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RESEARCH ARTICLE

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Genetic parameters for first lactation dairy traits in the Alpine and Saanen goat breeds using a random regression test-day model

Mathieu Arnal^{1,2*} , H el ene Larroque¹, H el ene Leclerc³, Vincent Ducrocq⁴ and Christ ele Robert-Grani e¹

Abstract

Background: Random regression models (RRM) are widely used to analyze longitudinal data in genetic evaluation systems because they can better account for time-course changes in environmental effects and additive genetic values of animals by fitting the test-day (TD) specific effects. Our objective was to implement a random regression model for the evaluation of dairy production traits in French goats.

Results: The data consisted of milk TD records from 30,186 and 32,256 first lactations of Saanen and Alpine goats. Milk yield, fat yield, protein yield, fat content and protein content were considered. Splines were used to model the environmental factors. The genetic and permanent environmental effects were modeled by the same Legendre polynomials. The goodness-of-fit and the genetic parameters derived from functions of the polynomials of orders 0 to 4 were tested. Results were also compared to those from a lactation model with total milk yield calculated over 250 days and to those of a multiple-trait model that considers performance in six periods throughout lactation as different traits. Genetic parameters were consistent between models. Models with fourth-order Legendre polynomials led to the best fit of the data. In order to reduce complexity, computing time, and interpretation, a rank reduction of the variance covariance matrix was performed using eigenvalue decomposition. With a reduction to rank 2, the first two principal components correctly summarized the genetic variability of milk yield level and persistency, with a correlation close to 0 between them.

Conclusions: A random regression model was implemented in France to evaluate and select goats for yield traits and persistency, which are independent i.e. no genetic correlation between them, in first lactation.

Background

Random regression test-day (TD) models (RRM) are widely used in genetic evaluations of TD milk production in dairy cows but also in other species such as dairy goats [1–6]. RRM increase the accuracy of breeding value predictions and fit the variability of environmental effects throughout lactation more accurately [7].

Besides describing the variability of genetic parameters along the lactation curve, RRM serve to predict estimated breeding values (EBV) for lactation persistency, based on the variation of EBV throughout lactation. A persistent

animal is defined as producing on average less milk at the beginning but more at the end of the lactation period than animals with a similar overall production [8]. Lactation persistency is of interest for dairy producers because the shape of the lactation curve can affect an animal's nutritional needs, and consequently its health, as well as the distribution of the farm's milk output during the year [9, 10]. In a previous study, Arnal et al. [11] demonstrated the phenotypic variability of the shape of the dairy goat lactation curves in France, and showed that the main environmental factors influencing curve shapes are breed, kidding month, age at kidding, gestation stage, and length of the dry period.

Various functions have been used to model fixed, genetic and permanent environment effects on TD records, including the Wilmink function [12], Legendre polynomials [13, 14], and splines [15, 16]. Given that

*Correspondence: Mathieu.arnal@idele.fr

¹ GenPhySE, Universit e de Toulouse, INRA, ENVT, Toulouse INP, 31326 Castanet-Tolosan, France

Full list of author information is available at the end of the article



RRM are computationally expensive, Legendre polynomial functions have the computational advantage of reducing the correlations between estimated regression coefficients, which impact the convergence [7] of the iterative algorithms used for variance component estimation and genetic evaluation. They are also sufficiently flexible to fit differently shaped curves.

With eigen decomposition, it becomes possible to reduce the rank of the resulting variance-covariance matrix by ignoring the contribution of the smallest eigenvalues and eigenvectors [14, 17]. This decreases computing time by reducing the number of genetic and permanent environment regression coefficients that are estimated.

To date, the implementation of RRM in dairy goats has not been studied in France. In this paper, we estimate genetic parameters for milk yield, fat yield, protein yield, fat content, and protein content using RRM with Legendre polynomial functions of different orders, with or without rank reductions, to obtain TD EBV for French dairy goats in first lactation.

Methods

Data

The data consisted of 193,226 milk TD records from 30,186 first-lactation Saanen goats (234 herds) and 205,841 milk TD records from 32,256 first-lactation Alpine goats (198 herds) from northwestern France, collected between 1995 and 2015. The pedigree consisted of 66,716 and 67,159 Saanen and Alpine goats, respectively. Each lactation included at least four TD between the 7th and 270th day in milk (DIM). Lactation had to last between 180 and 350 days. Goats were milked twice a day and their records were summed to obtain their daily production. More than four animals per herd × test date combination were required. The sires of these goats were artificial-insemination bucks with at least 20 progeny in the dataset, with 379 and 324 sires from the Alpine and Saanen goat breeds, respectively. The dams of the goats had to be known. The traits analyzed were milk, fat and protein yields and fat and protein contents.

Random regression models (RRM)

Five RRM were tested separately for each breed and for all five traits and all had the form:

$$\begin{aligned}
 y_{ijkldgn} = & HTD_{ti} + A_{ijk} + M_{ijl} + \sum_{r=1}^6 \theta_{tkr} N_{(r,d)} \\
 & + \sum_{r=1}^6 \tau_{tlr} N_{(r,d)} + \sum_{s=1}^4 \gamma_{ts} M_{(s,g)} \\
 & + \sum_{o=0}^q a_{tno} \varphi_{(o,d)} + \sum_{o=0}^q p_{tno} \varphi_{(o,d)} + e_{ijkldgn},
 \end{aligned}$$

where $y_{ijkldgn}$ is the observation for trait t (milk yield, fat yield, protein yield, fat content, protein content) on day in milk (DIM) d (from 7 to 270 day) of goat n that was pregnant for the past g days ($g = 0$ if not pregnant), was in production year j (from 1995 to 2015), belonged to kidding age class k (7 classes [in months]: 9–11, 12, 13, 14, 15, 16, 17 or more), to kidding period class l (8 classes: January, February, March, April–May–June, July–August–September, October, November, December), and to herd × test-date class i ; HTD_{ti} is a herd-test-date fixed effect; A_{ijk} is the fixed effect of year-age at kidding; M_{ijl} is the fixed effect of year-kidding period; θ_{tkr} , τ_{tlr} and γ_{ts} are the fixed regression coefficients for age at kidding, kidding period and gestation stage, respectively; $N_{(r,d)}$ is covariate r at time d of a cubic natural spline function with six knots at $d = 7, 20, 50, 110, 190, 270$; $M_{(s,g)}$ is covariate s at time d of a cubic natural spline function with four knots at $d = 31, 53, 76, \text{ and } 100$ (at g between 0 and 30 day, the coefficients were assumed to be equal to 0, implying that there was no effect of gestation on production; due to scarcity of records, the effect of gestation was considered to be constant after $g = 100$). a_{tno} is the set of random additive genetic regression coefficients that follow a normal distribution ($Var(a_{to}) \sim N(0, \sigma_{a_{to}}^2)$) for the o th term of a Legendre polynomial of order q , p_{tno} is the set of random permanent environmental regression coefficients that follow a normal distribution ($Var(p_{to}) \sim N(0, \sigma_{p_{to}}^2)$) for the o th term of a Legendre polynomial of order q , $\varphi_{(o,d)}$ is the value of the o th term of the Legendre polynomial at time d and $e_{ijkldgn}$ is the normally-distributed residual term. The model assumed a heterogeneous residual variance along lactation that was considered to be constant within nine classes of DIM (7–36, 37–66, 67–96, 97–126, 127–156, 157–186, 187–216, 217–246, 247–270), without between-class correlations.

Legendre polynomials of order 0 to 4 (leg0 to leg4) were tested assuming the same order for both the random genetic and environmental effects. The EBV of a goat for a complete lactation milk yield was obtained by summing the EBV for each DIM (called SUM_leg q) [18]. Given that the RRM can give an estimation of the shape of the lactation curve for each goat, EBV for persistency (denoted PERS_leg q) was computed as the cumulative deviation in genetic contribution to yield from the DIM 40 to DIM 240 relative to an average animal having the same yield at DIM 40 [18].

The next step was to test various rank reductions of the genetic variance-covariance matrices based on an eigen decomposition of these matrices. To keep a consistent definition of persistency, the second eigenvector was set, by multiplying by -1 , to be systematically negative at the beginning of lactation, and then positive. For a reduction to

rank z , the first z eigenvectors obtained with the complete model were multiplied by the terms of Legendre polynomial as in Druet et al. [14] to obtain the z first eigenfunctions. These eigenfunctions were used to model the genetic and permanent environment effects. The Legendre polynomial function of order q with a covariance matrix reduced (RRM*) to rank z is written $\text{leg}q\text{Rz}$. The reduced models were of the form:

$$y_{ijklgdn} = HTD_{ti} + A_{tjk} + M_{tjl} + \sum_{r=1}^6 \theta_{tkr} N_{(r,d)} + \sum_{r=1}^6 \tau_{tjr} N_{(r,d)} + \sum_{s=1}^4 \gamma_{ts} M_{(s,g)} + \sum_{o=1}^z b_{tno} \chi_{(o,d)} + \sum_{o=1}^z c_{tno} \chi_{(o,d)} + e_{ijklgdn}$$

where b_{tno} is the random additive genetic regression coefficient that follows a normal distribution ($\text{Var}(b_{tno}) \sim N(0, \sigma_{b_{tno}}^2)$) for the o th eigenfunction of the genetic (co)variances matrix obtained with a q th-order Legendre polynomial model reduced to rank z , c_{tno} is the random permanent environmental regression coefficient that follows a normal distribution ($\text{Var}(c_{tno}) \sim N(0, \sigma_{c_{tno}}^2)$) for the o th eigenfunction of the genetic (co)variances matrix obtained with a q th-order Legendre polynomial model reduced to rank z , $\chi_{(o,d)}$ is the value of the o th eigenfunction of the genetic (co)variances matrix obtained with a q th-order Legendre polynomial model reduced to rank z ($z = \{2, 3\}$) at DIM d .

