

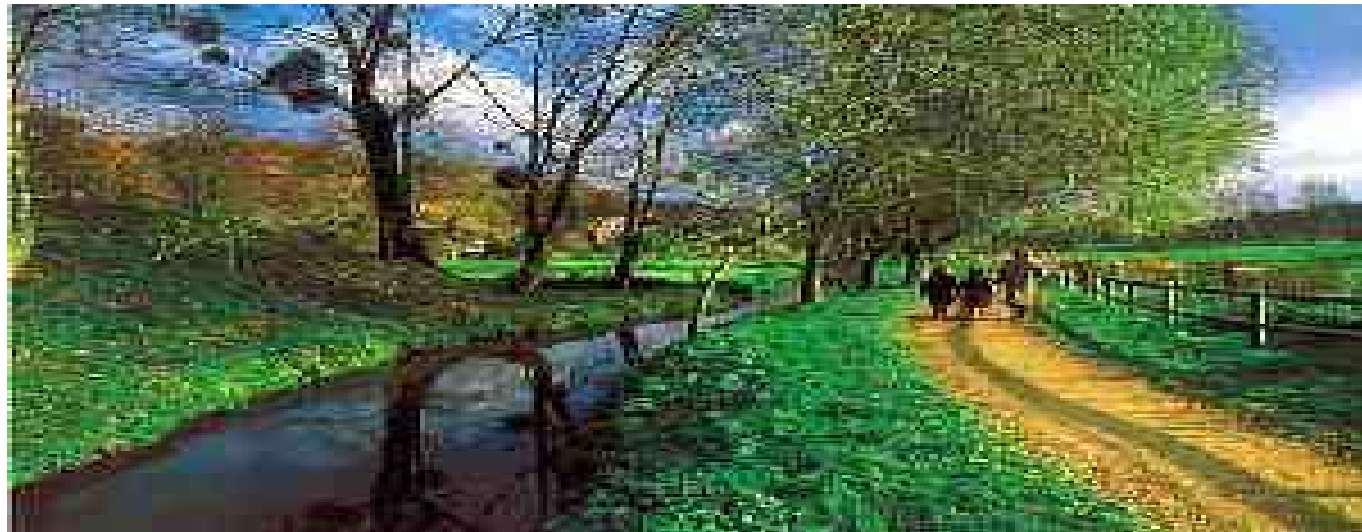
Semi-semi-Markov processes : a new class of processes for formalizing and generalizing state-dependent individual-based models

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Origin : epidemiological problem

Study the propagation of a disease (BVD : Bovine Viral Diarrhoea) in a dairy herd

3 types of individual transitions :

- Branching transitions (births, deaths)

- Group changes according to physiological status and age :

 - 4 groups : calves, heifers before breeding, heifers after breeding, dairy cows

- Health status changes which depend on the population infection

⇒ individual-based stochastic model : the individual transitions are population-dependent and semi-Markovian

⇒ propagation of a random process (disease) on a random graph (vertices : individuals)

Individual-based models

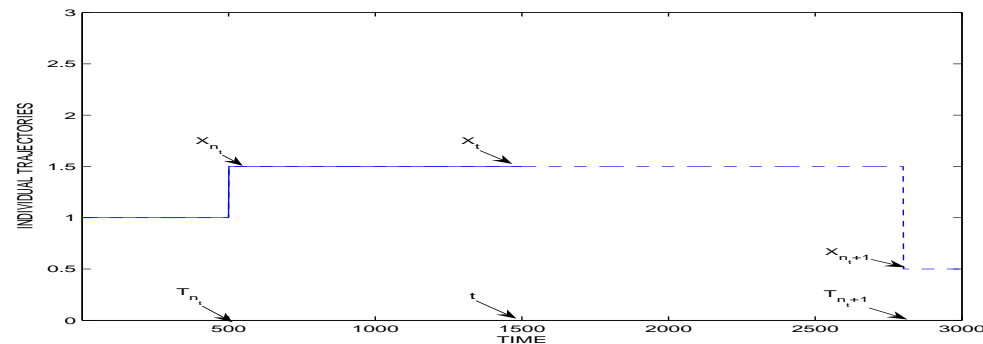
- used in population dynamics when individuals are marked by personal characteristics or when the next state-change is driven by complex rule decisions which depend on the current state of the population
- ω – calculate empirical distributions at the scale of the population from simulated individual trajectories (“bottom-up approach”)
- literature on individual-based models has considerably increased thanks to the increase of the computers capacity and the popularization of informatics
- no mathematical formalism : validation of this approach ? how is the process at the level of the population ?

