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- 1 Strong genetic structure among populations of the tick *Ixodes ricinus* across its
- 2 range
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20 Abstract: Ixodes ricinus is the most common and widely distributed tick species in 21 Europe, responsible for several zoonotic diseases, including Lyme borreliosis. 22 Population genetics of disease vectors is a useful tool for understanding the spread 23 of pathogens and infection risks. Despite the threat to the public health due to the 24 climate-driven distribution changes of I. ricinus, the genetic structure of tick 25 populations, though essential for understanding epidemiology, remains unclear. 26 Previous studies have demonstrated weak to no apparent spatial pattern of genetic 27 differentiation between European populations. Here, we analysed the population 28 genetic structure of 497 individuals from 28 tick populations sampled from 20 29 countries across Europe, the Middle-East, and northern Africa. We analysed 125 30 SNPs loci after quality control. We ran Bayesian and multivariate hierarchical 31 clustering analyses to identify and describe clusters of genetically related 32 individuals. Both clustering methods support the identification of three spatiallystructured clusters. Individuals from the south and north-western parts of Eurasia 33 34 form a separated cluster from northern European populations, while central 35 European populations are a mix between the two groups. Our findings have 36 important implications for understanding the dispersal processes that shape the 37 spread of zoonotic diseases under anthropogenic global changes.

38 Keywords: gene flow; infection risks; range shift

39 Introduction

40 Ixodes ricinus (Acari, Ixodidae) is the most widespread tick species occurring across 41 Europe and an important vector of multiple tick-borne diseases, both to humans 42 and livestock. Commonly reported pathogens transmitted by I. ricinus include: the 43 bacteria Borrelia burgdorferi sensu lato, responsible for the Lyme borreliosis, which 44 is the most prevalent tick-borne disease in temperate Europe (ECDC, 2015); 45 arboviruses (genus Flavivirus) causing tick-borne encephalitis (TBE) and louping-ill 46 (LI); the protozoan Babesia microti, responsible for the babesiosis; and the 47 bacterium Candidatus Neoehrlichia mikurensis, responsible for neoehrlichiosis, an 48 emerging tick-borne pathogen (Portillo et al., 2018; Welinder-Olsson et al., 2010). 49 The current climate-driven redistribution of hematophagous arthropods such as 50 ticks and mosquitoes may lead to severe challenges to public health and husbandry, 51 by carrying a wide range of vector-borne diseases to new areas (Dantas-Torres, 52 2015; Pecl et al., 2017). For instance, many studies have demonstrated that the 53 range of *I. ricinus* is already shifting northward and to higher elevations (e.g. 54 Hvidsten et al., 2020; Jore et al., 2011; Lindgren and Gustafson, 2001) and those 55 shifts are expected to continue in the future (Alkishe et al., 2017; Medlock et al., 56 2013).

57 Despite the threats of emerging infectious diseases following the 58 redistribution of *I. ricinus*, little is known about the genetic structure of tick 59 populations across the entire species range. Population genetic differentiation and 60 spatial structuring can, however, impact the vector fitness and distribution, and 61 therefore disease transmission (Blanchong et al., 2016; Wonham et al., 2006).

62 Population genetics approaches such as individual genetic clustering and 63 assignment methods enable inference on migrants (exchange of genes between 64 populations) and the risk of pathogen spread between populations (Kozakiewicz et 65 al., 2018). For example, Lang and Blanchong (2012) applied clustering and distance-66 based methods to assess gene flow and disease spread risk between populations 67 of white-tailed deer in the USA. Similarly, Van Zee et al. (2015) identified different 68 genetic clusters between the southern and northern range of the tick Ixodes 69 scapularis while the prevalence of borreliosis is known to be lower in the southern 70 range. The authors suggest that this pattern of spatial genetic structure might be 71 linked to differences in questing behaviour as ticks from the northern range would 72 be more likely to bite humans. Differences in several life history traits of I. ricinus -73 such as the temperature at which nymphs begin to quest – have been reported 74 along a latitudinal gradient (Gilbert et al., 2014), suggesting a spatially explicit 75 phenotypic plasticity or adaptation. Yet, such basic knowledge about the 76 distribution of genetic variation in *I. ricinus* and the migration processes involved in 77 disease transmission remain largely unknown, albeit being essential to design 78 better vector control strategies (Araya-Anchetta et al., 2015; Gooding, 1996; 79 Tabachnick and Black, 1995).

The genetic structure of parasites' populations is known to be influenced by the distribution of the hosts (Kempf et al., 2009; Wessels et al., 2019). In general, it is assumed that generalist parasites relying on a wide range of hosts tend to show weak or no genetic structure, as shown in many studies on various parasite species (e.g. Archie and Ezenwa, 2011; Wessels et al., 2019). The tick species *I. ricinus* is a

85 generalist ectoparasite infesting a wide range of hosts, such as reptiles, mammals, 86 and birds (Casati et al., 2008; Norte et al., 2012). It has been proposed that tick 87 abundance and population genetic structure are dependent on the species' biology 88 (such as reproduction strategies and life cycle), but also on the host distribution 89 and behaviour (Kempf et al., 2011; McCoy et al., 2001; Rizzoli et al., 2009; Norte et 90 al., 2012). Large ungulates, such as deer, bovidae, and wild boar may be highly 91 efficient carriers of ticks for long distances, as long as there are no severe barriers 92 to their migration (Handeland et al., 2013; Kriz et al., 2014). By contrast, 93 transportation of ticks by migrating birds seems to be less efficient across 94 contiguous landmasses (Hasle et al., 2009; Røed et al., 2016). Based on these 95 findings, it is expected that *I. ricinus* populations should show a weak spatial genetic 96 structure.

97 Regarding previous works on population structure and dispersal of I. ricinus, Noureddine et al. (2011) found a clear differentiation between European 98 99 and African populations using sequences from three nuclear and three 100 mitochondrial markers. Regarding the results from that study, it was later 101 suggested by Estrada-Peña et al. (2014) that those northern African samples could 102 correspond to Ixodes inopinatus, a sibling species of the I. ricinus complex within 103 the Ixodes subgenus. Considering only European populations, some studies showed 104 weak to no differentiation, but an extensive genetic diversity was observed within 105 each local population (Casati et al., 2008; Noureddine et al., 2011; Porreta et al., 106 2013; Carpi, 2016). Other investigations analysing the frequency of mitochondrial 107 haplotypes showed a marked phylogeographical structure in northern Europe,

108 notably when considering populations from the north of the UK (Scotland) and 109 Scandinavia (Al Khafaji et al., 2019; Dinnis et al., 2014; Røed et al., 2016). Although 110 none of the mitochondrial haplotypes was exclusive to any of those populations, 111 their frequencies varied significantly between populations from different regions. 112 Interestingly, the British clade identified by Røed et al. (2016) coincides with the 113 occurrence of a particular subtype of the louping-ill virus, which is closely related 114 to other Irish and Spanish subtypes. Other studies focusing on the genetic structure 115 of *I. ricinus* populations were based on microsatellite loci (Kempf et al., 2009; 116 Kempf et al., 2011). Microsatellite variations have led to the identification of 117 significant levels of genetic structure at different spatial scales, deviation from 118 panmixia in I. ricinus populations likely due to assortative mating and patterns of 119 host use (see Araya-Anchetta et al., 2015 for a review). However, those studies have 120 also assessed patterns of genetic variation from localised samples that cover only a 121 subset of the species range and thus likely do not capture the entire species genetic 122 structure at the continental level.

123 Here, we aim to elucidate the population genetic structure of the tick I. 124 ricinus based on single nucleotide polymorphisms (SNPs). To the best of our 125 knowledge, no other study on the population genetic structure of I. 126 ricinus throughout the Eurasian continent was based on the variation detected by 127 this type of marker. Although generally having a weaker mutation rate than 128 microsatellites, SNPs offers the possibility of building a larger range of markers and 129 have been suggested to be more reliable markers for population genetic studies 130 (Helyar et al., 2011; Smouse, 2010). Our main objective is to describe the genetic

structure of *l. ricinus* populations to infer the geographical and environmental factors shaping this structure. Particularly we hypothesized that (i) *l. ricinus* from the western parts of Europe might have genetic similarities to the Great Britain lineage (Røed et al., 2016) while (ii) there should be a pronounced genetic differentiation between ticks south and north of the extensive mountain areas covering central Europe (i.e., the Eastern Alps, the Western Alps, the Carpathian Mountains, and the Balkan Mountains).

138

139 Materials and Methods

140 Sampling

A total of 28 tick populations from 20 countries were sampled covering most of the 141 142 species' range, including populations close to the northern (Norwegian, Sweden, 143 Ireland, and England) and southern (Iran, Spain, and northern Africa) range limit 144 of I. ricinus (Figure 1). Samples were collected by flagging inside or near forest 145 fragments from the ground vegetation and were preserved in alcohol. A significant 146 subset of the sampled populations we used, covering 8 regions across Europe 147 (southern and northern France; Belgium; western and eastern German; southern and central Sweden; and northern Estonia), originated from a single project 148 149 (smallFOREST, **BiodivERsA** 2010-2011 Joint 150 call: https://www.biodiversa.org/491/download) and was sampled by the same 151 person during the same year 2013 (See Ehrmann et al., 2018 for details). The 152 remaining samples were collected for different projects (for details on those

projects see Røed et al., 2016 for the Norwegian samples and Noureddine et al.,
2011 for the remaining samples). The coordinates of the sampled populations are
provided in Table S1 (see Supporting Information). Aside from smallFOREST
samples, sampling dates varied among the sampled populations (Table S1).

157 Ticks sampled for those projects were identified at the laboratory using 158 standard morphological keys provided in Babos (1964), Hillyard (1996), or Perez-159 Eid (2007). As most samples we used were identified before the description of I. 160 inopinatus (Estrada-Peña et al., 2014) and considering that it was impossible to re-161 evaluate the identification of samples based on morphological features, we 162 conducted an *a posteriori* evaluation of the potential presence of *I*. 163 inopinatus among our samples. To fulfil this aim, northern African I. 164 inopinatus samples analysed by Noureddine et al. (2011) were included in the 165 present study.

166 DNA extraction and SNP genotyping

167 Since ticks and DNA samples analysed in this study had different origins and 168 therefore different storage methods, three different methods were used to ensure 169 DNA extraction. Ticks were either: (i) frozen and crushed with a pestle in individual 170 tubes before extracting DNA using DNeasyTM Tissue Kit (Qiagen); (ii) disrupted 171 using a Tissue Lyser (Qiagen) before DNA extraction using the Wizard Genomics 172 DNA Purification Kit (Promega, USA); or (iii) crushed with Lysing matrix H (MP 173 Biomedicals, Santa Ana, USA) before extracting DNA with MagNA Pure LC Total 174 Nucleic Acid Isolation Kit (Roche, Basel, Switzerland).

