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1 **Multifractal and Multiscale Entropy scaling of in-situ soil moisture**
2 **time series: study of SMOSMANIA network data, southwestern**
3 **France**

4

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9

10

11 **1 Introduction**

12 Soil moisture is a key variable in hydrology, meteorology and biosphere science. It is the result of a
13 strongly heterogeneous atmospheric forcing, i.e. rainfall, and of the interaction of other processes
14 like evaporation, transpiration, or infiltration. Consequently, soil moisture is subject to strong and
15 possibly quick variations and is heterogeneous in space and time. This results in potential difficulties
16 in sampling strategies since sensors have their own limitations in terms of observable scales
17 (Vereecken et al., 2014). For instance, operating passive L-band radiometers like those of SMOS (Kerr
18 et al., 2010; 2016) and SMAP (Entekhabi et al., 2010) missions can provide continental daily coverage
19 of soil moisture variability but their spatial resolution is limited to pixels of about 50x50 km².
20 Airborne missions can provide more details about spatial variability but data are available over
21 limited areas and a very limited amount of time (i.e. when the mission is operated). Conversely, in-
22 situ sensors have the ability to give a very local estimate of soil moisture over small (hourly or sub-
23 hourly) time steps but their observations are difficult to upscale or to extrapolate to other locations.
24 This resolution mismatch may be somewhat reduced by applying downscaling methods to satellite
25 data (e.g., Merlin et al., 2008; 2011), but the problem is still not fully solved since downscaled
26 products have still coarse (kilometric) resolutions in comparison to the small spatial domain seen by
27 an in-situ sensor.

28 An alternative approach is to try to characterize the spatial and temporal variability of soil moisture
29 with intrinsically multiscale statistical formalisms. The idea behind such formalisms relies on the fact
30 that a complex, heterogeneous signal or random process may hide remarkable symmetries. In this
31 context, we may look for statistical estimators that could be related to the process resolution by
32 simple analytical laws. Such scaling properties are accurately followed by various geophysical
33 variables (Lovejoy, Schertzer et al., 2008; Lovejoy and Schertzer, 2006, 2010; Gagnon et al., 2006;
34 Tessier et al., 1993). In many cases, these properties are even related to the concepts of scale
35 invariance and fractality (some of these concepts will be defined in Section 2).

36 In geophysics, the characterization of scaling laws may be useful for several kinds of applications such
37 as: (i) comparing the variability of measurements collected by sensors operating at different
38 resolutions, (ii) evaluating a numerical model (such as a General Circulation Model or a Land Surface
39 Model) that should be able to reproduce some scaling features of the observations (e.g., Lovejoy et
40 al., 2013; Verrier et al., 2014), (iii) constraining statistical downscaling methods (Gires et al., 2012) to
41 respect the observed scaling laws.

42 In the context of soil moisture, existing studies already evidenced the existence of fractal-like scaling
43 in space and in time. In particular, remote sensing data have been used by several authors in order to
44 investigate fractal properties across space scales. Kim and Barros (2002a, 2002b) analyzed airborne
45 soil moisture estimates from the Southern Great Plains 1997 experiment and proposed a
46 monofractal approach for downscaling soil moisture from 10 km to 1km resolution. Multifractal
47 variants of the latter approach have been proposed later by Mascaro et al. (Mascaro and Vivoni,
48 2010, 2012; Mascaro et al., 2010, 2011) based on (partly similar) airborne remote sensing data sets.
49 Several analyses have also been performed on satellite images. For instance, Lovejoy, Tarquis et al.
50 (2008) have shown that both MODIS radiances and surface soil moisture index exhibit scaling and
51 multifractal properties (based on a case study over central Spain).

52 In the time domain, Katul et al. (2007) also identified the existence of spectral scaling properties
53 based on relatively long-term in-situ acquisitions (8 years of hourly data) at one specific station. The
54 authors also proposed a physical approach to justify the existence of these properties, especially
55 taking into account the integrative nature of soil water content with respect to its meteorological
56 forcing.

57 In the following, scaling analysis tools are applied to surface soil moisture time series collected by a
58 network of probes, SMOSMANIA. Overall, the SMOSMANIA network consists of about 20 stations
59 located in South-Western and South-Eastern France. Capacitive measurements of hourly soil
60 moisture have been done continuously during several years at stations belonging to the Météo-

61 France operational observation network. This makes these measurements suitable for scaling
62 analyses due to the wide range of time scales covered by the data. Moreover, the network design is
63 based on a compromise between a large spatial coverage and a moderate distance between probes.
64 Finally, all probes are located at places with relatively normalized conditions (flat areas, fallow
65 vegetation, low or moderate altitude) since they are co-located with operational meteorological
66 stations operated by Météo-France (Calvet et al., 2007; Albergel et al., 2010).

67 In this paper, a scaling analysis of eight-year long SMOSMANIA time series is proposed, based on
68 spectral and multifractal analysis tools. The objective will be to try to identify whether scaling
69 properties are present and over which scale range. Scaling will also be investigated in terms of
70 complexity analysis by computing the multiscale entropy (MSE) of the different time series.

71 The rest of the paper is structured as follows. In Section 2, theoretical notions related to spectral
72 scaling, multifractal scaling, and MSE are recalled. Then the dataset is presented in Section 3.
73 Spectral analysis results are then presented in Section 4, followed by the multifractal analysis of the
74 dataset in Section 5. MSE analysis of the data is detailed in Section 6. Then, in the subsequent Section
75 7, the expected relationship between MSE and multifractal scaling parameters is investigated with
76 the help of numerical simulations and discussed. Finally, I conclude in Section 8.

77

78 **2 Theoretical background**

79 **2.1 Spectral scaling**

80 Scale invariance consists of a family of generic properties exhibited by many geophysical datasets
81 (time series, 2D/3D fields). The concept is closely related to the concept of fractality that has been
82 investigated and popularized by B. Mandelbrot (e.g., Mandelbrot, 1983). In the context of stochastic
83 modeling of geophysical processes, scaling may be typically detected by investigating the scale
84 dependence of some classical or less classical statistical indicators that often vary as a power-law of
85 the time (or spatial) resolution or lag. The exponents of such power-laws can therefore characterize a
86 physical process over given range of scales. For instance, the approaches proposed by Kolmogorov
87 (1941) for turbulence or by Hurst (1951) in hydrological science have led to a vast scientific literature.

88 Among various statistical tools that can be used to investigate scaling properties, spectral analysis
89 offers a relatively simple and convenient way to study the distribution of energy across a wide range
90 of temporal or spatial frequencies. In particular, power spectral densities have been used as a tool to
91 investigate the temporal scaling properties of various hydrologic variables such as rainfall (Fraedrich

92 and Larnder, 1993; Olsson, 1995 ; Verrier et al., 2011), river flows (Pandey et al., 1998) or soil
93 moisture (Katul et al., 2007).

94 In the spectral sense, a time series is scaling if its power spectral density $E(f)$ is proportional to a
95 power law of the form:

96
$$E(f) \sim f^{-\beta} \quad (\text{Eq.1})$$

97 where f is the frequency and \sim denotes equality within the limits of slowly varying factors that do not
98 affect the main scaling behavior. (Eq. 1) is valid for frequencies f comprised between two limit
99 frequencies defining the limits of the scaling range in which the scaling properties have been found.
100 For some processes like rainfall, several distinct scaling ranges may be found in the series, with
101 possible meteorological interpretation (e.g., Fraedrich and Larnder, 1993). In the case of soil
102 moisture, Katul et al. (2007) used theoretical and empirical arguments to separate two ranges, one
103 with a steep spectrum (behaving like $1/f^2$ spectrum or steeper) at high frequencies (equivalent to
104 time scales from hours to about one hundred hours) and another one with a flatter spectrum at
105 lower frequencies (i.e. larger time scales).

106

107 **2.2 Multifractals and cascades**

108 Multifractality may be viewed as a generalization of classical scaling properties to systems that do
109 not exhibit a unique scaling law per scaling range but a whole set of scaling laws with specific
110 parameters that nonlinearly depend on the order of statistical moments used in analysis (Schertzer
111 et al., 2002). Physically, it is interesting to investigate different orders of statistical moments since
112 they are representative of various levels of intensities of the process (e.g. moderate intensities vs.
113 extreme intensities).

114 Let us denote $\lambda = T/\tau$ the resolution factor consisting of the ratio of the series time length T and a
115 given analysis time scale τ (comprised between the time lag between two measurements and T). In
116 the strict sense, a time series $\Phi(t)$ is said to be multifractal if its statistical moments of real positive
117 orders q follow power-laws of the resolution λ :

118
$$\langle \Phi_\lambda^q \rangle \sim \lambda^{K(q)} \quad (\text{Eq.2})$$

119 In the previous equation, Φ_λ represents a low-resolution version of the original times series, obtained
120 by direct aggregation over disjoint intervals of length $\tau = T/\lambda$. Other approaches could have been

121 used for defining aggregated products (moving averages, wavelets...) but these approaches will not
122 be considered in the following.

123

124 In theoretical terms, stochastic processes following (Eq. 2) may be built by iterative multiplicative
125 procedures called multiplicative cascades (Schertzer et al., 2002). These procedures consist in a
126 series of resolution refinements where a given coarse scale series is modulated by random
127 multiplicative factors at each step (i.e. at each resolution refinement). Resolution refinements can be
128 done by explicitly dividing coarse pixels in smaller subpixels, thus causing a discretization of the
129 scales. Alternatively this construction can be extended to consider a continuum of scales, leading to
130 continuous in scale multiplicative cascades (Schertzer and Lovejoy, 1997). For the latter, random
131 multiplicative factors between two arbitrary scales should follow a log-infinitely divisible distribution.

132 In the general case described by (Eq. 2), the scaling exponent $K(q)$ is a convex function of the
133 moment order q . However, the exact form of $K(q)$ is subject to additional constraints when taking
134 into account the continuity of the scale space and the related constraints such as the distribution of
135 random multiplicative factors described above. Based on these considerations, some authors
136 proposed cascade models where the function $K(q)$ can be expressed in closed form. In particular,
137 two-parameter models have been proposed by Schertzer and Lovejoy (1987) and by She and Levêque
138 (1994).

139 In the following, the parameterization of the $K(q)$ function will be investigated with the help of the
140 Universal Multifractal (UM) model proposed by Schertzer and Lovejoy (1987). The choice of this
141 model was motivated by several factors such as its mathematical pertinence (the model is one of the
142 natural attractors of the class of continuous in scale log-infinitely divisible cascades), its wide use in
143 geophysical literature, and its physically interpretable parameterization.

144 In this framework, the function $K(q)$ is parameterized as follows:

145
$$K(q) = \frac{C_1}{\alpha-1} (q^\alpha - q) \quad (\text{Eq. 3})$$

146 C_1 is a dispersion parameter comprised in the interval $[0,1]$ for time series. For instance $C_1 = 0$ would
147 correspond to a homogeneous time series. Mathematically, C_1 is the fractal co-dimension associated
148 to the set of points exceeding a specific threshold specifically related to the mean value of the field .
149 The parameter α is an index of multifractality and belongs to the $[0, 2]$ interval ($\alpha = 0$ is the pure
150 monofractal case where $K(q)$ becomes linear while $\alpha = 2$ corresponds to log-normal cascades for
151 which scaling exponents $K(q)$ vary in a quadratic way with q).

152

153

154 **2.3 Fractionally integrated cascades**

155 By construction, multiplicative cascades are potentially appropriate for modeling time series whose
156 power spectrum follows a scaling law with a scaling exponent $\beta < 1$. More precisely the scaling
157 exponent should follow the relationship $\beta = 1 - K(2)$ where $K(2)$ is necessarily positive (see, e.g., Tessier
158 et al., 1993, §3.b).

159 However, most geophysical variables are characterized by steeper power spectra ($\beta > 1$) in space
160 and/or time domains. This is typically the case for variables closely related to turbulence (where β is
161 generally close to $5/3$ for both wind speed and passive scalars coherently with the predictions of
162 Kolmogorov (1941), Obukhov (1949), and Corrsin (1951)), but also for many other variables, including
163 topographic fields (Gagnon et al., 2006). Anticipating on Sect. 4, the subsequent analyses of
164 SMOSMANIA soil moisture will show evidence of power spectra that are closer to $1/f^2$ behavior.

165 In the latter cases, multifractality can be investigated by testing the validity of (Eq. 2) for a derivative
166 of the process, called the flux and denoted Φ . This derivative can be a fractional derivative of the
167 original process. It is recalled that fractional derivative and integrations are useful notions of
168 fractional calculus that extend classical definitions of derivatives and integrals to non integer orders
169 (Loverro, 2004). These concepts are useful for modeling processes with various spectral scaling
170 exponents, e.g. integrations can represent effects leading to steeper power spectra (Schertzer and
171 Lovejoy, 1991, appendix B2; Gagnon et al., 2006). In the context of our study, the observable variable
172 Y (e.g., soil moisture) is modeled as the fractional integration of order H of a multifractal flux Φ that
173 follow Eq.2. Combined with UM cascades, this defines the Fractionally Integrated Flux (FIF) model
174 (Schertzer and Lovejoy, 1987) which depends of the three parameters α , C_1 , and H . When H is
175 positive, its effect is that of a smoothing parameter. On the contrary, the case
176 $H < 0$ corresponds to additional roughening effects. The case $H = 0$ corresponds to multiplicative
177 cascades described in the previous paragraph (i.e. without any additional smoothing or roughening).

178 In practice, this parameter H may also be estimated by computing the first-order structure function
179 $S(\Delta t)$ defined by the first-order absolute increments of the time series as a function of the time lag
180 Δt :

$$181 \quad S(\Delta t) \stackrel{\text{def}}{=} \langle |Y(t + \Delta t) - Y(t)| \rangle \quad (\text{Eq.4})$$

182 In the case of a FIF model, $S(\Delta t)$ follows a scaling law similar to the one exhibited by fractional
183 Brownian motions processes:

184
$$S(\Delta t) \sim \Delta t^H \quad (\text{Eq. 5})$$

185 where \sim again means equality within the limits of slowly varying factors. It should be mentioned that
186 the definitions above can be extended by using other definitions of fluctuations in time series. For
187 instance, Lovejoy and Schertzer (2012) have proposed a variant based on the replacement of
188 difference fluctuations used in Eq. 4 by Haar fluctuations. Interestingly, such variants may differ from
189 each other in terms of range of validity of H estimates. For instance, structure functions estimates
190 based on Eq. 4 will allow the retrieval of H only when $0 < H < 1$. The use of Haar fluctuations can lead
191 to correct estimates on a wider range of H values ($-1 < H < 1$). In the rest of the paper, analyses will
192 nevertheless be based on the more classical structure functions defined by (Eq. 4). This is not a
193 limitation here since H estimates will be found near the middle of the $[0, 1]$ interval (see section 6.2).