Goal

- 4 Build a rigorous mathematical formalism of these models at the population level (top-down)
 - ⇒ good readability of the different model components independently of the programming language
 - ⇒ validate and supplement the empirical distributions by analytical results

Homogeneous Semi-Markov Process (SMP) for one individual ω

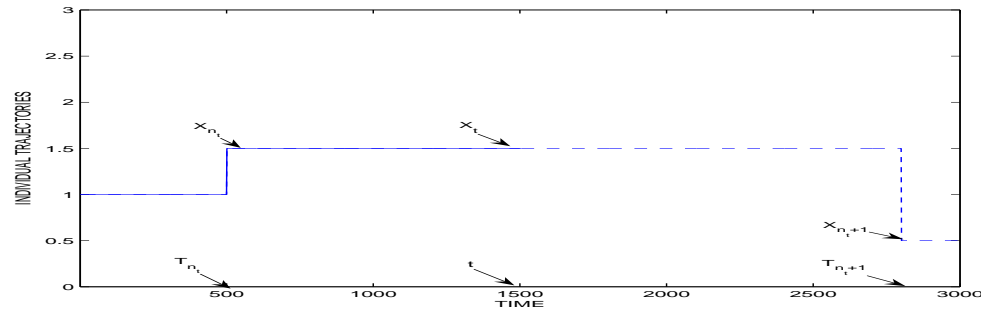
Feller W. (1964), Çinlar (1975), Kulkarni, V. (1995), Becker G. *et al.* (1999), Iosifescu M. (1999)



Countable state space, jumps at random times

$\{X_t\}_{t \in \mathbb{R}^+}$ is an homogeneous semi-Markov process if

$$X_t = X_{n_t} 1_{\{n_t = \sup\{n: T_n \leq t\}\}}, \quad T_n : n\text{th jump time}, X_n : \text{state at } T_n, \{X_n, T_n\} \text{ MRP}$$



Law of $\{X_t\}_t \iff$ law of $\{X_n, T_n\}_n \iff$ transition laws of $\{X_n, T_n\}_n$

Assumption : $\{X_n, T_n\}$ is a MRP

$$\begin{aligned} \circ \quad P(X_{n+1} = j, \Delta T_{n+1} \leq \tau | X_n, \dots, X_0, T_n, \dots, T_0) &= P(X_{n+1} = j, \Delta T_{n+1} \leq \tau | X_n) \\ &\stackrel{\text{homog.}}{=} P(X_1 = j, \Delta T_1 \leq \tau | X_0) \\ \Delta T_{n+1} &\stackrel{\text{defin.}}{=} T_{n+1} - T_n \text{ (waiting time between 2 jumps)} \end{aligned}$$

Kernels (transition laws)

$$\begin{aligned} Q_{i,j}(\tau) &\stackrel{\text{defin.}}{=} P(\Delta T_1 \leq \tau, X_1 = j | X_0 = i) \\ &= P(\Delta T_1 \leq \tau | X_1 = j, X_0 = i) P(X_1 = j | X_0 = i) \\ &= F_{i,j}(\tau) P(i, j) \end{aligned}$$

$F_{i,j}(\tau)$: cdf of the sojourn time in i before jumping in j ; $P(i, j)$: transition probability of $\{X_n\}$

Kernels \Longleftrightarrow Transition rates

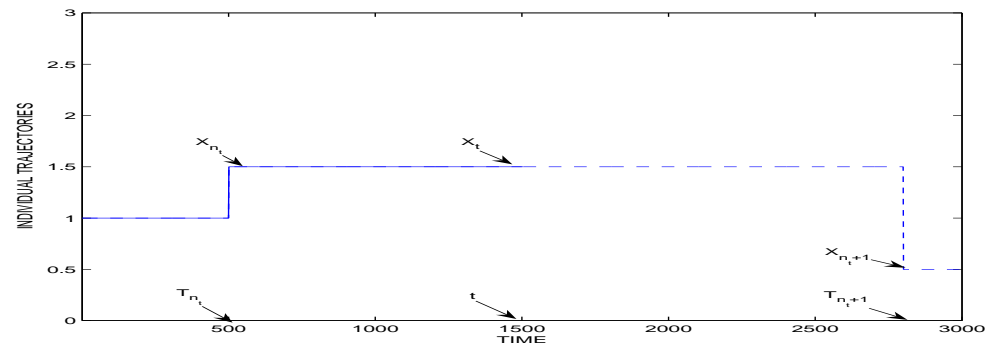
$$\lambda_{i,j}(\tau) \stackrel{def.}{=} \lim_{\Delta\tau \rightarrow 0} \frac{P(\Delta T_{n+1} \in (\tau, \tau + \Delta\tau), X_{n+1} = j | X_n = i, \Delta T_{n+1} > \tau)}{\Delta\tau}$$

$$\lambda_{i,j}(\tau) = \frac{\dot{Q}_{i,j}(\tau)}{1 - \sum_j Q_{i,j}(\tau)} \Longleftrightarrow Q_{i,j}(\tau) = \int_0^\tau \lambda_{i,j}(u) \exp\left(-\int_0^u \sum_j \lambda_{i,j}(s) ds\right) du$$

