the variables. No significant differences of predictor estimates were revealed by rank tests performed on subsamples of sequences showing increasing diversity of geographic origins. To evaluate the impact of virulent isolates on our results, we also compared the estimates of the ω ratio obtained with sequence samples including or not the virulent isolates for BNYVV-*RzI*, BaYMV-*rym4* and PMMoV-*L*³ which included sufficient numbers of virulent and avirulent isolates (See additional material Table 2). We also conducted our analyses with a dataset excluding the sequences of virulent isolates to get rid of any bias on the correlation between the ω ratio of avirulence factors and resistance

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durability. This dataset was made up of 16 virus species-plant resistance gene combinations satisfying the condition of a minimum of 8 sequences of the avirulence gene (BYMV-*w/v*, SMV-*RsvI*, PVY-*pvr2*¹ and PVX-*Nb* were excluded).

Estimation of the durability of resistance

The durability of resistance genes or alleles was assessed according to the timescale of the protection efficiency before the appearance of crop damages due to resistance-breaking isolates and according to the abundance of resistance-breaking isolates in the agro-ecosystem (Table 1). A majority of these data was collected from García-Arenal and McDonald (2003) and Harrison (2002) and additional references are available in supplementary Table 1. All the resistances in the dataset have been used in the field for more than 20 years or were broken down rapidly (Table 1). Resistances were considered as “broken-down” if the resistance was defeated by resistance-breaking isolates and lead to important economic losses at least in some growing regions. Resistance durability was categorized in four discrete durability classes as follows.

Class 1: Rapid resistance breakdown (less than 12 years after initial deployment). Frequent virulent isolates in the field prevent the efficiency of the resistance at least in some growing regions.

Class 2: Slow resistance breakdown (more than 12 years after initial deployment). Frequent virulent isolates in the field prevent the efficiency of the resistance at least in some growing regions.

Class 3: Resistances deployed for more than 20 years with rare virulent isolates in the field. Virulent isolates are restricted to small areas and did not economically compromise the use of the resistant cultivars.

Class 4: Resistances deployed for more than 20 years and no virulent strains described in the field.

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In our dataset (Table 1), the distinction between the durability classes 1 and 2 was clear cut since breakdown of resistance in class 1 occurred in one to eleven years while breakdown of resistances in class 2 occurred mostly in more than 20 years. The SMV-*Rsv-1* gene combination only had an intermediate durability (breakdown in about 15 years). However, similar results were obtained when it was classified in class 1 or 2. Due to the small number of resistances in class 4, classes 3 and 4 were grouped for further analyses.

Statistical analysis

Statistical analyses were used to evaluate the correlation between the response variable “durability of resistance” and four potential predictor variables (nucleotide and amino acid diversities of avirulence factors, ω ratios of avirulence factors and evolutionary potential of viruses). We first used a non-parametric Kruskal-Wallis test for the null hypothesis: “The average values of the considered explanatory variable are similar for all resistance durability classes”. We also used a parametric approach based on ordinal multinomial models. Multinomial models extend logistic regression to the case where the response variable has three (or more) discrete outcomes (*i.e.* the response variable follows a multinomial distribution) (McCullagh and Nelder, 1989). In our case, the durability of resistances is an ordinal response variable because resistance durability increases from class 1 to class 3+4. Ordinal multinomial models allow both (i) to test for a correlation between predictor and response variables by taking into account the ordinal and multinomial nature of the response variable and (ii) to model the probability to belong to a particular durability class as a function of a set of predictor variables. More precisely, Y being the response variable “durability of resistance”, let's denote $P_r = p(Y \leq r)$ the probability of Y being in the r or lower class of durability ($r \in \{1, 2, 3+4\}$). In the proportional odds model, P_r is modelled with a logit link as

$$P_r = \frac{\exp(\alpha_r + \beta.X)}{1 + \exp(\alpha_r + \beta.X)} \quad \text{for } r=1,2 \quad \text{and} \quad P_3 = 1 \quad (\text{Eq.1})$$