175 We genotyped 192 SNPs as described by Quillery et al. (2014). The list of 176 SNPs, variant basis, and primers are presented in Table S2. All samples were 177 amplified by whole genome amplification (WGA) before genotyping. The PEP-PCR 178 WGA kit (LGC-Biosearch Technologies) was used for whole genome amplification of 179 each sample. The WGA protocol associated with KASP genotyping has already been 180 tested by Quillery et al. (2014) and showed a reduced number of "no-call" data 181 (missing values) during genotyping. The WGA and genotyping steps were 182 subcontracted by the GENTYANE platform (INRA, Clermont-Ferrand, France: 183 http://gentyane.clermont.inra.fr/). The GENTYANE platform is an INRAE (French 184 National Research Institute for Agriculture, Food and Environment) research facility 185 located in Clermont-Ferrand (France) which offers sequencing and genotyping 186 services. Genotyping was conducted in a Biomark HD System (Fluidigm) and KASPar 187 assays. The KASPar method is a KBiosciences competitive allele-specific PCR 188 amplification. A PCR mix containing two allele-specific forward primers and one 189 common reverse primer was carried out. Each forward primer had a 5' tail sequence 190 homologous to universal secondary oligos labelled with a fluorophore (FAM or 191 HEX). If a particular locus is homozygous, only one fluorescent signal is generated. 192 Bi-allelic loci generate both fluorescent signals.

193 *Quality control*

Data was filtered after genotyping and before statistical analysis. First, all invariant
SNPs were removed. After this first filtering step, all individuals with more than 20%
of non-amplified sites (missing data) were removed. Finally, all remaining SNPs with
more than 20% missing data were also removed. The remaining dataset consisted

198 of both individuals and SNPs with less than 20% missing data. After quality control

199 steps, 125 SNP loci and 497 individuals were kept for further analyses.

200 Cluster analysis and genetic structure

201 Two complementary clustering methods were used to access the genetic structure 202 of I. ricinus populations. First, we investigated the genetic clustering by performing 203 a discriminant analysis of principal components (DAPC, Jombart et al., 2010) with 204 the package 'adegenet' (Jombart, 2008) in R (R Core Team, 2019). The 205 optimal k number of clusters was identified by the k-means algorithm using 206 the find.cluster() function based on BIC values. A maximum of 28 clusters was 207 allowed, i.e. the total number of sampled populations. Next we performed a 208 Bayesian analysis in STRUCTURE (Pritchard et al., 2000) with the parameter K, i.e. 209 the optimal number of clusters, varying from 1 to 10, according to the results from 210 the DAPC. We used a non-admixture model with the sampling locations as prior. 211 Twenty repetitions of 80,000 MCMC iterations with a burning length of 20,000 212 iterations were run for each value of K. The results were analysed with Structure 213 Harvester (Earl and vonHoldt, 2012). The best K value for the optimal number of 214 clusters was identified by comparing the estimates of log probabilities of the data 215 (i.e. $\ln[\Pr(X|K)]$) for each K value as well as Evanno's delta K method (Evanno et al., 216 2005). Pritchard et al. (2007) suggested aiming for the smallest value of K that 217 captures most of the genetic structure in the data. Assigning probabilities for 218 individuals and populations across repetitions were then averaged in CLUMPP 219 (Jakobsson and Rosenberg, 2007). We applied a hierarchical clustering analysis (e.g. 220 Vähä et al., 2007) in each identified cluster to detect more refined patterns of

genetic structure. Hierarchical analysis in STRUCTURE was realised with ten repetitions and the same other parameters as the first round of analysis. We realised a similar analysis for each cluster identified by DAPC.

224 To test our data for isolation by distance (IBD), pairwise F_{ST} values were 225 estimated with the package 'hierfstat' (Goudet and Jombart, 2018) in R (R Core 226 Team, 2019) as Weir and Cockerham unbiased parameter θ (Weir and Cockerham, 227 1984). The IBD pattern was first tested across all pairs of Eurasian samples and 228 second only between pairs of samples collected during the same year to avoid 229 potential biases due to temporal variability in dispersal and genetic structure. Those 230 corresponded to samples from southern and northern France, Belgium, western 231 and eastern German, northern Estonia, southern and central Sweden, a total of 8 232 samples (28 pairs). Since the 25 Eurasian samples are distributed across a large 233 continental extent, pairwise geographical distances were calculated with the 234 'geosphere' package (Hijmans, 2017) in R (R Core Team, 2019) to account for the 235 curvature of the Earth. The strength of the IDB was evaluated as the relationship 236 between $\theta/(1-\theta)$ and the natural logarithm of the geographic distance as 237 described by Rousset (1997). In a two dimensions population, the slope parameter b of the linear regression $\theta/(1-\theta) = a + bD_{Geo}$ is inversely proportional to the 238 239 average neighbourhood size Nb = 1/b, and $b = 1/(4D_e \pi \sigma^2)$, where D_e is the sub-240 population density and σ^2 is the averaged square axial distances between adults 241 and their parents and σ is half the average adult-parent distance (Séré et al., 2017). In this case, a proxy of dispersal can be calculated as $\delta \approx 2 * \sqrt{(4\pi Deb)}$ 242 (Manangwa, 2018). The population density was calculated as $D_e = N_e/S\pi$, where 243

244 S is the smallest distance between sites considered and included in the IBD analysis. 245 We used NeEstimator version 2.1 to calculate effective population sizes (Ne) by 246 applying two different methods, one based on linkage disequilibrium and another 247 based on molecular co-ancestry (Do et al, 2014). We calculated the mean of Ne 248 estimated with these two methods after the exclusion of 'infinity' results. The 249 obtained mean value was weighted by the number of times one of the two methods 250 generated a non-infinity value. The significance of the IBD pattern was assessed by 251 Mantel tests as implemented in the 'vegan' package (Oksanen et al., 2019) in R (R 252 Core Team, 2019).

253 *Genetic diversity*

254 For each locus, we estimated the observed heterozygosity (Ho), the gene diversity 255 (Hs), and Wright's fixation indices F_{IS}, F_{ST}, and F_{IT}. Wright's statistics measure 256 inbreeding in three levels of population structure: F_{IS} is the inbreeding coefficient 257 of individuals relative to subpopulations; F_{ST} is the inbreeding coefficient of 258 subpopulations relative to populations; and F_{IT} is a measure of the inbreeding of 259 individuals relative to populations. All metrics were calculated with the package 260 'hierfstat' (Goudet and Jombart, 2018) in R (R Core Team, 2019). A Monte-Carlo 261 permutation test (999 replicates) was conducted to test for the significance of the 262 differences of mean gene diversity and F_{IS} values over loci between pairs of genetic 263 clusters identified. For each replicate, individuals were randomly assigned to one 264 genetic cluster and the simulated statistics were calculated. We ran 265 the randtest() function from the 'ade4' package (Dray and Dufour, 2007) to access 266 the significance of the observed differences.

267	To investigate null alleles and possible Wahlund effect on genotype
268	frequencies, we followed the procedure proposed by De Meeûs (2018). According
269	to that study, the presence of null alleles could be identified by a suit of
270	comparisons of F_{IS} , F_{ST} , and the number of missing data. In case of null alleles, we
271	would observe: (i) a high positive correlation between F_{IS} and F_{ST} ; (ii) high variation
272	of both F_{IS} and F_{ST} across loci; (iii) F_{IS} standard errors (StrdErrFIS) much bigger than
273	F_{ST} standard errors (StrdErrFst); and (iv) F_{IS} values mainly explained by the presence
274	of missing data. For the Wahlund effect, the correlation between F_{IS} and F_{ST} should
275	approximate zero, a small variation of F_{ST} and a moderate variation of F_{IS} should be
276	observed across loci, F_{IS} standard errors (StrdErrFIS) should be higher than F_{ST}
277	standard errors (StrdErrFst) and no or rare missing data should be obtained. To test
278	those relations, values of F_{IS} , F_{ST} , StrdErrFst, and StrdErrFIS were calculated in the
279	FSTAT software version 2.9.4 (Goudet, 2003), the latter values calculated by
280	Jackknife. The Spearman's rank correlation test was applied to test for correlations.
281	Finally, De Meeûs (2018) suggested a linear regression between F_{IS} and missing data
282	to quantify, using the R^2 value, the contribution of missing data in F_{1S} values.
283	Because the Wahlund effect can produce between-locus dependencies, we also
284	tested linkage disequilibrium for each pair of loci by using G-based tests
285	implemented in FSTAT 2.9.4. Since <i>p</i> -values from each test are not independent,
286	we applied the procedure described by Benjamini and Yekutieli (2001) to calculate
287	the false discovery rate (FDR) and correct <i>p</i> -values.

289 Results

290 Clustering analysis, genetic differentiation and isolation by distance

291 The DAPC analysis identified two possibilities for the number of clusters, one 292 suggesting three different genetic clusters and the other suggesting four genetic 293 clusters (the BIC difference is 0.842 between K = 3 and K = 4, Figure S1). Choosing K 294 = 4 clusters created two overlapping groups, while K = 3 grouped individuals into 3 295 well-separated clusters (Figure 2). Hence, we decided to set the number of clusters 296 to K = 3 with the DAPC approach. Bayesian analysis performed with STRUCTURE 297 also identified a K = 3 differentiated genetic clusters (Figure 2b and Figure S2) 298 whose compositions are very similar to the three clusters retained with the DAPC 299 approach. In both analyses, northern African (yellow colour in Figures 2 and 3) and 300 Eurasian populations (the other colours) were highly differentiated. Two main 301 groups were identified within Eurasia, one corresponding mainly to northern and 302 continental middle European populations (grey colour in Figures 2 and 3), the other 303 corresponding mainly to southern and western populations in Eurasia (blue colour 304 in Figures 2 and 3). The DAPC approach separated northern African populations 305 from Eurasian ones along the first axis, while Eurasian clusters were mostly 306 separated along the second axis (Figure 2a). Regarding clustering analyses with 307 STRUCTURE, individual probabilities of different K values ranging from 2 to 10, 308 excepted for K = 3 which is already depicted in Figure 2b, are presented in the 309 Supporting Information (see Figure S3).

310 Finer genetic structure was identified from our hierarchical analyses 311 (Figures S4 and S5). These analyses, either carried out with DAPC (Figure S4) or the 312 STRUCTURE approach (Figure S5), were able to isolate Iran and/or Turkey from the 313 other sampled sites within the southern Eurasian cluster. Atlantic sites (Spain, 314 southern and western France, Ireland, and England) were further isolated from the 315 remaining sites in this group (Italy, Romania, Hungry, and Slovakia). The northern 316 European sites showed a more admixture structure, and separation in further 317 clusters varied between the DAPC and STRUCTURE approaches (see the 318 'Hierarchical analysis' section in the Supplementary Information for more details).

319 A pattern of isolation by distance (IBD) was observed across all sampled 320 populations (Mantel r = 0.726, p < 0.001). Restricting the IBD analysis to the set of 321 sites sampled during the same year, we found an even stronger pattern of IDB 322 (Mantel r = 0.870, p < 0.0001, Figure 4). In the latter case, the coefficient estimate 323 of the slope parameter (b) in the regression was b = 0.01 with a 95% confidence 324 interval (CI) ranging from 0.007 to 0.013. Neighbourhood size (Nb) reached Nb = 99 325 individuals, on average (95% CI = [71-140]), and immigration rate ($N_e m$) was 326 estimated to reach $N_em = 16$ (95% CI = [11-22]) individuals per generation and 327 subpopulation.