The corresponding EBV for whole lactation performance of goat n from the RRM was obtained by summing its EBV throughout the DIM period (from DIM 7 to DIM 270) (SUM_legxRz).

Multiple-trait model (MT)

The evolution of the estimated heritabilities throughout the DIM period and the genetic correlations between DIM that were obtained with RRM were compared to those obtained with a MT model, in which lactation was subdivided into six periods (1 = [7, 45], 2 = [46, 90], 3 = [91, 135], 4 = [136, 180], 5 = [181, 225], 6 = [226, 270]). TD yields and contents for each period r ($r=1$ to 6) were considered as different traits. The frequency of TD recording (on average every 28 or 35 days) implied that only one TD record was performed normally per period but if two TD were carried out, the measures were averaged. The MT model was:

$$y_{ijklnro} = \pi_{tir} + \delta_{tjr} + \gamma_{tkr} + HY_{tir} + a_{tnr} + e_{ijklnro}$$

where $y_{ijklnro}$ is the observation for trait t (milk yield, fat yield, protein yield, fat content, protein content) of goat n for period r ($r=1$ to 6), in subclass i for π , the age at

kidding effect, subclass j for δ , the month at kidding effect, subclass k for γ , the gestation stage effect at DIM 270 (the last DIM studied), and class l for the herd-year (HY) effect. The fixed effects classes are the same as in the previous RRM. a_{nr} is the additive genetic breeding value of animal n for period r and followed a normal distribution ($\text{Var}(a_r) \sim N(0, \sigma_{MT_r}^2)$).

Lactation model (LACT)

We implemented a lactation (LACT) model that was close to that routinely used for official genetic evaluations in France. The total milk, fat and protein yields throughout lactation were calculated as in routine genetic evaluations using the Fleischmann method [19]. Fat and protein contents were derived from the total lactation yields. The genetic evaluation model based on these phenotypes was:

$$y_{ijkln} = \pi_i + \delta_j + \gamma_k + HY_l + a_n + e_{ijkln}$$

where y_{ijkln} was the lactation observation for trait t (milk yield, fat yield, protein yield, fat content, protein content) of goat n , in subclass i for π , the age at kidding effect, subclass j for δ , the month at kidding effect, subclass k for γ , the gestation stage effect at DIM 270 (the last DIM studied), and class l for herd-year effect (HY). Fixed effects were defined as in the previous RRM. a_n was the additive genetic breeding value of animal n and followed a normal distribution (mean: $\mu_{LACT} = 0$, variance is σ_{LACT}^2).

All the genetic parameters were estimated using the WOMBAT software [20].

Estimation of genetic correlations and heritabilities with the RRM

For each trait, the heritability of the o th regression coefficients (h_b) was calculated as in Schaeffer [21], by dividing the genetic variance of the o th regression coefficient by the sum of the genetic variance of the o th regression coefficient, the permanent environment variance of the o th regression coefficient, and the mean square error.

For each trait, the genetic variance–covariance matrix between all DIM was obtained following Druet et al. [14] as $G_{264} = QK_gQ'$, where G_{264} is a 264-by-264 genetic variance–covariance matrix, Q is a 264-by- q matrix with the (daily) values of the q terms of the Legendre polynomial, and K_g is the q -by- q genetic variance–covariance matrix. The same approach was used to obtain the permanent environmental variance–covariance matrix W_{264} . The phenotypic variance–covariance matrix between all DIM, P_{264} , was obtained by summing G_{264} , W_{264} and the residual variance for the relevant DIM.

Heritabilities for each test day were obtained by dividing the diagonal elements of G_{264} by the corresponding diagonal elements of P_{264} .

The genetic correlations between DIM d and the other DIM were derived from \mathbf{G}_{264} .

Genetic variances for each trait throughout the whole lactation (g_{wl}) were obtained following Hammami et al. [22] as $g_{wl} = \mathbf{sG}_{264}\mathbf{s}'$, where \mathbf{s} is a summation vector (vector of 1 s) of length 264. The same approach was used to obtain the permanent environmental (w_{wl}) and phenotypic variances for the whole lactation p_{wl} . The heritability of each trait on the lactation scale (h_{wl}) was obtained by dividing g_{wl} by p_{wl} .

Criteria for comparing the models

The goodness-of-fit of the models for each trait was assessed by comparing the Bayesian information criterion (BIC) and the Pearson correlation coefficients (ρ) between observed and predicted phenotypes for each model. Pearson correlation coefficients were also used to compare the EBV of the bucks obtained from different models.

Results

Rank reduction of the variance–covariance matrix of the most complex model

In the Eigen decomposition of the genetic matrix from the leg4 model, the first two principal components (PC) represented on average more than 97% of the total genetic variance (88 and 9%, respectively), and Additional file 1: Table S1 shows the percentage of variance represented by each PC for the five traits and the two breeds. The proportion of variance accounted for by the first eigenvalue was higher for yield traits and fat content in Saanen than in Alpine goats (from +2% for fat content to +4.5% for fat yield). In contrast, the proportion of variance accounted for by the third eigenvalue was higher in Alpine than in Saanen goats, although for all the traits it represented a very small fraction of the total variance (less than 4.2%).

The first eigenfunction was almost constant throughout the DIM period, which suggests that the first PC can be regarded as linked with the average production level throughout lactation, whereas the second PC varied almost linearly, which indicates extreme production levels at the beginning and end of lactation independently of the average production level, and thus it is associated with persistency. The correlations between these measures of average production level throughout lactation and persistency are equal to 0 by construction. The third eigenfunction showed contrasted production characteristics between those measured in the middle of lactation and those measured at the beginning and end of lactation. We found no significant between-trait differences in the shapes of eigenfunctions, see Additional file 2: Figure

S1 that shows the eigenfunctions of each PC for the five traits in the Saanen breed.

Model fitting

Table 1 shows the BIC values for each trait and both breeds with the five complete models (leg0 to leg4) and the five reduced models (leg x R z with $x = \{2, 3, 4\}$ and $z = \{2, 3\}$).

Regardless of trait and breed, BIC decreased rapidly as the order of the Legendre polynomials increased from 0 to 2. The decay was smaller between orders 2 and 4 although 18 additional parameters had to be estimated. However, the differences between all models were significant (difference in BIC > 10), which indicates that, for all traits and both breeds, the fit to the data improved as the order of the Legendre polynomial used increased. As rank decreased, BIC decreased when the first three PC were kept instead of just the first two, without significant differences between leg3R3 and leg4R3. Regardless of the order of the Legendre polynomial, the BIC obtained by using two eigenfunctions of the genetic (co)variances matrix with model leg x R2 were close and often better (i.e. smaller) than with model leg1, with the same number of estimated parameters. The BIC obtained by using three eigenfunctions (leg x R3) was similar and better than that with leg2, again with the same number of estimated parameters.

Pearson correlation coefficients (ρ) between observed data and predicted values were calculated to compare adjustment to data between traits and breeds and between RRM and MT models; Additional file 3: Table S2 presents the evolution of these correlations under the different models, for each trait in both breeds. For RRM, the conclusions drawn were similar to those for BIC. For the fourth-order Legendre polynomial, ρ were high for most traits (~ 0.96), but slightly lower for fat content in both breeds (0.93), which highlights a less satisfactory modeling of this trait than for the others. The MT model was the worst model for all traits, with ρ values ranging from 0.80 to 0.90, which can be explained by the genetic effect and certain fixed effects being constant throughout each period. With random regression models other than leg0, these effects gradually change with DIM.