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where α_r are intercept parameters that depend only on the categories such that $\alpha_1 < \alpha_2$ and β is a vector of the regression coefficients for the set of predictor variables X . Inherent in this model is the proportional odds assumption, which states that the odds (*i.e.* the quantity $p(Y \leq r) / p(Y > r)$) of being in a category less than or equal to r is $\exp[\beta \cdot (x_1 - x_2)]$ times higher at $X = x_1$ than at $X = x_2$, x_1 and x_2 being two particular values of X . The PROC LOGISTIC of the SAS software (version 8, SAS Inc., Cary, NC) was used to estimate α_r and β by maximum likelihood for five models differing by the set X of predictor variables included (amino acid diversity, nucleotide diversity, evolutionary potential, ω ratio and the combination of ω ratio and evolutionary potential). Before fitting these models, a score test for the proportional odds assumption was performed, a non significant test supporting the evidence that the assumption is valid.

The predictive performance of the ordinal multinomial model was assessed with cross-validation tests. Cross-validation tests were performed with the R software environment (<http://cran.r-project.org/>) by (i) randomly leaving out 3 pathosystems (*i.e.* 15% of the dataset), (ii) fitting the model to the remaining dataset and (iii) checking if the true class of durability for each left-out resistance gene-avirulence factor combination corresponded to the highest probability class predicted by the model (if so, the prediction is correct). The predictive performance of the model was assessed as the percentage of correct assignments and estimated by iterating 1,500 times the previous three steps.

Acknowledgements

We acknowledge Dr. D. Fargette (IRD, Montpellier, France) for sharing information about the durability of the *Rymv-1* gene, C. Marchal (Clause-Tézier) for providing TEV isolates, and K. Ezzaier (INRAT, Tunisia) and Dr. G. Marchoux (INRA, Avignon) for providing PVY isolates. We also acknowledge the ANRT foundation and the companies Gautier Semences,

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452 Clause Vegetable Seeds, Vilmorin SA, Sakata Seeds Europe and Rijk Zwaan for their
453 financial support.

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Tables

Table 1. Dataset used for analysis of resistance durability and diversity and evolutionary constraint of avirulence factors.