We found a mean effective population size of 62 individuals. The closest sampled sites were North France and Belgium, separated 119 km from one another. We found surface and population densities to reach, on average, $S^2 = 11.3$ km² and

331 De = 5.4 individuals/m², respectively. We found the dispersal rate to reach, on 332 average, $\delta \approx 76$ km/generation (95% CI = [65-90]).

333 Genetic diversity

334 The observed heterozygosity (Ho), gene diversity (Hs), and F_{IS} were highly variable 335 across loci (Table S3). The observed F_{ST} values were, however, more constant than 336 F_{IS} ones. For most loci, gene diversity was higher than the observed heterozygosity. 337 Consequently, the overall gene diversity across all loci was significantly higher than 338 the observed heterozygosity (Wilcox Signed-Rank Test, V = 6959, p < 0.0001). The 339 mean gene diversity per sampled population was still higher than the observed 340 heterozygosity (Wilcox Signed-Rank Test, V = 406, p < 0.0001) and mean F_{IS} was 341 always positive. Mean values of observed heterozygosity, gene diversity, and F_{IS} for 342 each population are shown in Figure S6 (Supporting Information). The highest mean 343 gene diversity and F_{IS} values over loci were identified in the southern Eurasian 344 cluster (Hs = 0.355, F_{IS} = 0.275), followed by the northern European cluster (Hs = 345 0.340, F_{IS} = 0.2708) and the cluster from northern Africa (Hs = 0.171, F_{IS} = 0.191) 346 (Figure 5). The Monte-Carlo test showed a significant difference in gene diversity 347 values for all pairs of clusters (p = 0.001 for all three comparisons), but none for F_{IS} 348 values (p = 0.199 and 0.239 when comparing northern Africa to the northern 349 European cluster and northern Africa to the southern Eurasian cluster, respectively; 350 while p = 0.644 when comparing the southern Eurasian cluster to the northern 351 European cluster). Populations from northern Africa showed a high deficit in 352 heterozygosity, of which 71 out of 125 loci with Hs values of zero.

353 After *p*-value correction (Benjamini and Yekutieli, 2001), no pair of locus 354 showed significance values of linkage disequilibrium. No correlation was found 355 between F_{IS} and F_{ST} ($\rho = -0.0206$, p = 0.8198) and missing data were positively 356 correlated to F_{IS} values ($\rho = 0.5804$, p < 0.001). The linear regression of F_{IS} against 357 the number of missing data estimated an adjusted R^2 of 0.19, suggesting that 358 around one-fifth of F_{IS} variance is explained by the number of missing data. Finally, 359 StrdErrFIS was around 4 times bigger than StrdErrFst (0.033 and 0.008, 360 respectively).

361 Discussion

362 We investigated the genetic structure of populations from the tick *I. ricinus* in much 363 of its range, i.e. in Eurasia and in northern Africa. In addition to a strong and 364 expected divergence between northern African and Eurasian populations, the two 365 Eurasian genetic clusters described here showed clear spatial patterns. The 366 isolation by distance patterns we found, either throughout the entire dataset or 367 restricted to samples from the same period, suggest an association between the 368 genetic structure of *I. ricinus* populations and the geographical location of these 369 populations. Hierarchical analyses confirmed the genetic affinity between western 370 European populations, from the UK and Ireland in the north to Spain in the south, 371 supporting our first hypothesis regarding genetic similarities in western continental 372 Europe and the British Isles. Also consistent with our second hypothesis stating a 373 genetic signature of central European mountains, we found a clear differentiation 374 between populations from southern Eurasia and populations from northern

Europe. Indication of migration of individuals between the two clusters is suggested
by the different degrees of affinity from central Europe with one cluster or another
(e.g. in Romania, Hungary, Slovakia, and Moldova).

378 Ixodes ricinus and I. inopinatus have recently been suggested to be 379 sympatric both in northern Africa (Younsi et al., 2020) and in Europe (Estrada-Peña 380 et al., 2014; Chitimia-Dobler et al., 2018). Our results are clear concerning the 381 genetic identity of northern African samples. According to the results from both the 382 DAPC and STRUCTURE analysis, there is no possibility of any individuals from those 383 populations to belong to any other genetic clusters. Also, no individual from Eurasia 384 had any probability of identity with the northern African cluster. Converging results 385 of both analyses indicate with a great deal of certitude that: (i) all samples from 386 northern Africa belong to the same species and have the same ancestry; (ii) no 387 sample in Eurasia share ancestry with northern African ones. Northern African 388 samples were also a particular case as more than half loci were monomorphic 389 across all three populations, which was not found in Eurasian populations. Again, it 390 is important to note that individuals from the three northern African populations 391 analysed here were identified before the description of I. inopinatus (Estrada-Peña 392 et al., 2014). If I. inopinatus was present in the Eurasian samples, we would expect 393 at least small probabilities of identity of Eurasian samples with the northern African 394 cluster, which was not the case. The clear-cut genetic differentiation we obtained 395 between Eurasian and northern African populations strongly suggests that all the 396 individuals from the three northern African populations analysed here correspond

to *I. inopinatus.* Those results also illustrate the potential of using some of the SNPs
analysed here to differentiate at a molecular level the two *Ixodes* species.

399 Two previous studies covering a large spatial extent of *I. ricinus'* range 400 (Noureddine et al., 2011; Porreta et al., 2013) did not find such a clear geographical 401 structure between Eurasian populations. Several reasons may explain this 402 difference. First, a somewhat reduced number of individuals per population 403 (sometimes a single individual per population in Nourredine et al., 2011) may 404 explain a lack of spatially structured signal in former studies. Second, those former 405 studies were based on mitochondrial and nuclear sequences. This said, a marked 406 genetic differentiation into two distinctive clades has already been reported (Dinnis 407 et al., 2014; Røed et al., 2016), suggesting a split in *I. ricinus* populations between 408 northern continental Europe and Great Britain. Our results confirm and extend this 409 pattern to most of the Eurasian range of the species by suggesting that 410 Scandinavian populations are genetically closer to the populations from the north-411 eastern continental parts of Europe. Although there is a certain degree of gene flow 412 between the two clusters, the north vs. south-eastern exchange may be hampered 413 by mountain areas in central Europe. This reinforces the argument that large 414 animals efficiently maintain high gene flow between tick populations across 415 contiguous and permeable landscapes, while intense transportation by birds, 416 during spring and autumn migration across sea or mountains (Hasle et al., 2009; 417 Røed et al., 2016), may not be as sufficient to break down boundaries between 418 established genetic entities.

419 Surprisingly, we found a close genetic affinity between all Atlantic samples 420 (i.e. Ireland, England, western and southern France, and Spain) and the 421 geographically separated populations from Turkey and Iran. This genetic affinity 422 among distant populations in Eurasia was supported by the two different clustering 423 methods we used (DAPC and STRUCTURE). Besides these results, the refined 424 hierarchical analyses isolated Iran and Turkey in their particular clusters in the first 425 (DAPC) and second (STRUCTURE) round of hierarchical clustering analyses. This 426 suggests that an east-west transport of ticks across southern Eurasia must be 427 sufficient to maintain a genetically identifiable cluster across this extensive area. 428 Interestingly, louping-Ill like viruses are also known from Greece and Turkey (Gao 429 et al., 1993; Marin et al., 1995), which might further support our findings and a link 430 between tick lineages and *Flavivirus*, although the causation is not known.

431 Since migratory birds carry I. ricinus across long distances, different 432 migratory routes could also contribute to the north-south genetic differentiation 433 we observed (Hasle et al., 2009; Røed et al., 2016). However, birds mainly carry 434 larvae and nymphs. Since surviving rates between development states are low, the 435 overall reproductive success of tick transported by birds is likely smaller than that 436 of adult ticks carried by large mammals. This may explain the maintenance of 437 genetic differentiation e.g. between the UK and Norway despite massive transport 438 of ticks' larvae in both directions (Røed et al., 2016).

Regarding the population structure observed within samples, the deviation
from Hardy-Weinberg equilibrium we found is in agreements with previous studies
on population genetics of *I. ricinus* based on SNPs (Quillery et al., 2014) and

442 microsatellites (Kempf et al., 2009; Kempf et al., 2011; Røed et al., 2006), as well as 443 other tick species (Dharmarajan et al., 2011). Possible causes of the observed 444 deviation from the Hardy-Weinberg equilibrium are assortative mating (or 445 assortative pairing), Wahlund effect, or errors in the genotyping. A tendency of 446 mating between phenotypically or genetically similar individuals may effectively 447 increase the inbreeding and thus heterozygote deficiency within populations (Jiang 448 et al., 2013). Kempf et al. (2009) suggested that assortative mating might occur in 449 *I. ricinus*, mostly via host selection (Kempf et al., 2011). Inbreeding in ticks could be 450 a result of host infestation by related individuals, which leads to high breeding 451 success of sibling groups (Araya-Ancheta et al., 2015). The highly aggregated egg 452 masses in I. ricinus (1000 to 3000 eggs) and the limited active dispersal of larvae 453 and nymphs may lead to a high likelihood of mating between related individuals 454 and thus inbreeding. Finally, the parasite-host relationship specificities could also 455 play an important role in establishing or maintaining population structure in I. 456 ricinus. If different host populations are present locally and exhibit behaviours 457 favouring mating within (and not between) each host population, this may induce 458 a Wahlund effect and explains the heterozygote deficiency observed. The existence 459 of such a host population behaviour has been characterized in I. uriae, a tick 460 associated with sea birds (Mc Coy et al., 2001) but also suggested in I. ricinus (Kempf 461 et al., 2009, 2011). Even though we did not conceive this study to test for such a 462 hypothesis, our results support at least partially non-random mating in I. ricinus 463 populations and the consequent Wahlund effect. Dharmarajan et al. (2011) facing 464 a similar result for the American species I. texanus showed that subdivided

breeding groups and high variance in individual reproductive success can correctlyexplain Hardy-Weinberg equilibrium deviation.

It is widely acknowledged that more or less isolated populations could 467 468 develop particular adaptations in response to environmental differences between 469 habitats. Nonetheless, very few studies to date have clearly observed phenotypic 470 variations among *I. ricinus* populations from different geographical origins. In 471 Estrada-Peña et al. (1996 and 1998), differences in cuticular hydrocarbon 472 composition among European populations of *I. ricinus* were observed according to 473 the geographical origin of those populations. Interestingly, the multivariate 474 phenotypic analysis presented in those studies showed a somewhat similar pattern 475 to our hierarchical genetic clustering analysis, notably concerning what the authors 476 call 'peripheral populations'. Aside from chemical differentiation, behavioural 477 differences between ticks' populations have also been documented, such as 478 mismatches in questing peaks (Schulz et al., 2014) and questing responses to 479 temperature (Gilbert et al., 2014; Tomkins et al., 2014). In controlled conditions, 480 Gilbert et al. (2014) and Tomkins et al. (2014) showed that I. ricinus nymphs from 481 cooler climates begin questing at lower temperatures than nymphs from warmer 482 climates. They also start questing sooner when the temperature was kept constant. 483 In any case, local adaptations could impact the spatial redistribution of the species 484 range in response to changes in abiotic conditions. In a global changing context, 485 such consequences could be explored by environmental niche modelling to identify 486 areas of potential future expansion. It remains to be investigated if the different

487 clusters we identified here could pose different threats for human health and the488 potential risk of tick-borne disease transmission to humans.