The evolution of the residual variance with DIM is another criterion for characterizing the effect of this variation on the quality of the adjustments throughout lactation. In the French dairy goat breeding program, the most important trait for cheese production is protein yield, which can also be expressed as protein content with respect to the level of milk yield. The residual variances for milk yield and protein content for different lactation periods are shown in Fig. 1 and Additional file 4: Figure S2, respectively. To keep the figures easy

Table 1 Bayesian information criterion^a (BIC) for the five complete and five reduced models

	Saanen					Alpine				
	Milk yield	Fat yield	Protein yield	Fat content	Protein content	Milk yield	Fat yield	Protein yield	Fat content	Protein content
Complete model										
leg0	19165	11543	13089	4229	14490	21314	14748	13014	4921	18671
leg1	7008	2631	4253	1749	4555	6539	2578	3392	1648	4913
leg2	1472	525	1023	747	1773	1739	835	975	826	2113
leg3	293	81	201	231	788	293	197	220	285	947
leg4	0	0	0	0	0	0	0	0	0	0
Reduced model										
leg2r2	3962	1967	3382	1737	4852	4575	2617	3259	1647	4769
leg3r2	3715	2258	3439	1825	4889	3827	2843	3352	1668	4858
leg4r2	3699	2307	3442	1765	4931	3794	2822	3292	1632	4849
leg3r3	1285	499	936	438	1521	952	612	906	570	1748
leg4r3	1040	500	915	251	1426	873	619	895	481	1898

^a Values are expressed as deviations from the best (smallest) value (leg4)

to interpret, the results for the leg2 and leg3 models are not presented, but these were very close to those of the leg4 model. The between-model differences in residual variance are larger at the beginning and end of lactation. As for BIC and ρ , the higher is the order of the polynomial, the better is the adjustment, as indicated by the smaller residual variances. Figure 1 and Additional file 4: Figure S2 illustrate the better adjustment obtained with leg4 compared to the other models, at both ends of lactation. legxR3 resulted in smaller residual variances than legxR2, particularly at the beginning (DIM 7–45) and end of lactation (DIM 225–270) (results not shown). For clarity, the residuals of the reduced models derived from leg2 and leg3 were not included in the figures but were close to those from leg4. Overall, the RRM* adjustments were systematically much better than for the model that assumes a constant genetic value throughout lactation (leg0). For model fit, similar results were found for the other traits in both breeds.

Choice of final models

Model leg4 was chosen as the reference model for RRM because it resulted in a better fit to the data, but it is also the most complicated model in terms of number of genetic parameters to estimate. However, leg0 and leg1 appear too simplistic. In order to choose a robust model that is sufficiently manageable for large-scale routine evaluations (after extension to all lactations), we chose the leg4R2 model as an attractive compromise, since it derived from the best RRM and its interpretation was straightforward, i.e. it summarized two important lactation characteristics (average level and persistency). The third PC was considered less relevant because it accounted for much less variance. These different RRM (leg0, leg1, leg4R2 and leg4) were compared to a lactation model (LACT) and a MT model for the estimation of genetic parameters.

Heritability estimates

Heritabilities for each trait on the whole lactation (h_wl) scale were estimated as in Hammami et al. [22] with the LACT and different RRM models and were close to 0.3

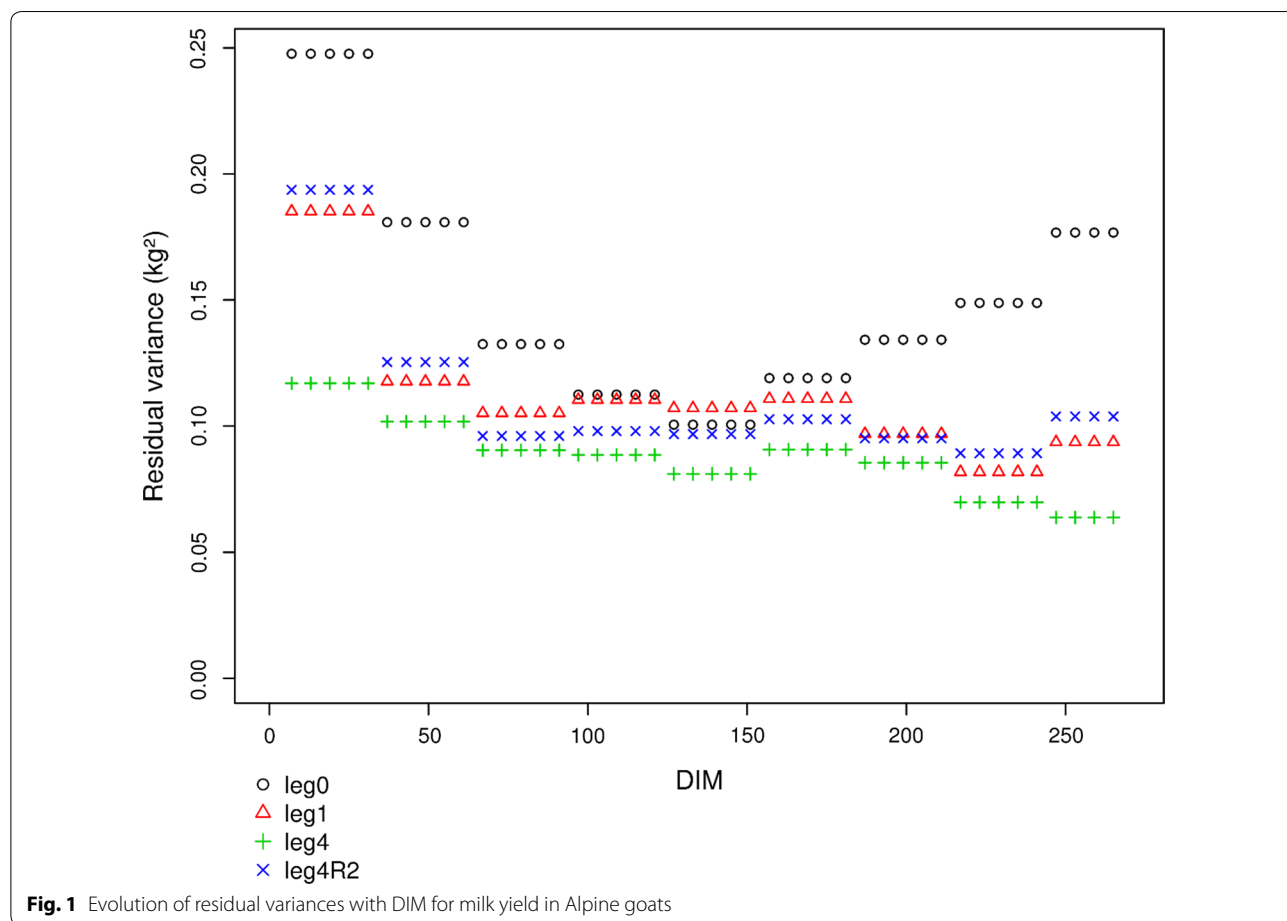


Fig. 1 Evolution of residual variances with DIM for milk yield in Alpine goats

for milk yield in both breeds and for fat and protein yields in Alpine goats, and slightly higher i.e. 0.34 for protein yield and 0.37 for fat yield in Saanen goats (see Additional file 5: Table S3). Heritability estimates reached 0.68 for fat content and 0.65 for protein content in Saanen, and 0.75 for fat content and 0.70 for protein content in Alpine goats. We observed no differences in heritability between the RRM for any of the traits. Furthermore, heritabilities estimated with the RRM* were close to those estimated with RRM. For all the traits, heritabilities estimated with the RRM were slightly higher, especially for fat content, than those estimated with the LACT model

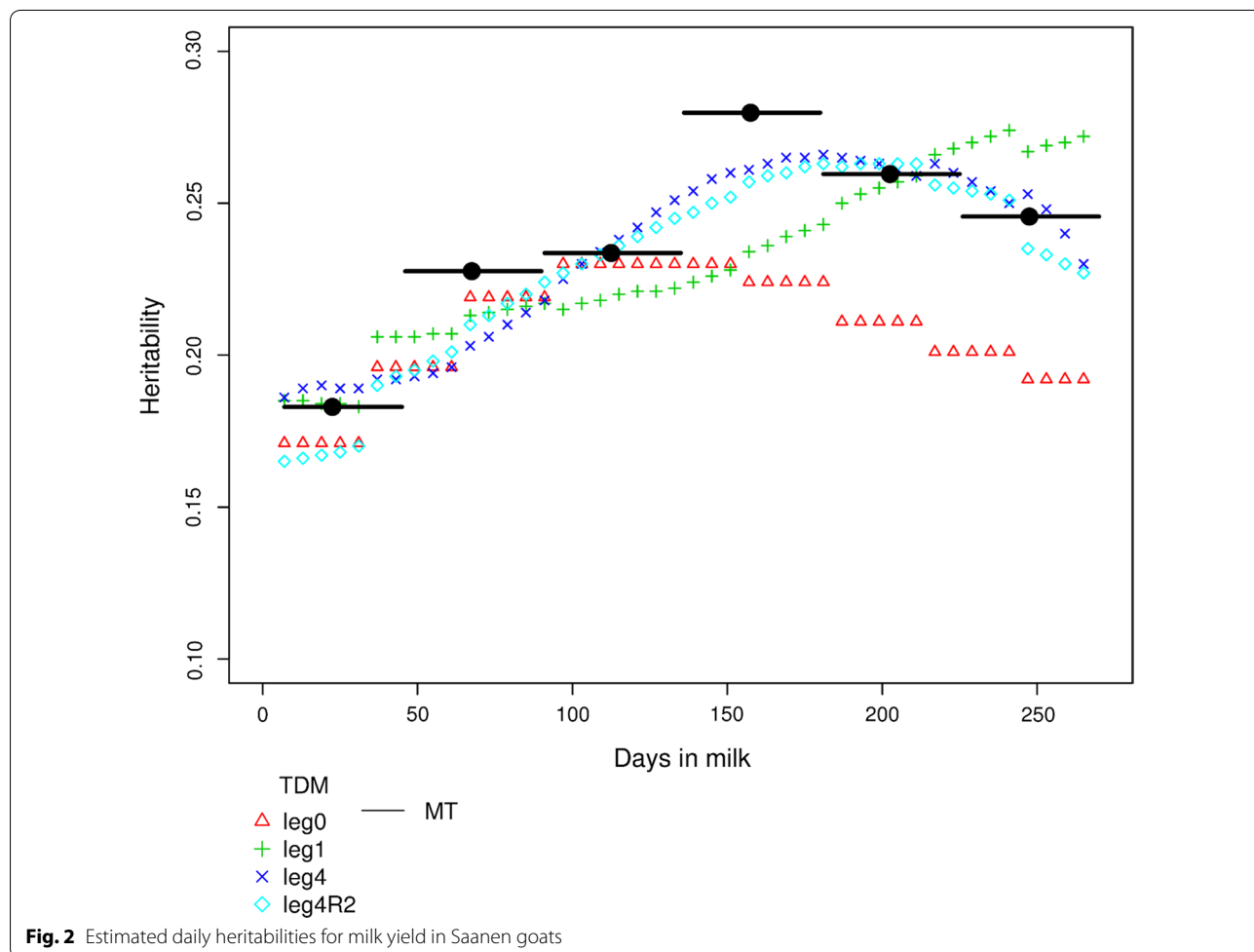
(0.65 for LACT and 0.75 with leg4 for fat content in the Alpine breed).