194 The value of H directly affects the distribution of energy across scales: namely, a higher value of H
195 implies a steeper spectrum. For a FIF series, both values are related by $\beta = 1 - K(2) + 2H$ (Tessier et al.,
196 1993). Note that for a monofractal signal such as a fractional Brownian motion (characterized by a
197 unique scaling exponent H), a more or less similar relationship ($\beta = 1 + 2H$) still holds. The case of FIF
198 processes differs from the fBm case by the multifractal correction term $K(2)$.

199

200 **2.4 Multiscale entropy**

201 Other signatures of scale-invariance may be found by using analysis methods more closely related to
202 information theory. For instance, the notion of information dimension has been defined to
203 characterize the density of points in fractal geometric sets and strange attractors (e.g., Hentschel and
204 Procaccia, 1983; Farmer et al., 1983). Zhang (1991) suggested computing Shannon entropies of
205 coarse-grained time series as a prerequisite to define a complexity measure of a signal. A similar
206 approach has been proposed by Costa et al. (2002, 2003, 2005) in the context of the analysis of
207 biological signals. Their approach, called Multiscale Entropy (MSE) is based on the estimation of the
208 sample entropy (SampEn) (Richman and Moorman, 2000) applied on time series that are coarse-
209 grained at various time resolutions. The sample entropy provides an approximation of the entropy
210 information of any discrete time series $y[i]$ of length N by computing the conditional probability that
211 two similar sequences of m consecutive points will remain similar when considering extended
212 sequences of $m+1$ points.

213 If $u_m[i]$ denotes a sequence (vector) of m points starting at index i , i.e. $u_m[i] = \{y[i], \dots, y[i + m -$
 214 $1]\}$, and d is a given distance between vectors (here taken as the maximal absolute distance between
 215 their components), then we may define the probability of similarity between two different vectors
 216 with respect to a threshold T :

$$217 \quad A(m, T, N) = \Pr(d(u_m[i], u_m[j]) \leq T) \quad (j \neq i) \quad (\text{Eq. 6})$$

218 Then the sample entropy may be computed by estimating the negative logarithm of the conditional
 219 probability that two matching vectors of length m will still match while taking into account an
 220 additional point within all vectors (Costa et al., 2005):

$$221 \quad S_E(m, T, N) = -\ln\left(\frac{A(m + 1, T, N)}{A(m, T, n)}\right) \quad (\text{Eq. 7})$$

222 Due to the negative logarithm, the sample entropy tends to increase when “unexpected” changes in
 223 the series tend to occur, i.e. when they are less repetitive and contain more information. In the
 224 following, the parameters m and T will be fixed to values close to those used in the literature, i.e.
 225 $m = 2$ and T taken as 15% of the standard deviation of the time series y .

226 Within this context, the Multiscale Entropy is a procedure consisting at estimating the sample
 227 entropy not only from given a data time series $Y[i]$ but also from coarse-grained time series $Y^{(\tau)}[i]$
 228 that may be estimated by averaging Y over non-overlapping windows of length τ . This provides an
 229 entropy estimate $MSE(\tau)$ that depends on the time resolution τ .

230 Recent results tend to prove that scale invariance has a specific signature on the Multiscale Entropy:
 231 as recalled by Nogueira (2017), $MSE(\tau)$ should scale as a power-law with exponents closely related
 232 with the spectral slope., i.e.

$$233 \quad MSE(\tau) \sim \tau^{H'} \quad (\text{Eq. 8})$$

234 This has been verified empirically in the case of colored noise/fractional Brownian motion (fBm) time
 235 series (Courtiol et al., 2016). Analytical arguments based on fBm theory have also been proposed by
 236 Gao et al. (2015) to understand this relationship. These authors also confirmed their findings on
 237 empirical studies of biological signals.

238 Nogueira (2017) compared the scaling laws of power spectra and MSE approach for both geophysical
 239 data (near-surface wind) and synthetic data (based on monofractal fBm series). In the case of
 240 monofractal fBm signals, the results were tending to confirm that the MSE scaling exponent H'
 241 should be identical to the scaling exponent H determined by the first-order structure function. On

242 the contrary, the case of multifractal signals has not been explicitly analyzed. It may be noticed that
243 Nogueira's empirical results on turbulent (i.e., likely multifractal) wind processes suggest that a
244 scaling law if the form of Eq. 8 could still hold for multifractal signals.

245 In the following, both MSE tools and multifractal analysis tools will be applied to the same
246 SMOSMANIA surface soil moisture time series. This will provide two types of multiscale information
247 that could eventually be compared.

248

249

250

251 **3 Data sets**

252 SMOSMANIA consists of a network of about 20 measurement stations distributed across
253 Southwestern and Southeastern France in grassland and agricultural areas (Calvet et al., 2007;
254 Albergel et al., 2010, 2011). This network has been set up in relation with the much broader HyMeX
255 campaign (<http://www.hymex.org>) which was mainly devoted to the monitoring and modeling of
256 water cycle within the Mediterranean region. The SMOSMANIA probes are located at existing
257 operational Météo-France weather stations (RADOME network). The average distance between two
258 SMOSMANIA stations is of the order of 45 km (Albergel et al., 2010). These locations may be viewed
259 at ISMN website at the following address: www.geo.tuwien.ac.at/insitu/data_viewer/ISMN.php. All
260 stations are located at places with a fallow vegetation cover.

261 The SMOSMANIA network provides operational measurements of volumetric soil moisture and soil
262 temperatures at various depths. Concerning the former, soil moisture is estimated on a hourly basis
263 with the help of Thetaprobe sensors based on a capacitive technology. Measurements are available
264 at four references depths ranging from 5 cm to 30 cm. All the time series may be downloaded for
265 free at the International Soil Moisture Network (ISMN) page ([http://ismn.geo.tuwien.ac.at/data-](http://ismn.geo.tuwien.ac.at/data-access/)
266 [access/](http://ismn.geo.tuwien.ac.at/data-access/)). For more information about ISMN, the reader is also referred to the presentation by Dorigo
267 et al. (2011).

268 The present study is focused on time series located in Southwestern France, i.e. on time series of soil
269 moisture collected at 12 places reported in Table 1. Since scaling-related analyses must generally rely
270 on data that have a small number of missing data, time series presenting long missing periods of
271 several dozens of days have been discarded. We therefore restricted the study to nine of the twelve
272 initial series (discarded series are marked with a star in Table 1). All remaining missing data have
273 been replaced by linear interpolation.

274 Several differences between locations should be noted. For instance, Urgons, Sabres, and Créon
275 d'Armagnac are located closer to the Atlantic ocean and therefore are subject to a more oceanic
276 climate. They also differ from most of the other stations by a different soil texture with a much larger
277 sand proportion. Stations that are closer to the Mediterranean Sea (Narbonne, Lézignan-Corbières)
278 typically have a silt loam texture. Finally, stations that are distant from both seas tend also to have
279 the largest clay fraction (Condom, Peyrusse Grande, Lahas). We can notice that all stations (except
280 Mouthoumet station) are located at low altitudes.

281 The present study is focused on the 2007-2014 period that covers most of the available time series
282 from ISMN. Thus, all soil moisture measurements comprised between 2007/01/07 6:00 and
283 2015/01/01 0:00 (local time) have been selected in order to cover approximately 8 years.
284 Measurements before 2007/01/07 were only available at a few stations and thus were discarded.
285 Since all sensors provide measurements on an hourly basis, all selected time series have about
286 70,000 data points which provides the possibility of investigating potential scaling properties over
287 four orders of magnitude in scale.

288 Finally, some data at the beginning of the data series have been removed in order to reduce all the
289 time series to 65536 measurements points each. All the retained series still cover the same dates,
290 ending on 2015/01/01 0:00. The purpose of this restriction is to work with 2^n length time series
291 (here $n = 16$) which is preferable for technical reasons. Indeed, multifractal analysis requires the
292 estimation of statistical moments at different aggregated resolutions and the use of a dyadic cascade
293 (i.e. 2^i resolutions with integer i) is very convenient for this. Moreover, FFT-based spectral analysis
294 techniques are also faster when applied to series with 2^n data points.

295

296

297 **4 Spectral analysis of SMOSMANIA series**

298 **4.1 Surface soil moisture data (5 cm depth)**

299 Spectral analysis has been carried out on the data set of soil moistures measured at 5 cm depth in
300 order to identify scaling properties and possible scaling ranges. Since scaling properties generally
301 take the form of a power-law spectrum, i.e. $E(f) \sim f^{-\beta}$, the results are presented in logarithmic
302 coordinates. In such a graph, scaling laws should indeed appear as linear. However, this
303 representation may have less desirable counterparts when analyzing long time series. In particular,
304 this usually results in a much denser sampling of frequencies near the high frequencies and often in a
305 less readable graphic representation. To avoid these problems, an average of the spectra over
306 equally spaced logarithmic bins (of length 0.02 here) was performed.

307 The power spectra of the nine selected series are shown on Figure 1. The figure is presented in
308 logarithmic coordinates. In order to make it more readable, the spectra are translated upwards from
309 one station to the next one. The spectrum of the Condom series is the only one that has not been
310 vertically translated. Time scale tick labels are displayed on the horizontal axis (i.e. high frequencies

311 are still on the right of the figure but the horizontal graduations specify the time scale equivalent to
312 this frequency, i.e. $\tau = 1/f$.

313

314 The main feature exhibited by the spectra on Figure 1 is the presence of scaling laws over a large
315 scale range covering a large range of scales. All spectra seem nevertheless slightly steeper at the high
316 frequencies than at low frequencies. A more oscillatory behavior at low frequencies (e.g. time scales
317 larger than several thousands of days) may be noticed. This feature is probably related to the coarser
318 sampling of frequencies when analyzing large scale structures over time series that is not much
319 longer than these structures.

320 It may also be noticed that all spectra appear roughly parallel to each other which suggests that all
321 spectra should have very similar scaling laws with exponents β that are not expected to vary much.
322 This may be checked by calculating these scaling exponents.

323 While the spectra look scaling over three orders of magnitude at first glance, several additional
324 factors must be taken into account to determine the scale range considered in the fits: (i) the low-
325 frequency part of the spectrum has a poorer sampling and is less robustly estimated, (ii) the
326 application of a segmentation algorithm (D'Errico, 2017) has been used to identify possible
327 transitions between ranges with different spectral exponents β , but failed to notice very significant
328 transitions, (iii) however, some spectral estimations (for instance at Lahas station) exhibit a very local
329 drop at time scales of the order of 1000-3000 h, (iv) the retained fit range must permit linear
330 regressions with a good coefficient of determination, for example $R^2 > 0.98$.

331 Taking these elements into account and also anticipating on Sect. 5 where a narrower scaling range
332 will be found, the estimation of scaling exponents has been limited to the range of scales 1000h – 1h.
333 The linear regressions are shown on Figure 1 and the scaling coefficients β are reported in Table 2
334 (for the nine soil moisture time series collected at depth 5 cm). It may be noticed that this choice
335 permits one to obtain a very good coefficient of determination for each of the nine spectra
336 presented. All spectral slopes β are found within or very close to the range [1.9, 2] and thus are very
337 similar to each other when considering that each exponent is given with an uncertainty of ± 0.04 . This
338 confirms the apparent parallelism of the spectra on the figure.

339 **4.2 Other depths**

340 By repeating the above procedure for the three other depths, similar results were obtained with
341 scaling holding to a good approximation for scales smaller than 1000 h (all error bars $\Delta\beta$ were found

342 smaller than 0.04 for a linear fit within this range). For all time series, the surface (5 cm) data
343 spectrum remains above the other spectra, meaning that the surface series contains more energy (in
344 the signal processing sense) than the other ones. Physically, this is also coherent with the greater
345 variability amplitude of soil moisture at the surface due to the meteorological forcing.

346 However, significant differences have been found between stations with typical cases illustrated on
347 Figures 2a and 2b. At Peyrusse Grande station (Fig. 2a), the spectra are close to each other and the
348 spectral slopes do not change considerably with depth (here from 1.93 to 2.08) while at Montaut the
349 spectra are clearly distinct with different spectral slopes (ranging from 1.97 to 2.35). It may be noted
350 that the spectra displayed on Figure 2 do not include any vertical shifts between spectra.

351 The spectral slopes β estimated for all stations at all considered depths are summarized on Figure 3.
352 The estimates of β are represented on the vertical axis while the stations names are distributed on
353 the abscissa axis (the stations have been sorted from the most western position to the most eastern
354 position). More precisely, the abscissa axis refers to the three first letters of the names of the
355 stations. Error bars are omitted but are of the order of ± 0.03 or 0.04 for all estimates.

356 As illustrated on Figure 2 and Figure 3, it seems that the spectral slope tends to increase as we
357 consider deeper layers. However, β estimates become more variable across stations as the depth
358 increases. This could possibly be related to the differences in subsurface processes and soil texture,
359 the latter being heterogeneous across the stations.

360 While similar studies in the region of interest do not seem to be reported in the literature, several
361 published studies may help to put our findings in perspective. In the space domain, lower scaling
362 exponents are generally reported. For instance, Pelletier et al. (1997) found a value of $\beta = 1.8$, while a
363 value closer to 1.2 might be deduced from the findings presented by Kim and Barros (2002a), based
364 on remote sensing data. More recently, Neuhauser et al. (2018) reported a spectral exponent closer
365 to 1.0 at spatial scales > 40 km based on SMOS data.

366 In the time domain, higher estimates of β have been reported by Katul et al. (2007) based on
367 measurements of soil moisture in a more subtropical eastern US region (Durham, NC). These authors
368 considered an eight-year long time series of integrated (0-30 cm) soil moisture. They found a steep
369 scaling $\beta = 2.75$ in the scale range 90 h – 1h and a much flatter spectrum at larger scales.
370 Interestingly, they also proposed a simplified physical model to show that soil moisture time series
371 may behave as a time integration of precipitation time series at relatively small time scales (at least
372 in first approximation). Since power spectra of precipitation data are either flat or decrease in a
373 scaling way (i.e. f^{-a} where $a > 0$, cf. Fraedrich and Larnder, 1993), the expected spectrum of soil

374 moisture could be of the form $f^{2-\alpha}$, which was verified in their case. The results obtained by Katul et
375 al. (2007) therefore suggest that it is not surprising that soil moisture has a steeper scaling in the
376 time domain than in the space domain, due to this time integration effect. However, in the present
377 study estimates of β are found closer to 2.0, i.e. a red-noise like scaling. In this sense, our results may
378 differ from those of Katul et al. (2007) since the model proposed by the latter authors does not seem
379 compatible in practice with values of β smaller than 2.

