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## Further investigations on the relationship between the O-PLS preprocessing and the NAS.

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## Abstract

O-PLS is a preprocessing that was presented as an improvement of the PLS algorithm it was issued from. Nevertheless, the bibliography did not confirm, neither for prediction, nor for interpretation. To contribute to a better understanding, we investigated the relationship between O-PLS and the NAS net analyte signal. For four numerical applications, the matrix obtained after the O-PLS deflation tended towards a matrix of rank 1 when the number of removed dimensions increases. Therefore, the line-vectors of this matrix are colinear to the NAS, and so the usual one-latent-variable PLS1 regression after the O-PLS preprocessing can be replaced by almost any regression taking into account from one to all of the variables.

**Keywords:** PLS, O-PLS, OSC, NAS, prediction, interpretation

# 1 Introduction

2 Chemometrics is largely based on the Partial Least Squares or Projection  
3 to Latent Structures (PLS) method. Hereafter, all the discussion will ap-  
4 ply to the prediction of a single variable. Let  $\mathbf{X}$  and  $\mathbf{y}$  be a matrix and a  
5 column-vector containing the values of variables measured on the same ob-  
6 servations. PLS regression was developed to predict the values of  $\mathbf{y}$  using  $\mathbf{X}$ ,  
7 specially when the variables in  $\mathbf{X}$  are highly correlated. For example, PLS  
8 regression allowed the prediction of gluten in a wheat sample from a simple  
9 near infrared spectrum. Many PLS algorithms have been proposed. Nine  
10 of them were compared in Andersson [1]. We proposed a tenth algorithm  
11 more focused on the geometry of PLS [2]. Countless applications of PLS can  
12 be found in industrial and scientific applications. Nevertheless, in certain  
13 fields such as metabolomics, another method is being replacing PLS. This  
14 method is based on a preprocessing called Orthogonal Projection to Latent  
15 Structures (O-PLS). The O-PLS algorithm was obtained by a slightly mod-  
16 ification of the PLS algorithm. To understand it, let us come back to the  
17 origin.

18 The principle of PLS is to extract the subspace in  $\mathbf{X}$  that contains the  
19 information related to the explanation of  $\mathbf{y}$ . The idea of the Orthogonal  
20 Signal Correction (OSC) preprocessings is to help PLS by removing from  $\mathbf{X}$   
21 all information not related to  $\mathbf{y}$ . OSC are followed by a regression, usually  
22 a PLS. A first OSC algorithm, derived from the NIPALS-PLS algorithm,  
23 was proposed by Wold et al [3]. This algorithm was latter improved by  
24 Trygg and Wold, yielding to a new method called O-PLS [4, 5, 6]. But  
25 while other OSC algorithms were proposed [7, 8], the ability of OSC to  
26 improve the predictive ability of a regular PLS was being discussed [9, 10].  
27 More attention was paid to O-PLS. It was observed that one O-PLS factor

28 removed spared one PLS latent variable, without gain in prediction errors  
29 [7, 11, 9]. Then Verron et al [12] mathematically confirmed this observation,  
30 and Kemsley & Tapp [13] showed a direct relationship between O-PLS and  
31 PLS scores. As a conclusion, O-PLS followed by a PLS regression yields  
32 exactly the same prediction than a single PLS regression. Then O-PLS  
33 benefits were focused on interpretation rather than prediction, e.g. [14],  
34 but even the added value of interpretation was challenged [15], conclusions  
35 we agree with. Nevertheless, there is a gap between O-PLS users on the  
36 one hand, and mathematical considerations on O-PLS on the other hand.  
37 Maybe a geometrical approach could contribute to a better understanding  
38 of O-PLS? In particular, Goicoechea et al [10] and Ni et al [16] studied  
39 the relationship between OSC methods and the Net Analyte Signal (NAS)  
40 introduced by Lorber [17]. We propose to further investigate this property.  
41 Notations are presented in the figure 1.

## 42 **2 Theory.**

43 The name O-PLS let think to a PLS-like method, i.e. a regression method.  
44 Certain so-called "O-PLS" functions associate an O-PLS preprocessing fol-  
45 lowed by a PLS regression with one latent variable. These names are mis-  
46 leading, because O-PLS as introduced in the original papers of Trygg and  
47 Wold is clearly a preprocessing, not a calibration method. In order to obtain  
48 predictions, O-PLS needs to be followed by a calibration step, e.g. a classical  
49 least squares or a PLS regression [10] usually processed with a single latent  
50 variable.