The heritabilities for the regression coefficients (h_b) estimated with the leg4R2 model as in Schaeffer [21] are in Table 2. These can be interpreted as heritability of average yield (PC1) and persistency of yield (PC2), respectively. The heritabilities for persistency (PC2) were close to 0.18 for protein content and ~0.10 for the other traits.

Figure 2 shows the evolution of the estimated heritabilities of test day yield throughout the lactation period obtained with the MT model and different RRM models

Table 2 Heritabilities of the regression coefficients (h_b) according to the PC derived from the leg4R2 model

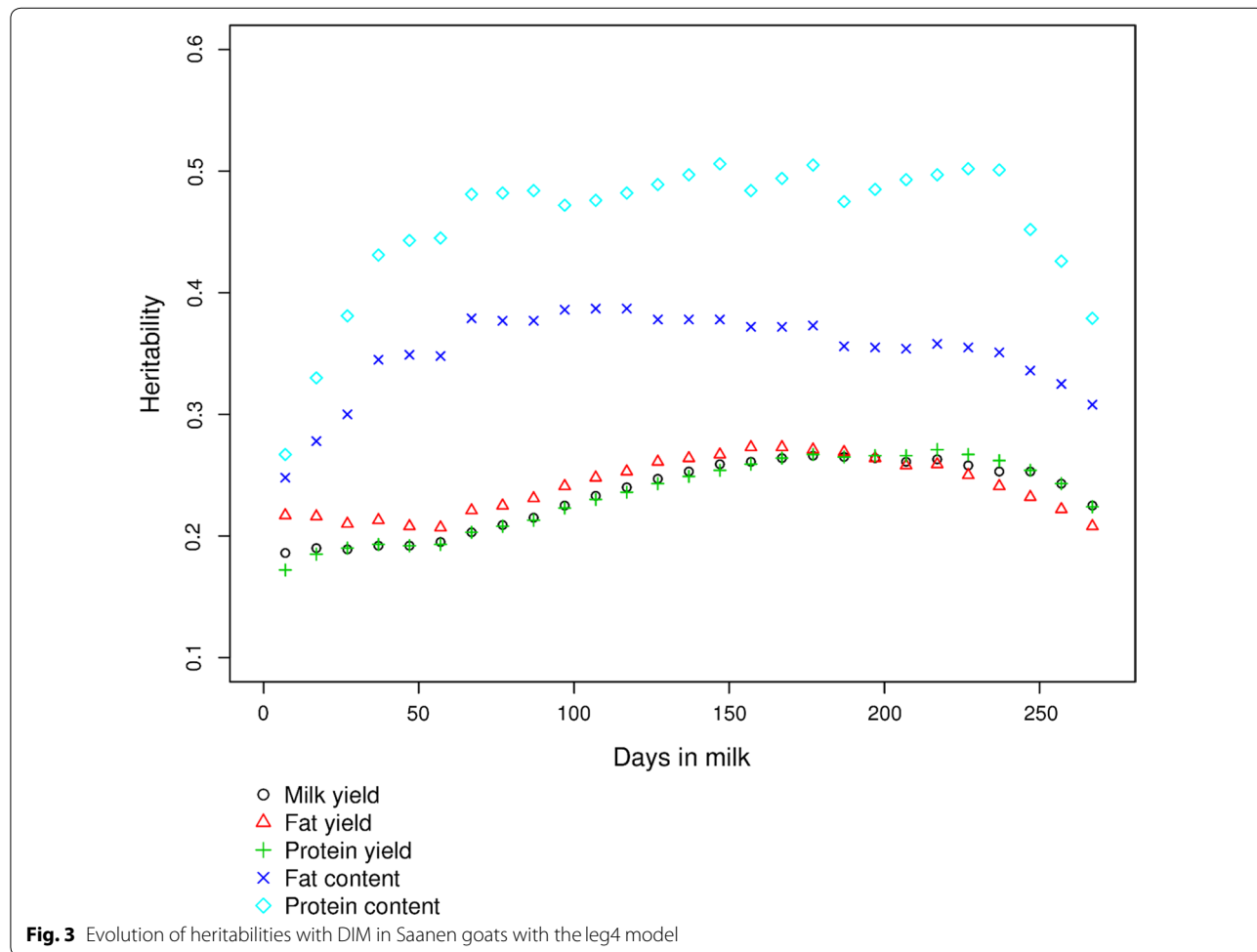
	Saanen					Alpine				
	Milk yield	Fat yield	Protein yield	Fat content	Protein content	Milk yield	Fat yield	Protein yield	Fat content	Protein content
PC1	0.28	0.31	0.29	0.50	0.56	0.27	0.24	0.25	0.54	0.62
PC2	0.13	0.12	0.11	0.07	0.18	0.14	0.10	0.09	0.10	0.19



for milk yield in the Saanen breed, i.e. leg4 and leg4R2, which were the models without and with reduction, leg1, which was a model that had the same number of parameters as leg4R2, and finally leg0. Models leg4 and leg4R2 yielded the highest heritabilities. The discontinuous curve of the estimated heritability throughout the DIM period is due to the residual variance that was different for each of six 30-day periods in the different RRM. The heritability estimated with the MT model first increased up to a peak in mid-lactation (0.26) and then decreased. However, leg0 yielded a lower heritability after DIM 150 for milk yield (due to a larger residual variance), whereas leg1 resulted in a different pattern with low heritabilities at the beginning but high heritabilities at the end of lactation. Additional file 6: Figure S3 shows the evolution of the estimated heritabilities throughout the DIM period for protein content in Saanen goats. The comparison between models resulted in similar results for milk yield and protein content, except with leg1 for protein content, which showed a profile closer to that of leg4. Similar results were also observed for the Alpine breed and the

other traits. Estimated heritabilities obtained with RRM* were close to those calculated with MT and leg4 models, because their eigenfunctions are neither constant nor linear throughout lactation, unlike those of the leg0 and leg1 models.

Figure 3 for the Saanen breed and Additional file 7: Figure S4 for the Alpine breed present the heritabilities estimates for test days with the leg4 model throughout the lactation period for all five traits and show the same within-breed patterns for the three yield traits. The highest heritability estimated for milk yield was 0.26 in both breeds, and was reached on ~DIM 175 and ~DIM 125 for the Saanen and Alpine breeds, respectively. For component yield traits in the Alpine breed, the highest heritability of 0.23 was reached on ~DIM 125, similar to that for milk yield. For component yield traits in the Saanen breed, the estimated heritabilities were highest, i.e. 0.27 on ~DIM 157 for fat yield and ~DIM 211 for protein yield. For protein and fat contents, the evolution of the heritabilities throughout the DIM period followed the same pattern in both breeds but values were higher in the