380

381 **5 Multifractal analysis**

382 **5.1 Surface soil moisture data (5 cm depth)**

383 Results of the previous section show that soil moisture power spectra estimated above are rather
384 steep (similar to a f^{-2} red-noise). This usually means that the series cannot be directly described by a
385 multifractal multiplicative cascade which requires $\beta < 1$. However, as explained in Section 2.3, it is
386 possible to check whether a multifractal cascade structure may describe some derivative (or
387 integration) of the process. Fractionally integrated stochastic processes (e.g. fractional Brownian
388 motions, fractionally integrated flux...) are models that may be appropriate for scaling processes with
389 steeper spectral slopes. Therefore, a possible way to investigate multifractality within the series is to
390 consider a “flux” estimated from a derivative of the time series and to test whether the moments of
391 this flux follow a multifractal law analogous to that described in Section 2.2 (Eq. 2).

392 In the following, the fluxes Φ are estimated at 1h resolution by taking the absolute value of the finite
393 difference derivative of the 5cm moisture series, following the process originally described by
394 Lavallée et al. (1991, 1993). For each series, the flux is aggregated over non-overlapping intervals of
395 length $\tau = 2^i \tau_0$ using a dyadic cascade approach (i.e. at each step the resolution ratio $\lambda = T/\tau$ is
396 divided by a factor two and the data are aggregated on successive couples of points). This leads to an
397 ensemble of aggregated fluxes series $\Phi_\lambda(t_i)$ indexed by the resolution $\lambda = 2^j$ for $j = 0, 1, \dots, 16$ (each
398 aggregated series thus containing λ data points).

399 Then the empirical moments are used as an estimate of the (unknown) statistical moments. For each
400 series, and for each chosen moment order q , the aggregated fluxes are elevated to power q and
401 averaged over the whole series length. The computation is done for all averaging resolutions $\lambda = 2^j$
402 ($j = 0, 1, \dots, 16$) and for various orders q comprised between 0 and 2.5 (with a spacing of 0.1 between
403 two successive orders). This provides the empirical moments $M(\lambda, q)$ that approximate the statistical

404 moments $\langle \Phi_\lambda^q \rangle$, the statistical average being replaced by the empirical average. This procedure
405 corresponds to the Trace Moment approach presented by Tessier et al. (1993).

406 The empirical moments are represented in logarithmic coordinates as a function of the resolution λ .
407 The process is repeated for each series. The result is displayed in a larger figure for the case of
408 Peyrusse Grande data, and in a composite figure for other stations (see Figures 4 and 5).

409 Figure 4 may be read as follows. First, each curve is associated to one specific moment order. The
410 higher moment orders $q > 1$ are associated with the upper, increasing curves on the figure, while the
411 slightly decreasing curves are associated to orders q such that $0 < q < 1$. Similarly to Figure 1, finer
412 resolutions are located on the right of Figure 4.

413 At the coarsest time scales, typically > 1000 h, the curves are undistinguishable meaning that no
414 significant multiscaling may be found. However, moments are much more variable towards finer
415 resolutions. Log-log linear behavior is found to a good approximation within the range of resolutions
416 comprised between 128 h and 1 h (green fit lines). Since the figure is in logarithmic coordinates, each
417 regression is representative of Eq. 2 for a fixed order of moment q and its slope corresponds to the
418 scaling exponent $K(q)$.

419 Comparable figures are found for the eight others time series with very similar scaling ranges
420 (Figure 5). Thus, and in order to facilitate the comparison between the different series, it has been
421 chosen to perform the regression over the same scaling range (128 h – 1 h). It can be noted that this
422 range is smaller than the one observed on power spectra presented in the previous section.

423 Based on Figure 4, we also remark that all linear regressions are convergent towards a point while
424 being extrapolated towards coarser resolutions. The abscissa of this point may correspond to the
425 “external scale” (Lovejoy et al., 2008) at which the multifractal cascade is initiated. For the nine time
426 series, the exact position of the external scale may vary a bit but is comprised in the range 200-600
427 hours. For most of the time series, the external scale would be close to 200-300 h, but for some of
428 them higher values seem obtainable (e.g., Urgons, Narbonne, or Créon d’Armagnac). For the latter
429 station, the external scale seems indeed closer to 600 h. Due to the uncertainty of this estimate, it is
430 not possible to clearly tell if the latter difference is really significant and related to meteorological
431 forcing. Nevertheless, the position of the external scale is plausibly related to the
432 weather/macroweather transition scale that has been evidenced for numerous atmospheric fields
433 (Lovejoy et al., 2013; de Lima and Lovejoy, 2015). The transition between the weather and
434 macroweather regimes occurs at the lifetime of planetary meteorological structures, i.e. at time
435 scales of the order of a couple of weeks. For meteorological variables, the transition separates

436 scaling regime respectively characterized by a small intermittency (i.e. large-scale/macroweather)
437 and a stronger (multifractal) intermittency at smaller scales (i.e. weather regime).

438 As explained above, the slopes of the green lines on Figure 4 provide the estimates of the moment
439 scaling function $K(q)$ for the orders q that have been used in previous computations. The graph of the
440 estimated $K(q)$ function at Peyrusse Grande station is provided on Figure 6. The round dots
441 correspond to the empirical estimates. We may notice that the values of $K(q)$ are also given for
442 higher moment orders (up to 3.0) that have not been represented on Fig. 4 for readability purposes.

443 As expected, the $K(q)$ estimates are distributed along a convex curve. The latter may then be fitted by
444 the two-parameter universal form recalled in Section 3. The green curve is obtained by nonlinear
445 optimization on the interval defined by $q \in [0, 2.5]$. More precisely, the Nelder-Mead method is
446 used to find the couple of parameters that minimizes the (quadratic) distance between fixed $K(q)$
447 estimates and the theoretical values predicted by (Eq. 3). In this case, the universal parameters
448 estimates for Peyrusse Grande station are $C_1 = 0.28$ and $\alpha = 1.51$.

449 For other stations, more or less similar graphs may be obtained but some variability still exists in both
450 estimates of α and C_1 (Fig. 9). All the different $K(q)$ function estimates were fitted within the interval
451 $[0, 2.5]$ (see Figure 5). The obtained parameters are reported in Table 3.

452

453 As shown on Table 3, some variability may be found within parameter estimates. Nevertheless,
454 C_1 remains always close to 0.25-0.3 and the observed variability does not seem related with the
455 positions of the stations and/or soil texture. Such estimates are significantly larger than the values
456 (closer to 0.1 or even less) reported for other geophysical fields (e.g., Lovejoy and Schertzer, 2013).
457 However, some studies report large (> 0.3) estimates in the case of rainfall (e.g., Lovejoy, Schertzer et
458 al., 2008), at least at time scales larger than 1 h (Verrier et al., 2011). Thus, the relatively high value of
459 C_1 that has been observed in this study for soil moisture could be a consequence of the large
460 intermittency of the rainfall process. Concerning the index of multifractality α , most estimates are
461 found within the range 1.6-1.8. In this aspect, the stations of Créon d'Armagnac and Urgons seem to
462 differ significantly from the others with much lower α estimates closer to 1.3-1.4. Interestingly, these
463 stations are relatively close to each other (about 55 km) and both located in the French department
464 of Landes, at the most western part of the region covered by SMOSMANIA, i.e. closer to the Atlantic
465 Ocean and with significantly more sandy soil texture. This could suggest a possible impact of these
466 features on α . However, the parameter α characterizes the degree of nonlinearity of the $K(q)$ curve

467 and is known to be more difficult to estimate than C_1 . This could also explain part of the observed
468 difference despite the relative homogeneity of α estimates for the other stations.

469

470 **5.2 Other depths**

471 By repeating the previous analysis to other depths, a similar multifractal scaling was found in the
472 range 128 h – 1 h for all stations. This is illustrated on Figure 8 where moments of Peyrusse Grande
473 (10, 20, 30 cm) data follow a scaling law that could be extrapolated to external scales of the order of
474 several hundreds of hours. From the slopes of the green lines, scaling exponents $K(q)$ may be
475 computed and then plotted on Figure 8 (bottom right). On this figure, each color is associated to one
476 depth (note that 5 cm depth estimates are also plotted on the same graph). While the different
477 curves globally follow the same shape, they differ significantly from each other for low and high
478 moment orders. In this case, high-order moments are higher for 10 cm depth data than for other
479 data, while 20 cm depth data differ from the others for low moment orders. These differences affect
480 both estimates of multifractal parameters, which can be checked by comparing the couples (α, C_1) at
481 depths 5 cm and 20 cm. Indeed, the latter are respectively equal to (1.51, 0.28) and (0.79, 0.40). Part
482 of the change may be explained by the partial “linearization” of the $K(q)$ curve as the depth increases
483 (the limit case of linearity can theoretically be attained at $\alpha = 0$, i.e. monofractality case).

484 For all time series, the parameters α and C_1 have been estimated at the different depths. The results
485 are summarized on Figures 9a and 9b where estimates are given as a “function” of the station (once
486 again, stations are sorted from the most western to the most eastern and are identified by their first
487 three letters). As deeper data are considered, smaller values of α and larger values of C_1 are
488 systematically obtained, with a globally larger variability from one station to one another. Similar to
489 Fig. 8, this change seems to be related to a less pronounced convexity of the $K(q)$ curves as depth
490 increases.

491 We may conclude that soil moisture at these depths still follows multifractal properties but with
492 different $K(q)$ scaling exponents, especially at low and high order moments. This is illustrated by a
493 change in multifractal parameter estimates that is noticeable when compared to the relative
494 homogeneity of the parameters for 5 cm depth data across different stations. A possible explanation
495 could reside in the impact of soil texture on soil moisture at the deeper layers. For instance,
496 subsurface processes could cause a change of the scaling of some moment orders, this change
497 propagating to the multifractal parameters.

498

499 **6 Multiscale entropy properties**

500 **6.1 MSE analysis**

501 The nine surface (5 cm depth) soil moisture time series are now analyzed within the MSE framework.
502 For each series, sample entropy is estimated at various coarse resolution time series obtained by the
503 means of averages performed over non-overlapping intervals of various durations τ chosen between
504 1 h, i.e. the finest resolution available, and 690 h. We cannot degrade the time series to coarser
505 resolutions since the sample entropy calculation can only be performed over coarse time series
506 containing at least 100 points (Nogueira, 2017). The estimated entropies are represented in
507 logarithmic coordinates on Figure 10 for the different time series. Once again, MSE curves are
508 regularly shifted upwards to make the figure more readable. MSE curves are steeper at lower time
509 scales and a bit flatter at larger scales. The segmentation tool proposed by J. D'Errico (2017) was
510 used to estimate the position of the transition between the two behaviors, with a result close
511 to 30 h. Due to the better sampling of the MSE on the right of the figure, only time scales within the
512 range 650 h – 30h have been considered for fitting. The exponents H' are determined by linear
513 regressions over this interval of time scales. The values of H' estimates and error bars may be found
514 in Table 4. Estimates are comprised between 0.31 and 0.66 for a mean value of $H' = 0.43$ for all time
515 series. All the regressions have a very good coefficient of determination except in the case of the
516 Mouthoumet station.

517 Despite the narrower scale range, the results are partially coherent with the findings previously
518 obtained by spectral analysis, with often similar scaling exponents across stations (see Table 2). Once
519 again, all time series have rather similar scaling properties which could be related to the sampling
520 strategies (all sensors being located in grassland on flat areas). Nevertheless, it could be noted that
521 the dispersion of the values of H' is greater than the dispersion of spectral slopes. This dispersion is
522 nevertheless probably not representative of a better sensitivity of MSE to local differences: indeed,
523 H' is estimated over a significantly narrower scale range than β which could lead to more dispersion
524 in the estimates.

525 Furthermore, the soil moisture series collected at 10 cm, 20 cm and 30 cm depths have been
526 analyzed in a similar way. The scaling exponents H' obtained by fitting the MSE curves are displayed
527 on Figure 11. All these exponents have been obtained by linear fits on the log-log graph of $MSE(\tau)$ for
528 time scales in the range 30h – 650 h (all linear regressions had an R^2 greater than 0.94 except in the
529 case of Peyrusse Grande where the fits are poorer). Figure 11 also confirms that the scaling

530 exponents may depend on depth. Coherently with what was reported on Figure 3, the scaling seems
531 in general a bit steeper at greater depths.

532 The results of this section confirm that the scaling properties of soil moisture may be retrieved not
533 only in terms of spectra and moments, but also in terms of information content. This was expectable
534 since the signal information (occurrence of unexpected structures) is of course strongly constrained
535 by the probability distribution of the signal at various resolutions. Nevertheless, this result shows
536 that multiscale entropy scaling can be found in non-biological natural signals, consistently with the
537 findings by Nogueira (2017). Additionally, our results show that MSE scaling may be found in natural
538 signals that follow multifractal statistics. This could extend the findings of the literature that mainly
539 focuses on monofractal Brownian processes or natural processes assimilated to the former ones (e.g.
540 Gao et al., 2015).

541 Complementarily, we may investigate which degree of redundancy may exist between the MSE
542 approach and the structure function approach, i.e. to check if they estimate the same scaling
543 properties or not. Indeed, first-order structure functions can be used to estimate the third parameter
544 H of the FIF model (cf. Section 2.3). In the case of our dataset, is H indeed close to the MSE scaling
545 parameter obtained above?

546 **6.2 Comparison with first order structure functions**

547 In order to investigate this issue, the first-order structure functions (defined by Eq. 4) have been
548 estimated based on the surface soil moisture data from the nine stations. The structure functions are
549 displayed on Figure 12, where they are regularly shifted in the vertical direction to improve
550 readability. It should be noted that on Figure 12, the large time increments are on the right of the
551 figure. Very small and very large time increments are not taken into account as they might be
552 affected by sampling limitations. On the middle of the Figure 12, a scaling regime may be identified
553 for $30 h \leq \Delta t \leq 1000 h$. This regime is representative of a scaling behavior of the form given by Eq.
554 5. In order to make the comparison with Sect. 6.1 easier, the linear fit range was restricted to $30 h \leq$
555 $\Delta t \leq 650 h$. Good scaling results are obtained, with an very small error bar on the slope estimate (
556 $\Delta H \approx \pm 0.01$ or smaller). The scaling exponent does not vary much from one series to one another,
557 the mean value of H being equal to $H = 0.41$ (see values in Table 5).