Our findings on isolation by distance suggest small population densities and large dispersal distances among the sampled populations. The large dispersal distance is not a surprising result since ticks can parasitize highly mobile species. In a changing climate context, this result indicates that ticks could easily colonize new suitable habitats outside the current limits of the species geographical range in a few generations.

495 Despite being a generalist ectoparasite, our results highlight geographically 496 distinct and genetically structured populations in *I. ricinus*. More research on host 497 preference and dispersal capacity is needed to better understand those patterns. 498 The differentiation of Eurasian populations into two geographically distinct clusters 499 (northern Europe vs. southern Eurasia) could have important implications for the 500 redistribution of *I. ricinus* in response to anthropogenic climate change. Ticks from 501 a given genetic cluster could be more or less prone to increase in abundance in 502 some regions. Combining tick and pathogen population genetics with knowledge 503 on host distribution could help in the early detection of the spread of tick-borne 504 diseases and thus improve the responsiveness of public authorities to limit major 505 public health concerns.

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530 Conflict of Interest

531 The authors declare that they have no conflict of interest.

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Figure 1. Distribution of the sampled populations of *Ixodes ricinus* across its
putative range. The range of *I. ricinus* is displayed in dark orange on the map and
was adapted from the European Centre for Disease Prevention and Control – ECDC
(January 2019). The size of each blue dot on the map is proportional to the sample
size of each sampled population.



814 Figure 2. Cluster assignment analysis results based on either the DAPC scatter plot 815 of individual memberships for K = 3 (a) or the STRUCTURE individual membership 816 probabilities for K = 3 as described by Evanno et al. (2005) (b). The sampled 817 populations are coded as follows: MAR: Morocco; DZA: Algeria; TUN: Tunisia; ESP: 818 Spain; IRN: Iran; TUR: Turkey; FRA-W: West France; IRL: Ireland; FRA-S: South 819 France; GBR-BP: England Blue Pool; GBR-BR: England Bristol; ITA-D: Italy 820 Domodossola; ITA-V: Italy Varese; ROU: Romania; HUN: Hungary; SVK: Slovakia; 821 MDA: Moldavia; FRA-N: North France; DEU-W: West Germany; BEL: Belgium; EST-822 S: South Estonia: DEU-E: East Germany; DEU-S: South Germany; SWE-S: South 823 Sweden; SWE-C: Central Sweden; NOR-So: Norway Søgne; NOR-Gr: Norway



Grønnsundfjellet; EST-N: North Estonia. Coordinates of sampled populations are



Figure 3. Distribution of the relative importance of each cluster on each sampled population (see Figure 2 for the groups which colors are matching). Results are provided for both the DAPC (a) and the STRUCTURE (b) analysis.



Figure 4. Isolation by distance between all Eurasian samples. Red triangles represent the pair of samples from the same year: South and North France, Belgium, West and East German, North Estonia, South and Central Sweden. The regression line (plain line), 95% confidence interval (CI) calculated by bootstrap (dashed lines), Mantel test significance and regression equation corresponds only

to red triangles pairs of samples are also shown. Black points correspond to all other

837 pairs of samples not used for further IBD analysis.



Figure 5. Values of gene diversity (a) and F_{1S} (b) for each of the three genetic clusters identified by DAPC. Yellow: northern Africa cluster; Blue: southern Eurasia; Grey: northern Europe. Permutation test (Monte-Carlo test, 999 replicates) between all pairs of clusters was significant for gene diversity (p = 0.001) but no significance was identified for F_{1S} . Eurasian clusters show a more pronounced heterozygote excess than the northern African one. A variation of F_{1S} values across loci was observed in

- the three clusters, even though this variation was much larger in the northern
- 846 African cluster.

Supporting Information for:

Strong genetic structure among populations of the tick Ixodes ricinus across its range: insights from population genetics

Pedro Poli, Jonathan Lenoir, Olivier Dr. Plantard, Steffen Ehrmann, Knut H. Røed, Hans Petter Leinaas, Marcus Panning, Annie Guiller

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Materials and Methods

Table S1. Sample coordinates. The Reference column indicates from which source samples were made available. PC = Personal collection.

Sample locality	Code	Longitude	Latitude	Number of samples	Sample Date	Reference
Morocco	MAR	4221933.21	1519759.51	10	Before 2010	Dr. Plantard, PC
Algeria	DZA	4165854.78	1520079.18	8	Before 2010	Dr. Plantard, PC
Tunisia	TUN	4287083.09	1370080.62	13	Before 2010	Dr. Plantard, PC
Spain	ESP	3292343.37	2302053.84	19	Before 2010	Dr. Plantard, PC
Iran	IRA	7920535.19	2511813.36	13	Before 2010	Dr. Plantard, PC
Turkey Istambul	TUR	5907775.11	2200447.26	9	Before 2010	Dr. Plantard, PC
North France	FRA-N	3872010.67	2994279.45	40	2013	Erhmann et al., 2018
West France	FRA-W	3465235.38	2853298.78	15	2016	Dr. Degeilh, PC
South France	FRA-S	3593881.21	2296634.56	17	2013	Erhmann et al., 2018
Ireland	IRL	3013710.61	3385835.15	20	Before 2010	Dr. Plantard, PC
England Blue Pool	GBR-BP	3470079.25	3130233.33	19	Before 2010	Dr. Plantard, PC
England Bristol	GBR-BR	3450947.31	3224484.53	19	Before 2010	Dr. Plantard, PC
ltaly Domodossola	ITA-D	4188665.99	2556599.15	11	Before 2010	Dr. Plantard, PC
Italy Varese	ITA-V	4229419.76	2523525.45	10	Before 2010	Dr. Plantard, PC
Romania	ROU	5643875.12	2813096.13	9	Before 2010	Dr. Plantard, PC
Hungary	HUN	5064737.95	2796444.23	18	Before 2010	Dr. Plantard, PC
Slovakia	SVK	5008087.64	2900574.08	13	Before 2010	Dr. Plantard, PC
Moldavia	MDA	5711169.6	2856440.17	10	Before 2010	Dr. Plantard, PC

Sample locality	Code	Longitude	Latitude	Number of samples	Sample Date	Reference
West Germany	DEU-W	4257417.83	3352915.67	24	2013	Erhmann et al., 2018
East Germany	DEU-E	4462732.5	3348531.08	38	2013	Erhmann et al., 2018
South Gemany	DEU-S	4440340.3	2784710.43	14	2013	Dr. Plantard, PC
Belgium	BEL	3924610.12	3095109.35	18	2013	Erhmann et al., 2018
North Estonia	EST-N	5186688.83	4032319.73	49	2013	Erhmann et al., 2018
South Estonia	EST-S	5313297.86	3950296.69	14	Before 2010	Dr. Plantard, PC
South Swqeen	SWE-S	4533959.53	3622513.31	20	2013	Erhmann et al., 2018
Central Sweden	SWE-C	4720133.45	4047795.89	19	2013	Erhmann et al., 2018
Norway West	NOR-W	4186225.49	3886420.36	15	2006	Dr. Leinaas, PC
Norway East	NOR-E	4389275.28	4003811.98	13	2006	Dr. Leinaas, PC

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
1133	T/C	GCTTGGCCACTTCCACTGCTTT	GCTTGGCCACTTCCACTGCTTC	ACAACAGAGAAGGCAGCCCACA
3705	A/C	AGCATGGCGCACTGTGAAAGCTC	AGCATGGCGCACTGTGAAAGCTA	TCCTAGTCGGCTGGCTGGAG
6283	T/C	AATGAGGCGTCAGTGACAGCATAAC	AATGAGGCGTCAGTGACAGCATAAT	CGTGACGTCAAGGCAGAATGCTAT
6363	A/G	TCGTCCTCCGTCACGTAGCCG	TCGTCCTCCGTCACGTAGCCA	CCATTGAACCCTGGTGGGTCATCA
10041	A/G	GTTGTTCCCTTGGCAGACG	GTTGTTCCCTTGGCAGACA	AACATACCCGAGACTGTCAAC
19998	A/G	CAGAAGTGGAGATTGTTGCGTGTG	CAGAAGTGGAGATTGTTGCGTGTA	TACATACATTGAGCATCGACCAA
				AGGCACGTAGATCACGAGAATTATT
21130	C/T	GCTGCTGCAACCGGTTTATCTTC	GCTGCTGCAACCGGTTTATCTTT	TC
30736	C/G	GCTAGGTGACGAGGACTGGACG	GCTAGGTGACGAGGACTGGACC	GTTGTTCCACCTTTCGCAGGAGAT
31200	A/G	CGTTCAGGTTGACCGAGAAGTAA	GTTCAGGTTGACCGAGAAGTAG	GCCTCTCGTTACTGTCGTATC
		GACTAATCACCAGGAAATCCATTCTG	GACTAATCACCAGGAAATCCATTCTG	
32114	C/T	С	Т	GGCTATACTCGGACGTATGTTGA
32551	T/C	TTCGGTGGCAACAGCTCGTCCATC	TTCGGTGGCAACAGCTCGTCCATT	CCAGCCTCATAGCCGAGCACCA
34502	G/A	CGGATTCGAACCAGTTATCAATGGG	CGGATTCGAACCAGTTATCAATGGA	GCCTCTCTAGAAAACAGTTGCTCTC
42351	A/G	CTTGTAGGAATGGAGGTCATCTTCG	CTTGTAGGAATGGAGGTCATCTTCA	CTTCTGTGTCGCAGGTGGCATCAT
				ACGTGACAACACTTACACGGCATTT
57206	C/G	GCACTATGAGCCATCGAAGCCAAG	GCACTATGAGCCATCGAAGCCAAC	С
60684	C/T	TGCACATAGTCGCGCAATACGTTC	TGCACATAGTCGCGCAATACGTTT	CGAGCCGTTGCAACCGATCCG
61606	G/A	ACATAGGACATCTCAAGGTCATTCG	ACATAGGACATCTCAAGGTCATTCA	GAAGAAACCGAGGATGAGTGTCATG
66390	C/T	GCCGAACAGCCGTGCAACCC	GCCGAACAGCCGTGCAACCT	TCGCTGCTGTATACCCATTG
				TAGAGGTTTCCCAAGTATTTATCGT
68328	G/C	CAGGCAGTTTGCGGTTCACAG	CAGGCAGTTTGCGGTTCACAC	A
68391	A/G	CAGCGTCAAGTTGTGGTGTT	CAGCGTCAAGTTGTGGTGTC	GCATCGCGTGACATTAGTTACA
72226	G/A	GAGGTTCCTGACATGCAGGAAACG	GAGGTTCCTGACATGCAGGAAACA	GCTCTGCAGATGCAAGTTCCAA
77668	G/A	GGAACGTCGTGACAGCCGTAG	GGAACGTCGTGACAGCCGTAA	GGATGGCTTCGAGTTGGACTACTA
78934	G/C	AAAGAAGCGTTTCCCGGTCG	AAAGAAGCGTTTCCCGGTCC	TCTGGCAAAGCAAGCACTCACC
81501	T/C	GTCCTTTCGAAGGTGTATGCATTC	GTCCTTTCGAAGGTGTATGCATTT	ACGATGCTAGTTTGTCAAATAGTG