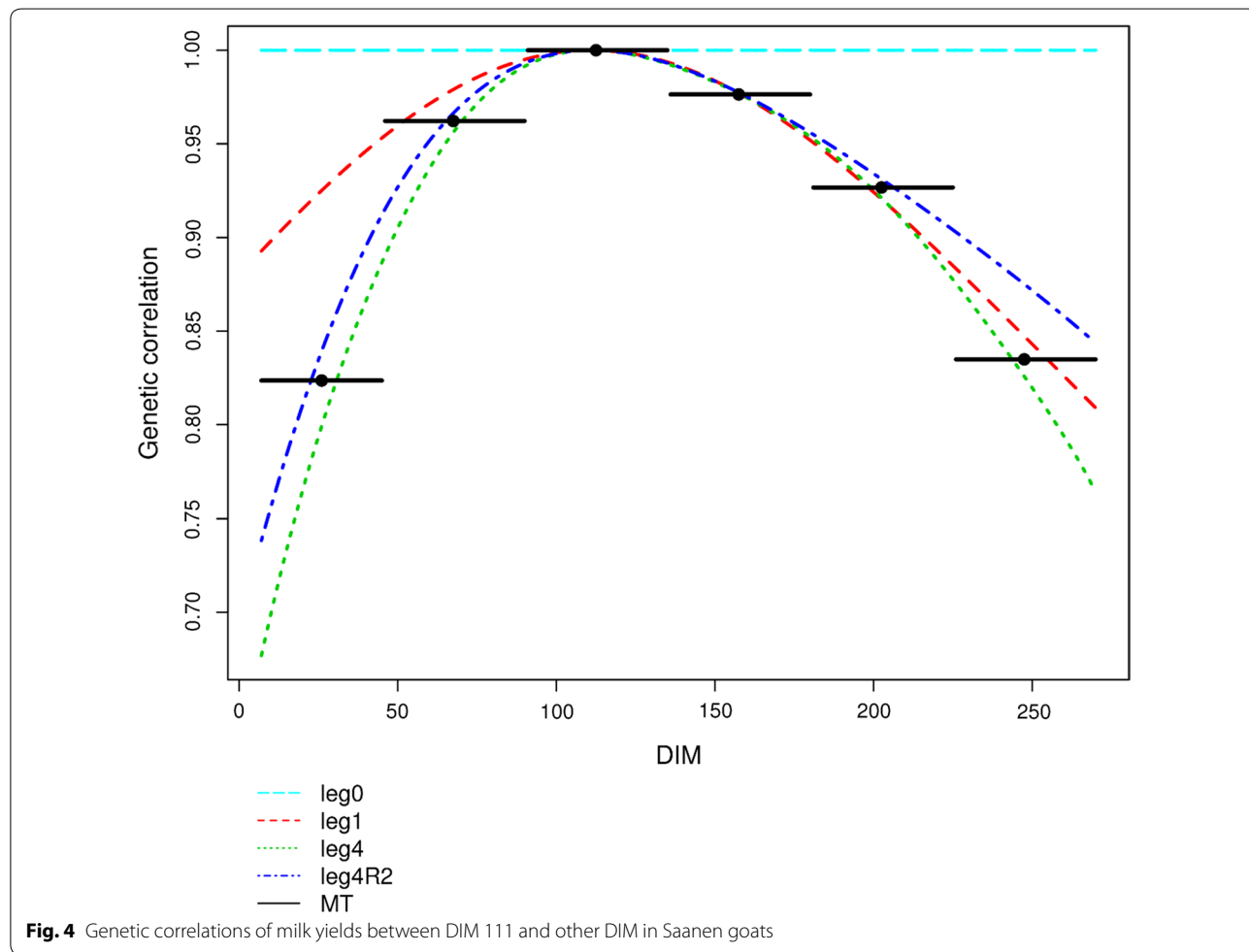
Alpine breed i.e. 0.45 and 0.57, respectively for fat and protein content on ~DIM 150 versus 0.38 for fat content on ~DIM 106 and 0.51 for protein content on ~DIM 150 in the Saanen breed.

Genetic correlations between DIM from different models

In order to compare the genetic correlations between periods obtained with the MT model and the genetic correlations between DIM obtained with the different RRM models, we selected the median day for each period of the MT model. Figure 4 and Additional file 8: Figure S5 show the genetic correlations of milk yield and protein content, respectively, at DIM 111 (i.e. the middle of the third period) with those at other DIM in the Saanen breed with the same models as those used for Fig. 2. The results are similar for both breeds. With model leg0, which assumes a constant genetic effect throughout lactation, the genetic correlation is equal to 1 between all DIM. This is also implicitly the case for the LACT model, which gives a global EBV for the whole lactation assuming a correlation of 1 between all DIM. The most similar

genetic correlations are those estimated with the leg4 and MT models. As observed earlier with BIC, leg4 was the best model among all RRM. For all traits, the genetic correlations between any two periods estimated with the MT and RRM models were positive. The lowest correlation i.e. 0.44 was found with the MT model for milk yield in the Alpine breed between periods 1 and 6. The genetic correlations between DIM from the RRM* were fairly aligned to the pattern from the best full-rank model (leg4) for all traits and both breeds.

Additional file 9: Figure S6 and Additional file 10: Figure S7 show the genetic correlations, estimated with the leg4 model, for all production traits between TD at DIM 40, which corresponds to the DIM at lactation peak, and at the other DIM, in both the Saanen and Alpine breeds. Genetic correlations for milk yield between the peak and the end of lactation were positive, i.e. ~0.5 in the Saanen and ~0.4 in the Alpine breed. In the Alpine breed, daily genetic correlations for fat and protein contents were higher (e.g. a correlation of ~0.7 for protein content between DIM 40 and at the end of lactation) than those



for yield traits (e.g. a correlation of ~0.4 for milk yield between DIM 40 and at the end of lactation). For each trait, the genetic correlations between DIM were higher in the Saanen than the Alpine breed (e.g. correlations of ~0.4 (Alpine) and ~0.5 (Saanen) for milk yield between DIM 40 and end of lactation).

Correlations between EBV for level of production and persistency

Within-trait

Table 3 presents the correlations between EBV for milk yield of bucks that were sires of the phenotyped goats (379 in Alpine and 324 in Saanen) in Saanen under the leg1, leg4 and leg4R2 models.

We were able to evaluate the level of milk produced throughout lactation using the LACT model, or by considering the SUM_legx and SUM_leg4R2 values, or the first coefficient of RRM (a₀) or RRM* (b₁). Correlations between these values reached a value of at least 0.99 for milk yield (Table 3) and all other traits, confirming that they characterize the same trait.

For the study of persistency, the leg0 and LACT models were unsuitable because the resulting EBV were not DIM-dependent. Persistency could be evaluated by considering the PERS_leg4 value, or the second coefficient of RRM (a₁) and RRM* (b₂). A correlation close to 1 was found between the a₁ coefficient of leg4 and PERS_leg4 for milk yield in the Saanen breed. The correlation between b₂ and PERS_leg4 was 0.98 for milk yield and 0.88 for protein yield.

The correlation between SUM_leg4 and PERS_leg4 was low for milk yield (i.e. 0.15, see Table 3) in Saanen and close to 0 in Alpine, but higher for protein content in both breeds (0.45 in Saanen and 0.58 in Alpine) (results not shown). The correlations between a₀ and

a₁ from leg1 and leg4 were equal to 0.44, whereas those between b₁_leg4R2 and b₂_leg4R2 were equal to 0.01. The aim is to have a low genetic correlation between the level and the persistency because it allows selection of animals with a desired persistency throughout the lactation without changing the net-total lactation output.

Correlations between milk yield and other production traits

Table 4 reports the correlations between EBV of SUM_leg4 (i.e. milk production level) and PERS_leg4 (i.e. persistency) for milk yield and of SUM_leg4 and PERS_leg4 for the other traits for Saanen and Alpine bucks.

Table 4 Correlations between EBV for milk yield and the other traits for bucks

	Milk yield			
	Saanen		Alpine	
	SUM_leg4	PERS_leg4	SUM_leg4	PERS_leg4
Fat yield				
SUM_leg4	0.75	0.15	0.73	-0.13
PERS_leg4	-0.07	0.86	-0.17	0.85
Protein yield				
SUM_leg4	0.90	0.18	0.88	-0.04
PERS_leg4	0.16	0.94	0.02	0.93
Fat content				
SUM_leg4	-0.20	0.06	-0.33	-0.14
PERS_leg4	-0.36	-0.10	-0.32	-0.17
Protein content				
SUM_leg4	-0.33	0.01	-0.44	-0.10
PERS_leg4	-0.29	-0.45	-0.28	-0.58

Table 3 Correlations between EBV for milk yield of buck sires of the recorded goats in Saanen

	LACT	SUM_leg4	PERS_leg4	a ₀ _leg4	a ₁ _leg4	SUM_leg1	a ₀ _leg1	a ₁ _leg1	SUM_leg4R2	b ₁ _leg4R2
SUM_leg4	0.99									
PERS_leg4	0.06	0.15								
a ₀ _leg4	0.99	1.00	0.15							
a ₁ _leg4	0.11	0.20	1.00	0.20						
SUM_leg1	0.99	1.00	0.15	1.00	0.20					
a ₀ _leg1	0.99	1.00	0.15	1.00	0.20	1.00				
a ₁ _leg1	0.10	0.19	1.00	0.19	0.99	0.19	0.19			
SUM_leg4R2	0.99	1.00	0.15	1.00	0.20	1.00	1.00	0.19		
b ₁ _leg4R2	0.99	1.00	0.18	1.00	0.23	1.00	1.00	0.22	1.00	
b ₂ _leg4R2	-0.10	-0.01	0.98	-0.01	0.97	-0.01	-0.01	0.97	-0.01	0.02

LACT, lactation model; SUM_legx, sum of the daily EBV of legx; PERS_leg4, persistence calculated as in [18] from leg4; a₀_legx, regression coefficient for the 0th term of Legendre polynomial; a₁_legx, regression coefficient for the 1st term of the Legendre polynomial; b₁_leg4R2, regression coefficient for the first eigenfunction; b₂_leg4R2, regression coefficient for the second eigenfunction

For levels of production, the correlations between SUM_leg4 for milk yield and SUM_leg4 for the other traits averaged 0.75 for fat yield, 0.89 for protein yield, -0.25 for fat content and around -0.4 for protein content. The correlations between SUM_leg4 for milk yield and persistencies of the other traits were moderate (all < 0.36 in absolute value). The correlations between milk yield persistency and SUM_leg4 for the other traits were close to 0 (between -0.14 and 0.18) for both breeds. The correlations between milk persistency and the other persistencies were very high (> 0.85) for protein and fat yields but very low for fat content (-0.13, on average). However, the correlation with persistency of protein content was negative and relatively high (-0.45 in Saanen and -0.58 in Alpine). These quite high correlations indicate that a higher milk persistency is positively correlated with a higher-than-average protein content at the beginning of lactation. Because the correlation between fat content persistency and milk yield persistency was low (-0.10 in Saanen and -0.17 in Alpine) and because protein content at the beginning of lactation was higher

than average, persistency is correlated with a reduced fat:protein ratio at the beginning of lactation.