558

559 We may now compare the estimates of H and H' obtained for the different series. Figure 13 displays
560 H' as a function of H for the different time series. For most stations, estimates of H and H' are very
561 similar despite some differences larger than the regressions error bars. However, significant

562 differences were found for two stations (Narbonne and Savenès) where the H' estimates are much
563 larger than their structure function counterparts. It remains unclear if this is due to greater
564 uncertainties in the estimates obtained by the MSE approach and/or an ability of the latter method
565 to catch specific features at some stations. This issue could be investigated in future studies since
566 MSE and structure function approaches may perform differently in the estimation of Hurst-like
567 scaling parameters.

568 While the detailed comparison of the two methods still requires additional theoretical work, we can
569 already notice that each method has qualitatively identifiable practical advantages and drawbacks.
570 Namely, classical structure functions are more easy to use to analyze a large scale range but do not
571 have direct interpretation in terms of signal information. Additionally, they cannot estimate negative
572 scaling exponents (unless modified variants based on wavelets are used, e.g., Lovejoy and Schertzer,
573 2012). Meanwhile, MSE algorithms are appropriate for information and complexity estimation but
574 they are applicable over a relatively limited scale range due to computational constraints (Nogueira,
575 2017). The MSE framework also provides equations that seem more difficult to relate analytically to
576 equations such as those governing the statistics of multifractal cascades for non-unity orders of
577 statistics (Eq. 2). It could be interesting to try to investigate possible links with other theoretical
578 notions such as the information dimension, which contains both scaling and (Shannon) entropy
579 characterization. The latter point is left as a perspective in this work. However, a better
580 understanding of the possible relationship between multifractal scaling and MSE scaling can already
581 be achieved with the help of numerical tests. This is the subject of the next section.

582

583

584 **7 Expected relationship between H' and multifractal parameters**

585 **7.1 Numerical experiments**

586 The analysis presented in Section 6 was aimed to test if H and H' estimates could have close values in
587 the empirical case of the study. While the hypothesis $H \approx H'$ has been tested numerically by Nogueira
588 (2017) in the case of fractional Brownian motion processes where the structure function H is the
589 unique scaling parameter, the extension to the case of multifractal processes has not yet been
590 investigated by theoretical or numerical approaches (at least to the author's knowledge). Multifractal
591 processes differ from fractional Brownian motions by the existence of more nonlinear properties that
592 are characterized by the parameters C_1 and α defined above. A priori, there is no theoretical reason
593 to rule out a possible effect of C_1 and α on the MSE scaling parameter. Therefore, we might look for a
594 possible functional relationship of the form $H' = f(H, C_1, \alpha)$. While it is still unknown if this function f
595 can be defined in closed form, numerical tests can be performed to check if there is a systematic
596 dependency of H' on the three multifractal parameters.

597 In the following, numerical tests are presented based on the multifractal simulation techniques of FIF
598 processes described by (Pecknold et al., 1993; Schertzer and Lovejoy, 2002; Macor, 2007; Verrier,
599 2011). The main steps are the following: (i) generation of a unit extremal Lévy-stable noise (with an
600 asymmetry parameter equal to -1 and a stability parameter α identical to the multifractal α defined
601 above); (ii) multiplication by a (normalizing) multiplicative constant $\sigma = \left(\frac{C_1}{\alpha-1}\right)^{1/\alpha}$; (iii) Convolution by
602 a kernel proportional to $|t|^{-\frac{D}{\alpha}}$ where D is the embedding space dimension (i.e. $D = 1$ for time series);
603 (iv) Exponentiation; (v) fractional integration of order H . Step (i) above is performed by applying the
604 procedure described by Chambers et al. (1976) to a couple of independent simple random variables
605 at each data point. It is worth noting the procedure allows to store numerical realizations of these
606 two simple "input" variables U and E (with uniform and exponential distributions respectively, cf.
607 Pecknold et al., 1993, § 3) and then to repeat the five-step procedure with different sets of
608 parameters (H, C_1, α) . Therefore, it is very easy to test different sets of parameters on the same
609 probabilistic events (i.e. numerical realizations).

610 Numerical tests have been implemented in the following way. Twenty independent realizations of
611 couples of U and E time series are first generated in order to build the experiment on twenty
612 different probabilistic events. Then, the procedure of simulation is applied for various sets of
613 multifractal parameters (H, C_1, α) – in other words 20 synthetic multifractal time series are obtained
614 for each set of parameters but the 20 random realizations used to generate the series are the same

615 for all sets. Since synthetic multifractal time series can have numerical discrepancies at the highest
616 frequencies, it has been chosen to perform all numerical simulations based on time series of 2^{18}
617 points and then to average each of them over a series of disjoint intervals of 16 points. This
618 procedure provides 20 time series of $2^{14} = 16384$ points for each set of multifractal parameters.

619 In a second step, MSE analysis is applied to each individual time series and then the $MSE(\tau)$ function
620 is averaged over the 20 individual realizations available for each set of multifractal parameters. The
621 H' exponent is then estimated *from the averaged $MSE(\tau)$* by performing a linear regression in log-log
622 coordinates over the whole range of available scales (i.e. from 1 to 163 data points since sample
623 entropy needs 100 data points to be computed).

624

625 7.2 Results

626 Examples of MSEs estimated from individual series are presented in log-log coordinates on Figures 14
627 and 15. Figure 14 shows the $MSE(\tau)$ obtained for a single random realization with $\alpha = 2$ and $C_1 = 0.05$.
628 The blue curve is obtained when $H = 0$ and the green curve when H has been fixed to 0.4. It may be
629 noticed that the effect of the fractional integration (i.e. transition from the blue to the green curve)
630 changes both the height and the slope of the $MSE(\tau)$ function. In this individual case, the hypothesis
631 $H' \approx H$ seems to hold when 0.4, but a moderate scaling is found in the $H = 0$. Figure 15 illustrate the
632 case where the parameter C_1 changes (while $\alpha = 2$ and $H = 0$) based on a different individual
633 realization. The increase of C_1 leads to a drastic decrease of the entropy and seems also to increase
634 the slope H' of the MSE curve.

635 These conclusions are globally confirmed when considering the H' estimates obtained from the
636 averaged MSEs (Figs. 16-17). Figure 16 shows the H' estimate as a function of H , with variable C_1
637 while α is fixed to 2. On the figure, each curve corresponds to one value of C_1 comprised between
638 0.05 (continuous blue curve) and 0.55 (dashed green curve) with a spacing of 0.10. H is sampled in
639 the range $[0, 0.7]$ with a spacing of 0.10. The first bisectrix is represented as a dashed black line. For
640 all curves, H' increases with H but does not follow the bisectrix (except when C_1 is very small). When
641 $H > 0.5$, all curves tends to get relatively close to each other meaning that when H is sufficiently high,
642 the influence of C_1 on H' is relatively small. On the contrary, when $H < 0.4$, the influence can become
643 quite large with a systematic increase of H' with C_1 (confirming the observation of Figure 15). In the
644 case of a non integrated multiplicative cascade ($H = 0$), this can lead to an H' estimate varying within
645 the range $[0.1; 0.3]$ depending on the C_1 value. Regression error bars on H' estimated from the
646 average $MSE(\tau)$ functions are not represented because they are very small (inferior to 0.01).

647 The possible influence of the α parameter (coupled with H) is presented on Figure 17, where H' is
648 represented as a function of H , with one curve for one value of α (C_1 being fixed to 0.25). The
649 parameter α has been chosen in the interval $[1.5, 2]$ with a spacing of 0.1 (i.e., values usually
650 encountered in geophysics, and also reported in this study). Again, H' is an increasing function of H
651 for all curves which also remain above the first bisectrix. However, H' is a decreasing function of α
652 such that the relationship between H' and H tends to become closer to the first bisectrix when $\alpha = 2$.
653 We may notice that the dependency of H' on α is very limited for high values of α ($\alpha > 1.7$) but
654 becomes more pronounced for smaller values. Finally, we may again notice that for high values of H ,
655 all curves become close to the first bisectrix and increasingly undistinguishable from each other (e.g.,
656 $H = 0.7$).

657 These numerical results show that the relationship between multifractal parameters and the MSE
658 scaling properties is more complex than originally anticipated and that a dependency on the three
659 multifractal parameters should be expected in the general case. However, we may distinguish two
660 basic cases:

- 661 • For relatively high values of H ($> 0.4-0.5$), the relationship $H \approx H'$ might still provide a useful
662 work assumption since our results suggest that the fractional integration washes out the
663 possible effects of the multifractal parameters α and C_1 . In this case, the MSE scaling
664 properties of a multifractal process are expectedly close to those of a fractional Brownian
665 motion which have been investigated by Nogueira (2017).
- 666 • For lower values of H , and in particular in the case of a non-integrated multiplicative cascade
667 ($H = 0$), the effect of α and C_1 is visible and results in a MSE scaling exponent that is higher
668 than H (i.e. $H' = H + g(H, C_1, \alpha)$). The positive offset modeled by the function g is higher when
669 C_1 increases or when α becomes significantly smaller than 2. Interestingly, the offset
670 becomes very small when C_1 takes small values, which illustrates that multifractal
671 intermittency plays an important role in the difference $H' - H$.

672 While more exhaustive numerical tests and theoretical work remain to be done to quantify and
673 understand the relationships between multifractal parameters and observed MSE scaling laws, some
674 qualitative arguments may help to interpret the present observations. First, the effect of the
675 fractional integration included in the FIF model is similar to a (scaling) low-pass filter that significantly
676 attenuates the variability at the smallest scales. It is therefore expectable that a fractionally
677 integrated signal ($H > 0$) possesses more entropy at coarse resolutions than at fine resolutions. This is
678 obvious on Figure 14 where we see that the effect of H is to strongly reduce entropy estimates at
679 small scales (i.e. on the left part of Fig. 14). Moreover, as already noted on Figure 15, high C_1

680 parameter values also contribute to strongly reduce entropy. This could seem a bit paradoxical since
681 C_1 is known as an intermittency/heterogeneity parameter. However, multifractal processes with high
682 C_1 correspond to processes that exhibit slow variations most of the time, while most of the variance
683 is due to very strong and sparse peaks (e.g., fig. 10 in Pecknold et al., 1993). In this case, sample
684 entropies tend to decrease since repetitive patterns are (or seem) more frequent due to the
685 existence of wide areas with small variations. This effect is even more pronounced due to the
686 increased standard deviation of the process when C_1 is large due to the contribution of extreme
687 peaks. Indeed, since the sample entropy builds its similarity criterion by using the standard deviation
688 of the series (see Eq. 6), it classifies events as similar more easily when large standard deviation and
689 C_1 are involved. We can notice that the decrease of entropy is expectably stronger at small scales
690 where peaks are both very intense and cover a very small fraction of the time.

691 **7.3 Application to SMOSMANIA empirical parameters**

692 Finally, we may come back to the results of Section 6 since we have just demonstrated that the initial
693 hypothesis $H \approx H'$ may be too simplistic for multifractal processes in some situations. By reporting the
694 parameters (H, C_1, α) estimated in Sections 5 and 6.2 to the graph presented on Figure 17, we obtain
695 an expected parameter H' between 0.4 and 0.5. However, the empirical H' estimates found in Section
696 6.1 (≈ 0.40) are less coherent with this expected parameter ($H' \approx 0.47-0.48$ for multifractal
697 parameters $\alpha = 1.7, C_1 = 0.25, H = 0.40$) than the one resulting from the naive $H \approx H'$ hypothesis.
698 Beyond possible estimation uncertainties, a plausible explanation for this observation is the fact that
699 the multifractal parameters C_1 and α are expected to be valid in the range of scales 128h – 1 h, while
700 H and H' are valid in the range 650 h – 30 h. At scales larger than 128 h (i.e., beginning of the
701 “macroweather” regime), the flux moments have much slower variation with scale (Figs. 4-5), which
702 is equivalent to a strong decrease of the C_1 parameter. This could explain why the empirical
703 estimates of H' are close to the structure function exponents since it is what is expected for low
704 values of C_1 (see Figure 16, continuous blue curve), especially when H is still relatively large (0.4 in
705 this case).

706

707

708 **8 Conclusions**

709 Like other geophysical processes, soil moisture is strongly variable over a rather wide range of time
710 (and space) scales. Some previous literature papers advocated the existence of scaling and fractal

711 properties over various ranges of scales. In this study, 8-year time series of surface soil moisture have
712 been analyzed with the help of various multiscale analysis tools in order to investigate the existence
713 of scaling laws. The twelve sensors are located in the same area of about 300 x 200 km² in
714 southwestern France and the average distance between stations is of the order of a few dozens of
715 kilometers. For nine of these stations, power spectra of surface soil moisture showed a red noise-like
716 $1/f^2$ scaling over a large scale range of time scales ($1 \text{ h} \leq \tau \leq 1000 \text{ h}$), without any strong break in the
717 scaling. The scaling exponents are rather homogeneous at the surface layer and exhibit some
718 variability at deeper horizons where spectra tend to become steeper.

719 In a second step, multifractal analysis has been applied in order to identify the possible existence of
720 scaling properties for non-quadratic moment orders. Multifractality has also been observed in the
721 range of scales comprised between 1 h and 128 h. The moment scaling functions have been
722 successfully fitted with the two-parameter functional form predicted by the Universal Multifractal
723 Model. At the surface (5 cm depth) layer, the obtained parameters showed a slightly greater
724 dispersion than that of spectral slopes, with $C_1 \approx 0.2-0.3$ and $\alpha \approx 1.3-1.8$ for surface data (5 cm
725 depths). However, a greater dispersion has been found for data collected at deeper horizons (10-30
726 cm). On average, the depth affects the multifractality parameters more strongly than spectral
727 slopes. At 30 cm depth, C_1 is closer to 0.4-0.5 while α often becomes smaller than 1. This shift and
728 this variability could be related to the soil composition while surface parameters are likely to be more
729 representative of the influence of the meteorological forcing.

730 Multiscale entropy has also been applied as a complementary scaling analysis tool devoted to the
731 multiscale analysis of information content present in the series. This approach had been originally
732 developed in Biology/medicine research literature and has a good potential for analyzing time series
733 in other scientific fields. As noticed by Nogueira (2017), a systematic relationship exists between MSE
734 scaling and spectral/first-order structure function scaling in the case of synthetic monofractal data
735 and seems to hold for scaling geophysical observations.