Table S2. List of SNPs, variant basis and primers used in the study (from Quillery et al., 2014)

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
81758	G/A	ACAAATCTGAAGCAGGCGCGAAAG	ACAAATCTGAAGCAGGCGCGAAAA	AGGACGTCGCCGAGTCGTAGAT
87199	T/C	GCTGGATTGCGTCGTCGCCT	GCTGGATTGCGTCGTCGCCC	CGGCTCTGGCCAGGACCTGATG
93695	G/A	GTCCTAGCCGCTGTCCCGTG	GTCCTAGCCGCTGTCCCGTA	CTGGGACAAACTCTTTCTCGAAGTG
		GCATAAGCAAACTTCAAAGCTTCCAC	GCATAAGCAAACTTCAAAGCTTCCAC	
96296	G/A	G	A	ACGAGGCGGCTCTCATGTACCA
105385	T/G	CCGCGAGCATTTTTGCCACATG	CCGCGAGCATTTTTGCCACATT	TTGACGTCACGACCTATTTGACGAA
113142	A/T	GAGCTCATAGTCCTGAAGACCACA	GAGCTCATAGTCCTGAAGACCACT	TTACGTTGGTCACTATGGGAACGCT
114791	G/C	CGCTGCTAGCAGACGGGAGG	CGCTGCTAGCAGACGGGAGC	GAGAGCGTACACGATTTGCCACGA
116335	A/C	GTGCGTCGAATGTCCAGGTTTATCC	GTGCGTCGAATGTCCAGGTTTATCA	CAAGTTGCGCAAGAGGTGGCAAA
125671	C/T	GTCTGCTTCTGCTATGCTCTGTTTC	GTCTGCTTCTGCTATGCTCTGTTTT	AGCGTCTGCTGCGGAACATCGTA
129322	T/A	CAAGGCAGCGCAGTTCTGACACT	CAAGGCAGCGCAGTTCTGACACA	ATCTGCGTAGCATAAGCCGTGCC
133049	G/A	ACGGGTCGTACAGCGACAAGAG	ACGGGTCGTACAGCGACAAGAA	CGAACATTACAAACGCCGCAAGAGG
137096	T/G	GTGAATGGCAATGCCAGAGTGTAT	GTGAATGGCAATGCCAGAGTGTAG	CTCGGTATTCTGCGGAGCACAA
143089	G/A	GGCACAGGATTTGCTGGTTATAGAGG	GGCACAGGATTTGCTGGTTATAGAGA	GGTGCTATGTGTACCTCACGCC
144259	C/T	GTTGAGTGTCGTGTCCTTCGCC	GTTGAGTGTCGTGTCCTTCGCT	AACAGCTCCTCGTAGACTGCGTAC
145634	C/T	CGGACGCGTGGACGTGACTC	CGGACGCGTGGACGTGACTT	TGGTGACCGTGTGTTGCGCAG
150669	T/C	TGTGCACAAGATGATTCCATAATT	TGTGCACAAGATGATTCCATAATC	GTCATCGGTGATTGTGTCAGTTTAT
		GAATGTGATCGTGGGAGAAGATATAG	GAATGTGATCGTGGGAGAAGATATAG	
155043	G/A	G	A	GCTGTGGAAGCTAAGTGCTCGTTG
159151	C/G	AGACAACGTACGCGCGATTTCAC	AGACAACGTACGCGCGATTTCAG	TGCTAACTGCCAGCGCGTGG
166766	A/G	ATCGACCGGCTGGCTGGCTA	ATCGACCGGCTGGCTGGCTG	GCCTGTTCTTCTGTAAGTCGCTCTA
167418	T/A	TGTCCGATACCTGCCTCCAATTTGTT	TGTCCGATACCTGCCTCCAATTTGTA	TTACCTCCACCGGGTGTCCCAT
175115	T/C	ATGGCAGTGTCAAGAAGGCCAAGT	ATGGCAGTGTCAAGAAGGCCAAGC	CAATGGCAGTGTCAAGGTCGATCTC
176991	C/A	AGAAGCTAGACGCAGAGTTAGGGC	AGAAGCTAGACGCAGAGTTAGGGA	AGGAAGAGTCCAATGTGTGCGCAA
180239	G/T	GTCCTGTGCTGTTGCCGCCG	GTCCTGTGCTGTTGCCGCCT	TGTTCCTGGACGCAAGTCACG
				TCTAAGGCTCCTGGTGTAAGCACAC
189207	T/A	TGGGCGTTGCAGTAATGCAACAGTT	TGGGCGTTGCAGTAATGCAACAGTA	G
197784	C/T	GTTCATTAGAAGCTGTCAGTTGACTC	GTTCATTAGAAGCTGTCAGTTGACTT	CAGTGGCGTAACACGAGAAACTAG
198227	C/T	GACAACATCCAGGGCGAGTTCTAC	GACAACATCCAGGGCGAGTTCTAT	TTGCTATAACCAGTCTTCGACGC

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
205578	G/A	GATGTAGCCCCAGATATACTCAAAG	GATGTAGCCCCAGATATACTCAAAA	ACAGGTACTAAACCAATTTCGGC
		CGAGGTAAGATTGCCACTTATCTTTC	CGAGGTAAGATTGCCACTTATCTTTC	
207995	A/C	С	A	ACCACCTGCCAGTGTTCGACGAT
208593	C/G	GGTCTGGTGCCTGGAAAGTGC	GGTCTGGTGCCTGGAAAGTGG	GGACGCAGTAAACAGAGCAGTCATA
209761	C/T	ACATCATAAGTCACGTGGCCTGAC	ACATCATAAGTCACGTGGCCTGAT	ACGCCGTGACGTCTCCTGAT
		GTGATTCTGCTGGTGATCTTTGTGAT	GTGATTCTGCTGGTGATCTTTGTGAT	
210654	T/C	С	Т	AGCACGCCCAACAAGATCAACGG
212829	C/G	GGCATCTGAACGACATCGTCCACC	GGCATCTGAACGACATCGTCCACG	CGTGTGTCAGGAATGAGAGATAATC
214684	T/C	GTAACGCCGTCACACGGTAAGAC	GTAACGCCGTCACACGGTAAGAT	CTGTCTGATCCAGGCTTTACGCAA
221603	T/C	AGTCGATCATACCTTACTGCTGTGT	AGTCGATCATACCTTACTGCTGTGC	TTCGCGAGTCCGAGTTGCACAGA
				CTATTCCCCTTTCGATCGAACATCG
224277	C/A	ACAGCTAGGAGCAAAGTCCAGTTCCC	ACAGCTAGGAGCAAAGTCCAGTTCCA	G
225377	G/A	TAAAGAGTCGCCTTGGGGAATCTGG	TAAAGAGTCGCCTTGGGGAATCTGA	CACGGACAACAACATTGAACGAG
230247	T/G	GTTTCCAGCTCGCGGTCGATT	GTTTCCAGCTCGCGGTCGATG	GACTGCGTAGAGTGCGCTTTTCAA
233961	A/C	GTCATGCATTTGACAAACTTTGTTA	GTCATGCATTTGACAAACTTTGTTC	GACACTACTAGGGCCTCAATCAA
234508	C/T	TGCTGTCGCTACGCTCGACC	TGCTGTCGCTACGCTCGACT	GAGAGCAGCTCCTGGGAGTCCTTG
236290	T/C	GATGCAATATGTTTACTGGATTCGC	GATGCAATATGTTTACTGGATTCGT	TAGAAATCGGGGCCCCAACGG
243436	T/C	CTTGTGCCTGGCGTCATCTGT	CTTGTGCCTGGCGTCATCTGC	AGGCCCGTGCTCGCTCG
251320	T/A	AGGATCACGTTATACGAAGGCAAGT	AGGATCACGTTATACGAAGGCAAGA	CAAGGATGACAGCACCGGTACGA
255757	T/G	TTCATCGGCGTATCCTTTGAGCGAT	TTCATCGGCGTATCCTTTGAGCGAG	ATGATGGCGACGTAGAGGTAGTTCA
259770	C/G	ACCCTTTTTGAAAGATGAACGTTGTC	ACCCTTTTTGAAAGATGAACGTTGTG	CGTTGCTCAAAGTCAAATGCCAGTG
		GACACTACTAGGGCCTCAATCAAGCA	GACACTACTAGGGCCTCAATCAAGCA	
281206	T/G	Т	G	CAGTCATGCATTTGACAAACTTTG
283680	T/A	GGCGAAACCTTTGAAGCGTTCTTCAT	GGCGAAACCTTTGAAGCGTTCTTCAA	GACAGCGTGATGACTGTTCTTGTG
287805	T/G	CTGCCGCCTGTAATTCCCGACT	CTGCCGCCTGTAATTCCCGACG	TAGGTTCACGACACGAGGTTGATTC
292025	C/T	AACGCCGTGAAAGCCGCGAAC	AACGCCGTGAAAGCCGCGAAT	GCACACCGTACATCACCGAAGCC
296275	C/A	CTGCGTAGAGTGCGCTTTTCAAGGTC	CTGCGTAGAGTGCGCTTTTCAAGGTA	TCGTTTGGTTTCCAGCTCGCGGT
		TTTGTTCAGTTGTCAGAGGTGGCAGT	TTTGTTCAGTTGTCAGAGGTGGCAGT	
298125	A/G	A	G	CCTTGTGGCATGCTCCAGTGATTC