As expected, the correlations of EBV between b_1_{leg4R2} and b_2_{leg4R2} for milk yield and b_1_{leg4R2} and b_2_{leg4R2} for the other traits (not shown) were very close to those obtained in Table 4. This result confirms that b_1_{leg4R2} is a measure of lactation production and b_2_{leg4R2} is a measure of persistency.

Relative contribution of b_1 and b_2 coefficients to lactation production

Figure 5 and Additional file 11: Figure S8 show the relative contribution to lactation production of a change of one standard deviation in the b_1 (level) and b_2 (slope) coefficients for milk yield and protein content, respectively, with leg4R2. An increase of one genetic standard deviation for b_1 leads to an increased production level throughout lactation, with a higher increase at the end of lactation for Saanen compared to Alpine goats. This study on the one-standard-deviation increase of b_1 and b_2 allows to analyze the potential result of a selection

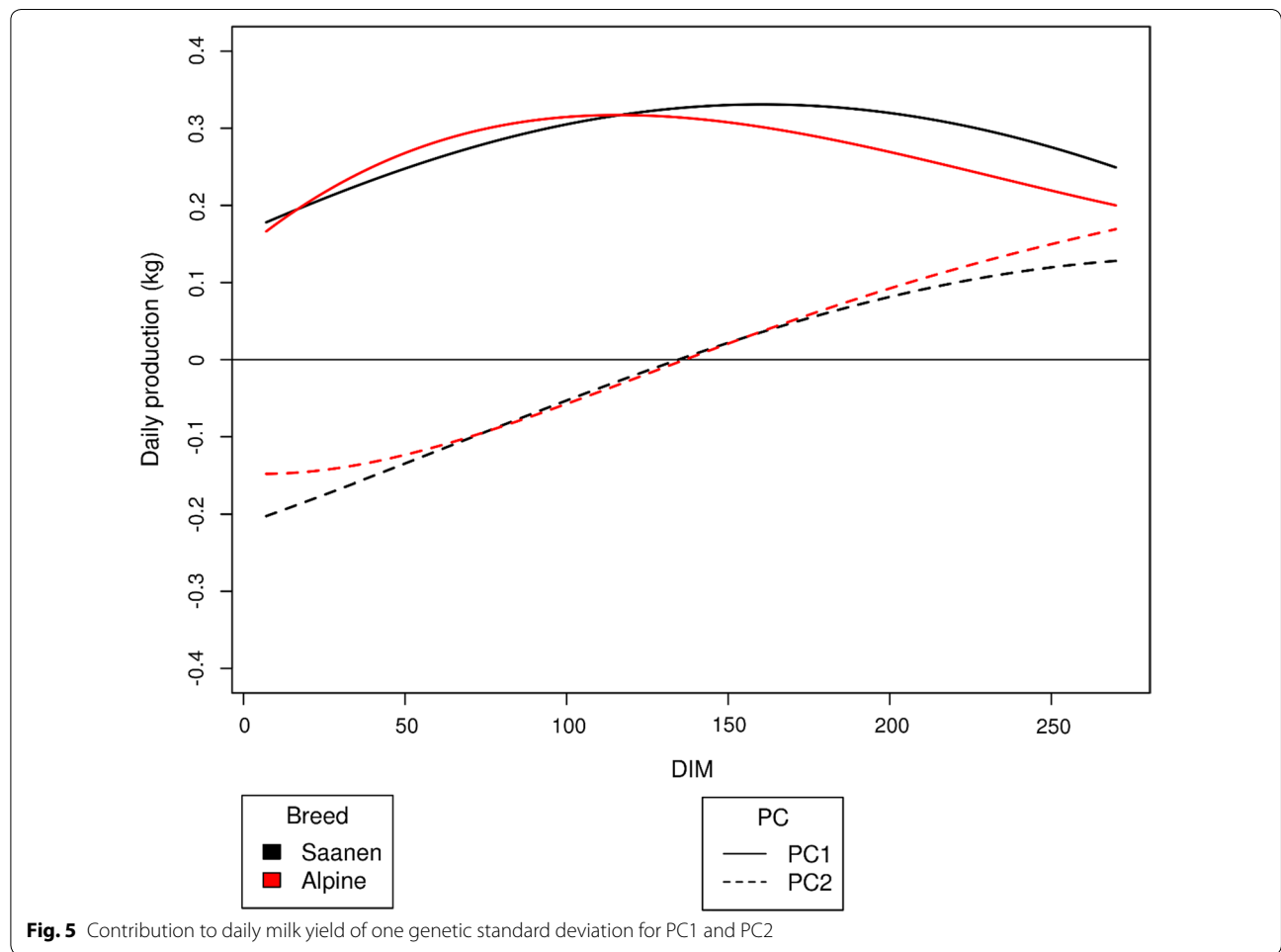


Fig. 5 Contribution to daily milk yield of one genetic standard deviation for PC1 and PC2

on these two traits. As a consequence of the one-standard-deviation increase of b_1 , the sum of the added daily milk productions represents 76 kg of milk for Saanen and 71.1 kg for Alpine goats, which is relatively consistent with the LACT model (genetic standard deviation equal to 92 kg in Saanen and 85 kg in Alpine goats). At the beginning of lactation, increasing the persistency (b_2) by one standard deviation reduces production by 12.9 kg until DIM 129 and then increases it by 10.5 kg at the end of the lactation for Saanen. For protein content, a gain of one standard deviation of b_1 increases average production more in Alpine than in Saanen goats. A loss of one standard deviation in protein content persistency leads to an increase of 1 g/kg at the beginning of lactation and a loss of 1 g/kg at the end of lactation.

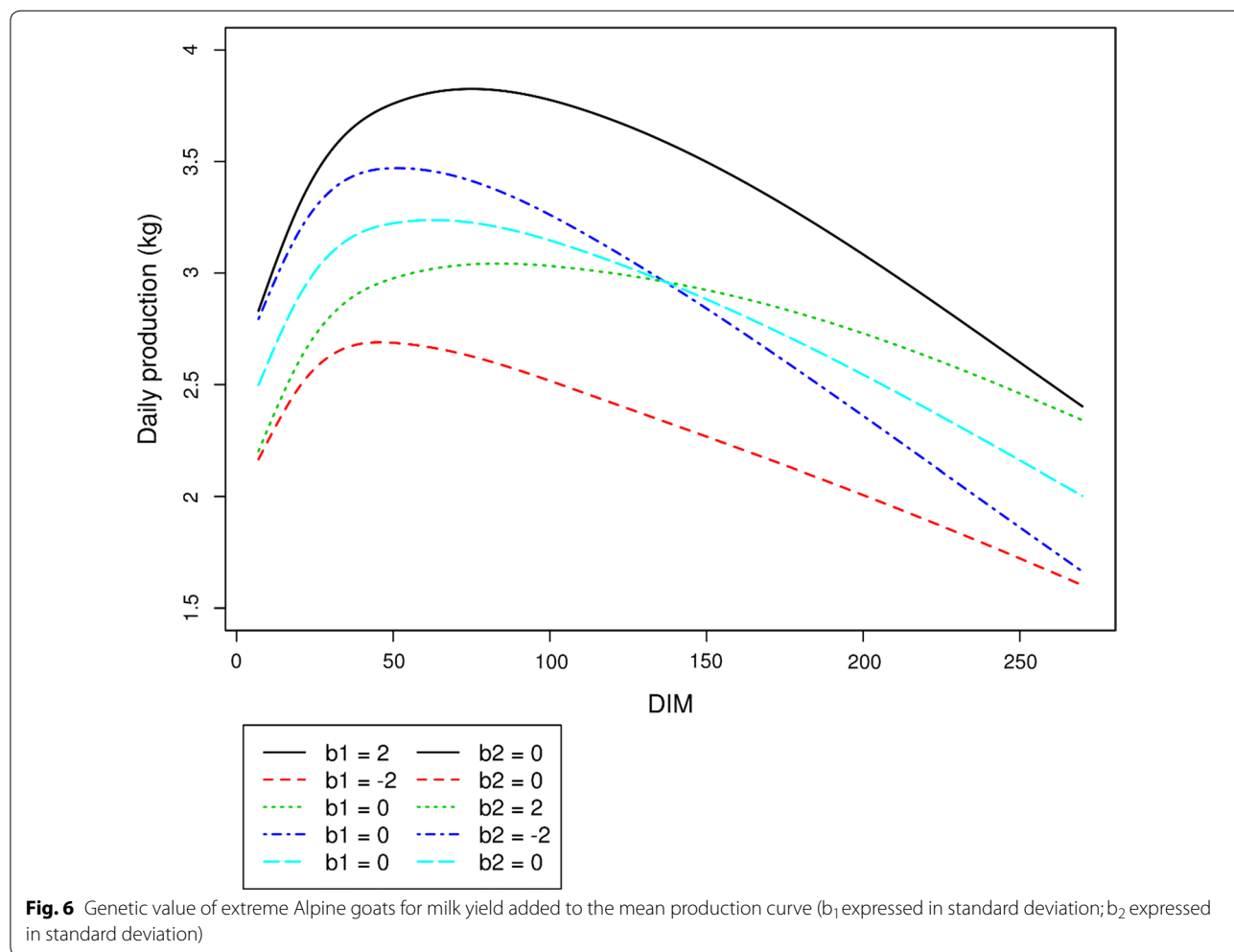
As an illustration, Fig. 6 displays the genetic daily EBV added to the average standard production curve, for two extreme-value goats for global production level and two extreme-value goats for milk yield persistency.