Plant		Virus					Resistance durability				Avirulence factor variables			Selected references ^(f)
Crop	Resistance gene ^(a)	Species	Genus	Avirulence factor ^(b)	Number of sequence	EP ^(c)	Timescale of protection efficiency	Status of the resistance	Virulent isolates	Resistance durability class	Nucleotide diversity ^(d)	Amino acid diversity ^(d)	ω ratio ^(e)	
Pepper	<i>Tsw</i>	TSWV	<i>Tospovirus</i>	NSs	28	7	1 year	Broken-down	Common	1	0,034	0,036	0,291	Margaria <i>et al.</i> , 2007
Tomato	<i>Tm-1</i>	ToMV	<i>Tobamovirus</i>	184 kDa protein	8	6	2 years	Broken-down	Common	1	0,007	0,005	0,169	Meshi <i>et al.</i> , 1988; Strasser and Pfitzner, 2007
Tomato	<i>Tm-2</i>	ToMV	<i>Tobamovirus</i>	MP	11	6	2-3 years	Broken-down	Common	1	0,005	0,006	0,314	Meshi <i>et al.</i> , 1989; Strasser and Pfitzner, 2007
Pepper	<i>L³</i>	PMMoV	<i>Tobamovirus</i>	CP	18	6	5 years	Broken-down	Common	1	0,029	0,020	0,204	Berzal-Herranz <i>et al.</i> , 1995; Hamada <i>et al.</i> , 2007
Sugar beet	<i>Rz-1</i>	BNYVV	<i>Benyvirus</i>	P25	33	4	6-11years	Broken-down	Common	1	0,022	0,037	0,969	Acosta-Leal and Rush, 2007; Chiba <i>et al.</i> , 2008
Soybean	<i>Rsv-1</i>	SMV	<i>Potyvirus</i>	P3-HcPro	8	7	15 years	Broken-down	Common	2	0,034	0,022	0,132	Eggenberger and Hill, 1997; Eggenberger <i>et al.</i> , 2008 Hajimorad <i>et al.</i> , 2006
Barley	<i>rym-4</i>	BaYMV	<i>Bymovirus</i>	VPg	15	4	>20 years	Broken-down	Common	2	0,053	0,050	0,258	Kühne <i>et al.</i> , 2003
Pepper	<i>pvr2¹</i>	PVY	<i>Potyvirus</i>	VPg	11	6	>20 years	Broken-down	Common	2	0,090	0,065	0,137	Moury <i>et al.</i> , 2004
Potato	<i>Nb</i>	PVX	<i>Potyvirus</i>	25 kDa protein	9	4	>20 years	Broken-down	Common	2	0,035	0,021	0,096	Malcuit <i>et al.</i> , 1999
Rapeseed	<i>TuRB01</i>	TuMV	<i>Potyvirus</i>	CI	29	9	>20 years	Broken-down	Common	2	0,138	0,030	0,029	Jenner <i>et al.</i> , 2000
Pepper	<i>pvr2²</i>	TEV	<i>Potyvirus</i>	VPg	15	7	>20 years	Broken-down	Common	2	0,068	0,050	0,194	Kang <i>et al.</i> , 2005a
Rice	<i>rymv-1</i>	RYMV	<i>Sobemovirus</i>	VPg	18	4	>20 years	Broken-down	Common	2	0,076	0,031	0,075	Hébrard <i>et al.</i> , 2006
Tomato	<i>Tm-2²</i>	ToMV	<i>Tobamovirus</i>	MP	11	6	>20 years	Not broken-down	Rare	3	0,005	0,006	0,314	Weber <i>et al.</i> , 1993; Strasser and Pfitzner, 2007
Potato	<i>Nx</i>	PVX	<i>Potexvirus</i>	CP	26	4	>20 years	Not broken-down	Rare	3	0,086	0,040	0,058	Santa Cruz and Baulcombe, 1995
Potato	<i>Rx</i>	PVX	<i>Potexvirus</i>	CP	26	4	>20 years	Not broken-down	Rare	3	0,086	0,040	0,058	Bendahmane <i>et al.</i> , 1995
Pepper	<i>pvr2²</i>	PVY	<i>Potyvirus</i>	VPg	11	6	>20 years	Not broken-down	Rare	3	0,090	0,065	0,137	Moury <i>et al.</i> , 2004
Bean	<i>wlv</i>	BYMV	<i>Potyvirus</i>	VPg	10	8	>20 years	Not broken-down	Rare	3	0,160	0,109	0,100	Bruun-Rasmussen <i>et al.</i> , 2007
Potato	<i>Ry</i>	PVY	<i>Potyvirus</i>	Nia-Pro	23	7	>20 years	Not broken-down	Not reported	4	0,128	0,053	0,068	Mestre <i>et al.</i> , 2000
Tobacco	<i>N</i>	TMV	<i>Tobamovirus</i>	183 kDa protein	10	6	>20 years	Not broken-down	Not reported	4	0,025	0,011	0,087	Padgett <i>et al.</i> , 1997
Pepper	<i>L¹</i>	ToMV	<i>Tobamovirus</i>	CP	17	6	>20 years	Not broken-down	Not reported	4	0,014	0,017	0,186	Dardick <i>et al.</i> , 1999; Hamada <i>et al.</i> , 2007

BNYVV (beet necrotic yellow vein virus); SMV (soybean mosaic virus); PMMoV (pepper mild mottle virus); ToMV (tomato mosaic virus); TSWV (tomato spotted wilt virus); BaYMV (barley yellow mosaic virus); PVY (potato virus Y); PVX (potato virus X); TuMV (turnip mosaic virus); TEV (tobacco etch virus); RYMV (rice yellow mottle virus); BYMV (bean yellow mosaic virus); TMV (tobacco mosaic virus).