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
		GGTATCCGCTCGCTCGATATGTATAT	GGTATCCGCTCGCTCGATATGTATAT	
299627	C/T	С	Т	CGTGTGCAGCTATCCAAAGACTCG
300752	T/G	AGATGCTGAACTGTCAGATGACGAAT	AGATGCTGAACTGTCAGATGACGAAG	ACCACTGTAGTTGTGTCTCGCTCTG
				CTTGGTTAGTTTCTGCTGGCGTTTT
303781	C/G	CTCCAATTAGCTTCAAATGAATGTTC	CTCCAATTAGCTTCAAATGAATGTTG	С
305888	C/A	GTTTCCTCCACGCAGAGCGAAAGA	GTTTCCTCCACGCAGAGCGAAAGC	CATGCGCTTCGCACTGTCG
307361	T/C	GCGGTATTTTCGGTCAGGC	GCGGTATTTTCGGTCAGGT	GACAAATGTTCGTCGTTCTCAACAG
313057	A/C	AATAGCGGCCAGCAGTTCCTCATA	AATAGCGGCCAGCAGTTCCTCATC	CGAATCCGATAGTGCCGTGAGAGA
		CAAATTTCGTGTTCGTCCATGGCGTG	CAAATTTCGTGTTCGTCCATGGCGTG	
320000	A/C	A	С	CGTGACTTGACGTGACGTGCCA
				CTTTCCCAGTTCAAGCACTCTTTTA
329834	T/G	TAGAAAGCCGGCCCGGATCTT	TAGAAAGCCGGCCCGGATCTG	G
	- / -	GCTCCTCCATGTCTTGTCGTCGTTTC	GCTCCTCCATGTCTTGTCGTCGTCTTC	
333882	T/C	Т	С	CACGGTGGCAGCGGGAA
336267	G/T	GCGTTGTCTGTACATCCGCCAT	GCGTTGTCTGTACATCCGCCAG	GAGCGCAGCGGATACTCTGTTCA
339272	A/G	CCGCACCGGCTTTTACGACA	CCGCACCGGCTTTTACGACG	TCTCGTCGCTGGAGGCGTCAT
	,			ACTGAGTGGTTCTAGTAACGATGGC
340581	C/T	CTGAACCCAACGTTGGCTGAACT	CTGAACCCAACGTTGGCTGAACC	Т
356074	G/A	AAGTATGGGGGAACCCGTGTGA	AAGTATGGGGGAACCCGTGTGG	TAGGAGTTGGAACACTGCGACG
	- /	CATTTGCGATAGGTCGATCACGATAT	CATTTGCGATAGGTCGATCACGATAT	
356395	G/A	G	A	CCGACTTCCGACGCATGTAAAATG
371093	A/G	AGCGATGGCGTCTACCAGCGGA	AGCGATGGCGTCTACCAGCGGG	TTCTGGACTAGCAGCGAGCGAC
374382	T/C	CATGCTTTGTCAACTTTCGAGAT	CATGCTTTGTCAACTTTCGAGAC	TTATGCTGTCAGCTGAGTCCCG
	,			TGTAGAGTGTAGATGCCAGCTTCCT
376474	T/C	AGGTGGCCACTCTGACATGGATC	AGGTGGCCACTCTGACATGGATT	C
380487	C/T	CAGCCGTTCGACGGGATC	CAGCCGTTCGACGGGATT	TCGCTCGTGTCCCTCGTGT
	,	CTGCATGTCTTGGCGTCTGATGTCTT	CTGCATGTCTTGGCGTCTGATGTCTT	GGTTCACTGGCCAAACGCTCCTCTA
393248	T/G	СТ	CG	C
399212	A/G	GTTCAATGGGGCTTCTGCTATCA	GTTCAATGGGGCTTCTGCTATCG	GCGTGAATTCAACGTTCGCTAAG
411541	G/A	AGTCGTTGTGGGCGCGCATGGG	AGTCGTTGTGGGCGCGCATGGA	GTCAGGCTGTTCGGCTTGACGTATG

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
419658	T/G	TGTCCTCGTACGTGCTCGTTGTGACT	TGTCCTCGTACGTGCTCGTTGTGACG	AGCAGATGGCCTGGTAGCGGTCC
428503	G/A	CATGCAGGATACCGTGTGAGTTCAG	CATGCAGGATACCGTGTGAGTTCAA	GATGCTGTGCGCGTTGGACTG
		GCACTGCAAACACCTCTGCTCAAGTA	GCACTGCAAACACCTCTGCTCAAGTA	CTATGAATGCTCTTGCTAGCAGGCT
438644	A/G	TG	ТА	ТТА
		GAATTCCAAACGCGGTTCATAAACCA	GAATTCCAAACGCGGTTCATAAACCA	TCGAAGATAGTGTGCTCAATGGCGG
441042	A/G	CG	CA	ТТА
		TTGTTGCGAACATAGAGTACAGAGGA	TTGTTGCGAACATAGAGTACAGAGGA	
446758	T/A	GCA	GCT	GCTACAACGTGGGAATTGCCGAGGA
450975	T/G	TGCGGTTACGCAGTCGAAGCTATT	TGCGGTTACGCAGTCGAAGCTATG	ATGGGCACTCAAGGTGCGCACG
465604	A/T	CCTAAACGTCTCGGCGCTAATA	CCTAAACGTCTCGGCGCTAATT	AACTAAGACCACATTCCCGACATTG
				CATGCTCTTTCCTGTTGTCCGGTTC
465892	G/A	CCCACTGACGAGCGTGCTGAAGA	CCCACTGACGAGCGTGCTGAAGG	A
		CATAACGCTGAATTATCTTCGCCGAC	CATAACGCTGAATTATCTTCGCCGAC	
468480	A/G	ТА	TG	GTAAGGGGCCCACAAGCCTGG
480915	A/G	CTAATTCTCGTTCTACTGCCGCATG	CTAATTCTCGTTCTACTGCCGCATA	GGACACATCTCAGAACCAGATTG
487540	T/C	CACGGGAACGACGGGCACT	CACGGGAACGACGGGCACC	GGCACGTGAAGCTCCGAGATTTCAT
		TAGTGGGTTCGCTGAAGAACTACAAG	TAGTGGGTTCGCTGAAGAACTACAAG	
493429	A/G	AA	AG	CGCGCAGCTTTCTGAAGTAGTTGT
				GTTCTGGACTAAGTATGATTCGCTC
552113	T/A	TCATAGTTGGTTCACAGGCGACCT	TCATAGTTGGTTCACAGGCGACCA	CA
558063	A/G	CAGCTCCTGGGAGTCCTTGAGA	CAGCTCCTGGGAGTCCTTGAGG	AGTGGCTGCTGTCGCTACGCT
561492	T/C	ATCTTGCGACTGCTCGAT	ATCTTGCGACTGCTCGAC	TTCTCGCCCAGGAATGCCAT
580716	T/C	TCGGCGTTCAGCAGGCTTGAC	TCGGCGTTCAGCAGGCTTGAT	GCACCAGACCGCCGGCGA
583125	T/G	TGTTCTGAGGAAATGAGATGACTGTT	TGTTCTGAGGAAATGAGATGACTGTG	CAACACGTCAACAGCAACAT
585284	T/A	GCTTCAGTTATCAGCTGTAAACCTA	GCTTCAGTTATCAGCTGTAAACCTT	TTCGGTAATGCGTGTATTACTCA
585318	A/G	GTACATCACCGAAGCCGAACAG	GTACATCACCGAAGCCGAACAA	TTAGCCGCAACGCCGTGAAA
				CAAGAAACGGCAACAGCGGACAATG
589219	C/T	ATGCCGCACGTGCTTGAGGTC	ATGCCGCACGTGCTTGAGGTT	AAC
627150	A/C	CAATACAGCGGTATTTGCACTA	CAATACAGCGGTATTTGCACTC	CAATGGAGCAGACGCATCT
751708	G/A	TTGAAGCACAGCTCTTAGAGAAGG	TTGAAGCACAGCTCTTAGAGAAGA	GACTCCGTCAGCTGGTTTATG

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
754496	C/G	GCCTCGGCGTCGGAACTCG	GCCTCGGCGTCGGAACTCC	TGGCTGAAACCAGGGACCTCAA
761047	G/A	CAACATGGACGTTTTCAAGATTGCCA	CAACATGGACGTTTTCAAGATTGCCG	GAGCCTCGCTCAGCACGGAA
763022	T/C	CACAAAGGGCACGATTTCCTCT	CACAAAGGGCACGATTTCCTCC	AGATGAGTCTGCCATCGTGTCT
764527	T/A	GGGCGTTGCAGTAATGCAACAGTA	GGGCGTTGCAGTAATGCAACAGTT	AAGGCTCCTGGTGTAAGCACACG
767569	A/G	AAACACACCTTGAACTCAGCCTCA	AAACACACCTTGAACTCAGCCTCG	GGACGACAGCTATCAACATTAGCC
768618	C/G	AC.	AG	TACACTCCCAAGTGAGTTGATGC
771828	т/С	GATCCAAAGTGATCATGCCGATAGT	GATCCAAAGTGATCATGCCGATAGC	ATATCACAGTATCACGTCACGG
775381	A/G	TGTGCAGCTATCCAAAGACTCGG	TGTGCAGCTATCCAAAGACTCGA	ATGGTATCCGCTCGCTCGATATGT
777961	C/G	CTCAGCACAAGTGAATGTCAAG	CTCAGCACAAGTGAATGTCAAC	GGGCATTTGTAAGCATCTTATCGC
	-, -	GGCTCTATGTAGAACCAAAGATAAGT	GGCTCTATGTAGAACCAAAGATAAGT	
781023	G/T	GAG	GAT	ATTCTGCGGCTTCAACGAATCA
783090	G/A	ACCCGTACAGCAAACCACTACG	ACCCGTACAGCAAACCACTACA	CGACTGATTTCTCGCAACCCA
792422	T/C	TGCCACGGTAGTTTTGCTTAGT	TGCCACGGTAGTTTTGCTTAGC	ATGTTCCACGAGGCCCGTTG
43247	C/T	AGTAGACTTAAAGGCCACGCTCGAC	AGTAGACTTAAAGGCCACGCTCGAT	CCTTATATTCTCTGTCAGCGTAAG
		CAATCGAAATCGTGACCAATGGGATT	CAATCGAAATCGTGACCAATGGGATT	
84140	T/C	С	Т	ACCAAGTGCCGCGCAAAGCAT
117944	C/T	CGAATTCGAAGGCGGAGATCCTC	CGAATTCGAAGGCGGAGATCCTT	CGGCTTGGCGAAGCGACG
316915	T/G	CGCTTCGCCGAGCACTCG	CGCTTCGCCGAGCACTCT	ACCGGTTGTGCTACGCGTAGGT
197588	T/G	CAAGCGCATCCCCATTCTGATCTT	CAAGCGCATCCCCATTCTGATCTG	CTTAGAAAGGCAAGACCTCCTTCA
2932	C/T	CTCCTACGAGGGGGGGCCTGT	CTCCTACGAGGGGTGCCTGC	TTGTGACGTTCCTCGTGCTCCCT
112567	C/T	GCTCATGCGCATTGGAAGC	GCTCATGCGCATTGGAAGT	TTGCACGTACTACGTGCCTCTG
207179	T/C	CGCACGGAGATGGCATTCCTC	CGCACGGAGATGGCATTCCTT	ACACGATCTTCGGCGAGAACGTCA
165428	G/C	GTCCGCCACGTCGGTTCCAGAG	GTCCGCCACGTCGGTTCCAGAC	AAGCGGGGCTCTGCTTCCGCCT
109194	C/T	AGGCCCACAACTCCACTCTTC	AGGCCCACAACTCCACTCTTT	TACGGTAGCTATGTAACAGACACTA
139650	C/T	TACGACGGCACCGAGATC	TACGACGGCACCGAGATT	ATCTCCGGCGAGGCGTACA
56083	T/C	CCAGGCGCTCCTCCTCGGTC	CCAGGCGCTCCTCCTCGGTT	CGCCGGAGTTGGCCCAGGA
143860	A/G	ACAGGTACACGAACGATCGCAGAA	ACAGGTACACGAACGATCGCAGAG	TGCGTTCGTGCTTGTGTCATGT
152555	A/G	GCTCCAGGACAACCGTTTACCTCA	GCTCCAGGACAACCGTTTACCTCG	ATGGAAACATCGCTACACATGG

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
51899	A/G	GAGGTGTACGAGTGTCACTCGAAG	GAGGTGTACGAGTGTCACTCGAAA	GTATCTAGGAGGCTCGGGCGAAA
225801	C/T	GACTTCTGACATTTGATAGAATGCTC	GACTTCTGACATTTGATAGAATGCTT	TGCGGGTCAGCCATCTTACAAGTA
190468	G/A	TGAACGAAGCTGAGAGGCGCTATGA	TGAACGAAGCTGAGAGGCGCTATGG	TACGCCCAGACACTCTTGTTCAGT
31277	C/G	ATCATAGACCAACTCGCCTGCATC	ATCATAGACCAACTCGCCTGCATG	GATTCTGGAAGACAGCTTTTTCGC
				GGATTTCCGAGAGAAGCCATTTTCA
455987	G/C	AATGTACGCGACGTACGCACAAG	AATGTACGCGACGTACGCACAAC	G
27147	T/G	CGCAATTGTGACACCACTAG	GCGCAATTGTGACACCACTAT	CGGCTTTTGATACTCCCATCA
		CCGCATTTCTTCACTGCTGTTTGAAA	CCGCATTTCTTCACTGCTGTTTGAAA	
751588	G/T	G	Т	TCGCAAATCCTGGCGCGGTAA
313642	T/A	GTGCAGTTGGCAATGGAGGTGA	GTGCAGTTGGCAATGGAGGTGT	CCGGACAACTGAAGGTGGTGC
182969	G/A	AAGACGCACTTGCCCTGGAAACATG	AAGACGCACTTGCCCTGGAAACATA	GGTCTGAGTCTTGGTTGTGTCGCAT
186625	A/G	GAGGAGCTGCGATGCAGAAGTGGTA	GAGGAGCTGCGATGCAGAAGTGGTG	ATGCTGATGACGCAACGCTGACTTC
191703	A/G	CCGCCGTCTTTGCAGCCTCA	CCGCCGTCTTTGCAGCCTCG	GGGGCCCCGATTTCTAGAAC
438440	A/G	GTTGAGCGCATGCGCAGGGAA	GTTGAGCGCATGCGCAGGGAG	ACTCCCTGACGTAGCCTTCGTAGGA
82163	T/C	TAAGGCTTCCAGGTGACTTC	CTAAGGCTTCCAGGTGACTTT	GGTGTGTTGCTTCTATATTG
788521	C/T	ACCCGAACTTTGCAGGCCAT	ACCCGAACTTTGCAGGCCAC	AATGAACGACCGAGCGAATCCAGA
233756	C/G	TCTACAAACCAGGCGGTTGTAAGC	TCTACAAACCAGGCGGTTGTAAGG	TCTGTTTGGGACTCCTTCCACCG
201653	G/A	GCAGTCATCAAACGTGATTTCGTCCG	GCAGTCATCAAACGTGATTTCGTCCA	AAATTGGAGAGATCACTTGACCCGC
259800	C/G	CGTGTGCCTCGCTGGCATC	CGTGTGCCTCGCTGGCATG	GCGCATTCCAGAGGCTTCC
		GACACCCTAGCAAAGCAAAGCGTTCT	GACACCCTAGCAAAGCAAAGCGTTCT	
370147	C/T	С	Т	TTTCGTTCACGGCTCCCGCAA
153000	G/A	CCTACCTGCTTCCAACATTCTTTAGG	CCTACCTGCTTCCAACATTCTTTAGA	TGCACATTAGGTCAGAGATGCGGA
				AGACGATTATTCGGCTGTGACACAT
500950	A/T	CCACAACTCATCGCACCGAAGACT	CCACAACTCATCGCACCGAAGACA	Т
170547	C/G	GGTGAATACGCGTCGCGTGAGTC	GGTGAATACGCGTCGCGTGAGTG	GTGACCTTTGGTAGGACGGCAGC
		GAATATTTATGATGTGACCACGGCAA	GAATATTTATGATGTGACCACGGCAA	
466967	C/G	AC	AG	AACGCCCTGCCGCATAGTCC
246408	T/C	GGAAACAGTTATAACTATCTAGAACT	GGAAACAGTTATAACTATCTAGAACC	CACACCGAGAAATCAGACGTACC
5630	G/A	CAGCAAGCAGAGAACGTCGTCGATG	CAGCAAGCAGAGAACGTCGTCGATA	TTCAGGGTGAGACCGTCGGC
561563	A/G	TGAAGGATCTCGTACACAATACACAG	TGAAGGATCTCGTACACAATACACAA	CGAGTACTTCACGACCACGCA

locus	variant	Primer1-specific to allele1	Primer2-specific to allele2	Primer3-common to both alleles
338495	T/C	GGTTCTCGAAGCCGCGTTTC	GGTTCTCGAAGCCGCGTTTT	TCTGCAGCTGCTGTAGAGTCCTG
166887	A/G	TCGGCCGCCAGCAGCGTCA	CGGCCGCCAGCAGCGTCG	CCCGTCGGGAGCAATGCAG
766292	T/C	TGCCGAAGCTGGGTTTCGT	TGCCGAAGCTGGGTTTCGC	CTGGGCTGCTCCGAGGACTA
176206	G/T	ACTGCGATTGAAGTGCGTCCCG	ACTGCGATTGAAGTGCGTCCCT	ATCCTCTTGAAATTTGCTGCGGGTG
245496	T/C	TTCCAGCGTGCACCGTACC	TTCCAGCGTGCACCGTACT	GAAAATGCAATTTTTGTGAGCCT
199727	T/A	GGCTTCTTGTCTCGTTATTATCGT	GGCTTCTTGTCTCGTTATTATCGA	CAGTGCCACTTTATGTGAGTTG
				ACTAATTCATTGTAACCCATTTCAC
524153	G/A	CTCTATCAAACGATGTGCTACTGTGA	CTCTATCAAACGATGTGCTACTGTGG	GAT
54140	A/G	GGTAGACACAATCTGCTCATAATGG	GGTAGACACAATCTGCTCATAATGA	ATGACTGTTACAATCTTTTGAATGC
18708	A/G	CTCCGCGTGTATGCGAGTGAA	CTCCGCGTGTATGCGAGTGAG	GGCGCGTATCATCCCAGAGC
546612	G/C	TTTCCCGCGCAGGCCGCTAG	TTTCCCGCGCAGGCCGCTAC	TCAAGGCCAACGGCGCGCA
523859	C/A	CTGGACCTGTGCTACCGTGAGTCC	CTGGACCTGTGCTACCGTGAGTCA	GCTCAGGATGTCGTACGCGCGG
160279	A/T	ATCAGCAGCGCACACGCTCA	ATCAGCAGCGCACACGCTCT	CGTCGACGGGCGATCGTGA
		TATCAGCTAAAGCCTCCTTCTCAGTC	TATCAGCTAAAGCCTCCTTCTCAGTC	
624322	A/G	A	G	GAACTGAAGCACCAGCGCCT
	,	GTCAGAGTAAGGATCTGCTAGATACC	GTCAGAGTAAGGATCTGCTAGATACC	TAAGAAGGTTGGCCCGAATTTGTGA
410904	C/G	G	С	A
71660	C/A	GAAATTAGAATGGTACCTGGATTACC	GAAATTAGAATGGTACCTGGATTACA	CCTTTGGGGTGCGCTTATGTAAT GGTTGTATTTACAACTGACTCCTCG
87165	G/A	GAATCCACGTGTCAGAGCCCTGG	GAATCCACGTGTCAGAGCCCTGA	G
61479	A/G	GGCTAATCCTGCTTCTTGGCCTT	GGCTAATCCTGCTTCTTGGCCTC	CGATCCTGAAATCGAGCAAAGCC
571455	T/A	GTTCTGCCAGCAATTCTATCACT	GTTCTGCCAGCAATTCTATCACA	GGATGGATGCAAAGTGATATTTTAG
				CCTTTTTACGGACACTCACTTTCCT
270863	C/T	GCAATTATAGGATCTCCGTAAACTCT	GCAATTATAGGATCTCCGTAAACTCC	G
185472	C/G	ATTCGCCAGACCACTTGGATTCTC	ATTCGCCAGACCACTTGGATTCTG	CGTTTTCAATGAGTCTTGATTCTCG
200386	T/C	GATGGAATTAGGTACGGTCATTTCAT	GATGGAATTAGGTACGGTCATTTCAC	GTTCAGCGCATACTATGACTGACAA
40367	A/G	CACATGTGGCAAGCATTCAA	CACATGTGGCAAGCATTCAG	GCAGCAACGTTTGCTTCAGA
494898	T/C	AGCGTTGCACGCCATACATTCTCT	AGCGTTGCACGCCATACATTCTCC	TCCACAGGGTCACGTGACGCA
14134	C/T	CATACATTCCCTGAATACCTAGAGC	CATACATTCCCTGAATACCTAGAGT	ATTAGCCAAGCGCCCCG
361495	T/G	ATAACACAGGCAGACATTGGAGGCAG	ATAACACAGGCAGACATTGGAGGCAT	GCTCACATGCATTGAAACTGATGTC

Results

Table S3. Basic statistics per locus.

	Observed	Cono divorcity	Ect	Fic
LOCUS	heterozygosity	Gene diversity	FSL	FIS
1133	0.1101	0.2324	-0.0040	0.5262
31200	0.4319	0.467	0.0629	0.075
66390	0.1269	0.291	0.0539	0.5639
129322	0.0700	0.0873	0.0061	0.1974
159151	0.2649	0.3446	0.1504	0.2312
198227	0.5741	0.4987	-0.0093	-0.1512
221603	0.2760	0.3557	0.0932	0.2243
251320	0.1778	0.3271	0.1478	0.4564
298125	0.5543	0.4636	0.0127	-0.1956
329834	0.0918	0.2123	-0.0187	0.5675
374382	0.176	0.4705	0.0317	0.626
3705	0.1782	0.4146	0.0673	0.5704
32114	0.0759	0.1787	-0.0022	0.5753
68328	0.2477	0.3736	0.0396	0.3369
93695	0.2975	0.4519	0.0335	0.3418
133049	0.1991	0.3770	0.1375	0.4718
255757	0.2072	0.2468	0.2394	0.1604
299627	0.1109	0.2400	-0.0107	0.5381
376474	0.2141	0.3006	0.0491	0.288
6283	0.3140	0.4164	0.0634	0.2459
32551	0.262	0.343	0.0092	0.2361
96296	0.2678	0.3734	0.2144	0.2829
137096	0.5322	0.4652	0.0382	-0.1438
207995	0.8246	0.45	0.0726	-0.8323
225377	0.2191	0.3855	0.097	0.4316
259770	0.1676	0.2408	0.0235	0.3041
300752	0.4814	0.3724	0.1811	-0.2928
336267	0.2115	0.2254	0.0116	0.0618
380487	0.3168	0.4498	0.0807	0.2957
6363	0.3144	0.4011	0.1326	0.2162
34502	0.1219	0.3181	0.0031	0.6166
105385	0.3326	0.3783	0.0686	0.1208
143089	0.7837	0.4849	0.0132	-0.6163
208593	0.1799	0.2333	0.0843	0.2288
230247	0.2052	0.3581	0.223	0.427
281206	0.2277	0.3757	0.2359	0.394
303781	0.1121	0.2912	-0.027	0.6152

Locus	Observed	Gono divorsity	Ect	Fic
Locus	heterozygosity	Gene diversity	F31	FIS
339272	0.1745	0.1785	0.0237	0.0222
393248	0.0722	0.1882	-0.0019	0.6163
176991	0.1966	0.3406	0.029	0.4228
144259	0.1649	0.2687	0.0127	0.3863
113142	0.2391	0.3714	0.2421	0.3563
77668	0.4137	0.4424	0.0191	0.0648
42351	0.128	0.1783	0.439	0.2821
10041	0.1733	0.4534	0.0484	0.6178
399212	0.0839	0.3201	0.1127	0.7377
340581	0.0682	0.0911	0.0379	0.2513
305888	0.1105	0.1722	0.0151	0.3583
283680	0.6714	0.4485	0.0759	-0.497
233961	0.2212	0.3625	0.2481	0.3898
209761	0.0828	0.3481	0.0662	0.7621
180239	0.094	0.151	0.0109	0.3773
145634	0.3787	0.4585	0.018	0.1741
114791	0.2429	0.3708	0.0262	0.3449
57206	0.1859	0.3233	0.0253	0.425
19998	0.0914	0.1196	0.0199	0.2359
411541	0.227	0.3092	0.0348	0.2658
356074	0.2633	0.3343	0.0754	0.2123
307361	0.0887	0.1801	0.0333	0.5074
287805	0.0435	0.0744	0.0317	0.4154
234508	0.2516	0.3103	0.059	0.1891
210654	0.228	0.2674	0.4378	0.1475
189207	0.1252	0.3037	0.0042	0.5879
150669	0.2338	0.4764	0.0025	0.5092
116335	0.1823	0.3998	0.0415	0.544
81501	0.2806	0.4505	0.101	0.377
60684	0.2221	0.442	0.0547	0.4974
21130	0.1723	0.412	0.0976	0.5818
356395	0.3341	0.4604	0.0791	0.2745
313057	0.1085	0.388	0.0497	0.7203
292025	0.1001	0.1208	-0.0134	0.1714
236290	0.0798	0.1775	-0.0126	0.5505
212829	0.2648	0.4684	0.0637	0.4347
197784	0.3144	0.2259	0.1932	-0.392
155043	0.1033	0.1762	0.5208	0.4136
125671	0.3398	0.4652	0.0539	0.2695

Locus	Observed	Cono divorcity	Ect	Fie
LOCUS	heterozygosity	Gene diversity	FSL	FIS
81758	0.1436	0.2886	0.1153	0.5023
61606	0.2479	0.3716	0.2166	0.3331
428503	0.3744	0.395	0.0206	0.0521
320000	0.2643	0.2579	0.1614	-0.0248
296275	0.1384	0.3744	0.2015	0.6303
243436	0.1853	0.2832	0.3311	0.3455
214684	0.2222	0.4963	-0.0018	0.5522
438644	0.1799	0.4129	0.149	0.5644
487540	0.0981	0.1404	0.288	0.3014
767569	0.1635	0.2148	0.0246	0.2389
165428	0.211	0.2398	0.0199	0.1199
191703	0.1966	0.3729	0.2489	0.4729
153000	0.0919	0.4793	0.0403	0.8082
166887	0.1921	0.3481	0.0653	0.4483
441042	0.3243	0.4491	0.0996	0.2777
84140	0.1825	0.3554	0.0587	0.4865
438440	0.0944	0.1089	0.1972	0.1325
766292	0.738	0.4745	0.0398	-0.5553
523859	0.3518	0.4397	0.1064	0.1999
270863	0.0571	0.23	0.0146	0.7516
446758	0.0327	0.1022	0.0336	0.6802
552113	0.144	0.219	0.0062	0.3423
627150	0.1918	0.4255	0.074	0.5491
117944	0.8044	0.4649	0.0642	-0.7304
139650	0.2503	0.2708	0.1388	0.0757
176206	0.3211	0.5036	-0.0175	0.3624
185472	0.1899	0.3965	0.0348	0.5211
450975	0.2154	0.3343	0.1091	0.3557
558063	0.2358	0.2929	0.0731	0.195
751708	0.0754	0.1006	-0.026	0.2506
775381	0.0911	0.2703	-0.0258	0.6629
27147	0.1022	0.3726	0.0636	0.7257
200386	0.1562	0.4361	0.1004	0.6418
777961	0.0977	0.3904	0.1195	0.7499
754496	0.3408	0.4748	0.051	0.2822
561492	0.0813	0.3268	0.0409	0.7512
465604	0.0536	0.2293	0.0085	0.7664
410904	0.1761	0.2763	-0.0034	0.3628
199727	0.1067	0.4443	-0.0179	0.7599

Locus	Observed heterozygosity	Gene diversity	Fst	Fis
751588	0.1149	0.3961	0.0719	0.7099
152555	0.1713	0.4355	0.058	0.6067
2932	0.0525	0.0604	0.0063	0.1315
781023	0.1682	0.4324	0.0954	0.6109
761047	0.1318	0.173	-0.0271	0.2378
580716	0.316	0.3822	0.0479	0.1731
465892	0.113	0.1621	-0.0162	0.3033
5630	0.7425	0.4783	0.0441	-0.5525
313642	0.1283	0.1442	0.0071	0.1102
783090	0.2976	0.3471	0.0268	0.1426
763022	0.1247	0.302	0.0606	0.5872
583125	0.534	0.4398	0.097	-0.2141
468480	0.3318	0.4527	0.0932	0.267
14134	0.1556	0.345	0.0144	0.549
259800	0.125	0.1671	0.0074	0.2517
182969	0.2349	0.2092	0.005	-0.1228
225801	0.1255	0.254	0.3865	0.5058
792422	0.402	0.4864	-0.0019	0.1736
764527	0.1107	0.2841	0.0287	0.6103
585284	0.1384	0.2335	-0.0063	0.4072
480915	0.0849	0.2483	0.032	0.6581
338495	0.2448	0.3455	0.0157	0.2915
186625	0.5284	0.4536	0.0852	-0.1649



Figure S1. Discriminant analysis of principal component (DAPC) of *Ixodes ricinus* based on 497 individuals using 125 SNPs. A. BIC values as a function of the number of clusters k. The difference in BIC values between k = 3 and k = 4 is 0.842. B. Scatterplot of individuals on the two principal components of DAPC. The graph represents the individuals as dots and the groups as inertia ellipses. Two of the clusters overlap, while when k = 3 we identify 3 well separated groups (figure 3). Red : North African cluster; yellow : only individuals from southern Eurasian cluster; green : only individuals from the Northern European cluster; blue: admixture cluster with mainly individuals from the northern European cluster in fugure 3.



Figure S2. Probabilities $\ln P(X|K)$ for each level of hierarchical analysis. First round of analysis: a); Second round: b) southern Eurasian cluster and c) northern European clusters; Third round: d) Southern European cluster without Iran, e) Central Sweden, Norwegian West and East and North Estonia, f) Moldavia, North France, West German, Belgium, South Estonia, East German, South German and South Sweden; Forth round: g) Atlantic samples (Spain, South and West France, Ireland and England, h) South-west samples (Italy, Romania, Slovakia and Hungary), i) and i): fourth round of analysis. Details of each level of Hierarchical analysis are present in the corresponding session.



Figure S3. STRUCTURE Individual probabilities for each value of *K* from 2 to 10.

Hierarchical analysis

Finer genetic structure was identified from hierarchical analysis (Figure S6 and S7 for STRUCTURE and DAPC analysis, respectively). The southern Eurasian cluster was further separated into two differentiated clusters, irrespective of the approach used (STRUCTURE or DAPC). The STRUCTURE approach separated Iran from the remaining samples, while the DAPC approach assigned most individuals from both Iran and Turkey samples to the same cluster (violet). The northern European cluster was further separated into two to three clusters depending on the methods, DAPC and STRUCTURE, respectively. Clusters identified by the DAPC approach were distributed almost equally among the different sampled locations. Of the three clusters identified by STRUCTURE, the orange and green ones showed a clear affinity to certain sample locations, while the grey cluster was represented in all sampled locations. No further structure was identified for the African cluster in both methods.

The DAPC's third round of analysis was unable to identify further genetic structure in the northern European cluster. It did however identify two groups inside the southern Eurasian cluster (without Turkey and Iran as a result of previous analyse). It appears that individuals from Spain, Western France and Ireland were mainly assigned to one (light blue) cluster. No other cluster was identified by the DAPC approach regarding refined hierarchical analysis. The STRUCTURE's third round of analysis was able to identify a K = 4 in the southern European cluster. Individuals from Turkey were assigned to an exclusive cluster (grey). Individuals from southwestern Europe and from Italy were mainly assigned to one cluster (orange), while those from Spain, West France and Ireland were grouped in a different cluster (blue). The fourth cluster (green) was distributed across all sampling locations with few individuals (11 out of 179) exhibiting more than 50% of assigning probability. In the northern European cluster, for this third

round of hierarchical analysis, individuals were regrouped according to population probabilities of the two almost exclusive clusters from last step, green and orange ones. From this third round until the last one, Evanno's method (Evanno et al. 2005) always identified two clusters, but the analyses of $\ln[Pr(X|K)]$ was not clear in identifying those clusters (Figure SX). Also, individual probabilities of inside those K = 2 clusters show very mixed populations. The results for those subsequent rounds with a K = 2 are presented in Supplementary Information (figure SXX). We did a fourth and last round of hierarchical analysis for the two main southern Eurasian clusters identified in the previous round: (i) one cluster composed of Spain, West and South France, Ireland and England samples and (ii) the other cluster composed of Italy, Romania, Hungary and Slovakia. For the first one, Evanno's method identified K = 6, but the analysis of ln[Pr(X|K)] does not indicate any structure. For the later, both methods clearly identified a K = 7 structuring. In both cases, clusters are mainly distributed in all sample sites and very rarely a single individual had ~100% probability of being assigned to a particular cluster. The exceptions were individuals from West France and Ireland for which probability values to be assigned to the same cluster reached one.



Figure S4. DAPC Hierarchical analysis. Each column corresponds to one level of analysis.







Figure S6. Mean gene diversity, observed heterozygosity, and Fis per population. Mean population gene diversity was always greater than the observed heterozygosity and Fis was always positive.