Additional file 12: Figure S9 shows the same results but for protein yield.

Discussion

We estimated various genetic parameters in a large population of goats that had been regularly measured for five major dairy traits throughout lactation since 1995. This estimation was done by using a random regression model for the first time in France for the two main goat breeds Alpine and Saanen, which represent 97% of the goats recorded in France. Connectedness between the large number of herds was ensured by using phenotypes from daughters of sires from artificial insemination (379 and 324 in the Alpine and Saanen breed, respectively).

The fixed effects, i.e. age at kidding, month of kidding, dry-period length and gestation stage, were modeled with splines in the RRM model. In a previous study [11], we showed that these factors had an impact on the shape of the lactation curve, and that there was no interaction between these effects and year of lactation. These



results were consistent with other studies in dairy cows [14, 23]. Unlike other studies in dairy goats [4, 5], we did not use an effect on litter size since this information was not available in our dataset, but Mucha et al. [1] found no impact of this effect on EBV.

We applied a principal component analysis of the variance–covariance matrix of the most complex model, which showed that it was possible to associate biological significance with each PC. As in our study, van der Werf et al. [17], Olori et al. [24], Druet et al. [14] and Togashi and Lin [25] highlighted that the first eigenfunction was linked with the average production level throughout lactation, the second eigenfunction was associated with persistency, and the third eigenfunction opposed production around the middle of lactation against production at the beginning and end of lactation. The high percentages of variances explained by the first three PC (and even the two first PC), and the desire to reduce the overall dimension of the model in routine evaluations, pointed at the need to test a rank reduction of the variance–covariance matrix.

Implementing a RRM for genetic evaluation requires cumbersome testing of the best tradeoffs between model fit and complexity. Pool et al. [26] compared the residual variances of different RRM on cow data and found, as we did here, that the most suitable model for TD milk production traits was leg4. The use of other criteria (BIC and Pearson correlation coefficients) also confirmed that leg4 was the most suitable model for TD milk production traits. The leg \times R3 reductions (for $x = \{3,4\}$) had a better fit to the data than leg \times R2 although leg \times R2 reductions led to a better adjustment to the data compared to other more concise models such as leg0 and leg1, especially for milk yield. This rank reduction facilitates the extension of the model to subsequent lactations 2 and 3 in order to construct a genetic evaluation. Indeed, this model extension has to take into account how genetic correlations between first and following parities will probably differ from 1 (around 0.7 for milk yield) as found by several studies in dairy goats [1, 4].

For all traits, heritability estimates from the LACT model and for the regression coefficients from RRM and RRM* were close to those reported by Rupp et al. [27] for the same five traits in French Alpine and Saanen goats in first lactation. The evolution of the heritability throughout lactation for each trait was similar between RRM and RRM*. In Norwegian goats, Andonov et al. [6] showed that the heritability of milk yield estimated with the MT model increased up to a maximum at mid-lactation (0.26) and then decreased, which agrees with our observations. The evolution of the heritability throughout the DIM period found here for milk yield differed from that reported in other studies using RRM

in goats. Menéndez-Buxadera et al. [4] found a maximum heritability of 0.24 at the beginning of lactation, then a decrease with DIM in Murciano-Granadina goats whereas Zumbach et al. [5] found a maximum heritability of 0.4 and then a decrease with DIM in six German breeds. Andonov et al. [6] found a maximum heritability of 0.33 at \sim DIM 155, in a Norwegian goat breed and Mucha et al. [1] found a maximum heritability of 0.45 at \sim DIM 220, in a crossbred population including three goat breeds: Alpine, Saanen, and Toggenburg. The mean heritability estimates for protein yield found here (\sim 0.2) was in agreement with that reported by Muños-Mejías et al. [3] for goats of the Florida breed. This was not the case for fat yield, for which we found a higher heritability (\sim 0.2) than that for the Florida breed (\sim 0.15). Moreover, Muños-Mejías et al. [12] showed that the estimated heritabilities for fat and protein yield for Florida goats were higher at the end of lactation, which was not the case in our study. For protein and fat contents, the estimated heritabilities reported in Muños-Mejías et al. [3] and Andonov et al. [6] were lower than those found here, and followed a different shape, i.e. they increased with DIM. The RRM* made it possible to calculate the heritability of persistency for each trait [21]. The heritabilities of persistency for milk yield calculated from RRM* were close to those reported by Cole and VanRaden [8] in cattle, who also calculated milk yield persistency ensuring that it was not correlated with milk yield level. Menéndez-Buxadera et al. [4] reported a heritability of milk yield persistency 0.208, but the correlation with lactation yield level is unknown.

As in our study, Mucha et al. [1] found that all the genetic correlations for milk yield between two periods in the trajectory fitted in the RRM models were always positive. This indicates that selection of animals based on any daily EBV will yield positive responses for all the other days in the lactation curve. Also in agreement with our results, Muñoz-Mejías et al. [3], Andonov et al. [6] and Menéndez-Buxadera et al. [4] found that genetic correlations between days were higher for fat and protein contents than for yield traits. The moderate genetic correlations for milk yield between the peak and end of lactation indicated genetic variability in the level of milk production between the peak and end of lactation, and therefore a genetic variability for milk persistency. For all traits, the between-day genetic correlations were similar with RRM or RRM*.

We showed that correlations between EBV from RRM and RRM* were close to 1 for both average and persistency of yield as reported in Leclerc et al. [28] who compared EBV from a complete and a reduced model. The advantages of the reduced model are the smaller size of the variance–covariance matrices and the zero

correlation between EBV for lactation yield (b_1) and persistency (b_2) by construction. This allows selection of animals with a desired persistency throughout the lactation without changing the net-total lactation output. Model leg1 does not have this advantage as the correlation between mean level and persistency is high in that model. The large difference in lactation trajectory between extreme animals suggests that it can be valuable to consider persistency in selection. The form of the first eigenfunction (Fig. 5) is interesting because it represents the pattern of how a goat produces milk throughout lactation from its genetic makeup. The eigenfunction coordinates were higher at the end of the lactation period for the Saanen breed than for Alpine breed, confirming the observations of Arnal et al. [11] who showed a better persistency for Saanen goats than Alpine goats. This is also evidenced by the higher correlation between SUM_leg4 and PERS_leg4 for the Saanen compared to Alpine goats (0.15 in Saanen vs. 0 in Alpine) as well as the higher genetic correlations between DIM in Saanen compared to Alpine goats.

The correlations of the production level for milk yield with the production level for the other traits were very close to those reported by Bélichon et al. [29] on total lactation traits (0.90 for protein yield and 0.76 for fat yield). For protein content and fat content, Bélichon et al. [29] found slightly lower correlations, i.e. -0.28 for protein content and -0.13 for fat content while we found 0.39 and -0.27 , respectively, in our study. The correlations between milk yield persistency and the production level for the other traits were weak or close to 0 for both breeds, indicating that milk yield persistency can be selected for with no impact on content-related traits. Furthermore, the difference in the correlations observed here between milk persistency and protein content persistency on the one hand, and between milk persistency and fat content persistency on the other hand, indicates that a goat with a high milk persistency will tend to have a lower fat:protein ratio in early lactation. Several studies in dairy cattle [13, 30, 31] showed that a high fat:protein ratio was associated with a negative energy balance, sub-clinical mastitis, and poor fertility. These results point out to the potential value of lactation persistency in breeding schemes.

Finally, the development of a test-day model for milk production opens up new perspectives. For example, the study of the genetic relationship between persistency and other traits such as longevity or fertility could help to explain the negative correlation between high production and fitness. The estimated fixed effects of such test-day models, especially the herd-test-day effect, also offer producers important clues on the impact of herd

management on these traits. For example, herd test-day estimates can be compared between farms under similar systems in terms of mean and variability throughout the year [23].

Conclusions

In this paper, we show that the genetic parameters obtained with a test-day model using a fourth-order Legendre polynomial (leg4) for the genetic and permanent environmental components and a multi-trait model are consistent. However, this kind of model is complex and computationally demanding. Given that the aim was to develop TD genetic evaluations for traits in selection schemes (milk production traits and somatic cell score) over several parities and possibly including genomic information, a simpler model is necessary. We found that reducing the genetic and permanent environment (co) variance matrices of leg4 to its first two PC (leg4R2) was a satisfactory compromise, which accurately approximates genetic parameters and EBV under the complete model. This reduced model can give EBV for total lactation milk yield and persistency that are nearly independent (correlation close to 0). This negligible correlation is appealing because it allows to select animals with a desired lactation shape independently from selection for total lactation production. Therefore, for the extension of the TD models to several lactations, we consider that leg4R2 is more appropriate for implementation than the complete model leg4.

Additional files

Additional file 1: Table S1. Proportion (%) of genetic variance explained by the first three principal components (PC) using leg4.

Additional file 2: Figure S1. Coordinates of the first three eigenfunctions of the test-day genetic covariance matrix computed with leg4 in Saanen goats.

Additional file 3: Table S2. Pearson correlation coefficients (ρ) between observed and predicted values with the complete and reduced models. MT: multitrait model; legx: Legendre polynomial of order x; legxRz: Legendre polynomial of order x reduced to order z.

Additional file 4: Figure S2. Evolution of residual variances with DIM for protein content in Alpine goats (full-rank models: leg0, leg1, leg2, leg4; reduced model: leg4R2).

Additional file 5: Table S3. Whole-lactation heritabilities from the LACT model and heritabilities (h_{wl}) calculated as in [22] in Saanen and Alpine goats.

Additional file 6: Figure S3. Daily estimated heritabilities for protein content in Saanen goats (full-rank models: leg0, leg1, leg4; best reduced model: leg4R2 and the MT model).

Additional file 7: Figure S4. Evolution of heritabilities with DIM in Alpine goats with the leg4 model.

Additional file 8: Figure S5. Genetic correlations of protein contents between DIM 111 and other DIM in Saanen goats.

Additional file 9: Figure S6. Genetic correlations between DIM 40 and other DIM in Alpine goats from the leg4 model.

Additional file 10: Figure S7. Genetic correlations between DIM 40 and other DIM in Saanen goats from the leg4 model.

Additional file 11: Figure S8. Contribution to daily protein content of one genetic standard deviation for PC1 and PC2.

Additional file 12: Figure S9. EBV of Alpine goats added to the mean production of the population for protein content (b_1 expressed in standard deviation; b_2 expressed in standard deviation).

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Authors' contributions

CRG, HL and VD proposed the models for study and test. MA performed the analysis and wrote the paper. MA, CRG, HL, VD and HL interpreted the results. CRG, HL, VD and HL revised and improved the manuscript. All authors read and approved the final manuscript.

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Availability of data and materials

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Ethics approval and consent to participate

Not applicable.

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

Author details

¹ GenPhySE, Université de Toulouse, INRA, ENVT, Toulouse INP, 31326 Castanet-Tolosan, France. ² Institut de l'Élevage, Chemin de Borde Rouge, 31326 Castanet-Tolosan Cedex, France. ³ INRA UMR1313 GABI, Institut de l'Élevage, 78350 Jouy-en-Josas, France. ⁴ UMR1313 GABI, INRA, AgroParisTech, Université Paris-Saclay, 78350 Jouy-en-Josas, France.

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References

1. Mucha S, Mrode R, Coffey M, Conington J. Estimation of genetic parameters for milk yield across lactations in mixed-breed dairy goats. *J Dairy Sci.* 2014;97:2455–61.
2. Brito LF, Silva FG, Oliveira HR, Souza NO, Caetano GC, Costa EV, et al. Modelling lactation curves of dairy goats by fitting random regression models using Legendre polynomials or B-splines. *Can J Anim Sci.* 2017;98:73–83.
3. Muñoz-Mejías ME, Menéndez-Buxadera A, Sánchez-Rodríguez M, Serradilla JM. Genetic progress attained in the selection program of Florida breed of goats in Spain. *Option Méditerranéennes Ser A.* 2013;108:134–9.
4. Menéndez-Buxadera A, Molina A, Arrebola F, Gil MJ, Serradilla JM. Random regression analysis of milk yield and milk composition in the first and second lactations of Murciano-Granadina goats. *J Dairy Sci.* 2010;93:2718–26.
5. Zumbach B, Tsuruta S, Misztal I, Peters KJ. Use of a test day model for dairy goat milk yield across lactations in Germany. *J Anim Breed Genet.* 2008;125:160–7.
6. Andonov S, Ødegård J, Svendsen M, Ådnøy T, Vegara M, Klemetsdal G. Comparison of random regression and repeatability models to predict

- breeding values from test-day records of Norwegian goats. *J Dairy Sci.* 2013;96:1834–43.
7. Schaeffer LR, Jamrozik J. Random regression models: a longitudinal perspective. *J Anim Breed Genet.* 2008;125:145–6.
8. Cole JB, VanRaden PM. Genetic evaluation and best prediction of lactation persistency. *J Dairy Sci.* 2006;89:2722–8.
9. Sölkner J, Fuchs W. A comparison of different measures of persistency with special respect to variation of test-day milk yields. *Livest Prod Sci.* 1987;16:305–19.
10. Gipson TA, Grossman M. Lactation curves in dairy goats: a review. *Small Ruminant Res.* 1990;3:383–96.
11. Arnal M, Robert-Granié C, Larroque H. Diversity of dairy goat lactation curves in France. *J Dairy Sci.* 2018;101:11040–51.
12. Schaeffer LR, Jamrozik J, Kistemaker GJ, Van Doormaal J. Experience with a test-day model. *J Dairy Sci.* 2000;83:1135–44.
13. Jamrozik J, Schaeffer LR. Test-day somatic cell score, fat-to-protein ratio and milk yield as indicator traits for sub-clinical mastitis in dairy cattle. *J Anim Breed Genet.* 2012;129:11–9.
14. Druet T, Jaffrézic F, Boichard D, Ducrocq V. Modeling lactation curves and estimation of genetic parameters for first lactation test-day records of French Holstein cows. *J Dairy Sci.* 2003;86:2480–90.
15. White IMS, Thompson R, Brotherstone S. Genetic and environmental smoothing of lactation curves with cubic splines. *J Dairy Sci.* 1999;82:632–8.
16. Misztal I. Properties of random regression models using linear splines. *J Anim Breed Genet.* 2006;123:74–80.
17. van der Werf JHJ, Goddard ME, Meyer K. The use of covariance functions and random regressions for genetic evaluation of milk production based on test day records. *J Dairy Sci.* 1998;81:3300–8.
18. Jamrozik J, Schaeffer LR, Dekkers JCM. Genetic evaluation of dairy cattle using test day yields and random regression model. *J Dairy Sci.* 1997;80:1217–26.
19. Sargent FD, Lytton VH, Wall OG. Test interval method of calculating dairy herd improvement association records. *J Dairy Sci.* 1968;51:170–9.
20. Meyer K. WOMBAT—A tool for mixed model analyses in quantitative genetics by restricted maximum likelihood (REML). *J Zhejiang Univ Sci B.* 2007;8:815–21.
21. Schaeffer LR. Random regression models; 2016. <http://animalbioscienc.es.uoguelph.ca/~lrs/BOOKS/rrmbook.pdf> Accessed 3 July 2019.
22. Hammami H, Reik B, Soyeurt H, Gara AB, Gengler N. Genetic parameters for Tunisian Holsteins using a test-day random regression model. *J Dairy Sci.* 2008;91:2118–26.
23. Leclerc H. Development of the French dairy cattle test-day model genetic evaluation and prospects of using results for herd management. Ph.D. thesis, AgroParisTech; 2008.
24. Olori VE, Hill WG, McGuirk BJ, Brotherstone S. Estimating variance components for test day milk records by restricted maximum likelihood with a random regression animal model. *Livest Prod Sci.* 1999;61:53–63.
25. Togashi K, Lin CY. Selection for milk production and persistency using eigenvectors of the random regression coefficient matrix. *J Dairy Sci.* 2006;89:4866–73.
26. Pool MH, Janss LLG, Meuwissen THE. Genetic parameters of Legendre polynomials for first parity lactation curves. *J Dairy Sci.* 2000;83:2640–9.
27. Rupp R, Clément V, Piacere A, Robert-Granié C, Manfredi E. Genetic parameters for milk somatic cell score and relationship with production and udder type traits in dairy Alpine and Saanen primiparous goats. *J Dairy Sci.* 2011;94:3629–34.
28. Leclerc H, Nagy I, Ducrocq V. Impact of using reduced rank random regression test-day model on genetic evaluation. *Interbull Bull.* 2009;40:42–6.
29. Bélichon S, Manfredi E, Piacère A. Genetic parameters of dairy traits in the Alpine and Saanen goat breeds. *Genet Sel Evol.* 1999;31:52934.
30. Negussie E, Strandén I, Mäntysaari EA. Genetic associations of test-day fat:protein ratio with milk yield, fertility, and udder health traits in Nordic Red cattle. *J Dairy Sci.* 2013;96:1237–50.
31. Buttcherer N, Stamer E, Junge W, Thaller G. Genetic relationships among daily energy balance, feed intake, body condition score, and fat to protein ratio of milk in dairy cows. *J Dairy Sci.* 2011;94:1586–91.

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