^a : Upper- and lower-case initials indicate dominant and recessive resistances respectively.

^b : P25 : 25 kDa protein; P3-HcPro : protein 3 and helper component protease; CP : coat protein; MP: movement protein; NSs : non structural S protein ; VPg: genome-linked viral protein; CI: cylindrical inclusion protein; Nia-Pro: nuclear inclusion a protein.

^c : Evolutionary potential overall risk defined by García-Arenal and McDonald (2003) or from bibliographic data (see supplementary materials) .

^d : Average pairwise p-distance estimated with the MEGA 3.1 software.

^e : Nonsynonymous/synonymous substitution rate ratio estimated by the SLAC method implemented in the HyPhy software.

^f : Selected references for avirulence gene identification and amino acid substitutions involved in the resistance breakdown (when these are known).

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Table 2. Statistical analysis of the relationships between the variable “durability of resistance” and the four potential predictor variables computed with the ordinal multinomial model and with the non-parametric Kruskal-Wallis test (the underrepresented durability classes 3 and 4 were grouped for the analyses).

Variable	Ordinal multinomial model				Kruskal-Wallis test
	Odds assumption	Likelihood ratio test			P-value
	Score test ^a	Chi-Square	DF ^b	P-value ^c	
ω ratio	0.120	7.008	1	0.008 **	0.031 *
ω ratio (Avirulent isolates only)	0.372	5.488	1	0.019 *	0.110 ^{ns}
ω ratio (Two durability classes: 1 and 2+3+4)	Not applicable	8.582	1	0.003 **	0.008 **
Nucleotide diversity	0.039	4.330	1	0.037 *	0.052 ^{ns}
Amino acid diversity	0.157	2.454	1	0.118 ^{ns}	0.216 ^{ns}
Evolutionary potential (EP)	0.976	0.006	1	0.935 ^{ns}	0.995 ^{ns}
Resistance type	0.051	0.264	1	0.607 ^{ns}	0.107 ^{ns}
ω ratio + EP	0.238	7.014	2	0.028 *	Not applicable
ω ratio + Resistance type	0.015	7.277	2	0.026 *	Not applicable

^a : A non significant test ($P > 0.05$) indicates that the proportional odds assumption is valid.

^b : DF: degree(s) of freedom.

^c : P-value of the global test ($\beta=0$) : a significant test supports the evidence that at least one of the covariates is correlated with the resistance durability.

^{ns} : $P > 0.05$

* : $0.01 < P < 0.05$

** : $P < 0.01$

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Figure legends

Figure 1. Estimates of the ω ratio of virus avirulence factors according to the size of the sequence set. Five large sequence sets were used to analyse random samples of sequences of increasing size (10 random samples for each size). Arrows indicate the smallest sequence set size for which the estimate of the ω ratio is not significantly different from the estimates obtained with larger datasets.

Figure 2. Distribution into the 3 resistance durability classes of the putative predictor variables related to virus avirulence factors or to biology of viruses (durability classes 3 and 4 have been grouped). (A) ω ratio of avirulence factors, (B) nucleotide diversity of avirulence factors, (C) amino acid diversity of avirulence factors, and (D) evolutionary potential of viruses (García-Arenal and McDonald 2003). Each virus species-resistance gene combination is represented by a diamond and grey bars indicate average values.

Figure 3. Probability (P_r) that a resistance gene belongs to the durability class $r \in \{1, 2, 3+4\}$ as a function of the ω ratio of the targeted avirulence factor based on Eq.1 (see the Experimental procedures) and our dataset (Table 1).