51 O-PLS calculation is iterative. Let  $A_{max}$  be a predetermined large num-  
52 ber of dimensions to be removed. Scores  $\mathbf{t}_i$  and loadings  $\mathbf{p}_i$  are computed  
53 for each value of  $i$  between 1 and  $A_{max}$ , then the correction is performed

54 by deflation[9]:  $\mathbf{X}_{\text{opls}} = \mathbf{X} - \sum_{i=1}^{A_{\text{max}}} \mathbf{t}_i \mathbf{p}_i'$ . Each dimension  $i$  is supposed  
55 to contain information unrelated to  $\mathbf{y}$ . But while  $i$  is increasing, there will  
56 be a breakpoint for which no more information unrelated to  $\mathbf{y}$  will be left  
57 in  $\mathbf{X}_{\text{opls}}$ . Let  $A$  be this value of  $i$ ;  $A \leq A_{\text{max}}$ . A definition of the NAS  
58 is *the part of the signal orthogonal to the other constituents* [17]. This  
59 definition was made for spectra, but it can be extended to other signals.  
60 Let  $\mathbf{s}_{\text{nas}}$  be the NAS associated to the compound of interest whose val-  
61 ues form  $\mathbf{y}$ . An interesting property of the NAS is that the product of  
62  $\mathbf{X}_{\text{opls}}$  by  $\mathbf{s}_{\text{nas}}$  should be proportional to  $\mathbf{y}$ . It derives from the following  
63 equation [10] :  $\mathbf{X}_{\text{opls}} = k \mathbf{y} \mathbf{s}'_{\text{nas}}$  with  $k$  a constant. This situation should oc-  
64 cur when an appropriate number of dimensions  $A$  has been removed by the  
65 OSC. Therefore, the corresponding matrix  $\mathbf{X}_{\text{opls}}$  should be of rank 1, all the  
66 line-vectors of  $\mathbf{X}_{\text{opls}}$  being collinear. Under these conditions, the classical  
67 one-latent-variable PLS1 regression which follows the O-PLS preprocessing  
68 could be replaced by *any* other regression without significant loss of pre-  
69 dictive ability, even with a simple linear regression. On the other hand, if  
70 the theory is not verified, if  $\mathbf{X}_{\text{opls}}$  is not close enough from a rank-1 matrix,  
71 then the predictions should strongly depend on the choice of the regression  
72 method. To assess this property, three methods based on O-PLS were com-  
73 pared. The first method noted *OPLS-classic* is the classical combination of  
74 an O-PLS preprocessing followed by a PLS regression with 1 latent variable.  
75 The second method noted *OPLS-ones* is a modified version of the previous  
76 *OPLS-classic*. The modification consisted in replacing the weight vector  $\mathbf{w}$   
77 which begins the PLSR algorithm by an arbitrary vector of ones of same  
78 dimension. *OPLS-classic* and *OPLS-ones* algorithms are summarized in Ta-  
79 ble 2. The third method noted *OPLS-univ* is a classical OPLS followed by  
80 an univariate regression. The selected variable was chosen as the one with

81 the largest variability in  $\mathbf{X}_{\text{opls}}$  after the O-PLS preprocessing removed  $A_{\text{max}}$   
82 components.