736 MSE analysis has been performed over the twelve SMOSMANIA surface time series, showing the
737 existence of scaling properties for the sample entropies estimated at various aggregation resolutions
738 (30-650 hours). Due to the strong constraints on sample entropy estimation, the validity of MSE
739 scaling at larger scales could not be investigated. This is a basic (methodological) limitation of the
740 MSE approach that is compensated by other advantages: contrary to the usual structure function
741 approach, MSE could be appropriate to analyze processes with negative Hurst scaling exponents (i.e.
742 with $H < 0$ or $\beta < 1$) and has a rather direct interpretation in terms of information theory.

743 More generally, our numerical results also demonstrated that the MSE scaling properties are not
744 limited to fBm-like *monofractals* investigated by previous authors (Gao et al., 2015; Nogueira, 2017)
745 but should hold for synthetic and natural signals characterized by *multifractal* properties. Numerical
746 tests were used to investigate the existence of a systematic relationship between the MSE scaling
747 parameter H' and the three multifractal parameters H , C_1 and α . It has been observed that the
748 relationship $H' = f(H, C_1, \alpha)$ is more complex than in the monofractal case (where $H' \approx H$) and that in
749 some cases the intermittency parameter C_1 and the multifractality parameter α can affect
750 significantly MSE scaling properties. In other cases (i.e. large H or low intermittency), the
751 approximation $H' \approx H$ could be acceptable. Finally, more theoretical work remains to be done to
752 understand the relationships between MSE scaling and fractal/multifractal scaling properties. This
753 could open interesting perspectives for a better physical understanding of many geophysical
754 processes that exhibit multifractal properties.

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756

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914 **Tables**

Station	GPS coordinates (Lat/Lon)		Altitude (m)	Sensor
Condom	43.97440	0.33610	174	ThetaProbe-ML2X
Créon d'Armagnac	43.99360	-0.04690	149	ThetaProbe-ML2X
Lahas	43.54720	0.88780	249	ThetaProbe-ML2X
Lézignan Corbieres*	43.17330	2.72830	60	ThetaProbe-ML2X
Montaut	43.19220	1.64360	295	ThetaProbe-ML2X
Mouthoumet	42.96000	2.53000	538	ThetaProbe-ML2X
Narbonne	43.15000	2.95670	112	ThetaProbe-ML2X
Peyrusse Grande	43.66640	0.22170	245	ThetaProbe-ML2X
Sabres*	44.14750	-0.84560	81	ThetaProbe-ML2X
St Félix de Lauragais*	43.44170	1.88000	337	ThetaProbe-ML2X
Savenès	43.82500	1.17670	158	ThetaProbe-ML2X
Urgons	43.63970	-0.43500	145	ThetaProbe-ML2X

915
916 **Table 1a : Identification of the 12 SMOSMANIA stations located in the region of interest. Names**
917 **marked with a star correspond to stations that are not used in the rest of the study.**

918
919
920

Station	Soil texture
Condom	Clay
Créon d'Armagnac	Sand
Lahas	Clay
Lézignan Corbieres	Silt loam
Montaut	Silt loam
Mouthoumet	Silt loam
Narbonne	Silt loam
Peyrusse Grande	Clay
Sabres	Sand
St Félix de Lauragais	Clay
Savenès	Silt loam
Urgons	Silt loam

921

922 **Table 1b : Soil texture classification (in terms of USDA texture classification), based on**
923 **ISMN/SMOSMANIA metadata information.**

924

925

Station	β	$\Delta\beta$	R^2
Condom	1.99	0.037	0.988
Créon d'Armagnac	2.01	0.037	0.988
Lahas	1.96	0.039	0.987
Montaut	1.97	0.034	0.990
Mouthoumet	2.02	0.036	0.989
Narbonne	1.91	0.034	0.989
Peyrusse Grande	1.93	0.031	0.991
Savenès	1.92	0.030	0.992
Urgons	1.99	0.036	0.989

926

927 **Table 2: Estimates of spectral scaling exponents β for the time series collected at 5 cm depth. $\Delta\beta$ is**
 928 **half of the length of the 95% confidence interval on β while the last column presents the**
 929 **coefficient of determination.**

930

Station	C_1	α	Validity conditions
Condom	0.22	1.82	$q \leq 2.5$; scales 1-128 h
Créon d'Armagnac	0.26	1.42	«
Lahas	0.27	1.82	«
Montaut	0.31	1.60	«
Mouthoumet	0.22	1.76	«
Narbonne	0.31	1.61	«
Peyrusse Grande	0.28	1.51	«
Savenès	0.27	1.68	«
Urgons	0.29	1.37	«

931

932 **Table 3: Universal multifractal parameter estimates for the nine SMOSMANIA time series collected**
 933 **at 5 cm depth.**

934

935

Station	H'	$\Delta H'$	R^2
Condom	0.31	0.02	0.96
Créon d'Armagnac	0.45	0.02	0.99
Lahas	0.35	0.02	0.98
Montaut	0.38	0.02	0.98
Mouthoumet	0.35	0.04	0.90
Narbonne	0.66	0.02	0.99
Peyrusse Grande	0.36	0.02	0.96
Savenès	0.61	0.01	0.99
Urgons	0.36	0.01	0.98

936

937 **Table 4: MSE scaling parameter estimates over the time scale range 30 h – 650 h for the nine series**
938 **collected at 5 cm depth.**

939

940

Station	H	ΔH	R^2
Condom	0.42	< 0.01	0.98
Créon d'Armagnac	0.37	"	0.97
Lahas	0.41	"	0.98
Montaut	0.41	"	0.99
Mouthoumet	0.39	"	0.96
Narbonne	0.38	"	0.93
Peyrusse Grande	0.39	"	0.98
Savenès	0.47	"	0.99
Urgons	0.42	"	0.99

941

942 **Table 5: Structure function scaling parameter estimates over the time scale range 30 h – 650 h for**
943 **the nine stations (surface moisture).**

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952 **Figures captions**

953 **Figure 1:** Power spectral densities $E(f)$ of the nine selected SMOSMANIA surface (5 cm
954 depth) soil moisture series in logarithmic coordinates. In order to facilitate the
955 understanding in terms of time scales, the horizontal axis exhibits values of $1/f$ (time scale
956 equivalent to a Fourier frequency). The horizontal axis is presented so that high frequencies
957 (i.e., small time scales) appear on the right of the figure. *The different curves have regularly*
958 *been translated upwards in order to improve readability.*

959 **Figure 2a:** Power spectra of soil moisture at Peyrusse Grande, one spectrum per depth.

960 **Figure 2b:** Power spectra of soil moisture at Montaut, one spectrum per depth.

961 **Figure 3:** Estimates of spectral slope β for all stations and depths. Each line is associated to a
962 depth. On the abscissa axis, the three letters identify the stations. Stations are sorted from
963 west to east.

964 **Figure 4:** Flux moments estimated for the Peyrusse Grande time series of surface soil
965 moisture. Logarithmic scales are used. Moment orders q vary between 0 and 2.5.

966 **Figure 5:** Flux moments estimates for the eight other surface soil moisture series. Again,
967 moments orders comprised between 0 and 2.5 are represented.

968 **Figure 6:** Moment scaling function $K(q)$ of Peyrusse Grande surface soil moisture series.

969 **Figure 7:** Moments scaling functions $K(q)$ estimated for the eight other SMOSMANIA surface
970 soil moisture series.

971 **Figure 8: (Upper figures and bottom left)** Moments estimations based on Peyrusse Grande
972 data at depths 10, 20 and 30 cm. **(Bottom right)** Moments scaling functions estimated from
973 depths 5, 10, 20 and 30cm.

974 **Figure 9a:** estimates of the index of multifractality α for the different stations and depths.
975 Each line is associated to a depth. On the abscissa axis, the three letters identify the stations.
976 Stations are sorted from west to east.

977 **Figure 9b:** estimates of the inhomogeneity parameter C_1 for the different stations and
978 depths. Each line is associated to a depth. On the abscissa axis, the three letters identify the
979 stations. Stations are sorted from west to east.

980 **Figure 10:** MSE analysis for the nine surface soil moisture series. For each series, the Sample
981 Entropy (SampEn) is estimated for different time resolutions obtained by time aggregation.
982 *Plots are shifted vertically from one station to one another to improve readability.*

983 **Figure 11:** MSE scaling parameter estimates over the time scale range 30 h – 650 h for the
984 nine stations and the four depths considered.

985 **Figure 12:** First-order structure functions obtained for the nine SMOSMANIA series. *Plots are*
986 *shifted vertically in order to improve readability.* The different curves are sorted in the same
987 order as in Figure 10 (from Condom at the bottom to Urgons at the top).

988 **Figure 13:** Comparison of scaling exponents H and H' provided by SF and MSE methods.

989 **Figure 14:** MSE analysis for two synthetic multifractal time series generated with $H = 0$ (blue
990 curve) and $H = 0.4$, for the same random realization. C_1 and α are respectively fixed to 0.05
991 and 2. Note that contrary to Figs. 1, 2 and 12, no vertical shift is applied here between both
992 curves.

993 **Figure 15:** MSE analysis for two synthetic multifractal time series generated with $C_1 = 0.05$
994 (blue curve) and $C_1 = 0.35$, for the same random realization. H and α are respectively fixed to
995 0 and 2.

996 **Figure 16:** Estimated scaling parameter H' from the MSE analysis averaged on 20 random
997 realizations. α is equal to 2 in these simulations. Each curve represents the evolution of H' as
998 a function of H for a fixed C_1 . The values of C_1 are the following: 0.05 (continuous blue
999 curve), 0.15 (dashed blue curve), 0.25 (continuous red curve), 0.35 (dashed red curve), 0.45
1000 (continuous green curve), 0.55 (dashed green curve). H is sampled from 0 to 0.7 with 0.1
1001 spacing.

1002 **Figure 17:** Estimated scaling parameter H' from the MSE analysis averaged on 20 random
1003 realizations. C_1 is equal to 0.25 in these simulations. Each curve represents the evolution of
1004 H' as a function of H for a fixed α . The values of α are the following: 1.5 (continuous blue
1005 curve), 1.6 (dashed blue curve), 1.7 (continuous red curve), 1.8 (dashed red curve), 1.9
1006 (continuous green curve), 2 (dashed green curve). H is sampled from 0 to 0.7 with 0.1
1007 spacing.