Particular case : Markov process :

$$\lambda_{i,j}(\tau) = \lambda_{i,j} = \lambda_i P(i, j), \text{ for any } \tau, F_{i,j}(\tau) = 1 - \exp(-\lambda_i \tau)$$

$$\mathbf{P}(t) = \exp(\Lambda t), \mathbf{P}[i, j](t) \stackrel{def.}{=} P(X(t) = j | X(0) = i)$$



∞ **Event-driven simulation algorithm based on $P(i, j)F_{i,j}(\tau)$**

Current jump time and jump state : $(t_n, i) \implies$ determine the next jump time and jump state (t_{n+1}, j)

Simulate j according to $\{P(i, j')\}_{j'}$

Simulate $\tau \stackrel{def.}{=} t_{n+1} - t_n$ according to $F_{i,j}(\cdot)$

Probability law of the process (renewal equations)

$$P(X_t = j | X_0 = i) \stackrel{notat.}{=} P_{i,j}(t) = [1 - \sum_{j'} Q_{i,j'}(t)] 1_{\{j=i\}} + \sum_k \int_0^t dQ_{i,k}(s) P_{k,j}(t-s)$$

Matrix process : $\mathbf{P}[i, j](t) \stackrel{def.}{=} P_{i,j}(t)$

$$\mathbf{P} = (\mathbf{I} - \mathbf{Q}^\Sigma) + \mathbf{Q} * \mathbf{P} \implies \mathbf{P} = \sum_{n=0}^{\infty} \mathbf{Q}^{n*} * (\mathbf{I} - \mathbf{Q}^\Sigma), \quad \mathbf{Q} * \mathbf{P}(t) = \sum_k \int_0^t Q_{i,k}(s) P_{k,j}(t-s)$$

Approximate solutions

Empirical distribution (simulations)

$$\mathbf{P} \simeq \sum_{n=0}^{n_t} \mathbf{Q}^{*n} * (\mathbf{I} - \mathbf{Q}^\Sigma)$$

Recursive solution of \mathbf{P} by discretization of time

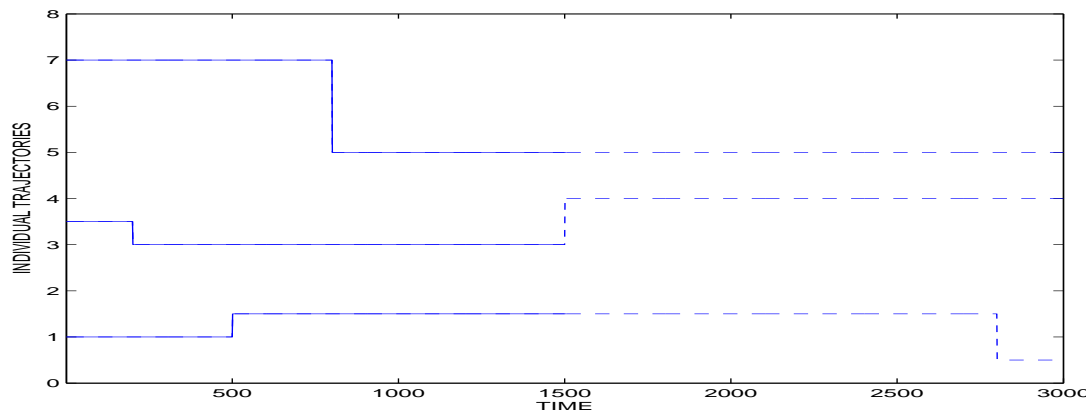
Upper and lower bounds of Li and Luo (2005)

Stationary law (all the states are recurrent) or quasi-stationary law (conditioned on staying in the nonabsorbing state)

Homogeneous Semi-Semi-Markov Processes for a closed population $\Omega = \{\omega_l\}_{l \leq N}$

Set of MRP : $\{\{(X_m^{(l)}(\omega_l), T_m^{(l)}(\omega_l))\}_m\}_{l=1,\dots,N}$, not synchronized and population-dependent

Goal : determine the distribution of the population process $\mathcal{X}_t = \{X_t^{(1)}, \dots, X_t^{(N)}\}$