### 83 **3 Material and methods**

84 Four datasets were selected. Two datasets were challenges proposed by  
85 Pierre Dardenne and Vincent Baeten, CRA-Wallonie, at the French annual  
86 chemometrics conference in 2007 and 2018. The goal of the 2007 challenge  
87 [18] was to predict gluten concentrations. The test dataset was picked up,  
88 since it contained a large number of observations (2000). The 2018 chal-  
89 lenge consisted in 3908 calibration plus 429 test samples of near infrared  
90 spectra. The compound to quantify was not identified. A third dataset  
91 was provided by Sylvie Bureau, INRAE, UMR408. Mid infrared spectra  
92 ( $4000 - 650\text{cm}^{-1}$ ) were acquired on 750 apricots, for which several reference  
93 analysis were performed including refractive index. The last dataset was  
94 provided by INRAE, UMR ITAP and SPO. Visible-near infrared spectra  
95 were acquired in transmittance and Brix degrees were also measured on 250  
96 grape berries. Processing needed a calibration and a test dataset. Thus,  
97 the 2007 challenge, the apricots and grape berries datasets were splitted,  
98 the 1500, 600 and 200 observations were assigned to the calibration dataset,  
99 the last 500, 150 and 50 observations were assigned to the test dataset, re-  
100 spectively. After centering, the three regression methods described above,  
101 *OPLS-classic*, *OPLS-ones* and *OPLS-univ* were processed on the calibra-  
102 tion datasets with 1 to  $A_{\text{max}}$  dimensions removed by OPLS,  $A_{\text{max}}$  being set  
103 to 50 for the apricot datasets and to 40 for the three other datasets. The  
104 variable selected for the *OPLS-univ* model was the one with the largest vari-  
105 ability after removal of  $A_{\text{max}}$  dimensions by OPLS. The predictive abilities  
106 of the three models were assessed by the calibration errors (RMSEC), the

107 cross-validation errors (RMSECV) and the prediction errors (RMSEP). The  
108 RMSECVs were the average of a random 2-blocks cross-validation repeated  
109 50 times.

## 110 4 Results

111 Results are presented as figures 3, 4, 5 and 6, one for each of the four  
112 datasets. After removal of the  $A_{max}$  OPLS components, the residual spec-  
113 tra in subfigures (a) present a similar shape, sometimes symmetrical along  
114 the axis  $y = 0$  which is consistent with a low rank of the corresponding  
115 matrix. These subfigures were used to choose the univariate variable of the  
116 *OPLS-univ* regression. RMSECs in subfigures (b) are similar for the four  
117 datasets. When  $i$  is low, a few differences between the three models can be  
118 observed. But when  $i$  increases, the curves overlay. RMSECV in subfigures  
119 (c) also present the same trend for all datasets, but with a different behav-  
120 ior according to the models. The RMSECV of *OPLS-univ* decreases steadily  
121 when  $i$  increases. On the other hand, *OPLS-classic* and *OPLS-ones* present  
122 the classical shape of a decrease, then an increase. When  $i$  is high, those  
123 two curves overlay. Then, RMSEP in subfigures (d) can present differences  
124 between models. But when  $i$  is high, the curves of the three models overlay.

## 125 5 Discussion and conclusion

126 The hypothesis that, by increasing the number  $i$  of removed dimensions, the  
127 rank of the resulting matrix  $\mathbf{X}_{\text{opls}}$  gets close to 1, still stands up after the  
128 benchmark on 4 datasets. Similar results, not presented, confirmed this ob-  
129 servation. When  $i$  increases, all the line vectors of  $\mathbf{X}_{\text{opls}}$  tend to be colinear.  
130 This situation occurred when  $i$  was higher than  $A = 7, 15, 40$  and 5 for the

131 2007 and 2018 challenges, the apricots and grape berries datasets respec-  
 132 tively. These values of  $A$  need to be compared to  $A_{opt}$  the optimal number  
 133 of dimensions to be removed, determined by the RMSECV curves, around  
 134 13–20, 15–20, 7–12 and 15–22 respectively. For the apricot dataset,  $A_{opt}$   
 135 corresponds to a situation where the three models *OPLS-classic*, *OPLS-ones*  
 136 and *OPLS-univ* do not yield the same prediction, and where prediction er-  
 137 rors remain high. Figure 5(d) suggests that a higher number of dimensions,  
 138  $A_{opt}$  around 20–22, would have yielded better predictions, but such values  
 139 remain before the convergence region obtained for  $A = 40$  ( $A > A_{opt}$ ). For  
 140 the three other datasets, convergence of the three models had already been  
 141 achieved:  $A < A_{opt}$ , and the models were found more robust when applied  
 142 to the test datasets. Maybe a clue to identify robust ( $A < A_{opt}$ ) from not  
 143 robust ( $A > A_{opt}$ ) models?