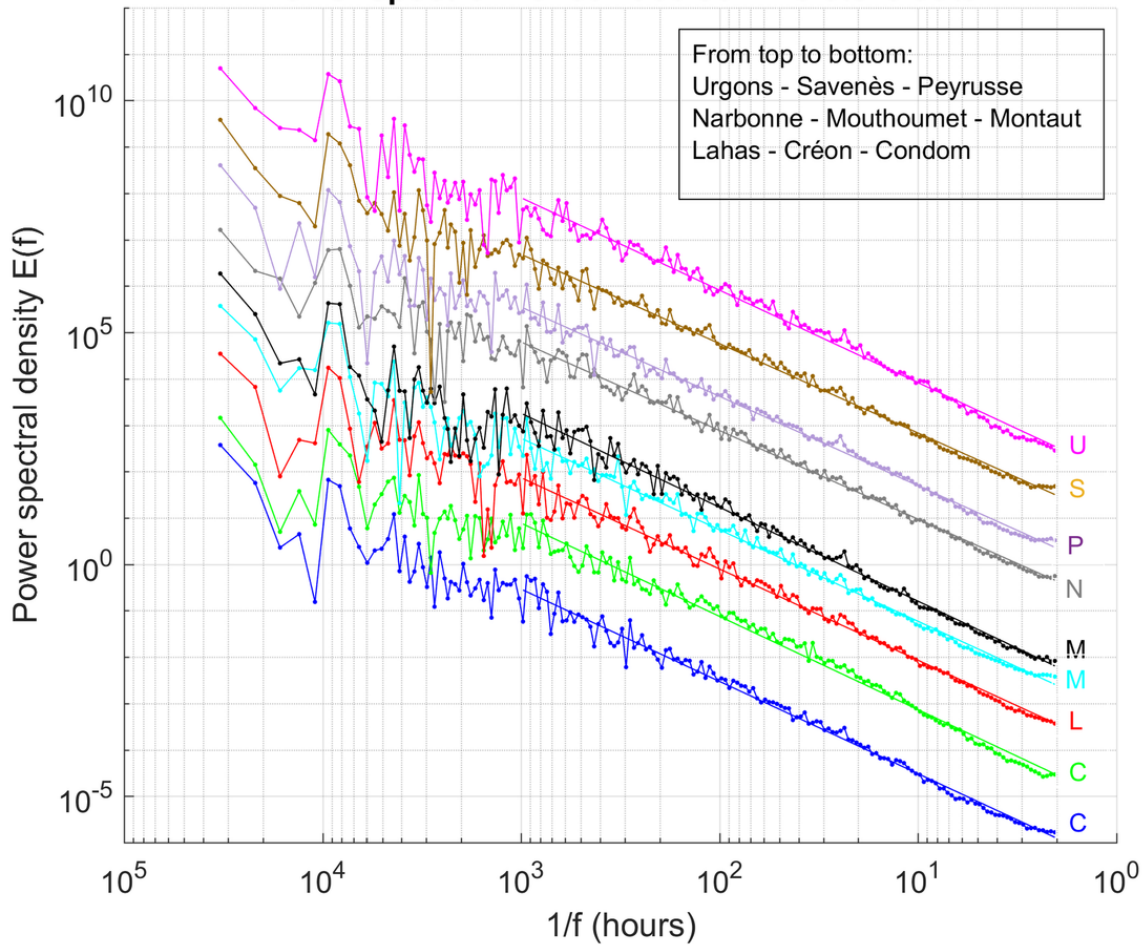
1008

1009

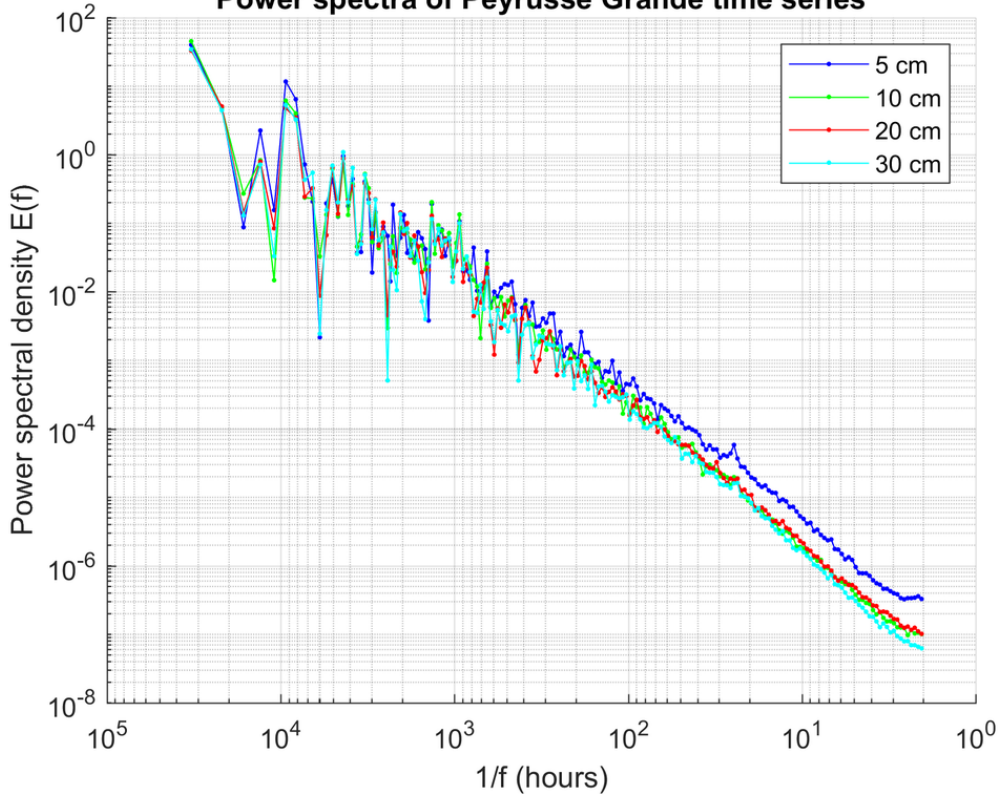
1010

1011

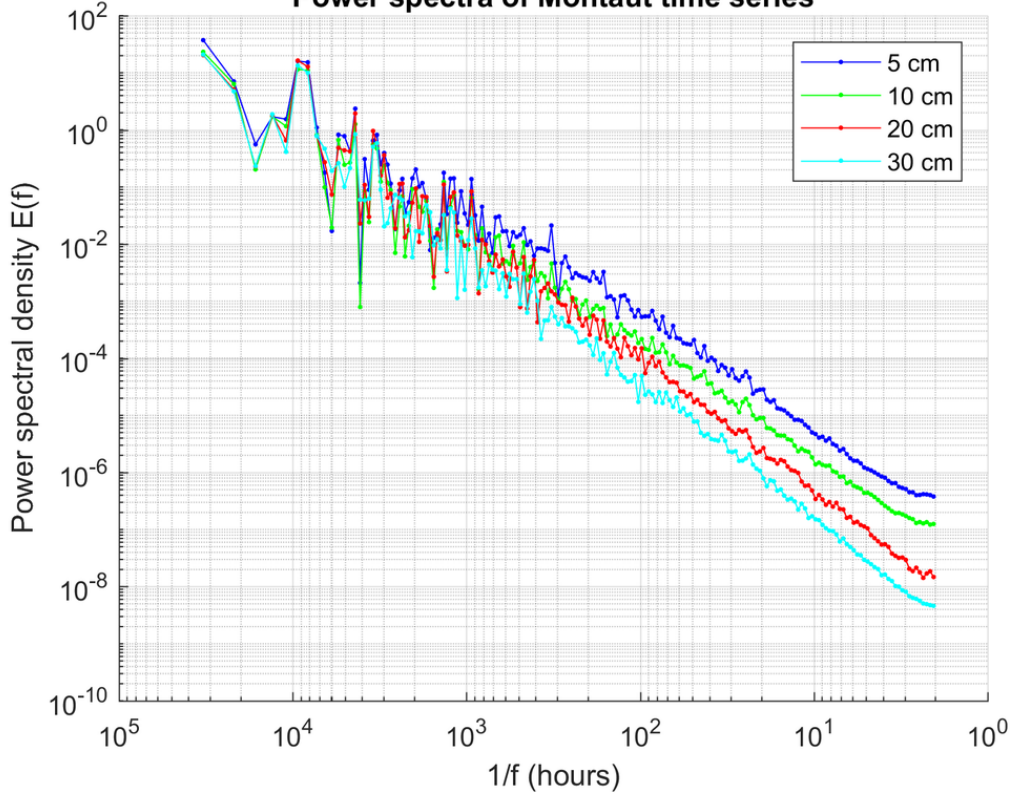
Power spectrum of surface soil moisture data



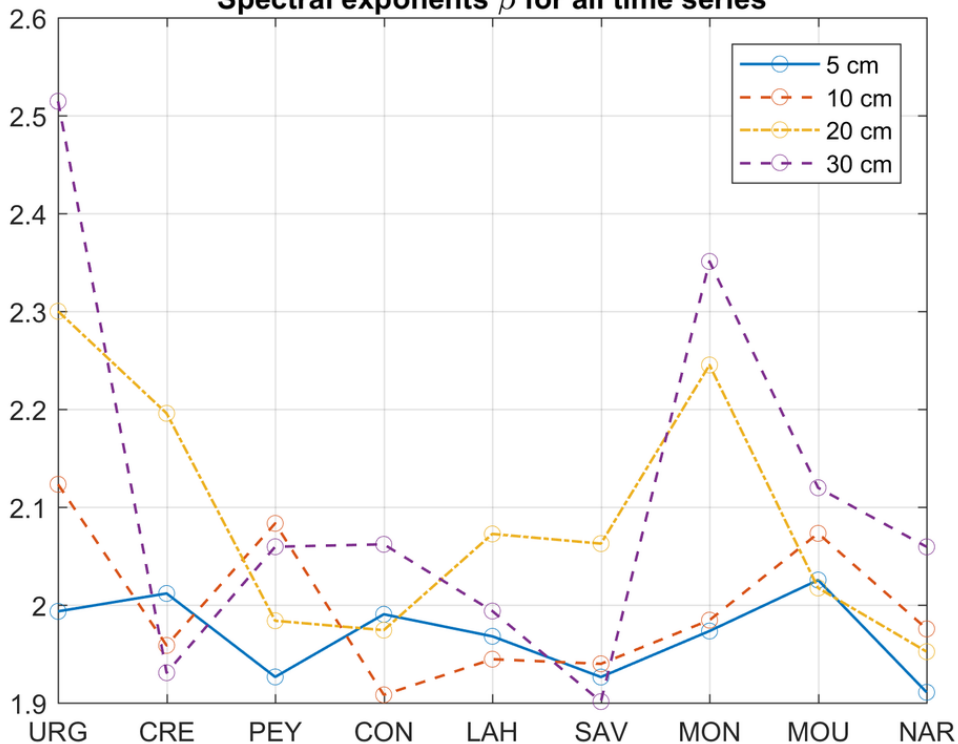
Power spectra of Peyrusse Grande time series



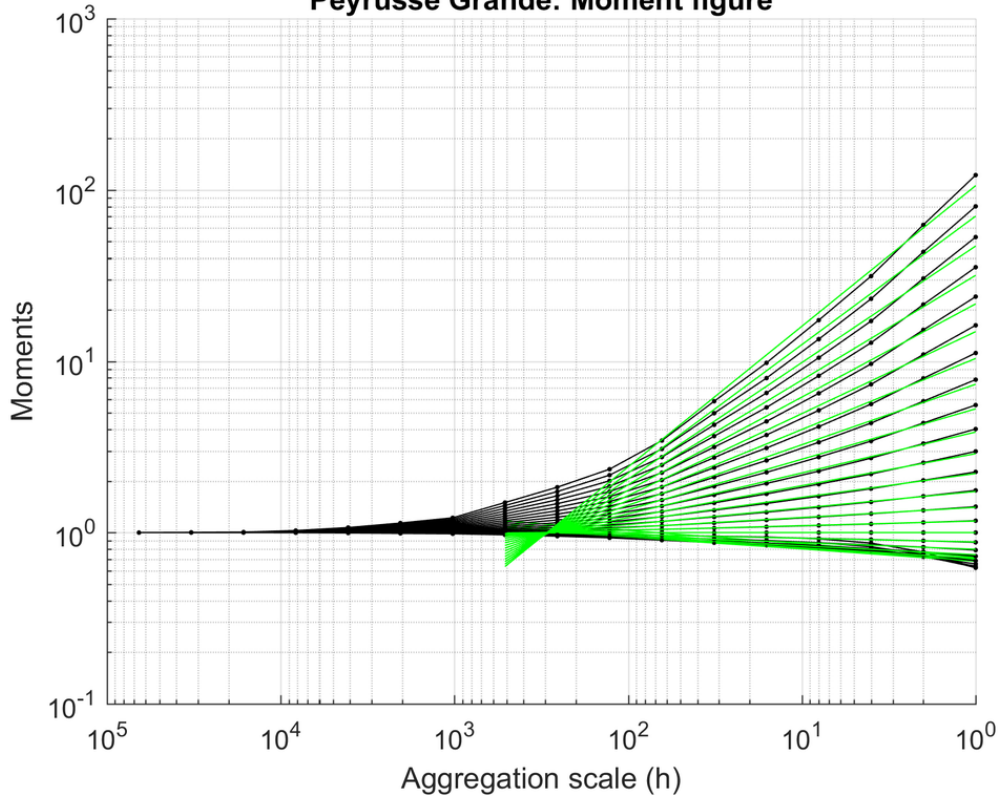
Power spectra of Montaut time series

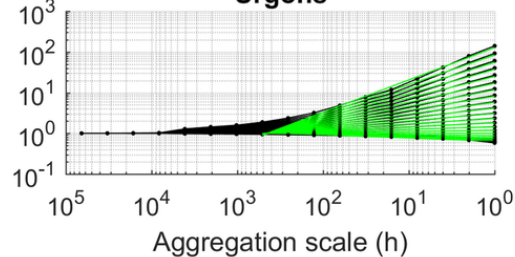
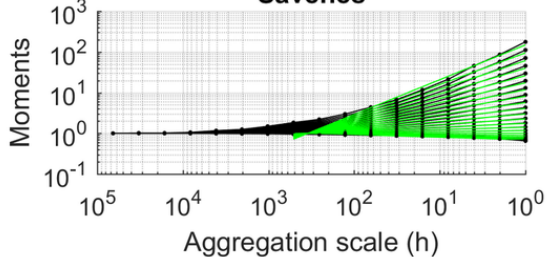
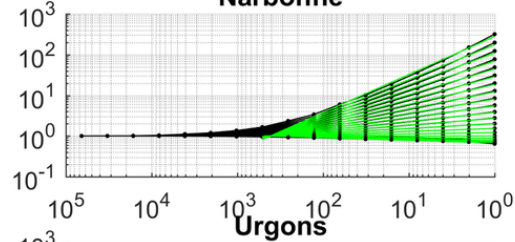
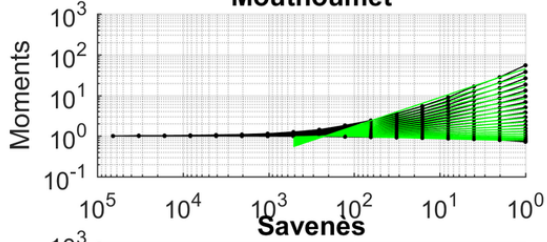
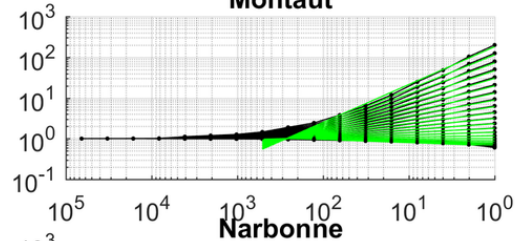
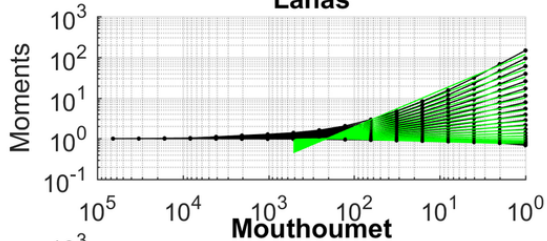
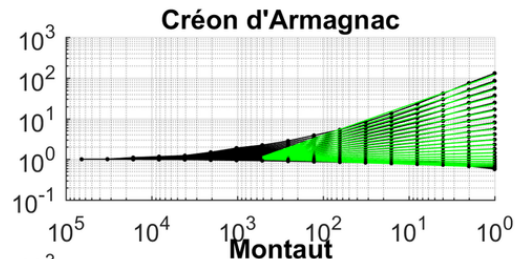
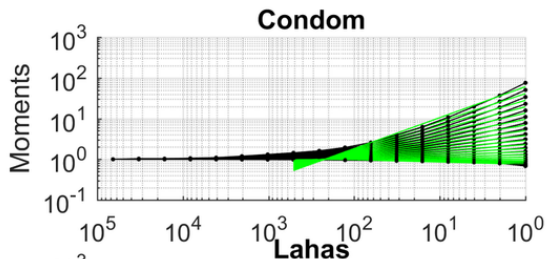


Spectral exponents β for all time series

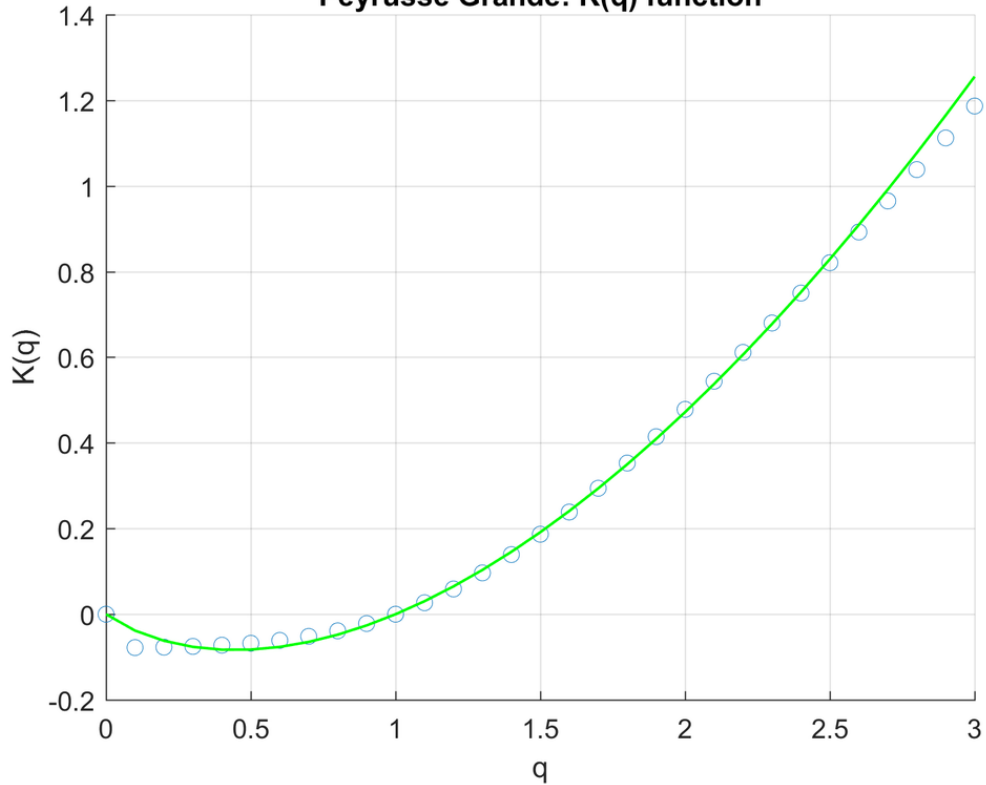


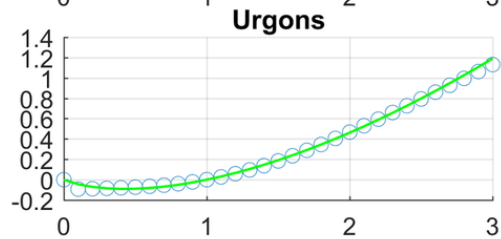
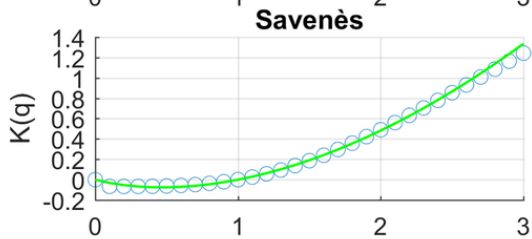
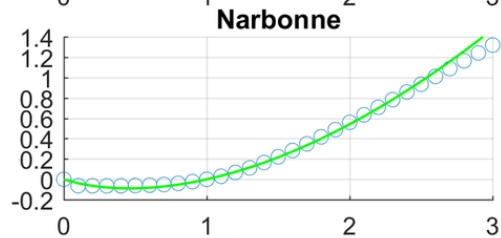
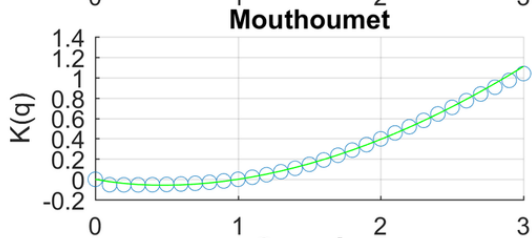
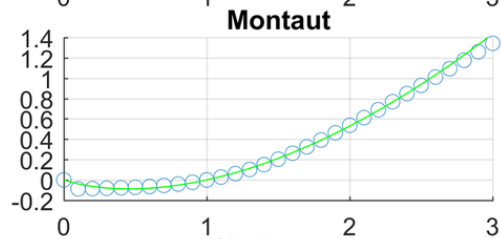
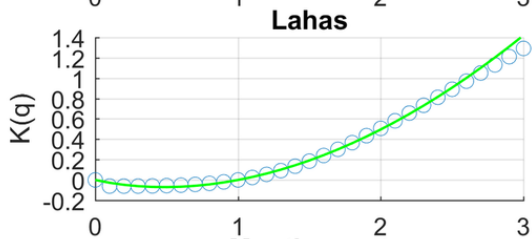
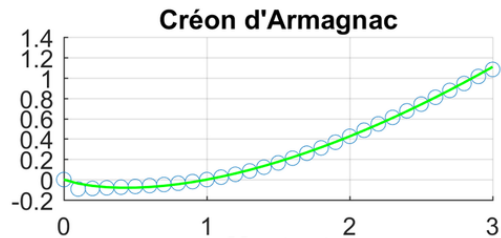
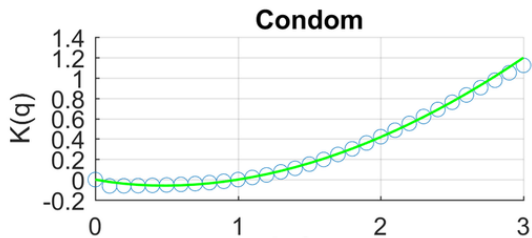
Peyrusse Grande: Moment figure





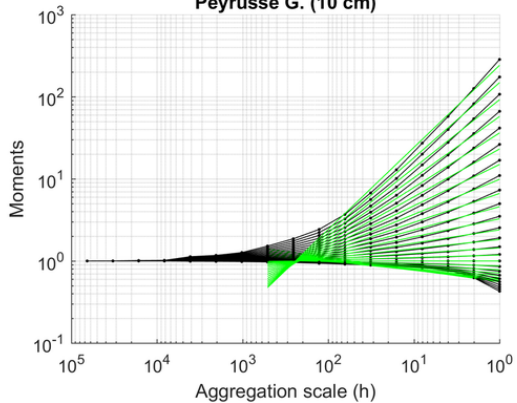
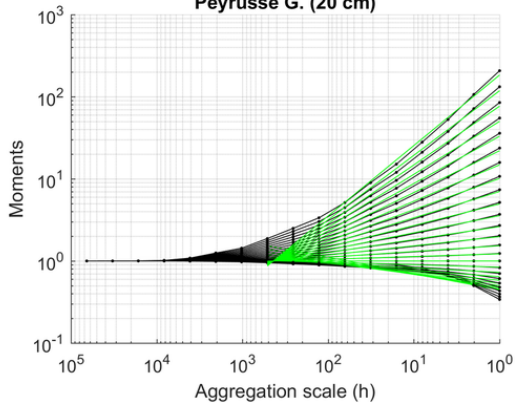
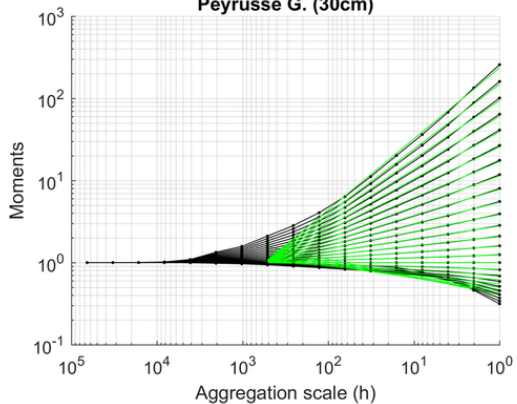
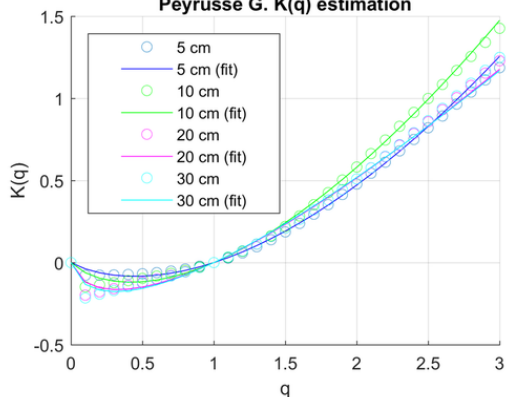
Peyrusse Grande: $K(q)$ function



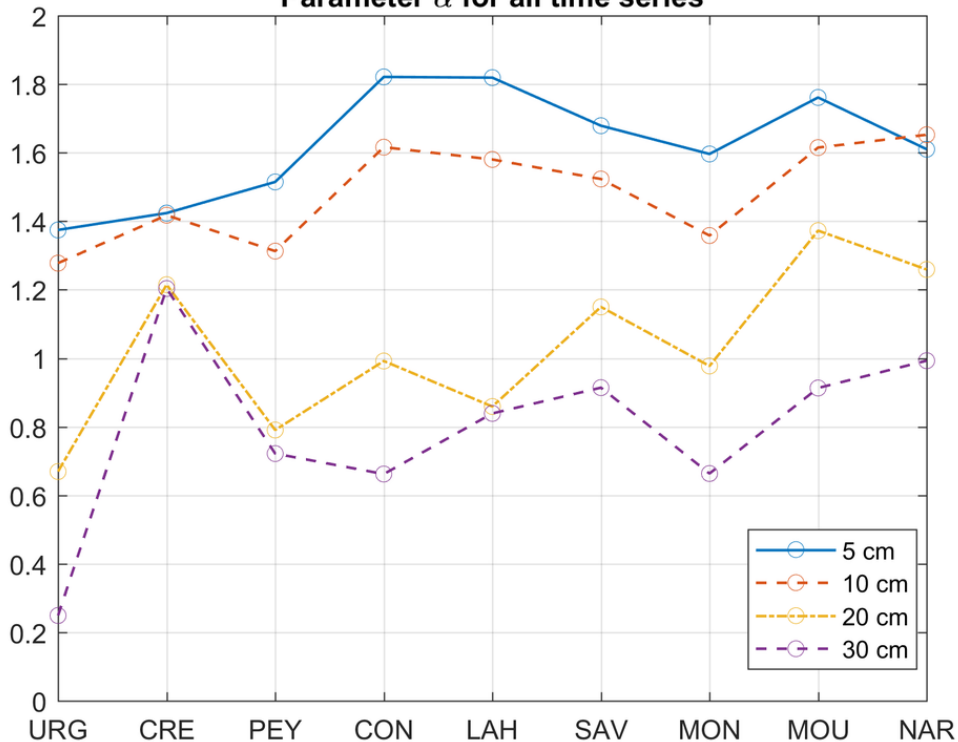


q

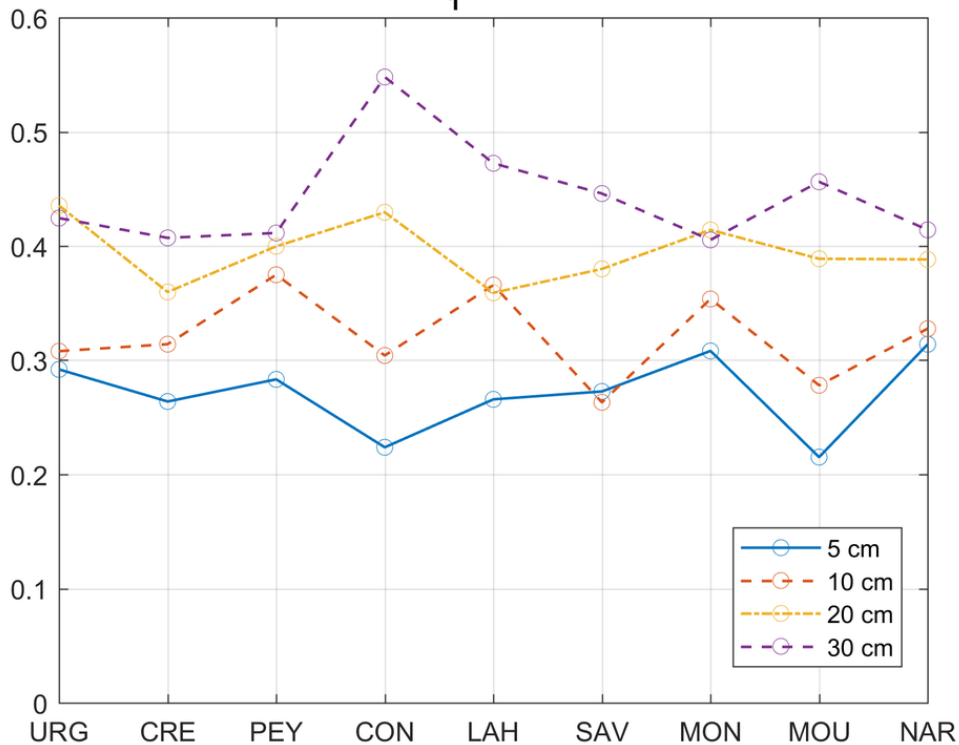
q

Peyrusse G. (10 cm)**Peyrusse G. (20 cm)****Peyrusse G. (30cm)****Peyrusse G. $K(q)$ estimation**

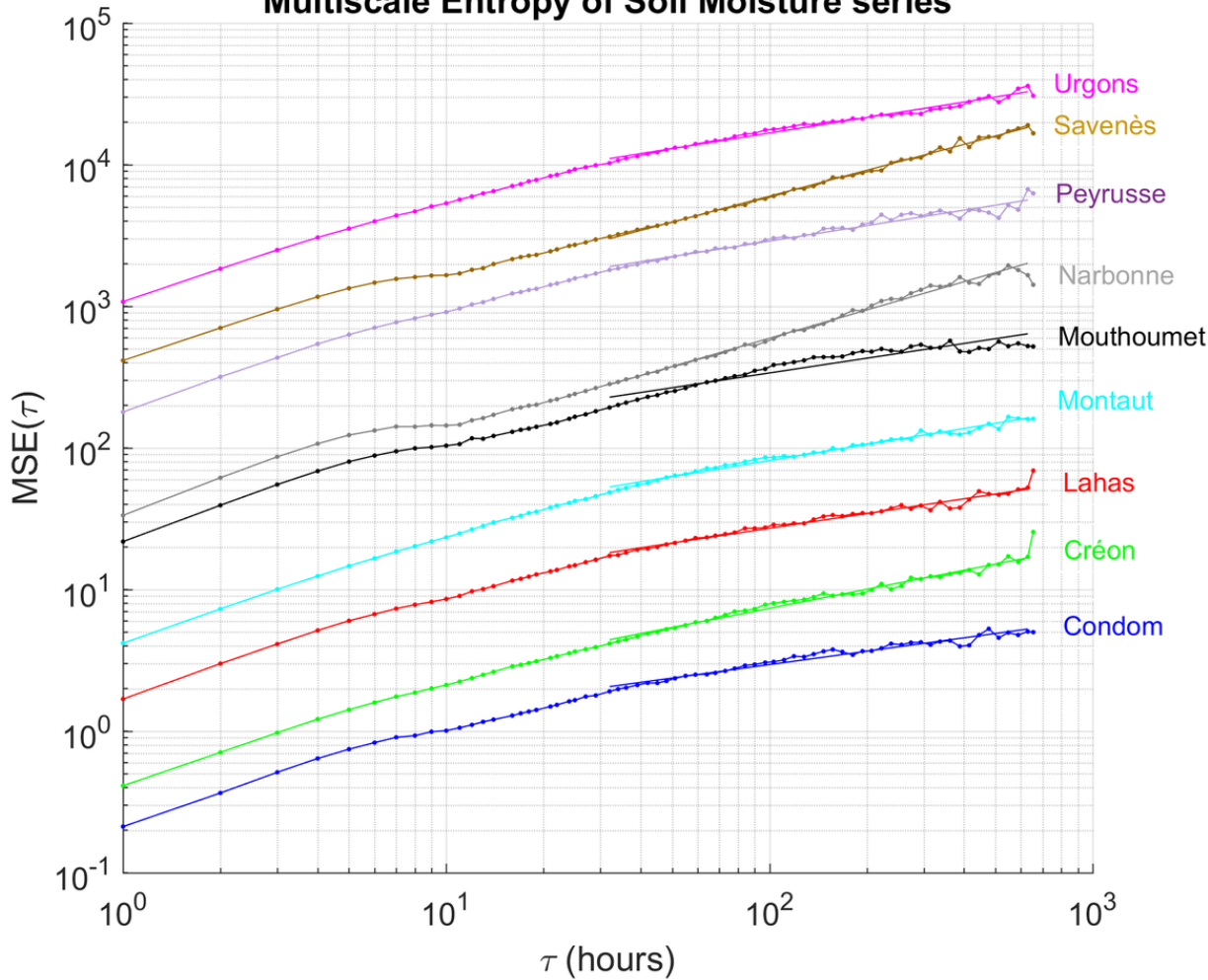
Parameter α for all time series



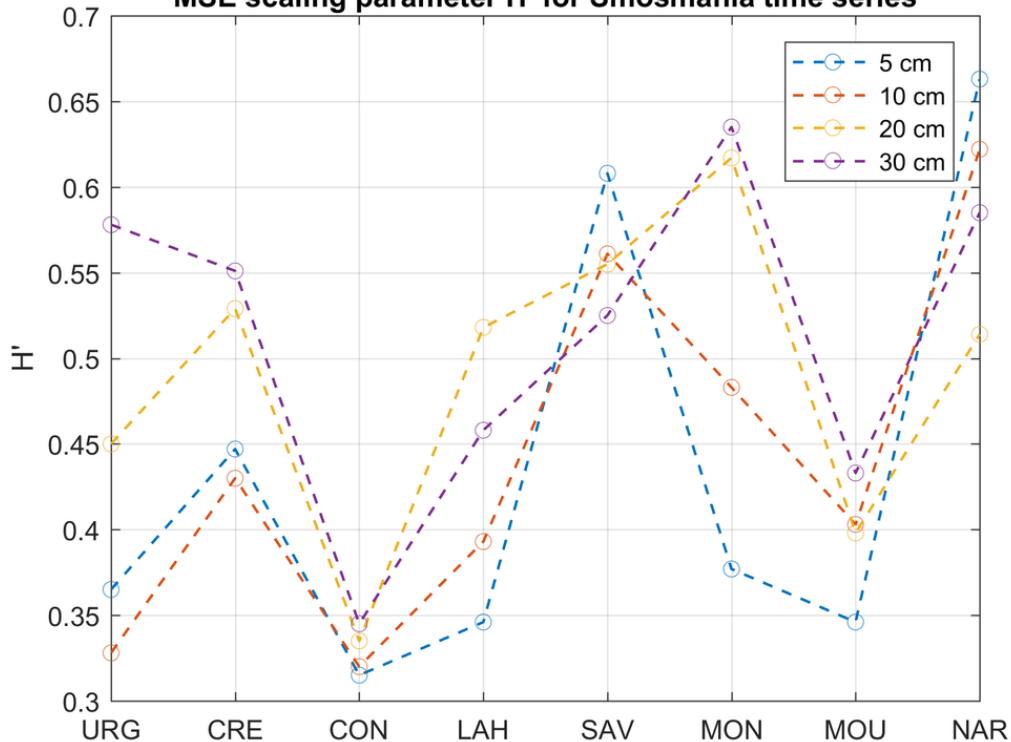
Parameter C_1 for all time series



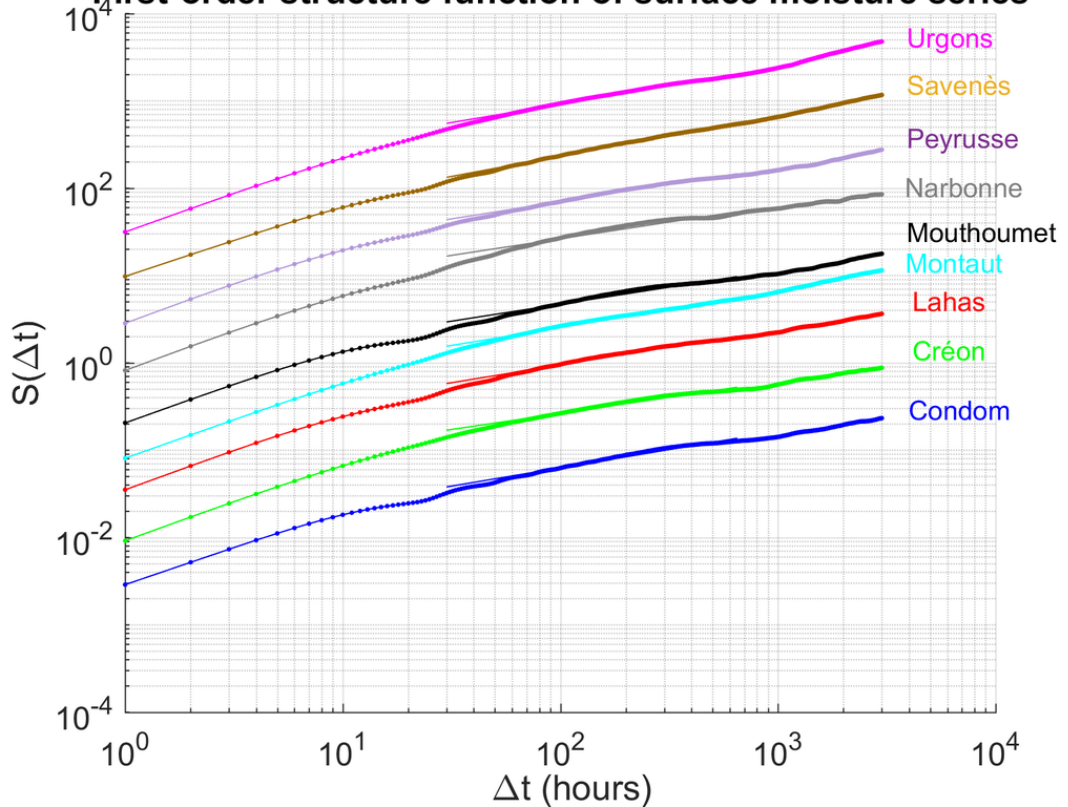
Multiscale Entropy of Soil Moisture series



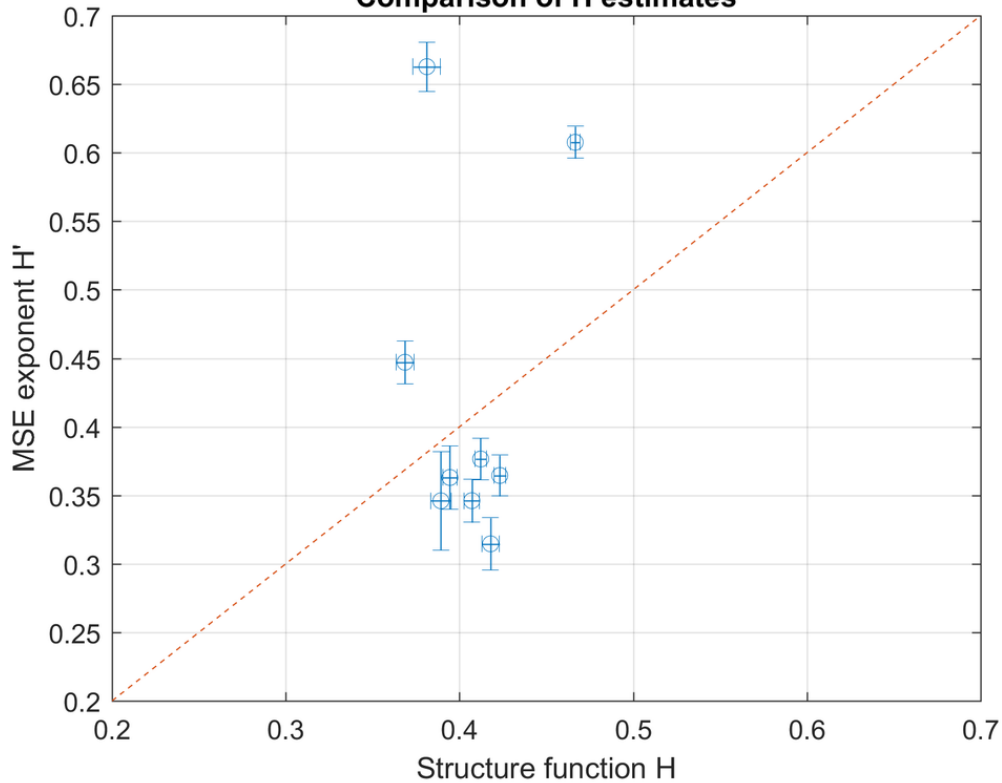
MSE scaling parameter H' for Smosmania time series



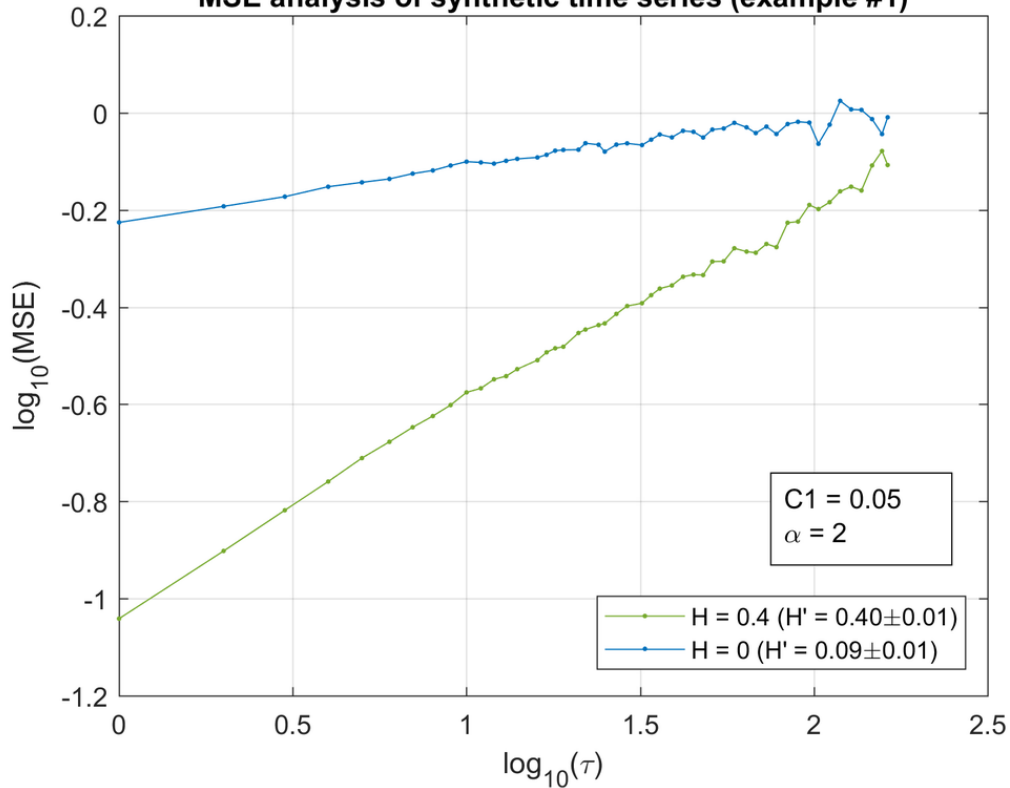
First-order structure function of surface moisture series



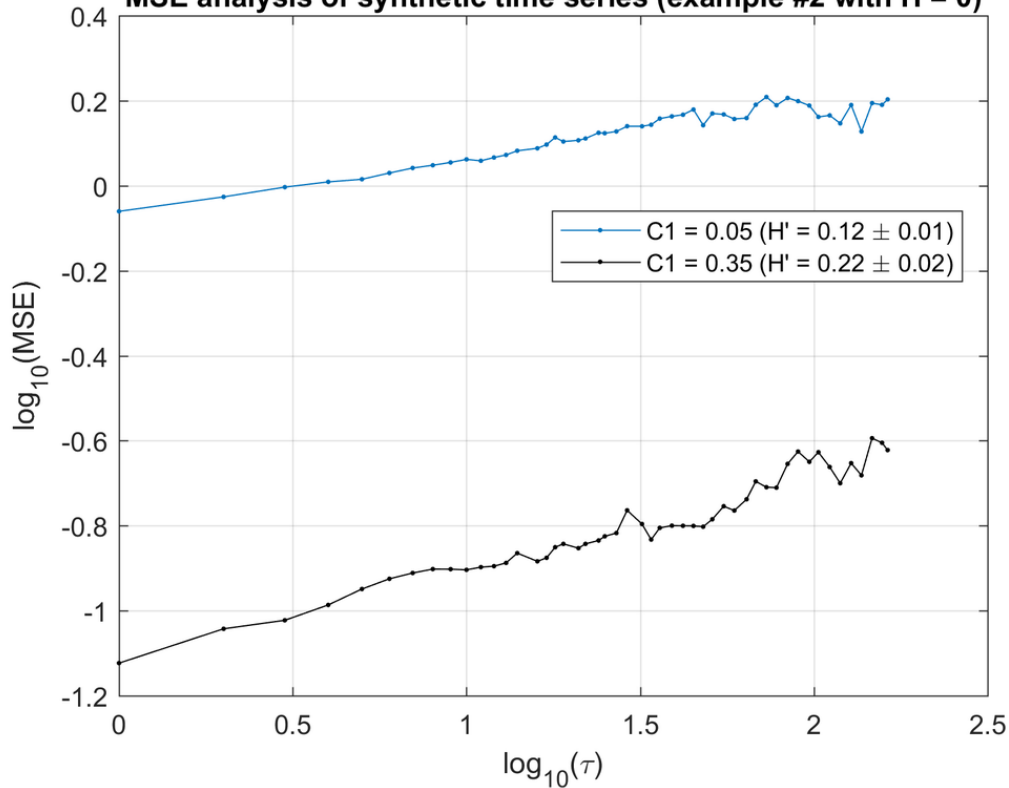
Comparison of H estimates



MSE analysis of synthetic time series (example #1)



MSE analysis of synthetic time series (example #2 with $H = 0$)



H' estimated from MSE analysis of synthetic time series