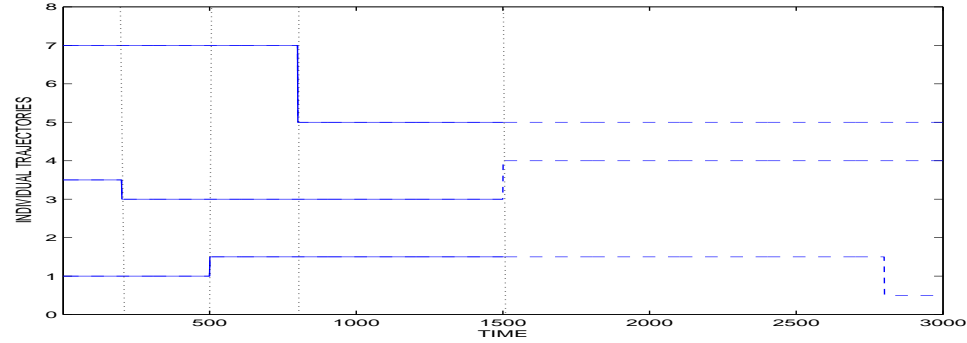
Particular case : the MRP are i.i.d.

communication networks, asymptotic distributions in the heavy-tailed case $1 - F(t) = t^{-\alpha}L(t)$
(Mikosh and Resnick, 2005, Mitov and Yanev, 2006)

Steps

- define $\{\mathcal{X}_t(\Omega)\}_{t \in \mathbb{R}^+}$ from $\{(\mathcal{X}_n(\Omega), \mathcal{T}_n(\Omega))\}_{n \in \mathbb{N}}$, defined itself from the MRP $\{\{(X_m^{(l)}(\omega_l), T_m^{(l)}(\omega_l))\}_{m \in \mathbb{N}}\}_{l \in \mathbb{N}}$
- deduce the kernel (transition law) of the population process from the individual kernels
- deduce transition rates, probability law, simulation algorithm, . . .

Definition of $\{\mathcal{X}_t(\Omega)\}$ from $\{(\mathcal{X}_n(\Omega), \mathcal{T}_n(\Omega))\}$, defined itself from the individual MRP



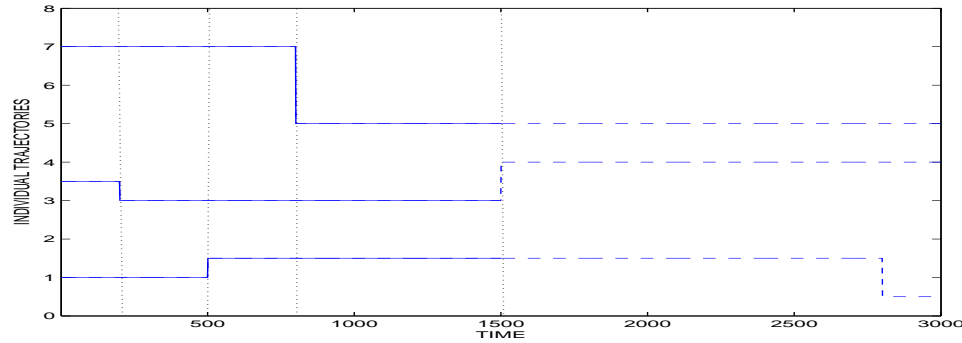
$$\mathcal{X}_t(\Omega) \stackrel{def.}{=} \mathcal{X}_{n_t}(\Omega) \stackrel{def.}{=} \{X_{m_{l,t}}^{(l)}(\omega_l)\}_l$$

$$m_{l,t}(\omega_l) \stackrel{def.}{=} \sup\{m : T_m^{(l)}(\omega_l) \leq t\}$$

$$n_t(\Omega) \stackrel{def.}{=} \sum_l m_{l,t}(\omega_l)$$

$$\mathcal{T}_{n_t} \stackrel{def.}{=} \sup_l \sup_m \{T_m^{(l)}(\omega_l) \leq t\}$$

Law of $\{\mathcal{X}_t(\Omega)\}_t \iff$ law of $\{\mathcal{X}_n, \mathcal{T}_n\}_n \iff \{P(\mathcal{X}_{n+1} = J, \Delta\mathcal{T}_{n+1} \leq \tau | \mathcal{F}_n(I))\}_{n,I,J}$ (kernels)



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$\mathcal{F}_n(I) = \{\mathcal{X}_n = I, \mathcal{X}_{n-1} = I_{n-1}, \dots, \mathcal{X}_0 = I_0, \mathcal{T}_n = t_n, \mathcal{T}_{n-1} = t_{n-1}, \dots, \mathcal{T}_0 = t_0\} : \text{past until } (t_n, I)$