144 Finally, all the information about the NAS is supported by the  $\mathbf{X}_{opls}$  ma-  
 145 trix, issued from the O-PLS preprocessing. The weights  $\mathbf{w}$  and the regression  
 146 vector  $\mathbf{b}$  of the PLSR in the *OPLS-classic* method are linear combination  
 147 of the lines of  $\mathbf{X}_{opls}$ , weighed by the  $\mathbf{y}$  values. Setting  $\mathbf{w}$  to a vector of  
 148 ones in *OPLS-ones* is another way to weight the lines of  $\mathbf{X}_{opls}$ , each line  
 149 has the same weight. When  $\mathbf{X}_{opls}$  is close to a rank-1 matrix, these calcula-  
 150 tions yield the same result. The  $\mathbf{X}_{osc}$  matrix or the regression coefficients  $\mathbf{b}$   
 151 can all be interpreted as an estimation of the NAS. According to its defini-  
 152 tion, the NAS is obtained after an orthogonal (or oblique) projection, and  
 153 therefore inherits of the properties of such calculations. In particular, the  
 154 NAS is expected to present features from the compound of interest (whose  
 155 concentrations are  $\mathbf{y}$ ), but more annoying it can present features from the  
 156 other coexisting constituents [19], and eventually it can have dropped fea-  
 157 tures that were present in the compound of interest. So the interpretation

158 of the NAS remains tricky and should be confirmed by other informations.  
159 These considerations also apply to the residual matrix,  $\mathbf{X} - \mathbf{X}_{opls}$ , which is  
160 the result of an orthogonal projection, too. Note that the same interpre-  
161 tation of the NAS calculated from O-PLS can also be performed with the  
162 regression coefficients of a PLS regression, already reported as an estimation  
163 of the NAS [20].

## 164 **6 Acknowledgements**

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167 manuscript.

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## 229 7 Captions

230 Figure 1: Notations

231 Figure 2: *OPLS-classic* a regular O-PLS followed by a PLS1 with one latent  
232 variable (left); *OPLS-ones* a regular O-PLS followed by a w-modified PLS1,  
233  $\mathbf{w}$  being set to a vector of ones (right)

234 Figure 3: Chimie 2007 challenge, NIR spectra: (a) spectra after 40  
235 OPLS components removed, the vertical line representing the variable se-  
236 lected for *OPLS-univ*; (b-d) RMSEC, RMSECV and RMSEP for the 3 mod-  
237 els: *OPLS-classic* (red), *OPLS-ones* (green) and *OPLS-univ* (blue)

238 Figure 4: Chimie 2018 challenge, NIR spectra: (a) spectra after 40  
239 OPLS components removed, the vertical line representing the variable se-  
240 lected for *OPLS-univ*; (b-d) RMSEC, RMSECV and RMSEP for the 3 mod-  
241 els: *OPLS-classic* (red), *OPLS-ones* (green) and *OPLS-univ* (blue)

242 Figure 5: Apricots MIR spectra: (a) spectra after 50 OPLS components  
243 removed, the vertical line representing the variable selected for *OPLS-univ*;  
244 (b-d) RMSEC, RMSECV and RMSEP for the 3 models: *OPLS-classic* (red),  
245 *OPLS-ones* (green) and *OPLS-univ* (blue)

246 Figure 6: Grape fruit NIR reflectances: (a) spectra after 40 OPLS compo-  
247 nents removed, the vertical line representing the variable selected for *OPLS-*  
248 *univ*; (b-d) RMSEC, RMSECV and RMSEP for the 3 models: *OPLS-classic*

249 (red), *OPLS-ones* (green) and *OPLS-univ* (blue)

250

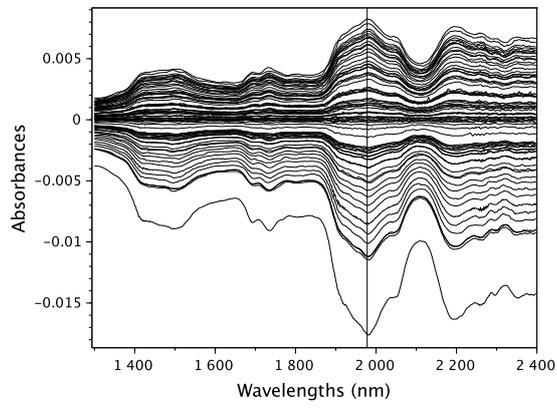
251 **8 Figures**

$N$	number of observations
$Q$	number of variables
$i$	number of O-PLS components removed
$A$	minimum value of $i$ for which $\mathbf{X}_{osc}$ can be approximated to a rank-1matrix
$A_{opt}$	value or range of values of $i$ for which the RMSECV is minimum
$A_{max}$	a value of $i$ larger than $A$ and $A_{opt}$
$\mathbf{X}$	matrix of $N$ lines and $Q$ columns
$\mathbf{y}$	column vector of $N$ elements
$\mathbf{X}_{osc}$	the matrix $\mathbf{X}$ after OSC correction
$\mathbf{X}_{opls}$	the matrix $\mathbf{X}$ after O-PLS correction
$\mathbf{t}_i$	$i^{th}$ score vector for OSC/OPLS
$\mathbf{p}_i$	$i^{th}$ loading vector for OSC/OPLS
$\mathbf{w}$	weight vector of PLS
$\mathbf{t}$	score vector of PLS
$\mathbf{b}$	regression vector of PLS
$\mathbf{1}_Q$	a vector of ones of length $Q$
$\mathbf{s}_{nas}$	the NAS net analyte signal

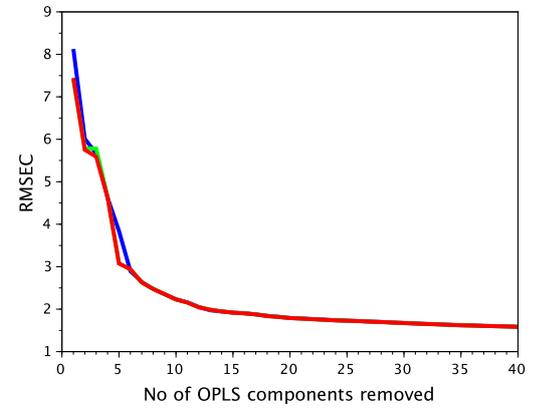
Figure 1:

$\mathbf{X}_{opls}$ calculated with O-PLS	
$\mathbf{w} = \mathbf{X}'_{opls}\mathbf{y}$	$\mathbf{w} = \mathbf{1}_Q$
$\mathbf{t} = \mathbf{X}_{opls}\mathbf{w}$	$\mathbf{t} = \mathbf{X}_{opls}\mathbf{1}_Q$
$c = \mathbf{y}'\mathbf{t}(\mathbf{t}'\mathbf{t})^{-1}$	$c = \mathbf{y}'\mathbf{t}(\mathbf{t}'\mathbf{t})^{-1}$
$\mathbf{b} = c\mathbf{w}$	$\mathbf{b} = c\mathbf{1}_Q$

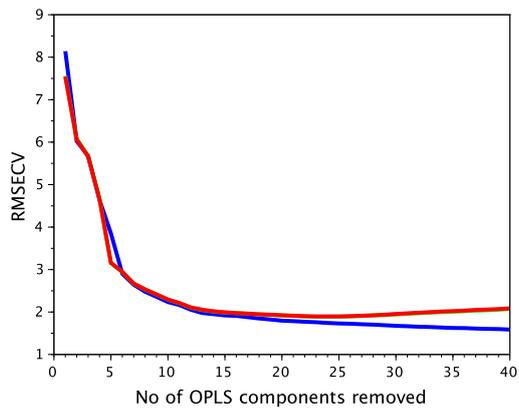
Figure 2:



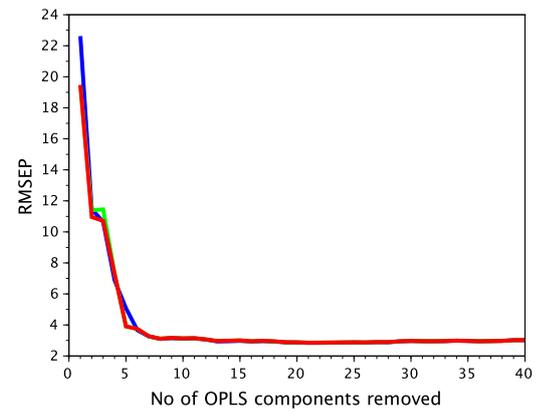
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(b)

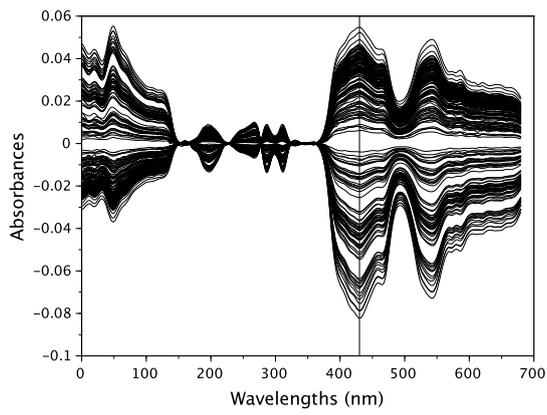


(c)