Kernel

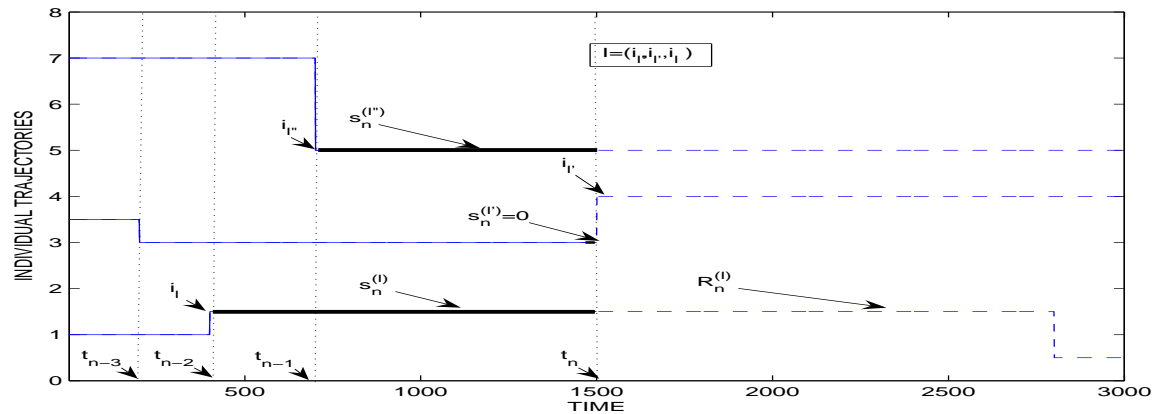
$$\begin{aligned} P(\mathcal{X}_{n+1} = J, \Delta\mathcal{T}_{n+1} \leq \tau | \mathcal{F}_n(I)) &= P(\Delta\mathcal{T}_{n+1} \leq \tau | \mathcal{X}_{n+1} = J, \mathcal{F}_n(I)) P(\mathcal{X}_{n+1} = J | \mathcal{F}_n(I)) \\ &\stackrel{\text{notation}}{=} F_{\mathcal{F}_n(I), J}(\tau) P(\mathcal{F}_n(I), J) \\ &\stackrel{\text{notation}}{=} Q_{\mathcal{F}_n(I), J}(\tau) \end{aligned}$$

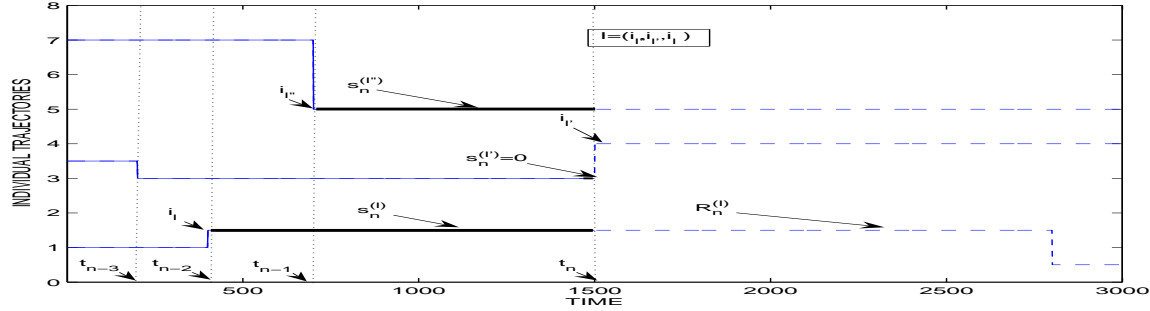
\implies calculate $F_{\mathcal{F}_n(I), J}(\cdot)$, $P(\mathcal{F}_n(I), J)$ from the individual transitions defined from (t_n, I)

Assumptions given the past until (t_n, I)

1. $A1$: the $\{(remaining\ waiting\ time\ R_n^{(l)},\ next\ jump\ state\ X_{m_n+1}^{(l)})\}_l$ are mutually independent
2. $A2$: for each l , the law of $R_n^{(l)}$ (remaining time in i_l before jumping in j_l) depends only on I and on $s_n^{(l)}$ (time already spent in i_l) and on I
3. $A3$: the probability for l to jump from i_l to j_l depends only on i_l, j_l , and I : $P^{(l)}(i_l|I, j_l)$
4. $A4$: the next population jump is defined by the individual jump which occurs the first

$$Q_{\mathcal{F}_n(I), J_l}(\tau) = P(\min_l \{R_n^{(l)}\} = R_n^{(l)}, R_n^{(l)} \leq \tau, X_{m_n+1}^{(l)} = j_l | \mathcal{F}_n(I))$$





Proposition. Let $I \rightarrow J_l : i_l \rightarrow j_l$. Then

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$$dF_{\mathcal{F}_n(I), J_l}(\tau) \stackrel{\text{def.}}{=} \frac{dQ_{\mathcal{F}_n(I), J_l}(\tau)}{P(\mathcal{F}_n(I), J_l)} = \frac{\int_0^\tau \Pi_{l' \neq l} (1 - \sum_{j_{l'} \in \mathcal{X}_{l'}(I)} Q_{i_{l'}|I, j_{l'}}^{(l')|s_n^{(l')}}(\tau)) dQ_{i_l|I, j_l}^{(l)|s_n^{(l)}}(\tau)}{\int_0^\infty \Pi_{l' \neq l} (1 - \sum_{j_{l'} \in \mathcal{X}_{l'}(I)} Q_{i_{l'}|I, j_{l'}}^{(l')|s_n^{(l')}}(\tau)) dQ_{i_l|I, j_l}^{(l)|s_n^{(l)}}(\tau)} \quad (\text{cdf of } \Delta \mathcal{T}_{n+1})$$

$$P(\mathcal{F}_n(I), J_l) \stackrel{\text{def.}}{=} \int_0^\infty dQ_{\mathcal{F}_n(I), J_l}(\tau) = \int_0^\infty \Pi_{l' \neq l} (1 - \sum_{j_{l'} \in \mathcal{X}_{l'}(I)} Q_{i_{l'}|I, j_{l'}}^{(l')|s_n^{(l')}}(\tau)) dQ_{i_l|I, j_l}^{(l)|s_n^{(l)}}(\tau),$$

$$Q_{i_l|I, j_l}^{(l)|s_n^{(l)}}(\tau) = \frac{F_{i_l|I, j_l}^{(l)}(s_n^{(l)} + \tau) - F_{i_l|I, j_l}^{(l)}(s_n^{(l)})}{1 - F_{i_l|I, j_l}^{(l)}(s_n^{(l)})} P^{(l)}(i_l|I, j_l) : \text{individual prior kernel}$$

Consequence : $Q_{\mathcal{F}_n(I), J_l}(\cdot)$ depends only on the current state I , the next state J_l and $\{s_n^{(l)}\}_l$

Time-driven simulation algorithm $\iff \lambda_{\mathcal{F}_n(I), J}(\cdot)$

Event-driven simulation algorithm $\iff Q_{\mathcal{F}_n(I), J}(\cdot) = F_{\mathcal{F}_n(I), J}(\cdot)P(\mathcal{F}_n(I), J)$

Determine the next jump (t_{n+1}, J) from (t_n, I) and $\{s_n^{(l)}\}$, $I = (i_1, \dots, i_l, \dots, i_N)$: for each individual l ,

1. choose j_l according to $\{P^{(l)}(i_l|I, j_l)\}_{j_l}$
2. simulate a remaining waiting time $r_n^{(l)}$ in i_l before jumping into j_l

Then $r_n^{(l)} = \min_{l'} \{r_n^{(l')}\}$ defines the next jump time and the next state $J_l = (i_1, \dots, j_l, \dots, i_N)$

Transition rates \Longleftrightarrow kernels

$$\lambda_{\mathcal{F}_n(I),J}(\tau) \stackrel{\text{def.}}{=} \lim_{\Delta\tau \rightarrow 0} \frac{P(\Delta\mathcal{T}_{n+1} \in (\tau, \tau + \Delta\tau), \mathcal{X}_{n+1} = J | \mathcal{F}_n(I), \Delta\mathcal{T}_{n+1} > \tau)}{\Delta\tau}$$

Proposition

$$\begin{aligned}\lambda_{\mathcal{F}_n(I),J}(\tau) &= \frac{\dot{Q}_{\mathcal{F}_n(I),J}(\tau)}{1 - \sum_J Q_{\mathcal{F}_n(I),J}(\tau)}, \tau \in \mathbb{R}^+ \\ Q_{\mathcal{F}_n(I),J}(\tau) &= \int_0^\tau \lambda_{\mathcal{F}_n(I),J}(u) \exp\left(-\int_0^u \sum_J \lambda_{\mathcal{F}_n(I),J}(s) ds\right) du\end{aligned}$$

Corollary. Assume $(Exp) : F_{i_l|I,j_l}^{(l)}(\tau) = 1 - \exp(-\lambda_{i_l|I} \tau)$

Then the SSMP is a MP, and for all I not absorbing

$$dF_{I,J_l}(\tau) = dF_I(\tau) = \left(\sum_{l'} \lambda_{i_{l'}|I} \right) \exp\left(-\sum_{l'} \lambda_{i_{l'}|I} \tau\right) d\tau$$

$$P(I, J_l) = P^{(l)}(i_l|I, j_l) \frac{\lambda_{i_l|I}}{\sum_{l'} \lambda_{i_{l'}|I}}.$$

$$\lambda_{I,J_l}(\tau) = \lambda_{i_l|I} P^{(l)}(i_l|I, j_l) = \lambda_{i_l|I, j_l}$$

$$\lambda_I(\tau) = \sum_J \lambda_{I,J}(\tau) = \sum_l \lambda_{i_l|I}$$

$$\mathbf{P}(t) = \exp(\Lambda t)$$

Consequence. Under (Exp) , if I is not an absorbing state, then

$$m_I = \left[\sum_{l'} \lambda_{i_{l'}|I} \right]^{-1} \text{ (mean time in } I \text{)}$$

Marginal probability law of $\{\mathcal{X}_t\}_t$: renewal equations

$$\begin{aligned}
 P(\mathcal{X}_t = J | I_0, \{s_0^{(l)}\}, t_0) &= P(\Delta\mathcal{T}_1 > t - t_0 | I_0, \{s_0^{(l)}\}) 1_{\{J=I_0\}} + \\
 &\quad \sum_{I_1 \neq I_0} \int_{t_1 \in (t_0, t)} dP(\mathcal{X}_1 = I_1, \Delta\mathcal{T}_1 = t_1 - t_0 | I_0, \{s_0^{(l)}\}) P(\mathcal{X}_t = J | I_1, \{s_1^{(l)}\}, t_1) \\
 &= P(\Delta\mathcal{T}_1 > t - t_0 | I_0, \{s_0^{(l)}\}) 1_{\{J=I_0\}} + \\
 &\quad \sum_{I_1 \neq I_0} \int_{t_1 \in (t_0, t)} dP((\mathcal{X}_1, S_1) = (I_1, \{s_1^{(l)}\}), \Delta\mathcal{T}_1 = t_1 - t_0 | I_0, \{s_0^{(l)}\}) P(\mathcal{X}_t = J | I_1, \{s_1^{(l)}\}, t_1) \\
 \mathbf{P} &= (\mathbf{I} - \mathbf{Q}^\Sigma) + \mathbf{Q}^\mathcal{Y} * \mathbf{P}; \quad \mathcal{Y} = (\mathcal{X}, S) \\
 \mathbf{P} &= \sum_{n \geq 0} \mathbf{Q}^{\mathcal{Y}*n} * (\mathbf{I} - \mathbf{Q}^\Sigma)
 \end{aligned}$$

$$\implies \text{Approximate solution : } \mathbf{P}(t) = [\sum_{n \geq 0}^{n_t} \mathbf{Q}^{\mathcal{Y}*n} * (\mathbf{I} - \mathbf{Q}^\Sigma)](t)$$

Approximate solution : discretization of time

$$s_h^{(l)} = [s_{h-1}^{(l)} + t_h - t_{h-1}] 1_{\{t_h \notin \{T_m^{(l)}\}_m\}} \implies \{s_h^{(l)}\}_l \stackrel{not.}{=} \Delta_{t_h - t_{h-1}}; \{s_0^{(l)}\}_l \stackrel{not.}{=} \Delta_0$$

$$\mathbf{P}_{\Delta_0}(t - t_0) = \mathbf{I} - \mathbf{Q}_{\Delta_0}^\Sigma(t - t_0) + \int_{t_1 \in (t_0, t)} d\mathbf{Q}_{\Delta_0}(t_1 - t_0) \mathbf{P}_{\Delta_{t_1 - t_0}}(t - t_1).$$

The discretization of the system using $t - t_0 = nh$, $t_1 - t_0 \in \{ih\}_{i \leq n}$, leads to the solution

$$\begin{pmatrix} \mathbf{P}_{\Delta_0}(nh) \\ \mathbf{P}_{\Delta_h}((n-1)h) \\ \dots \\ \mathbf{P}_{\Delta_{(n-1)h}}(h) \end{pmatrix} = \begin{pmatrix} \mathbf{R}_{\Delta_0}(0) & \mathbf{R}_{\Delta_0}(h) \dots \mathbf{R}_{\Delta_0}((n-1)h) \\ 0 & \mathbf{R}_{\Delta_h}(0) \dots \mathbf{R}_{\Delta_h}((n-2)h) \\ \cdot & \cdot \dots \cdot \\ 0 & 0 \dots \dots \mathbf{R}_{\Delta_{(n-1)h}}(0) \end{pmatrix}^{-1} \begin{pmatrix} \mathbf{B}_{\Delta_0, n} \\ \mathbf{B}_{\Delta_h, n} \\ \dots \\ \mathbf{B}_{\Delta_{(n-1)h}, n} \end{pmatrix}$$

$\mathbf{R}_{\Delta}(ih) = \mathbf{I} \delta_{0,i} - a_i \dot{\mathbf{Q}}_{0,\Delta}(ih)(1 - \delta_{0,i})$, $\delta_{0,i} = 1$ when $i = 0$ (and is 0 otherwise), $i = 0, \dots, n-1$

$\mathbf{B}_{\Delta_{jh}, n} = \mathbf{I} - \mathbf{Q}_{\Delta_{jh}}^\Sigma((n-j)h)$, $j = 0, \dots, n-1$

$\{a_i\}_i$ depends on the numerical integration scheme,

Semi-semi-Markovian branching process for individuals with a pregnancy period

$\mathcal{X}_t(\Omega)$ takes values in $\mathcal{X} = \{\{(P_l)\}_{l \in \mathcal{L}}\}_{\mathcal{L}}$, where $P_l \in \mathcal{P} = \{pregnant, not\ pregnant, R\}$
 $l \in \mathcal{L}$, l : (date of birth, number u of the individual among the individuals born at this date)

$$\mathcal{X}_t(\Omega) \stackrel{def.}{=} \mathcal{X}_{n_t}(\Omega) \tag{1}$$

$$n_t(\Omega) \stackrel{def.}{=} \sum_{l \in \mathcal{L}_{n_t-1}(\Omega)} m_{p,l,t} \tag{2}$$

$$m_{p,l,t} \stackrel{def.}{=} \sup\{m : T_m^{(p,l)} \leq t\}, l \in \mathcal{L}_{n_t-1}(\Omega) \tag{3}$$

$$\mathcal{X}_{n_t}(\Omega) \stackrel{def.}{=} \{X_{m_{p,l,t}}^{(B,l)}\}_{l \in \mathcal{L}_{n_t-1}(\Omega)} \tag{4}$$

$$X_{m_{p,l,t}}^{(B,l)} \stackrel{def.}{=} \{X_{m_{p,l,t}}^{(p,l)} \neq R, \{X_0^{(p,l')} \neq R\}_{l' \in \tilde{Y}_{n_t,l}}\} \tag{5}$$

$$T_0^{(p,l')} \stackrel{def.}{=} T_{m_{p,l,t}}^{(p,l)}, l' \in \tilde{Y}_{n_t,l}, l \in \mathcal{L}_{n_t-1}(\Omega) \tag{6}$$

$$\mathcal{T}_{n_t}(\Omega) \stackrel{def.}{=} \sup_{l \in \mathcal{L}_{n_t-1}(\Omega)} \{T_{m_{p,l,t}}^{(p,l)}\}. \tag{7}$$

$$\mathcal{L}_{n_t}(\Omega) \stackrel{def.}{=} \{l, \{labels\{\tilde{Y}_{n_t,l}\}\}\}_{l \in \mathcal{L}_{n_t-1}(\Omega)} \tag{8}$$

Example : if $i_l^p \rightarrow j_l^B : \text{pregnant} \rightarrow \text{not pregnant}$, then

$F_{i_l^p|I,j_l^B}^{(B,l)}(\cdot) = F_{\text{pregnant},\text{not pregnant}}^{(B,l)}(\cdot)$ (cdf of the pregnancy period),

$P^{(B,l)}(i_l^p|I, j_l^B)$ is the probability for l to give birth to $\tilde{Y}_{n,l}$ newborns at his next “jump” among the states $\{\text{alive with } \tilde{Y} \text{ newborns}\}_{\tilde{Y}, R}$

Spread of a disease in a branching population structured in groups

(4), (5), (6) replaced by

$$\begin{aligned}
 \mathcal{X}_{n_t}(\Omega) &\stackrel{\text{def.}}{=} \{(X_{m_{p,l,t}}^{(B,l)}, X_{m_{h,l,t}}^{(h,l)}, X_{m_{g,l,t}}^{(g,l)}) 1_{\{X_{m_{g,l,t}}^{(g,l)} \neq R\}}\}_{l \in \mathcal{L}_{n_t-1}(\Omega)} \\
 X_{m_{p,l,t}}^{(B,l)} &\stackrel{\text{def.}}{=} \{X_{m_{p,l,t}}^{(p,l)}, \{(X_0^{(p,l')}, X_0^{(h,l')}, X_0^{(g,l')}) 1_{\{X_0^{(g,l')} \neq R\}}\}_{l' \in \tilde{Y}_{n_t,l}}\}, \quad l \in \mathcal{L}_{n_t-1}(\Omega) \\
 T_0^{(c,l')} &\stackrel{\text{def.}}{=} T_{m_{p,l,t}}^{(p,l)}, \quad l' \in \tilde{Y}_{n_t,l}, \quad l \in \mathcal{L}_{n_t-1}(\Omega), \quad c \in \{p, h, g\}
 \end{aligned}$$

Conclusion

- Individual based models : empirical distributions based on individual simulated trajectories
- Population process : kernel, simulation algorithm, probability law, approximated probability law, asymptotic behavior ?

THANK YOU FOR YOUR ATTENTION!