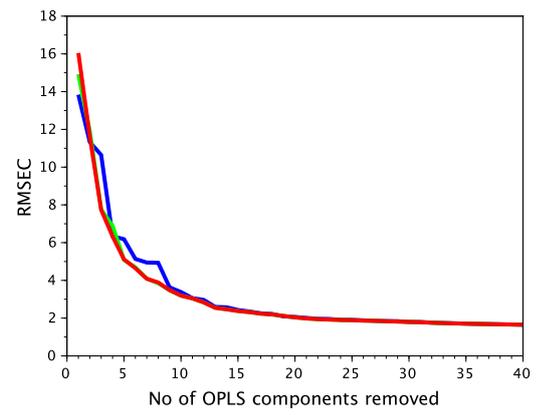


(d)

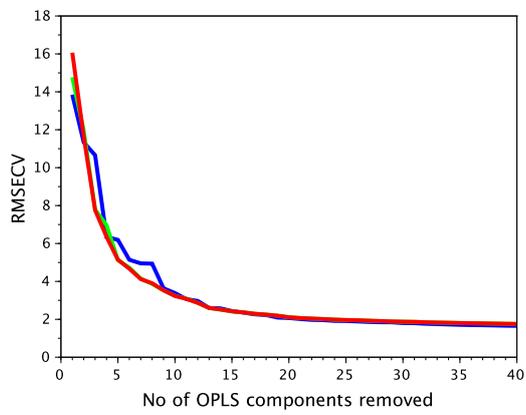
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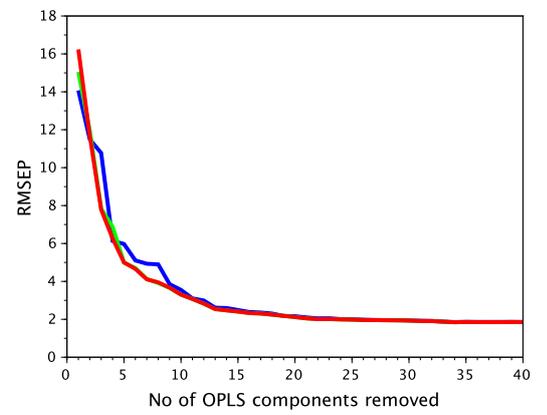
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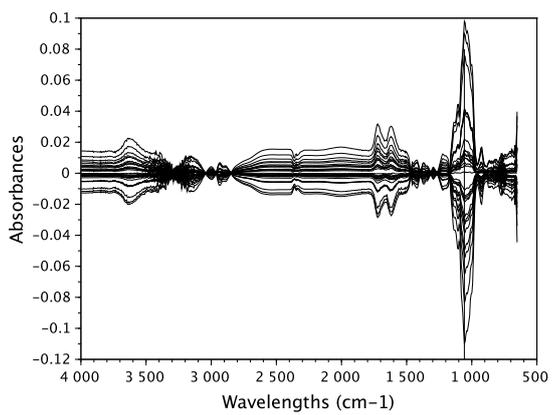


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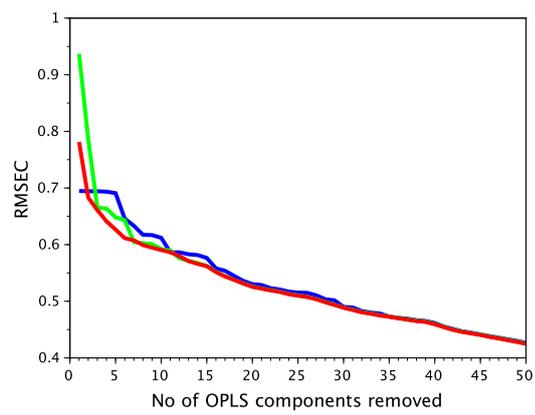


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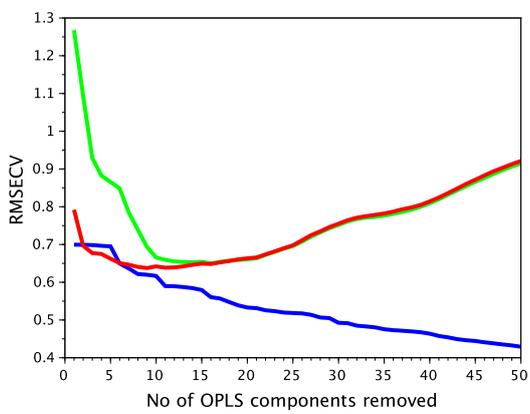
Figure 4:



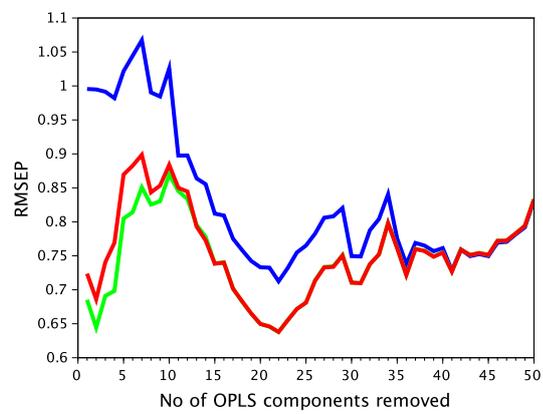
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(b)

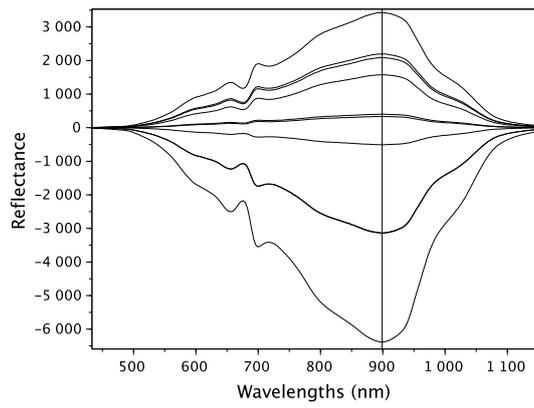


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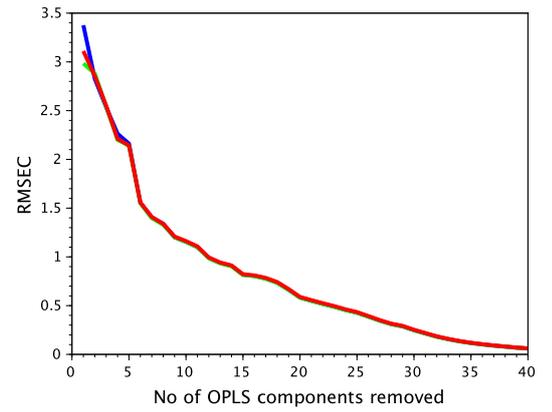


(d)

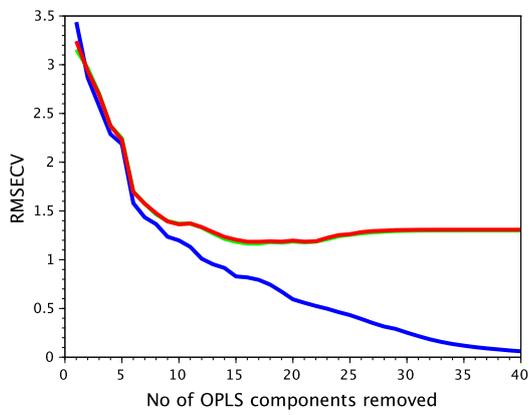
Figure 5:



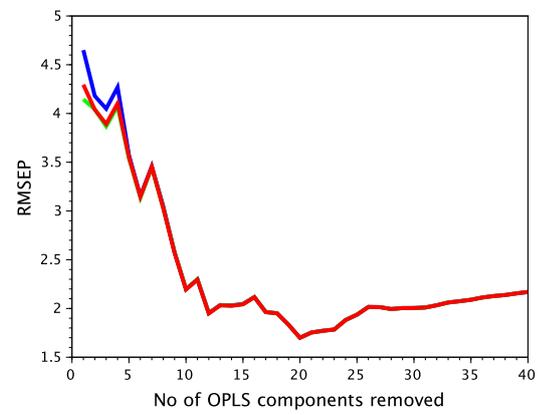
(a)



(b)



(c)



(d)

Figure 6: