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Linking intrinsic scales of ecological processes to characteristic scales of biodiversity and functioning patterns

Yuval R. Zelnik^{1,2*}, Matthieu Barbier³, David W. Shanafelt⁴, Michel Loreau¹, Rachel M. Germain⁵

¹Theoretical and Experimental Ecology Station, CNRS, Moulis, France

²Department of Crop Production Ecology, Swedish University of Agricultural Sciences, Uppsala, Sweden

³Plant Health Institute Montpellier, CIRAD, France

⁴Université de Lorraine, Université de Strasbourg, AgroParis Tech, Centre National de la Recherche Scientifique (CNRS), Institut National de Recherche pour l'Agriculture, l'Alimentation et l'Environnement (INRAE), Bureau d'Economie Théorique et Appliquée (BETA), Nancy, France

⁵Department of Zoology & the Biodiversity Research Centre, University of British Columbia, Vancouver, BC, Canada

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Abstract

Ecology is a science of scale, which guides our description of both ecological processes and patterns, but we lack a systematic understanding of how process scale and pattern scale are connected. Recent calls for a synthesis between population ecology, community ecology, and ecosystem ecology motivate the integration of phenomena at multiple levels of organization. Furthermore, many studies leave out the scaling of a critical process: species interactions, which may be non-local through movement or foraging and must be distinguished from dispersal scales. Here, we use simulations to explore the consequences of three different process scales (species interactions, dispersal, and the environment) on emergent patterns of biodiversity, ecosystem functioning, and their relationship, in a spatially-explicit landscape and stable equilibrium setting. A major result of our study is that the spatial scales of dispersal and species interactions have opposite effects: a larger dispersal scale homogenizes spatial biomass patterns, while a larger interaction scale amplifies their heterogeneity. Interestingly, the specific scale at which dispersal and interaction scales begin to influence landscape patterns depends on the scale of environmental heterogeneity – in other words, the scale of one process allows important scales to emerge in other processes. This interplay between process scales, i.e., a situation where no single process dominates, can only occur when the environment is heterogeneous and the scale of dispersal small. Finally, contrary to our expectations, we observe that the spatial scale of ecological processes is more clearly reflected in landscape patterns (i.e., distribution of local outcomes) than in global patterns such as Species-Area Relationships or large-scale biodiversity-functioning relationships. Overall we conclude that long-range interactions often act differently and even in opposite ways to dispersal, and that the landscape patterns that emerge from the interplay of long-ranged interactions, dispersal and environmental heterogeneity are not well captured by often-used metrics like the Species-Area Relationship.

Introduction

Scale is fundamental to ecology, from the spatial and temporal scales at which we observe and manage ecosystems [1, 2, 3] to the intrinsic scales at which processes occur within and across ecosystems [4]. Much of current research efforts describe ecological patterns across scales, such as Species-Area or Biodiversity-Ecosystem Functioning relationships [5, 3]. However, the scaling of ecological patterns is largely phenomenological – we can describe how patterns scale but not why [6, 5]. Although links between scales of patterns and processes have been explored in recent years [7, 8, 9], as we will discuss, a systematic and unified treatment of scale in ecology is incomplete. A critical question remains: how is the scaling of ecological patterns, such as patterns of biodiversity and ecosystem functioning, generated by scales of specific processes, and why?

In answering this question, a crucial process is often overlooked: the spatial scale of species interactions. While dispersal and environmental variation are often understood to operate at various spatial scales, existing research generally assumes that species only interact locally [10, 11, 12] (although exceptions exist, e.g., studies using multi-layer networks

* Corresponding author: zelnik@post.bgu.ac.il

59 to link interaction networks at local scales to their realization at the global scale [13, 14]).
60 Yet many species move, forage, or otherwise interact with each other at a range of spatial
61 scales [15, 16], even in the absence of dispersal. A simple distinction is that dispersing species
62 establish new “home” ranges when they move across the environment, while mobile species
63 always return to their “home” range. Many move daily across multiple habitat types, such
64 as seabirds connecting marine and terrestrial ecosystems [15], or predatory insects moving
65 between different habitats in the landscape [16]. Non-local competition can therefore arise
66 from foraging across multiple localities. Additionally, species interact indirectly across long
67 distances via intermediary species, (e.g., plants interacting indirectly via pollinators or her-
68 bivores), and many such intermediary interactions are not explicitly studied, thus being
69 best represented by long range interactions. As a result, scales of species interactions, such
70 as competition, likely have consequences for population persistence, affecting the spatial
71 distribution of biodiversity and ecosystem functioning in ways that are distinct from other
72 process scales [17, 18].

73 How do the spatial scales of dispersal, environmental heterogeneity, and species inter-
74 actions interactively influence ecological patterns? Answering this question is unlikely to
75 be achieved via observational studies, as different combinations of ecological processes may
76 generate identical patterns, but computational models can explore patterns that emerge as
77 processes interact across scales. Indeed, the scale of dispersal relative to the environment
78 has been studied most extensively, in particular within a metacommunity context [19, 7, 20].
79 These studies generally find that high rates of dispersal blur differences between local com-
80 munities, leading to losses of biodiversity and ecosystem functioning. Although there are
81 reasons to expect increased scales of dispersal and species interactions to have similar con-
82 sequences, as both processes are influenced by many of the same variables (e.g., animal
83 mobility) and serve to spread out the effects of species interactions, there are also reasons
84 to expect the opposite [21]. A key difference is that large dispersal scales can allow popu-
85 lations to permeate through whole landscapes over a few generations, whereas individuals
86 with large interaction scales are still bound to specific localities. As a result, increasing
87 scales of interactions may amplify spatial heterogeneity in an ecological system [22], counter
88 to the blurring effect of larger dispersal scales.

89 In addition to scales of species interactions, we will address an additional major gap which
90 prevents a complete knowledge of scaling in ecology: consideration of a wider range of ecolog-
91 ical patterns within a single study than has been examined previously. Two well-recognized
92 ecological patterns are Species-Area (SAR) and Biodiversity-Ecosystem Functioning (BEF)
93 relationships. The Species-Area relationship is the earliest and most widely-examined ecolog-
94 ical pattern to explicitly consider scale [5, 23]. Although SARs have been described as
95 one of “ecology’s few universal regularities” [24], accumulating evidence reveals consider-
96 able variation within and among biological systems [25, 5, 26]. Likewise, BEF theory has
97 revealed consistent patterns, typically a saturating relationship between community diver-
98 sity and biomass production [27], but most work has focused on BEFs at local scales, with
99 only recent work highlighting the importance of scale [3]. Previous studies have examined
100 how one pattern or the other are affected by process scales [28, 26, 29], but no study has
101 examined how SAR and BEF relationships change in tandem and if effects that are masked
102 through one pattern are apparent in the other. As a consequence, it is unclear how both
103 SAR and BEF relationships are affected by the interplay of processes acting at different
104 scales, making it difficult to assess how process scales affect the overall behavior of ecosys-
105 tems as different measures highlight different aspects of ecosystems. Resolving these issues
106 will be useful for both basic and applied biodiversity problems, for instance allowing us to
107 scale up to landscape scales our predictions of biodiversity loss and its effect of ecosystem
108 productivity, that are often based on local scales [30].

109 Here, we use a modified Lotka-Volterra metacommunity model to explore the conse-
110 quences of the scaling of ecological processes for biodiversity, ecosystem functioning, and
111 their relationship across spatial scales. Our simulations consist of species interacting in
112 a spatially-explicit landscape, with “patches” emerging from the environmental structure
113 of the landscape. Although metacommunities tend to be modelled as systems of discrete
114 patches embedded within an inhospitable matrix, Chase and Leibold [31] describe this ap-
115 proach as useful (easing computation and interpretation) but limited – they foreshadow a
116 “coming” in ecology in favour of models that allow “patches” to emerge from the structure
117 of the environment, which our model achieves. We first study the heterogeneity of local
118 outcomes across the landscape: patterns of patch biodiversity, patch functioning, and rela-
119 tionships between them (local BEF). We can then scale up to the whole landscape scale and
120 every scale in between. By varying the spatial scales over which metacommunity processes
121 (abiotic environment, competitive interactions, and dispersal) play out, we test the hypoth-
122 esis that ecological patterns depend on how processes interact across scales, including scales
123 of species interactions, and lead to different patterns from those generated by commonly-
124 assumed hierarchical process scales (i.e., scales of interactions < environment < dispersal;

125 Fig. 1).

126 Species-Area relationships depend on spatial turnover in species composition, and com-
 127 positional turnover is driven by ecological processes [32]. Thus, we would expect that
 128 ecological processes should strengthen SARs in scenarios where they increase compositional
 129 turnover. We predict that the strongest slopes of the SAR will occur when scales of dis-
 130 persal < environment < species interactions, because (i) interactions are not constrained to
 131 abiotically suitable patches, and (ii) weaker dispersal prevents the homogenization of species
 132 composition across the landscape. Additionally, we predict that the consequences for BEF
 133 relationships will differ between local and regional scales. On local scales, we expect BEFs
 134 to weaken as interaction scales increase relative to the others, given that species that are
 135 locally absent but present in nearby areas can affect local functioning. On regional scales, we
 136 expect BEFs to strengthen as interaction scales increase, since regional competition would
 137 keep only the most suitable species at a given location. Hence, more species would mean
 138 that multiple species are productive within a given region.

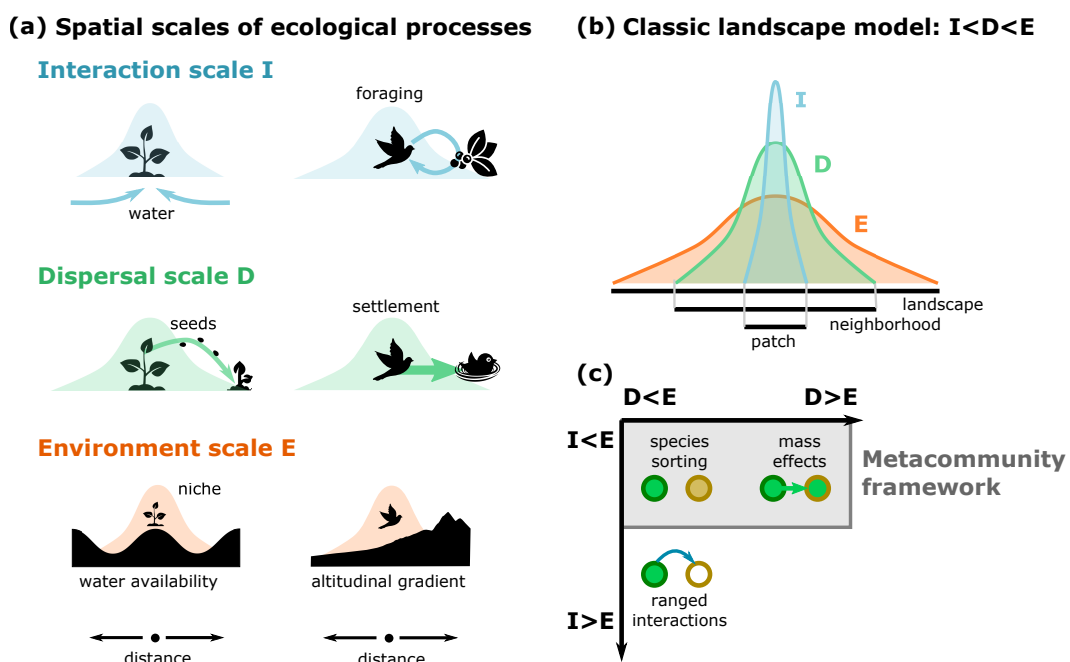


Figure 1: **Conceptual diagram of spatial scales of ecological processes.** (a) Illustration of the spatial scale of species interactions I , dispersal D and environmental heterogeneity E relative to the total size of the landscape (i.e., width of curves). (b) In the classic scenario, interactions take place within a patch, while dispersal is thought to act within a neighborhood and environmental factors vary broadly over the landscape. (c) Comparison of ecological scenarios along scales of I , D and E . Yellow and green represent two different species, with circle and its rim representing the resident species and the favoured species, respectively. Metacommunity theory has explored different scenarios for the relative scales of dispersal and environment (i.e., the ratio D/E), notably distinguishing “species sorting” (local environmental factors determine species distribution) and “mass effects” (population fluxes homogenize the landscape). Our work highlights the relative importance of species interactions scale (e.g., expressed through the ratio I/E , which was previously considered only in particular ecological settings (e.g., vegetation patterns or territoriality)). Ranged interactions may for instance induce exclusion of weaker competitors in a neighboring patch, even without a population flux of a stronger competitor into that patch.

139 Methods

140 Model

141 We use a modified Lotka-Volterra metacommunity model to explore the consequences of
 142 the spatial scaling of three ecological processes – abiotic environment, species interactions,
 143 and dispersal – for biodiversity and ecosystem functioning. Our specific assumptions and
 144 parameters are motivated by two important choices. First, we focus on a classic setting
 145 of ecological assembly, i.e., the patterns that arise when many species, originating from
 146 a regional pool, come together and reach an equilibrium state, with some species going
 147 locally or regionally extinct. Furthermore, we take species interactions in the pool to be
 148 disordered, that is, heterogeneous but without a particular functional group or trophic level
 149 structure [33]. We do not exclude that different patterns could emerge for more ordered
 150 interactions (e.g., a realistic food web) or for parameter values that lead to more complex
 151 dynamical regimes (e.g., population cycles or chaos, driven by stronger species interactions

152 or environmental perturbations). We note that our communities, in the chosen parameter
 153 regime of moderate competition, contain many species in a stable equilibrium (i.e., due to
 154 the assembly process). Our methodology thus differs from the extensive literature that has
 155 considered models with random interactions in order to study stability-complexity relation-
 156 ships [34], including more recent works in a spatial context [35, 36], as we rather focus on
 157 the abundance and diversity patterns arising from community assembly.

158 Second, we consider the possibility of species interacting over large spatial scales. Con-
 159 ventional metacommunity models describe discrete local communities of habitat patches
 160 connected by dispersal, within which species interact [37]. In doing so, they implicitly as-
 161 sume that the spatial range of species interaction is smaller than the scale of dispersal and
 162 contained within a patch, for all species and types of interactions [17]. To relax these as-
 163 sumptions, we construct a metacommunity model where populations of species can disperse
 164 and interact at different spatial scales, without specifying a mechanism underlying these eco-
 165 logical processes. Species interactions that manifest beyond local scales are abstracted from
 166 mechanisms such as individual foraging, vector species (e.g., pathogens) [38], and spatial
 167 resource fluxes [39, 17].

168 The model details the dynamics of S different species distributed across a spatially-
 169 explicit lattice landscape of 320x320 cells. The dynamical equation for the biomass N_i of
 170 species i at position \vec{x} in the landscape at time t is given by a generalized Lotka-Volterra
 171 equation of the form

$$\frac{\partial}{\partial t} N_i(\vec{x}, t) = N_i(\vec{x}, t) \left(r_i(\vec{x}) + \sum_j^S \int d\vec{y} A_{ij}(\vec{x}, \vec{y}) N_j(\vec{y}, t) \right) + \delta_i \Delta N_i(\vec{x}, t) \quad (1)$$

172 where \vec{x} and \vec{y} represent vectors of spatial (x, y) coordinates in the landscape. Equation
 173 (1) models the effects of three ecological processes on the biomass of species i : its intrinsic
 174 growth rate $r_i(\vec{x})$, which is influenced by abiotic environmental conditions at location \vec{x} ,
 175 dispersal to and from location \vec{x} , which is controlled by the diffusion coefficient δ_i , and
 176 interactions with all other species j , including when they are located elsewhere in the land-
 177 scape, $A_{ij}(\vec{x}, \vec{y})$. Although at face value cells in our model resemble patches in traditional
 178 metacommunity models, given that discrete populations are necessary to simulate Lotka-
 179 Volterra dynamics, here it is best to interpret cells as neighborhoods on a landscape. Each
 180 neighborhood may take on a unique environmental value and hold unique densities of in-
 181 dividuals of different species. Viewed in this way, landscape dynamics can be simulated
 182 more continuously, with the numerical limitation of needing to discretize dynamics at their
 183 finest resolution. While “patches” can emerge in autocorrelated environments (i.e., a spatial
 184 clustering of cells that are suitable to a given species), our model is also generalizable to
 185 landscapes with a diversity of environmental structures.

186 Environment

187 Abiotic conditions in each location are encoded by an environmental variable $V(\vec{x})$. This
 188 variable is continuous and varies smoothly over space, with parameters allowing one to tune
 189 the typical spatial scale of this variation [40]. For more details on the construction of the
 190 environment, see the Appendix section A2.

191 Each species has a Gaussian fundamental niche that determines its abiotic fitness in each
 192 location, with an optimal environmental value H_i and abiotic niche width ω_i

$$f_i(\vec{x}) = \exp \left[-\frac{(V(\vec{x}) - H_i)^2}{2\omega_i^2} \right] \quad (2)$$

193 Each fitness value is bound between 0 and 1 and reaches its maximum at an optimal envi-
 194 ronmental condition (i.e., when $V(\vec{x}) = H_i$). We take the growth rate as $r_i(\vec{x}) = f_i(\vec{x})$. In
 195 other words, $V(\vec{x})$ sets the actual structure of environmental conditions across the landscape,
 196 whereas $r_i(\vec{x})$ is how species experience the environment and its structure.

197 Interactions

198 We choose to limit ourselves to competitive interactions, defined by the matrix C_{ij} , which
 199 represents the per-capita competitive effect of species j on species i . The diagonal of the
 200 matrix (the impact of a species on itself) is set to 1, whereas all other interactions are
 201 taken independently from a random uniform distribution between 0 and \tilde{c} . We choose $\tilde{c} = 1$
 202 to allow for moderate interactions between different species (inter-specific competition is
 203 always weaker than intra-specific), suggesting that pairwise coexistence is often possible
 204 for species with different growth rates r_i , but the total impact of many competitors is
 205 still strong enough to allow for extinctions. Previous work has shown that, in disordered
 206 communities, the outcomes of ecological assembly are robust to many details such as the

207 nature of interactions (e.g., mutualism, predation), and depend only on a few statistical
 208 properties such as the mean and variance of interaction effects [33].

209 Furthermore, interactions are assumed to occur over a characteristic spatial scale encoded
 210 by a spatial kernel K . This scale may represent the distance an animal forages from its nest
 211 (without establishing a new nest), the scale at which trees gather resources with their roots,
 212 or the effective distance an immobile species interacts with its neighbors via an intermediary
 213 species (where the intermediary is not explicitly modeled). We use a Gaussian kernel whose
 214 standard deviation defines the interaction range such that

$$K(\vec{x}, \vec{y}|\gamma) = k_0 \exp\left[-\frac{\|\vec{x} - \vec{y}\|^2}{2\gamma^2}\right] \quad (3)$$

215 where $\|\vec{x} - \vec{y}\|$ indicates the norm of (distance between) the vectors \vec{x} and \vec{y} , and γ is the
 216 spatial range (scale) of the interactions. We note that while this modeling strategy is not
 217 physical as it implies that interactions occur instantaneously across distances, this is not
 218 expected to bias our results since we are focusing on the equilibrium state of the system,
 219 where hypothetical lag effects should be minimal.

220 We normalize the interactions by k_0 such that the overall effect of the kernel is always
 221 the same (i.e., the integral over K always equals 1). This normalization means that for large-
 222 scale interactions, local competition becomes weaker. However, some amount of (especially
 223 intra-specific) competition must remain locally strong to prevent species densities from
 224 growing exponentially and exploding. Therefore, we define interactions as partially local
 225 and partially regional, with β governing the fraction of interactions that are regional:

$$A_{ij}(\vec{x}, \vec{y}) = -\beta C_{ij} K(\vec{x}, \vec{y}|\gamma) - (1 - \beta) C_{ij}. \quad (4)$$

226 We choose β to ensure that the effect of interactions changes with their spatial scale (see
 227 scales subsection below), but local competition is never negligible (see more details in the
 228 Appendix, Fig. S12).

229 Dispersal

230 Finally, dispersal is modeled by the diffusion (Laplace) operator,

$$\delta_i \Delta N_i(\vec{x}, t), \quad (5)$$

231 where δ_i is the diffusion or dispersal coefficient of the species. For simplicity, we set the
 232 dispersal coefficient to be the same for all species.

233 Contrary to interactions, we do not use an explicit spatial kernel here, because intensity
 234 and spatial scale are unavoidably entangled in the case of dispersal (see Appendix section
 235 A1). The coefficient δ_i sets the spatial scale over which dispersal impacts ecological dynam-
 236 ics. Note that two aspects of our modeling choices mean that our choice of dispersal by
 237 diffusion is not qualitatively different from applying a large dispersal kernel: our focus on
 238 the equilibrium state, and having initial conditions where all species are introduced to every
 239 point in the landscape. The former aspect of equilibrium means that any potential non-
 240 equilibrium dynamics driven by species moving quickly across space due to a large dispersal
 241 kernels are not applicable. The latter aspect means that there is no limit to dispersal, i.e.,
 242 a short or long-ranged dispersal kernel does not affect which parts of the landscape can be
 243 reached by a species.

244 Scales

245 In this study we are concerned with spatial scales of three ecological processes:

- 246 1. E : environmental heterogeneity
- 247 2. D : dispersal
- 248 3. I : species interactions

249 To properly compare the interplay of different process scales, we must first compute their
 250 values for a given set of model parameters (Table 1). The scale of the environment com-
 251 bines two features often used in the literature to generate realistic, spatially-autocorrelated
 252 landscapes [41]: spectral color ρ , which indicates the relative importance of long-range and
 253 short-range variations in the environment, and spectral cutoff k_c , which indicates the finest
 254 grain of variation (Appendix section A2). The effective environmental scale E is controlled
 255 by these two parameters.

256 In the main text, we focus on a single value for the environment scale $E = 32$, and
 257 vary the other two scales on a logarithmic scale, with values of 1, 3.2, 10, 32 and 100,
 258 where the system itself has the scale (length) of 320 cells. Our distribution of I and D

Parameter	Interpretation	Baseline value [Range]
General		
S	species number	20
L	landscape size (cells) ($area = L^2$)	320
δ_i	dispersal coefficient	[0.01, 100]
Environment		
H_i	optimal environment value	\sim uniform(20, 80)
ω_i	abiotic niche width	\sim normal(10, 2)
ρ	spectral color	0.95
k_c	spectral cutoff	0.04
$K(\vec{x})$	local abiotic conditions	[0, 100]
k_0	normalization constant	-
Interactions		
\tilde{c}	max interaction strength	1.0
β	fraction of regional interactions	0.9
γ	spatial scale of interactions	[1, 100]
C_{ij}	interaction matrix	\sim uniform(0, \tilde{c})

Table 1: Parameters, default values and ranges.

259 are equally spaced along a log scale and allow us to have a clear separation between the
260 scales of each ecological process, while also being substantially smaller than the system
261 size (320 cells) and larger than the smallest scale in the system (1 cell). Details on the
262 construction of the environment are given in the Appendix section A2. We choose a value
263 of $E = 32$ specifically as it is the most straightforward to demonstrate our results (see
264 Appendix section A3 for other values). The scale of interactions is set by, and coincides
265 with, the width of the Gaussian kernel γ , such that $I = \gamma$. The scale of dispersal is mainly
266 determined by the diffusion coefficient δ_i , and it is expected to scale as $D \sim \sqrt{\delta_i}$ (see,
267 e.g., [42]). The normalization constant is, however, not trivial, and as we show in the
268 Appendix section A1, it is approximately 10. We therefore use: $D = 10\sqrt{\delta_i}$. Fixing the
269 environmental scale and varying the scale of interactions and dispersal allows us to isolate
270 the effects of interaction and dispersal scale without confounding the effects of different
271 landscape structures or differences between species.

272 Parameterization and simulations

273 To initialize our simulations, we first add environmental structure to a two-dimensional
274 landscape of size 320x320 cells (see the Appendix section A2 for details). We do not define
275 patches explicitly, but rather allow them to emerge from the spatial structure of the environ-
276 ment. We then seed $S = 20$ species onto the landscape, with initial biomass at each location
277 drawn from a uniform distribution between 0 and 1, resulting in roughly equal biomasses
278 at the landscape scale. For simplicity, we use periodic boundary conditions for the two-
279 dimensional system (i.e., a torus topology), for both dispersal and interactions. We do not
280 expect this choice to impact the results, due to the large size of the system considered.

281 We use 20 replicate landscapes, allowing environmental structure to vary among repli-
282 cates while keeping the environmental scale constant. Replicates with other values of envi-
283 ronmental scale are presented in the Appendix section A3. Each landscape replicate uses a
284 different set of species and their interactions, chosen at random. Each replicate landscape
285 was used to systematically vary the spatial scale of interactions I and dispersal coefficient
286 D , with 25 different combinations (5 values of D and 5 values of I , as given in Fig. 2), giving
287 a total of 500 simulations. We ascertain the generality of our findings by comparing across
288 replicates.

289 We run each simulation, where a simulation is defined as a model run with a unique
290 combination of process scales and replicate landscape, to a maximum time of $T = 1000$, or
291 until equilibrium is reached. For practical purposes, we define an equilibrium as when the
292 maximal change in biomass of any species in any location over a time-span of $T = 1$ is less
293 than 10^{-5} . A full list of parameter values can be found in Table 1. All simulations were
294 performed using MatLab 2019a.

295 Measurements

296 For each simulation we measure individual and total community biomass, species richness,
297 and sample the landscape to calculate Species-Area Relationships (SAR curves) as well as

298 Biodiversity-Ecosystem Functioning relationships (BEF curves). For species richness, SARs,
299 and BEFs, we define a species to be extinct at a given location if its biomass is below than
300 a threshold of 10^{-3} .

301 To calculate SAR curves, we sample at 40 different spatial scales from 1x1 (single cells)
302 to 320x320 (the entire landscape) on a logarithmic scale, and computed the species richness
303 at each. For a given scale, we randomly choose 100 locations in the landscape, and sampled
304 a region centered around the location chosen. We averaged over the 100 locations to obtain
305 the mean richness value for a given scale.

306 We calculate both local and regional BEF curves, based on random sampling of the
307 landscape. We do this in a similar way to the SAR curves, measuring species richness but
308 also total community biomass. For the local BEF, we use a 1x1 cell area with 102,400
309 random locations chosen, while for the regional BEF we use an intermediate area of size
310 10x10 with 1024 locations sampled. In this way the BEF measurement is done consistently
311 for different region sizes. For both local and regional BEF curves, we measure every cell on
312 average once.

313 A striking outcome observed in our results is that spatial patterns of biodiversity and
314 functioning in landscapes are not well captured by landscape summary measures, such as
315 SARs. To explain these patterns, we calculate how correlated the biomass is of a given
316 species as distance between sampling locations increases (i.e., spatial correlation), which can
317 be used to quantify the properties of spatial patterns we observe. To calculate species' spatial
318 correlations, we do the following: 1) we normalize the species' distribution by subtracting
319 its average biomass (taken over the whole system); 2) we obtain a correlation map by
320 calculating the convolution of a spatial distribution with itself, using a two-dimensional
321 Fast Fourier Transform; 3) we normalize the correlation map by dividing the resulting two-
322 dimensional map by its maximum value (i.e., we set a correlation value of 1 at the origin);
323 and 4) we define the one-dimensional correlation function as the average between a vertical
324 and horizontal transects through the correlation map. To define the scale of correlation
325 for a given species, we locate the distance at which the correlation function reaches half its
326 height, i.e., the distance from the origin where its value is the average of the maximum value
327 (which is always 1) and its minimal value (typically around 0). A step-by-step illustration
328 of calculating the spatial correlation is provided in the Appendix, Fig. S13.

329 Results

330 Local outcomes: functioning and diversity across localities

331 Our first major result is that, although they can arise from similar biological mechanisms
332 (e.g., individual mobility), dispersal and interaction scales have opposite impacts on biodi-
333 versity and functioning patterns across the landscape (Fig. 2 and S9). We start from the
334 case of weakly-connected communities with local interactions where all landscape patterns
335 result from environmental variation (top-left panel, Fig. 2). Increasing the spatial scale of
336 dispersal leads to a blurring of total community biomass over the landscape (from left to
337 right, Fig. 2). In contrast, increasing the scale of species interactions leads to a sharpening
338 of spatial patterns, amplifying underlying environmental heterogeneity (top to bottom, Fig.
339 2). The antagonism between these two effects can be seen by the fact that they counteract
340 each other when increasing both scales at once, leading to similar-looking outcomes (along
341 the diagonal, Fig. 2), but dispersal eventually wins out – the states along the right column
342 are virtually identical, whereas the same is not true across the bottom row. Critically, it
343 is not until the scales of dispersal or interactions exceed the scale of environmental hetero-
344 geneity (i.e., outside the dashed-lined boundary in Fig. 2) that the scale of either process
345 significantly alters spatial patterns in biomass (see also Fig. S4). Larger emergent scales
346 of total community biomass due to high D , and the opposite due to high I , can also be
347 seen in Fig. 5, which shows how quickly patterns among locations become dissimilar as the
348 distances between them increase.

349 We then focus on a subset of our scenarios above to show how process scales impact not
350 only total biomass but also individual species distributions (Fig. 3). We observe that in-
351 creasing dispersal scale predictably makes larger, more coherent domains (i.e., fairly defined
352 areas with similar characteristics) with typically higher local diversity. Increasing interaction
353 scale creates a more granular landscape with a broader range of diversities, including many
354 low-diversity patches and a few high-diversity ones. Indeed, large interaction scales lead to
355 more spotty species distributions, with rare species persisting in some locations where they
356 would not in other scenarios (Fig. 3 bottom row). Two notable examples include species 1
357 (red patches) persisting only when interactions are large and dispersal is small, and species
358 2 (individually green, but here cyan due to its coexistence with species 3, blue) taking on a
359 more clumped distribution with large interaction scales.

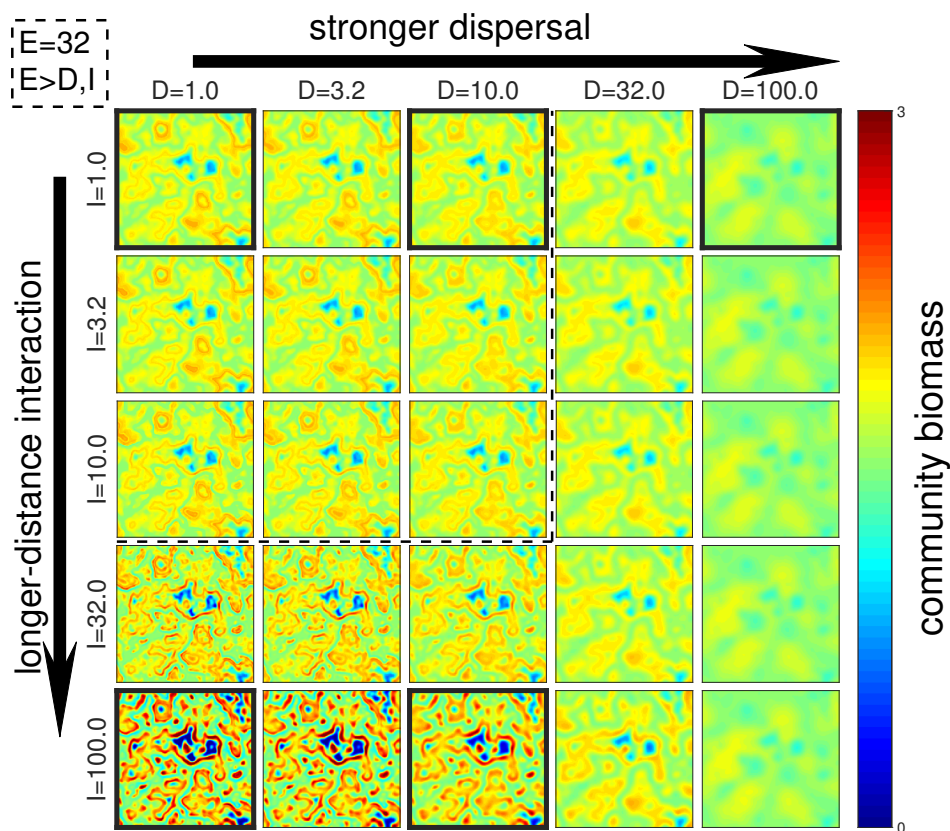


Figure 2: **Distribution of total community biomass** across the landscape as we change dispersal D (columns) and interaction I (rows) scales. Dashed black line shows where the environment scale $E = 32$ is larger than both D and I . Black frames around panels designate parameter values that we further examine in other figures. For better legibility, biomass levels above 3 are not shown.

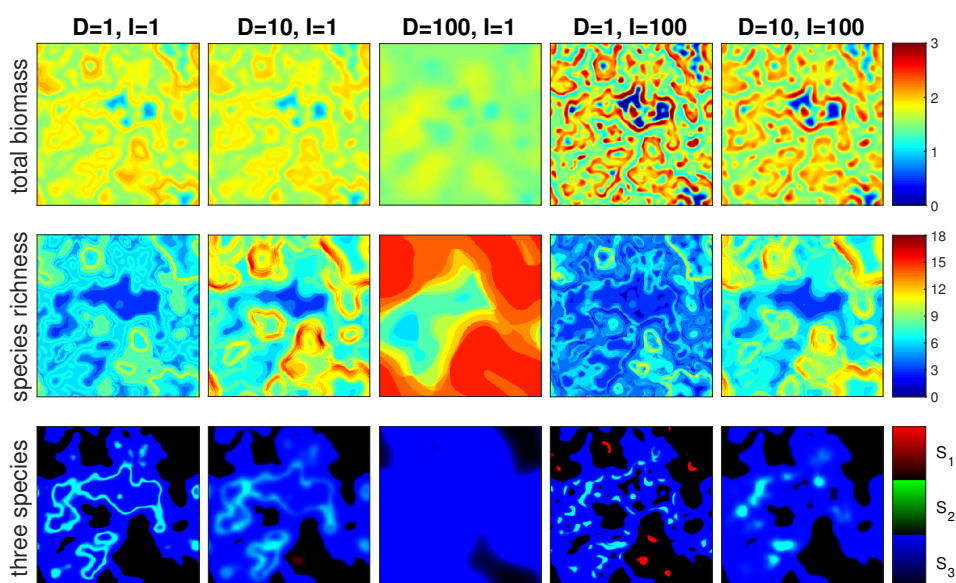


Figure 3: **Species distribution patterns** for five selected parameter sets, representing different scales of dispersal (D) and interaction (I), as designated in Fig. 2. **Top row:** total community biomass. **Middle row:** local species richness. **Bottom row:** distribution of three of the 20 species in original species pool (their biomass are encoded in the red, green and blue color channels, respectively; thus, cyan regions corresponds to coexistence of species 2 and 3). For better legibility, biomass levels above 3 are not shown.

360 Regional outcomes: functioning and diversity at the landscape scale

361 The outcomes described above allow us to identify spatial patterns in local outcomes in
 362 the landscape, but what are outcomes for the landscape as a whole? Given the additive
 363 nature of biomass across localities, two regions could have identical biomass at the land-
 364 scape scale even if one region has high variation among localities that span extremes of
 365 high and low values, whereas another varies little with biomass values that are interme-
 366 diate. Here, we see that biomass is highest when interaction scales are large (Fig. S10), an

367 effect that is quickly eroded as dispersal scales increase. Interestingly, these high-biomass
 368 landscapes had extreme variation in biomass among localities, including areas of extremely
 369 low biomass (dark blue in Fig. 2) and extremely high biomass (red in Fig. 2). Therefore,
 370 high biomass is driven by a disproportionate subset of local communities in a landscape.
 371 Furthermore, these high biomass landscapes were unremarkable in regional species richness
 372 in the landscape and actually had fewer species per locality on average than other scenarios
 373 (Fig. S11). For those who may be interested in comparing our findings to those typically
 374 reported in traditional metacommunity models more explicitly (e.g., [43]), we note that the
 375 left and right plots in Fig. S11 essentially show local (i.e., alpha) and regional (i.e., gamma)
 376 diversity, respectively, whereas compositional turnover among localities (i.e., beta diversity)
 377 is essentially differences between them.

378 Cross-scale outcomes: BEF and SAR

379 Next, we turn to two types of cross-scale outcomes (Fig. 4). First, we consider the relation-
 380 ship in BEF curves (i.e., total biomass vs. species diversity) at neighborhood (i.e., single
 381 cell) scales. In doing so, we find that BEF curves (Fig. 4, left panel) reflect underlying
 382 process scales. In particular, they exhibit a hump-shaped relationship for large interaction
 383 scales, suggesting that patches with the largest total biomass are not the most diverse, but
 384 rather have a few high-performing species. This result ties into our previous observation that
 385 the interaction scale tends to amplify environmental heterogeneity, and may thus put more
 386 weight on selection effects, where abiotic conditions select the best-performing species at the
 387 exclusion of others. We also examined BEF curves measured at larger scales, i.e., when spa-
 388 tially aggregating 100-cell neighborhoods, and found qualitatively identical patterns (Fig. 4,
 389 middle panel).

390 We also look at a pattern aggregated over continuously increasing spatial scales – the
 391 SAR (Fig. 4, right panel). We would expect that changes in the slope or shape of the SAR
 392 as the aggregation scale (x-axis) exceeds the spatial scales of our ecological processes, as has
 393 been demonstrated for the Stability-Area Relationships [8]. However, we do not observe a
 394 clear link between process and pattern scales, beyond the fact that the inflection point (in
 395 particular, for low D and I) corresponds to the environmental scale E (vertical gray line in
 396 Fig. 3). The main impact of process scale is that, by amplifying landscape heterogeneity, a
 397 large interaction scale I leads to a stronger SAR when large interaction scales are coupled
 398 with short dispersal scales. Specifically, as predicted, at the smallest scale the $D < E < I$
 399 scenario (magenta curve) yields the lowest species richness compared to all other scenarios,
 400 whereas at the scale of the entire landscape, its richness is very high.

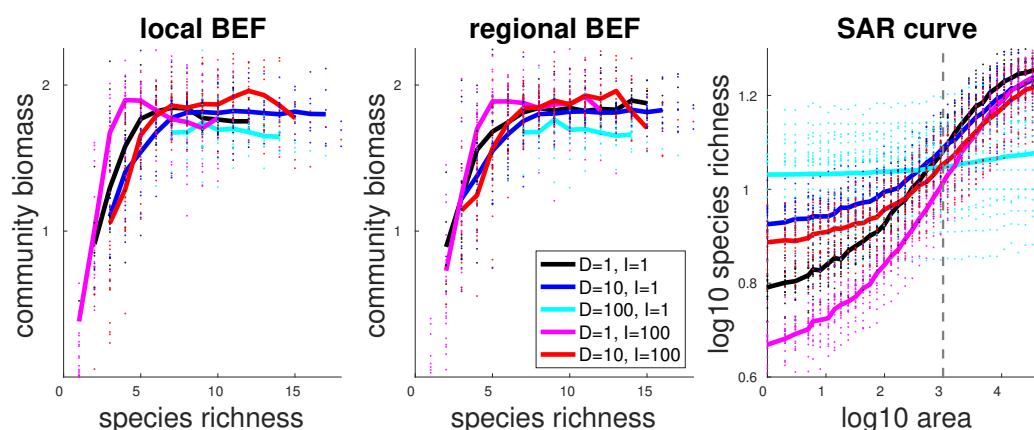


Figure 4: **BEF and SAR relationships.** Solid lines show average values over 20 replicates, small circles show values for individual replicates. Colors correspond to five selected parameter sets, representing different scales of dispersal (D) and interaction (I), as designated in Fig. 2. Local and regional BEF curves are measured at regions of size 1 and 100, respectively. Vertical gray line shows the area corresponding to the environmental scale $E = 32$. Although our model is deterministic (i.e., each replicate has only one possible outcome, given a specific set of parameter values and initial conditions), differences among replicates reflect differences in parameter values caused by sampling those values from distributions (Table 1).

401 Aggregated measures of biodiversity and functioning at regional scales miss much of the
 402 information captured by local measures, such as the distribution and turnover in biomass
 403 (Fig. 2 and Fig. 3). Yet these local patterns can be quantified. Figure 5 presents the
 404 results of the spatial correlation of species biomass distributions, which measures how the
 405 biomass of a species correlates over the distance between sampling. We observe clear trends
 406 in scale, with consistent patterns of growing (shrinking) correlation with higher dispersal
 407 (interaction) scales.

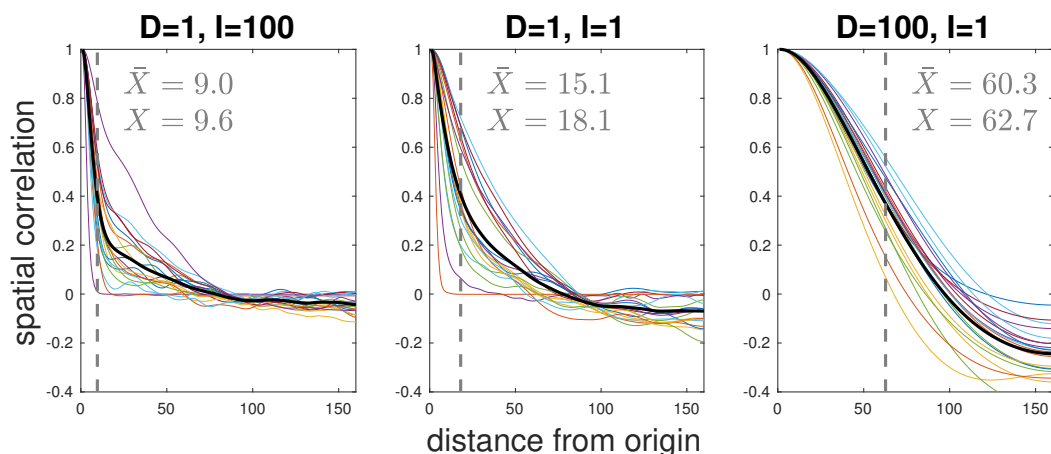


Figure 5: **Spatial correlation of each species’s biomass distribution**, for three scenarios. **Left:** $I = 100, D = 1$; **Middle:** $I = 1, D = 1$; **Right:** $I = 1, D = 100$. Recall that $E = 32$. Each of the 20 species is represented by a different color, with black showing the average correlation function, all for a single replicate. For this simulation run, the scale of correlation X is given, and is shown by gray vertical lines. The correlation scale averaged over the 20 replicates, \bar{X} , is also noted.

408 Discussion

409 This study focuses on a critical question: how is the scaling of ecological patterns, such as
 410 patterns of biodiversity and ecosystem functioning, related to scales of specific processes,
 411 and why? We have modelled how intrinsic scales of ecological processes align with the
 412 emergence of ecological patterns in a metacommunity, where we control the spatial scale
 413 of environmental heterogeneity, dispersal, and species interactions. In doing so, below, we
 414 highlight the following three take-home messages of our results:

- 415 • the scale of one process (here, environment) can cause the emergence of characteristic
 416 scales of other processes (dispersal, interactions)
- 417 • two interlinked ecological patterns (biodiversity and ecosystem function) and their
 418 relationship to each other are oppositely affected by two forms of organismal movement
- 419 • averaging ecological patterns at any one scale misses a rich patterning of spatial vari-
 420 ance that is closely tied to process scales

421 Below, we expand upon each finding and place them within existing knowledge, examine
 422 the mechanisms that underlie our findings, contrast results among ecological variables, and
 423 end by placing our results within a context of ecosystem preservation.

424 A main finding of our study is that the spatial scale of interactions amplifies environ-
 425 mental heterogeneity, sharpening observed spatial patterns, in contrast to dispersal scales.
 426 Importantly, observed spatial patterns did not reflect the absolute value of the spatial scale
 427 of each ecological process, but rather, their values relative to the environment; decreasing
 428 the spatial scale of the environment shifts the boundary of blurring/sharpening effects of
 429 dispersal and species interactions (Fig. S4). We find this effect because environmental con-
 430 ditions are quite literally the template upon which dispersal and species interactions mold
 431 species distribution. Large-scale (i.e., at scales above the template) processes are more
 432 important than small-scale ones in determining overall patterns, meaning that only when
 433 dispersal or interactions have large scales can they impact large-scale patterns.

434 We examined the impacts of process scales on two classes of patterns: first, on the spatial
 435 scaling of patterns (SAR and BEF), and second, on the spatial structure of species biomass
 436 in the landscape. Unexpectedly, the latter class of patterns appears to better reflect the scale
 437 of ecological processes, such as the distribution and turnover of biomass and biodiversity
 438 across the landscape. These patterns would be lost by examining mean biodiversity and
 439 function at specific aggregation scales (e.g., local vs. regional; Fig. S4), but were well
 440 captured via spatial autocorrelation (Fig. 5). From these analyses, one take-home message
 441 is that increasing the scale of species interactions actually amplifies variation on small scales.
 442 In other words, large-scale processes do not necessarily beget large-scale patterns.

443 The question of how process scales affect observed patterns can also be spun around:
 444 what information about process scales can be inferred from the various patterns we see?
 445 Considering the opposing effects that dispersal and interaction scales have on pattern scales
 446 (Fig. 2), it is not clear that such an inference is possible. However, given that patterns scales
 447 change differently (compare Fig. 2 with Fig. S3, for instance), combining several measures
 448 together may provide an answer, for instance by finding when changes in spatial correlations
 449 of biodiversity and biomass no longer behave similarly. In this context, it is perhaps to be

450 expected that no clear connection was found between well known patterns such as BEF
451 and SARs, and process scales. Over the past few decades, ecologists have been cautioned
452 from inferring processes from patterns [44]. Our results demonstrate exactly why this is
453 important: a lack of a 1:1 mapping between a pattern and any one specific process.

454 Indeed, our finding that the SAR curves did not exhibit transitions at particular spatial
455 scales, that would allow us to identify the typical scales of the underlying processes (other
456 than the environment), runs counter to other contexts, such as the invariability-area rela-
457 tionship [8]. In particular, we do not find a triphasic SAR curve that is often reported [45, 8].
458 This is the case since our model does not consider individual sampling and dispersal limi-
459 tation, which typically lead to stronger SAR slopes at small and large scales, respectively.
460 We thus see the strongest slopes at intermediate spatial scales, consistent with results under
461 similar settings [46], and hinting that we are largely seeing community dynamics typical of
462 species-sorting [37]. Centering on the average SAR slope itself, on the one hand, we found
463 that large interaction scales may enhance the SAR by amplifying landscape heterogeneity
464 and creating low-diversity strips along the edges of species ranges. On the other hand, this
465 spatial heterogeneity could also promote coexistence as a weaker competitor might thrive
466 in the margins [47]. This suggests that edge effects may play a prevalent role in the case
467 of long-range interactions, and deserves more extensive investigation. Overall, the scales of
468 biotic processes (interaction and dispersal) are mainly reflected inasmuch as they change
469 overall community properties, such as total diversity across the landscape.

470 Knowledge of the spatial scale of ecological processes is critical to understanding the
471 maintenance of ecosystems. To illustrate this argument, one can imagine a landscape man-
472 ager interested in preserving some baseline level of functioning in a landscape at a specific
473 spatial extent, for example, primary production. If the spatial scale of interest does not en-
474 compass the intrinsic scales of processes that govern functioning, then landscape alteration
475 beyond that scale might impact functioning in an unanticipated and undesirable manner;
476 these scales will differ among ecosystems based on how species' traits and the physical
477 landscape affect how organisms experience scales of E, D, and I. In other words, the scales
478 important to the maintenance of ecosystem function may be mismatched from the (typ-
479 ically small) spatial scales at which ecosystem functioning is observed and managed, but
480 the degree to which this is true depends on process scaling. Predictions of our model could
481 be best tested empirically in microcosm or mesocosm setups or using data syntheses, for
482 example, by examining the spatial structure of species richness and biomass depending on
483 process scales of focal taxa (e.g., small vs large-bodied animals using remotely sensed data,
484 experiments with insects where mobility is restricted).

485 Our results suggest that it will be difficult to manage landscapes to preserve biodiver-
486 sity and ecosystem functioning simultaneously, despite their causative relationship, for two
487 related reasons. First, the fact that increasing dispersal and interaction scales had opposing
488 effects on either ecosystem property presents a unique management challenge, given that
489 both scales are tied to organismal movement, albeit on distinct timescales (i.e., daily vs.
490 once-per-generation). Second, ecosystems attained the highest biomass in scenarios which
491 also led to the lowest levels of biodiversity, specifically, when interaction scales were large
492 and dispersal scales were small. We note that this second issue may only be relevant when
493 interactions are largely competitive, since our modeling, and thus results, did not consider
494 mutualistic interactions which would likely change the observed trade-off between biodiver-
495 sity and biomass. How would a manager plan a landscape to enhance interaction scales
496 (preserving function) while simultaneously minimizing scales of dispersal (preserving biodi-
497 versity)? This can, for instance, be relevant for managing predation of pest herbivores in
498 agricultural landscapes [16]. This type of intervention might be most successful in species
499 with body plans for long-distance movement, but that can remain philopatric for behavioural
500 reasons (which can be environmentally determined; i.e., territorial hunters).

501 Our metacommunity model differs from traditional metacommunity models in several
502 important ways. Traditional metacommunity models tend to include discrete local patches
503 embedded within an implicit inhospitable matrix, interconnected by rates of dispersal, often
504 from a spatially-implicit regional pool of dispersers. By contrast, "patches" in our model
505 emerge from the environmental template (Fig. 3), the structure of which may be viewed
506 differently by different species according to their fundamental niche. Further, these patches
507 may have fuzzy boundaries, within-patch heterogeneity, as well as different shapes and sizes.
508 Individuals may be lost to the matrix (i.e., habitat falling outside of the fundamental niche)
509 if they disperse there or may form stepping stone populations to reach new patches. In
510 doing so, dispersal limitation is more likely to emerge as the spatial grain of the environment
511 exceeds the scales at which species disperse, a major result of our study. These features align
512 with the recent calls [48, 31] to develop more realistic metacommunity models applicable to a
513 wider range of systems, beyond discrete, patchy, island-like systems. Given these strengths,
514 the next step is to extend a model like ours to multi-trophic systems, beyond "horizontal"
515 (sensu Vellend [49]) competitive communities. Our model is naturally amenable to multi-

516 trophic systems, as predators often perceive the landscape at a different scale than their prey
517 (i.e., a different interaction scale) and would perceive the scale of the environment via spatial
518 distributions of their prey—additionally, there is an opportunity to move beyond Lotka-
519 Volterra dynamics for modelling species interactions, towards more mechanistic consumer-
520 resource approaches [50]. Most metacommunity models have been applied to competing
521 species [17], with multi-trophic extensions becoming more common in recent years [51].

522 Our conclusions are twofold. First, we bring forward an important spatial scale – the
523 range of species interactions – that has been largely neglected in previous analyses (e.g.,
524 metacommunity theory). This interaction range can be derived from many of the same
525 ecological mechanisms as dispersal, such as individual mobility, yet these two processes
526 lead to opposite ecological effects. This suggests that we must carefully distinguish whether
527 mobility actually leads to population dispersal or to large-range interactions, and re-evaluate
528 possible consequences of evolution or environmental change in these processes. Finally, we
529 saw that the spatial scale of ecological processes might not appear clearly in the scale of
530 resulting patterns such as Species-Area or Biodiversity-Ecosystem Functioning relationships,
531 though they may sometimes be reflected in local outcomes. While we focused on a few
532 important biodiversity and functioning patterns, our study paves the way for future work
533 investigating systematically under which conditions various ecological pattern scales may or
534 may not reflect the spatial scale of underlying processes.

535 References

- 536 [1] L. Estes, P. R. Elsen, T. Treuer, L. Ahmed, K. Caylor, J. Chang, J. J. Choi, and
537 E. C. Ellis, “The spatial and temporal domains of modern ecology,” *Nature ecology &*
538 *evolution*, vol. 2, no. 5, p. 819, 2018.
- 539 [2] K. Henle, S. Potts, W. Kunin, Y. Matsinos, J. Simila, J. Pantis, V. Grobelnik, L. Penev,
540 and J. Settele, “Scaling in ecology and biodiversity conservation,” *Advanced Books*,
541 vol. 1, p. e1169, 2014.
- 542 [3] A. Gonzalez, R. M. Germain, D. S. Srivastava, E. Filotas, L. E. Dee, D. Gravel, P. L.
543 Thompson, F. Isbell, S. Wang, S. Kéfi, *et al.*, “Scaling-up biodiversity-ecosystem func-
544 tioning research,” *Ecology Letters*, vol. 23, no. 4, pp. 757–776, 2020.
- 545 [4] J. A. Wiens, “Spatial scaling in ecology,” *Functional ecology*, vol. 3, no. 4, pp. 385–397,
546 1989.
- 547 [5] M. V. Lomolino, “Ecology’s most general, yet protean 1 pattern: the species-area
548 relationship,” *Journal of Biogeography*, vol. 27, no. 1, pp. 17–26, 2000.
- 549 [6] D. L. Urban, “Modeling ecological processes across scales,” *Ecology*, vol. 86, no. 8,
550 pp. 1996–2006, 2005.
- 551 [7] S. P. Hart, J. Usinowicz, and J. M. Levine, “The spatial scales of species coexistence,”
552 *Nature Ecology & Evolution*, vol. 1, no. 8, pp. 1066–1073, 2017.
- 553 [8] R. Delsol, M. Loreau, and B. Haegeman, “The relationship between the spatial scaling
554 of biodiversity and ecosystem stability,” *Global Ecology and Biogeography*, vol. 27, no. 4,
555 pp. 439–449, 2018.
- 556 [9] C. P. Catano, E. Grman, E. Behrens, and L. A. Brudvig, “Species pool size al-
557 ters species-area relationships during experimental community assembly,” *Ecology*,
558 pp. e03231–e03231, 2020.
- 559 [10] N. J. Gotelli, *A primer of ecology*. Sinauer Associates, 1995.
- 560 [11] I. Hanski, *Metapopulation ecology*. Oxford University Press, 1999.
- 561 [12] M. Loreau, *From populations to ecosystems: Theoretical foundations for a new ecological*
562 *synthesis*. Princeton University Press, 2010.
- 563 [13] T. Poisot, E. Canard, D. Mouillot, N. Mouquet, and D. Gravel, “The dissimilarity of
564 species interaction networks,” *Ecology letters*, vol. 15, no. 12, pp. 1353–1361, 2012.
- 565 [14] S. Pilosof, M. A. Porter, M. Pascual, and S. Kéfi, “The multilayer nature of ecological
566 networks,” *Nature Ecology & Evolution*, vol. 1, no. 4, p. 0101, 2017.
- 567 [15] P. Fauchald, “Spatial interaction between seabirds and prey: review and synthesis,”
568 *Marine Ecology Progress Series*, vol. 391, pp. 139–151, 2009.

- 569 [16] T. A. Rand, J. M. Tylianakis, and T. Tscharntke, “Spillover edge effects: the disper-
570 sal of agriculturally subsidized insect natural enemies into adjacent natural habitats,”
571 *Ecology letters*, vol. 9, no. 5, pp. 603–614, 2006.
- 572 [17] M. L. Guzman, R. M. Germain, C. Forbes, S. Straus, M. I. O. OConnor, Gravel,
573 D. S. Srivastava, and T. P. L., “Towards a multi-trophic extension of metacommunity
574 ecology,” *Ecology Letters*, vol. 1, no. 1, pp. 1–1, 2019.
- 575 [18] R. Martinez-Garcia, C. H. Fleming, R. Seppelt, W. F. Fagan, and J. M. Calabrese,
576 “How range residency and long-range perception change encounter rates,” *Journal of*
577 *theoretical biology*, vol. 498, p. 110267, 2020.
- 578 [19] P. L. Thompson, B. Rayfield, and A. Gonzalez, “Loss of habitat and connectivity erodes
579 species diversity, ecosystem functioning, and stability in metacommunity networks,”
580 *Ecography*, vol. 40, no. 1, pp. 98–108, 2017.
- 581 [20] P. L. Thompson, L. M. Guzman, L. De Meester, Z. Horváth, R. Ptacnik, B. Vanschoen-
582 winkel, D. S. Viana, and J. M. Chase, “A process-based metacommunity framework
583 linking local and regional scale community ecology,” *Ecology Letters*, vol. 23, no. 9,
584 pp. 1314–1329, 2020.
- 585 [21] A. M. de Roos, E. McCauley, and W. G. Wilson, “Pattern formation and the spatial
586 scale of interaction between predators and their prey,” *Theoretical population biology*,
587 vol. 53, no. 2, pp. 108–130, 1998.
- 588 [22] R. E. Snyder and P. Chesson, “How the spatial scales of dispersal, competition, and
589 environmental heterogeneity interact to affect coexistence,” *The American Naturalist*,
590 vol. 164, no. 5, pp. 633–650, 2004.
- 591 [23] O. Arrhenius, “Species and area,” *Journal of Ecology*, vol. 9, no. 1, pp. 95–99, 1921.
- 592 [24] T. Schoener, “Patterns in terrestrial vertebrate versus arthropod communities: do sys-
593 tematic differences in regularity exist?,” *Community ecology*, pp. 556–586, 1986.
- 594 [25] S. Drakare, J. J. Lennon, and H. Hillebrand, “The imprint of the geographical, evo-
595 lutionary and ecological context on species–area relationships,” *Ecology letters*, vol. 9,
596 no. 2, pp. 215–227, 2006.
- 597 [26] M. Franzén, O. Schweiger, and P.-E. Betzholtz, “Species-area relationships are con-
598 trolled by species traits,” *PloS one*, vol. 7, no. 5, p. e37359, 2012.
- 599 [27] B. J. Cardinale, K. L. Matulich, D. U. Hooper, J. E. Byrnes, E. Duffy, L. Gamfeldt,
600 P. Balvanera, M. I. O’connor, and A. Gonzalez, “The functional role of producer diver-
601 sity in ecosystems,” *American journal of botany*, vol. 98, no. 3, pp. 572–592, 2011.
- 602 [28] S. M. Scheiner, A. Chiarucci, G. A. Fox, M. R. Helmus, D. J. McGlenn, and M. R.
603 Willig, “The underpinnings of the relationship of species richness with space and time,”
604 *Ecological Monographs*, vol. 81, no. 2, pp. 195–213, 2011.
- 605 [29] E. Ben-Hur and R. Kadmon, “Disentangling the mechanisms underlying the species–
606 area relationship: A mesocosm experiment with annual plants,” *Journal of Ecology*,
607 vol. 108, no. 6, pp. 2376–2389, 2020.
- 608 [30] J. M. Chase, B. J. McGill, D. J. McGlenn, F. May, S. A. Blowes, X. Xiao, T. M. Knight,
609 O. Purschke, and N. J. Gotelli, “Embracing scale-dependence to achieve a deeper un-
610 derstanding of biodiversity and its change across communities,” *Ecology letters*, vol. 21,
611 no. 11, pp. 1737–1751, 2018.
- 612 [31] M. A. LEIBOLD, J. M. CHASE, S. A. LEVIN, and H. S. HORN, *Metacommunity*
613 *Ecology, Volume 59*. Princeton University Press, 2018.
- 614 [32] A. Shmida and M. V. Wilson, “Biological determinants of species diversity,” *Journal*
615 *of biogeography*, pp. 1–20, 1985.
- 616 [33] M. Barbier, J.-F. Arnoldi, G. Bunin, and M. Loreau, “Generic assembly patterns in
617 complex ecological communities,” *Proceedings of the National Academy of Sciences*,
618 vol. 115, pp. 2156–2161, 2018.
- 619 [34] R. M. May, “Will a large complex system be stable?,” *Nature*, vol. 238, pp. 413–414,
620 1972.
- 621 [35] D. Gravel, F. Massol, and M. A. Leibold, “Stability and complexity in model meta-
622 ecosystems,” *Nature communications*, vol. 7, no. 1, p. 12457, 2016.

- 623 [36] J. W. Baron and T. Galla, “Dispersal-induced instability in complex ecosystems,” *Nature communications*, vol. 11, no. 1, p. 6032, 2020.
624
- 625 [37] M. A. Leibold, M. Holyoak, N. Mouquet, P. Amarasekare, J. M. Chase, M. F. Hoopes,
626 R. D. Holt, J. B. Shurin, R. Law, D. Tilman, *et al.*, “The metacommunity concept: a
627 framework for multi-scale community ecology,” *Ecology letters*, vol. 7, no. 7, pp. 601–
628 613, 2004.
- 629 [38] E. W. Schupp, “The janzen-connell model for tropical tree diversity: population impli-
630 cations and the importance of spatial scale,” *The American Naturalist*, vol. 140, no. 3,
631 pp. 526–530, 1992.
- 632 [39] J. Chelsea *et al.*, “Meta-ecosystems 2.0: rooting the theory into the field,” *Trends in*
633 *ecology & evolution*, 2017.
- 634 [40] D. A. Vasseur and P. Yodzis, “The color of environmental noise,” *Ecology*, vol. 85,
635 pp. 1146–1152, 2004.
- 636 [41] P. L. Thompson, S. Kéfi, Y. R. Zelnik, L. E. Dee, S. Wang, C. de Mazancourt,
637 M. Loreau, and A. Gonzalez, “Scaling up biodiversity–ecosystem functioning relation-
638 ships: the role of environmental heterogeneity in space and time,” *Proceedings of the*
639 *Royal Society B*, vol. 288, no. 1946, p. 20202779, 2021.
- 640 [42] Y. R. Zelnik, J.-F. Arnoldi, and M. Loreau, “The three regimes of spatial recovery,”
641 *Ecology*, vol. 100, no. 2, p. e02586, 2019.
- 642 [43] N. Mouquet and M. Loreau, “Community patterns in source-sink metacommunities,”
643 *The american naturalist*, vol. 162, no. 5, pp. 544–557, 2003.
- 644 [44] E. J. McIntire and A. Fajardo, “Beyond description: the active and effective way to
645 infer processes from spatial patterns,” *Ecology*, vol. 90, no. 1, pp. 46–56, 2009.
- 646 [45] M. L. Rosenzweig *et al.*, *Species diversity in space and time*. Cambridge University
647 Press, 1995.
- 648 [46] M. J. Crawley and J. Hurrall, “Scale dependence in plant biodiversity,” *Science*, vol. 291,
649 no. 5505, pp. 864–868, 2001.
- 650 [47] G. A. Maciel and R. Martinez-Garcia, “Enhanced species coexistence in lotka-volterra
651 competition models due to nonlocal interactions,” *Journal of Theoretical Biology*,
652 vol. 530, p. 110872, 2021.
- 653 [48] J. B. Logue, N. Mouquet, H. Peter, and H. Hillebrand, “Empirical approaches to meta-
654 communities: a review and comparison with theory,” *Trends in ecology & evolution*,
655 vol. 26, no. 9, pp. 482–491, 2011.
- 656 [49] M. Vellend, “The theory of ecological communities (mpb-57),” in *The Theory of Eco-*
657 *logical Communities (MPB-57)*, Princeton University Press, 2016.
- 658 [50] K. D. Lafferty, G. DeLeo, C. J. Briggs, A. P. Dobson, T. Gross, and A. M. Kuris, “A
659 general consumer-resource population model,” *Science*, vol. 349, no. 6250, pp. 854–857,
660 2015.
- 661 [51] P. L. Thompson and A. Gonzalez, “Dispersal governs the reorganization of ecological
662 networks under environmental change,” *Nature Ecology & Evolution*, vol. 1, no. 6,
663 p. 0162, 2017.
- 664 [52] M. H. Mohd, R. Murray, M. J. Plank, and W. Godsoe, “Effects of different dispersal
665 patterns on the presence-absence of multiple species,” *Communications in Nonlinear*
666 *Science and Numerical Simulation*, vol. 56, pp. 115–130, 2018.

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673 Conflict of interest disclosure

674 The authors declare they have no conflict of interest relating to the content of this article.

⁶⁷⁵ **Data accessibility**

⁶⁷⁶ Script files for simulations and analysis of results shown in the manuscript are available at
⁶⁷⁷ the open-access repository: <https://doi.org/10.5281/zenodo.5543190>.

678 Appendix

679 This appendix is made of four sections. A1: Measurement of scales; A2: Generating the
680 landscape; A3: Different environmental scales; A4: Additional plots.

681 A1 Measurement of scales

682 As explained in the main Methods section, we explicitly measure and compare three spatial
683 scales: environmental conditions (E), dispersal (D) and species interaction (I). We now
684 detail the definition of these three scales, and finally note the peculiarity of dispersal scale.

685 **Environmental scale E:** The environment itself is generated using a combination of a
686 spectral color and cutoff wavenumber (see next section), but this does not explicitly define
687 the scale. To measure the scale of the environment, we follow the same approach as for
688 the correlation function and measure the scale for a species biomass distribution (using a
689 convolution based on FFT), except that we do this for the value of intrinsic growth rate
690 $r_i(\vec{x})$, as it is directly set by the environment. For each of the 20 species, we can calculate
691 a correlation function (in the same manner as explained in the methods), and from this
692 we calculate the correlation scale (the point of middle height for the correlation function).
693 We average this value over all 20 species, to calculate the environment's scale for a given
694 system. Since this result depend on the randomization of the environment, we repeat this
695 for many replicates, and choose values of ρ and k_c that will on average give a value of E
696 we want to have.

697 **Dispersal scale D:** To estimate the dispersal scale D , we compare the effect of changing
698 the dispersal coefficient δ with changing γ . In Fig. S1 we show how changing δ and γ (and
699 thereby D and I) affects the community biomass distribution. As seen in the left panel, with
700 low δ and γ the difference from a null scenario of no dispersal and no interaction distance
701 is very small, but increasing either δ or γ changes the community biomass distribution
702 considerably. In the middle and right panels we see these differences, as we change only
703 δ (middle) or only γ (right). This clearly shows three things: 1) The effect of interaction
704 distance scales linearly with γ , as expected by its definition. 2) The effect of dispersal
705 coefficient scales with $\sqrt{\delta}$, as expected from dimensional considerations (e.g., [42]). 3) More
706 specifically, to make these two effects comparable, the dispersal scale is missing a factor of
707 10, i.e., $D = 10\sqrt{\delta}$. This can be seen by the fact that for both $\delta = 1$ in the middle panel
708 and $\gamma = 10$ in the right panel, the y-axis values are roughly the same ($10^{-1.2}$).

709 **Interaction scale I:** In our model, the species interactions are explicitly defined with
710 a distance over which they occur – via the Gaussian kernel function. This naturally gives
711 us the scale of interactions I , as the width of the Gaussian function, such that $I = \gamma$.

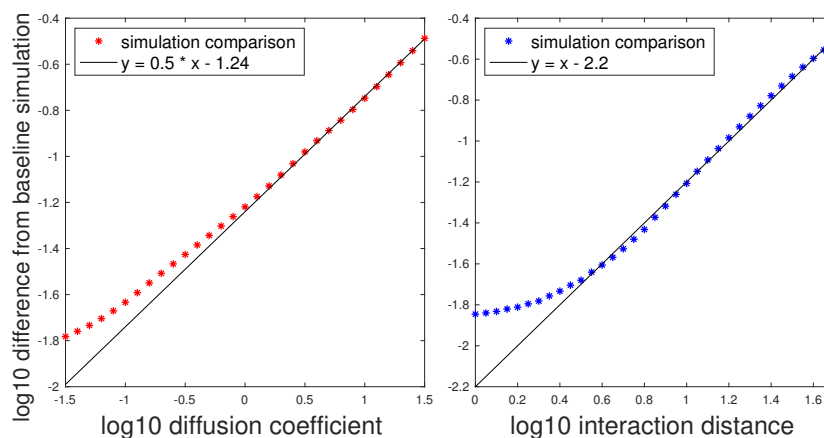


Figure S1: **Comparison of different diffusion coefficient and interaction distance scenarios to the case of no dispersal and local interactions alone.** Differences are squared, summed over all species, and averaged over domain. This is done along the diffusion coefficient (interaction distance) axis in the left (right) panel. Comparison shows that diffusion scales like a square root, and that a normalization factor of 10 should be applied to make it comparable to interaction distance (i.e., $d = 1$ is comparable to an interaction distance of 10).

712 **Peculiarity of dispersal scale:** An interesting problem we encountered, which is
713 worth expounding upon to aid future research in this area, is how to place dispersal on
714 comparable scales and strength to other processes. For both environmental factors and
715 species interactions, we could separate the intensity of variation and the scale over which it
716 takes place. We could do this, for instance, by modelling interactions with a spatial kernel
717 which defines the range of these interactions. For dispersal, however, this distinction does

718 not hold in the continuum limit nor in the stable equilibrium regime that we consider in this
 719 study. This can be understood intuitively in a single dimension: organisms who disperse
 720 from site x to site $x + 1$ at time t will be counted in those that disperse from site $x + 1$ to site
 721 $x + 2$ at a later moment in time. Therefore, dispersing twice as fast between neighboring
 722 sites can be equivalent to dispersing twice as far. This equivalence breaks down when the
 723 details of individual dispersal events matter, e.g., for very rare and long-ranged dispersal
 724 events [52]. But even then, the strength of each dispersal event would still play into the
 725 spatial scale over which dispersal impacts the dynamics over longer times. As a consequence,
 726 defining dispersal scale from a spatial kernel alone might seem more intuitive, but would
 727 actually hide the importance of intensity, and we prefer to simply model nearest-neighbor
 728 dispersal and acknowledge that intensity and scale are entangled.

729 A2 Generating the landscape

730 The landscape profile is defined by a spectral color (ρ) and cutoff (k_c). A spectral color close
 731 to 0 corresponds to “white” noise, i.e., noise that exhibits little or no spatial autocorrelation;
 732 a spectral color close to 1 indicates “red” noise – noise with high spatial autocorrelation [40].
 733 The spectral cutoff creates a point of truncation in the frequency profile that prevents
 734 high variation between adjacent cells, in effect smoothing the noise across the landscape.
 735 Together, color and cutoff control the degree of structural fragmentation of the landscape
 736 (see Fig. S3). More weight on higher frequencies (low ρ , high k_c) entails smaller and less-
 737 connected fragments of similar environmental conditions. Weight on lower frequencies (high
 738 ρ , low k_c) creates long bands of constant environmental conditions which can act as corridors
 739 for species favoring this value.

To generate the environmental landscape $K(\vec{x})$, we prescribe a frequency profile for the noise:

$$F(k \neq 0) = \frac{1}{k^\rho} e^{-k/k_w}, \quad F(0) = 1$$

740 which is a power-law with color ρ ($\rho = 1$ corresponds to red noise) and an exponential
 741 cutoff with wavenumber $k_w = k_c L/2$ which removes high spatial frequencies, smoothing the
 742 landscape and avoiding strong variations between adjacent cells. The construction process is
 743 demonstrated in Fig. S2. Note that the cutoff wavenumber is simply the normalization of the
 744 spectral cutoff by the number of different frequencies represented by the chosen resolution
 745 of the domain, $L/2$, with L the number of cells along the x and y axes, such that in the
 746 spectral domain it represents the resolution of the landscape.

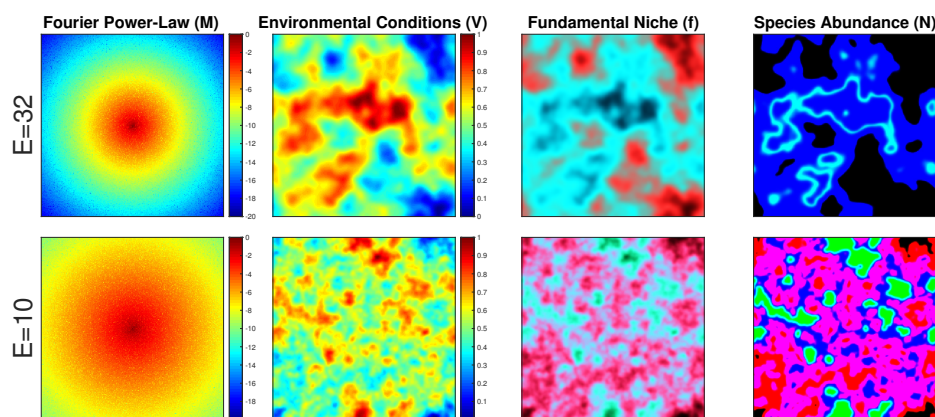


Figure S2: **Demonstration of landscape construction.** The steps of landscape construction are shown in the different columns, with the top (bottom) row corresponding to a landscape with $E = 32$ ($E = 10$). From left to right, the four columns correspond to: 1) The function M , which is a power-law function with exponential cutoff, on a two-dimensional spectral map (i.e., where each cell corresponds to a different spatial frequency), with the addition of random noise. 2) The environmental conditions V , which result from applying the Fourier transform on the previous step, and normalizing the values to range between 0 and 1. 3) The fundamental niches f_i of 3 species, where the value of f_i of each species are encoded in the red, green and blue color channels. 4) The spatial distribution of species biomass N_i at equilibrium, of the same 3 species and with the same color coding, as the previous column. Note that the top-right panel corresponds to the bottom-left column of Fig. 3.

Practically speaking, for a two-dimensional landscape, we generate a $L \times L$ matrix R_{ij} of uniform random numbers over $[-1, 1]$ corresponding to amplitudes for each wave vector (k_x, k_y) . We then multiply these random numbers by the profile above

$$M_{ij} = R_{ij} F(\sqrt{k_i^2 + k_j^2})$$

747 with $k_i = i - \frac{L}{2}$ where index i is a natural number running over $[1, L]$. We set the element
 748 $M_{L/2, L/2}$ corresponding to the uniform trend ($k_i = k_j = 0$) to 5. Finally, we apply a Fast
 749 Fourier Transform on the matrix M_{ij} to obtain the landscape matrix V . As explained in
 750 the main text, this landscape matrix V is used to define the growth rate r_i using a Gaussian
 751 function (see eq. 2), which in turn determines the species biomass distribution N_i (see eq. 1).
 752 We show in Fig. S3 the environment as a function of different values of ρ and k_c , to better
 753 visualize their effect.

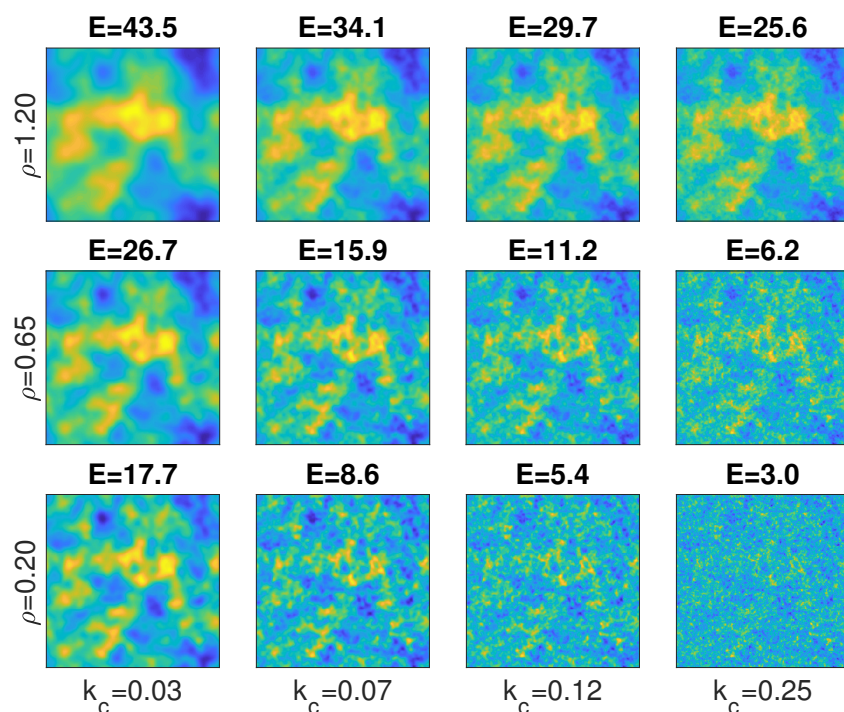


Figure S3: **How ρ and k_c shape the landscape structure**, shown by maps of the environmental conditions V . We show an example of how a landscape is affected by different values of ρ (rows) and k_c (columns). On top of each panel we also note the environmental scale E that corresponds to the combination of ρ and k_c . We can see that smaller k_c values lead to a landscape with less sharp transitions (i.e., smoother), whereas ρ has a more significant effect on the overall scale. In other figures and in the main text we choose ρ and k_c concordantly, with large ρ values together with small k_c values for a large E , and small ρ values together with large k_c values for a small E .

754 A3 Different environmental scales

755 We show below a few additional plots, which explore the impact of different values of
 756 environmental scale E . In Fig. S4 we show the overall difference in community state, between
 757 different sets of values of D and I to the case of no dispersal and local interactions, for two
 758 values of E .

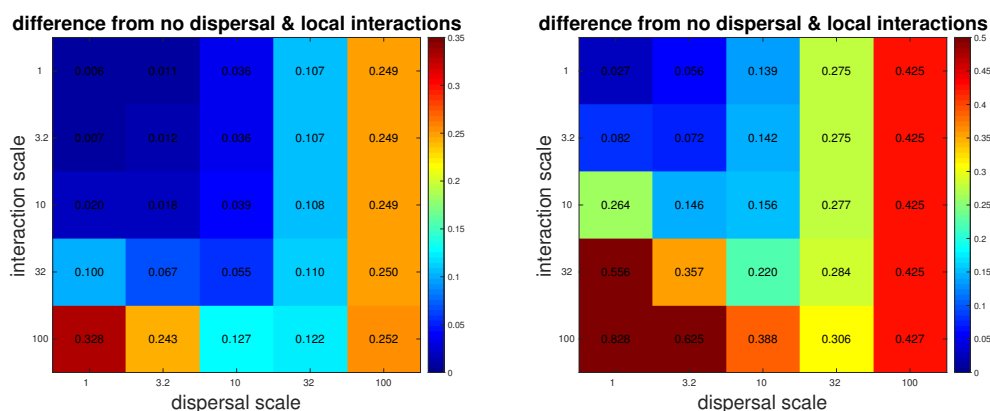


Figure S4: **Comparison of various scenarios to the case of no dispersal and local interactions alone**. Difference is measured by averaging over the squared sum of each cell for a given value of I and D , against the baseline of $D = I = 0$. This is done for for 5×5 different parameter sets with different values of D and I , for two different values of E , 32 and 10, in the left and right panels, respectively.

759 In Fig. S5 we consider different E values, and see how changing either I or D affects
 760 the overall change in system state (compared with the baseline of no dispersal and local
 761 interactions). In both figures we can see that big differences in the state of the system due
 762 to higher I or D (seen as dark blue region in Fig. S4, and region below the dotted line in
 763 Fig. S5) occur for lower values for I and D , and only when E is sufficiently high. This
 764 demonstrates that the environmental scale E determines the threshold scale of I and D in
 765 which they can have a substantial effect on the community.

766 We also test how the inflection point of SAR (measured in the same way as in the main
 767 text), changes along a range of E values (Fig. S6). We can see that as long as dispersal is not
 768 too high (i.e., the three cases where $D = 1$), the inflection point follows the environmental
 769 scale E (seen by the roughly parallel lines to the 1:1 line).

770 In Fig. S7 and Fig. S8 we show the spatial distributions of biomass and species richness,
 771 for a different landscape, one that has an environmental scale of $E = 10$.

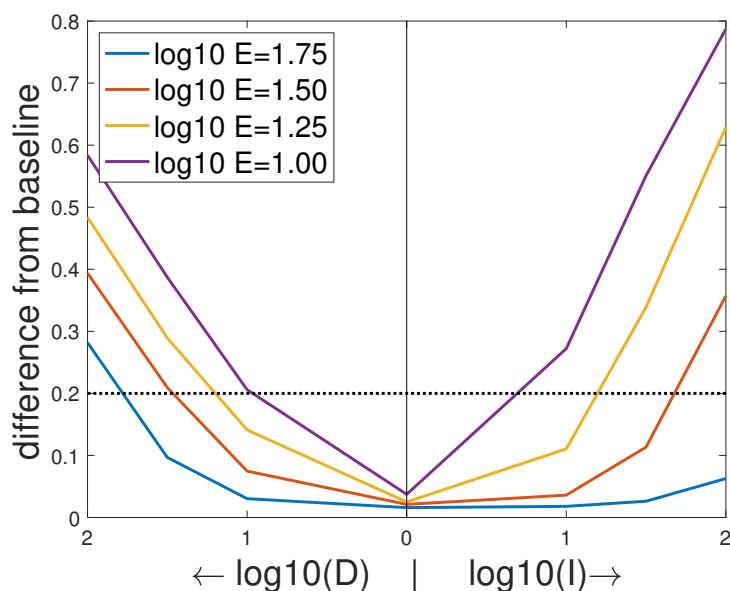


Figure S5: **Average difference from a community with no dispersal and local interactions only.** Difference is measured by averaging over the squared sum of each cell for a given value of I and D , against the baseline of $D = I = 0$. Left half shows the effect of D alone, while right half shows the effect of I alone.

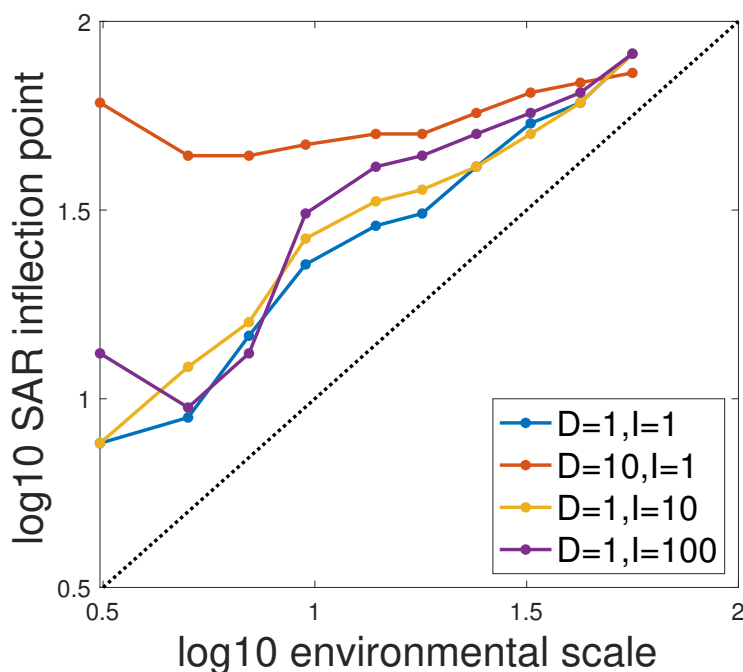
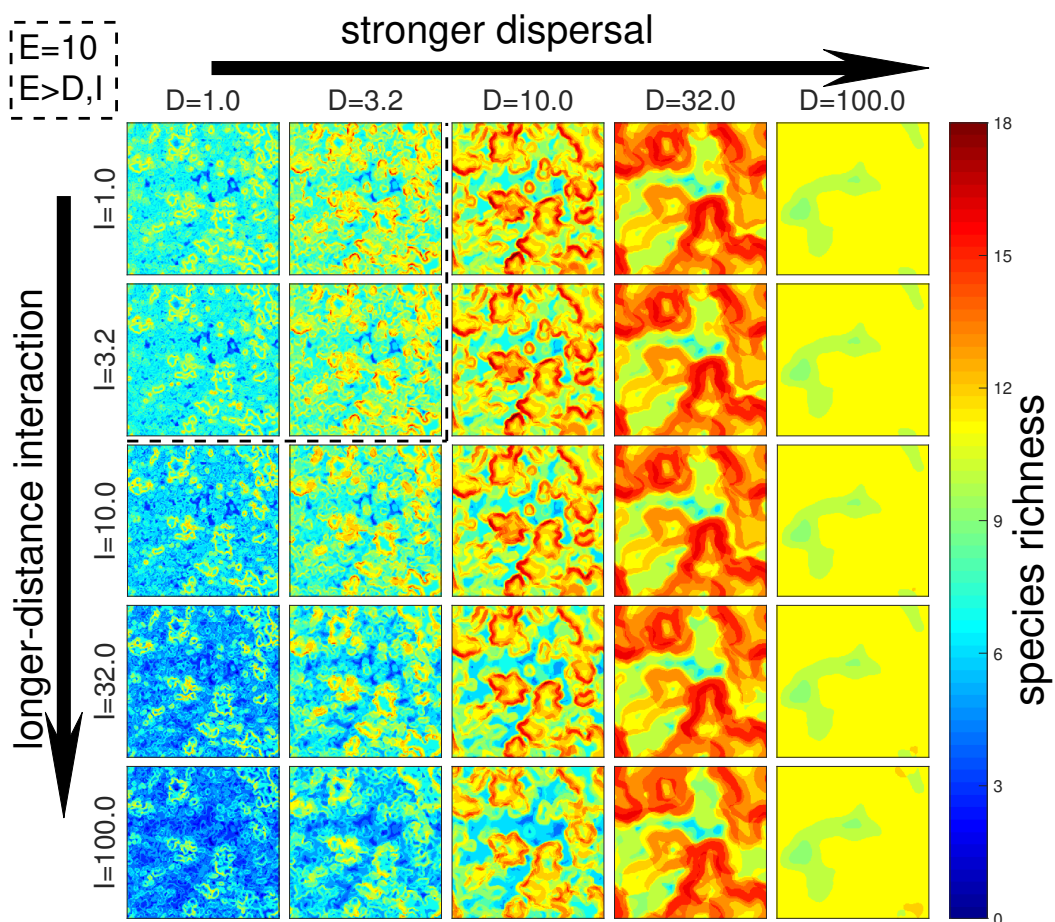
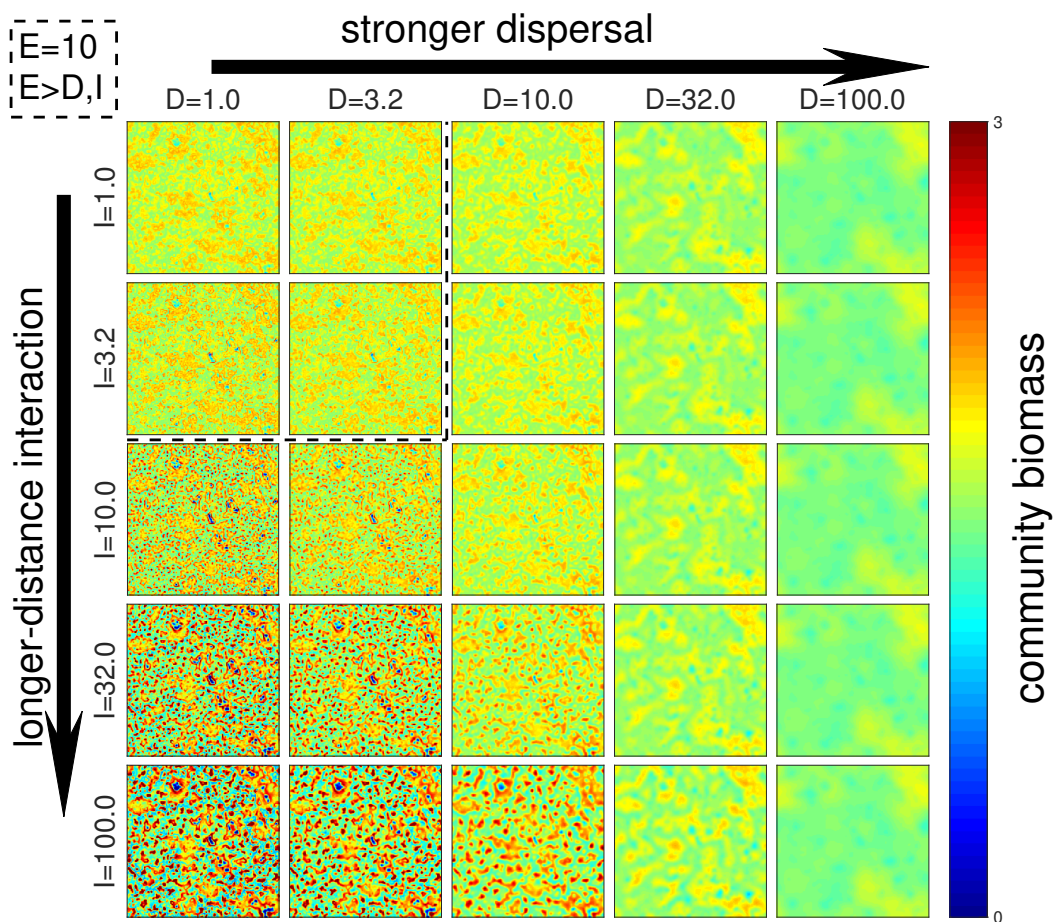


Figure S6: **Inflection point of SAR** for different combination of scales. For four sets of values of I and D ($D = 1, I = 1$; $D = 1, I = 10$; $D = 1, I = 100$; $D = 10, I = 1$), we show how the inflection point of SAR changes along a range of 10 values of E (with values between 56 and 3).



772 A4 Additional plots

773 We show below a few additional plots.

774 In Fig. S9 we show the spatial distribution of species richness, for 5x5 different parameter
 775 sets with different values of D and I , corresponding to Fig. 2. In Fig. S10 and Fig. S11 we
 776 show summary statistics for each of these 5x5 parameter sets, of total community biomass,
 777 average local diversity, and total diversity.

778 Finally, we explore in Fig. S12 the sensitivity of our results to the parameter β , and
 779 demonstrate using Fig. S13 the calculation of species' spatial correlations, which is used to
 780 estimate the environmental scale E .

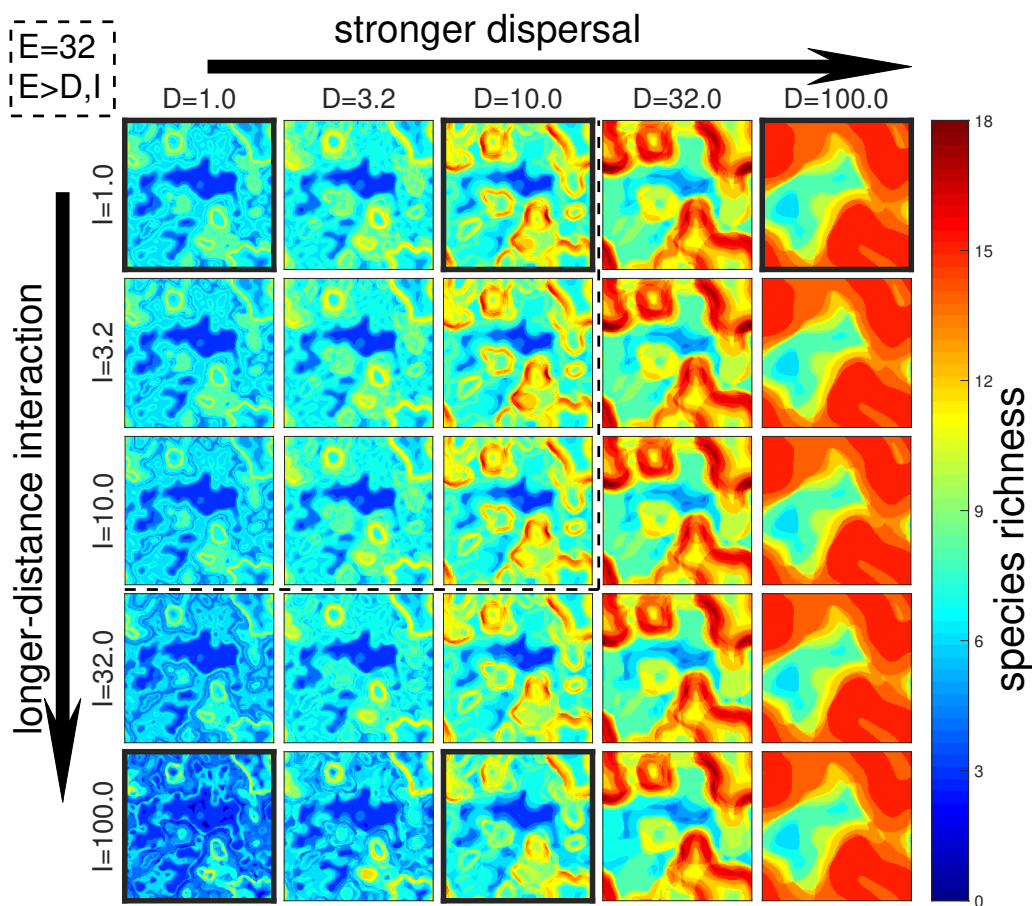


Figure S9: *Species richness plots*, corresponding to Fig.2, for the 5x5 scenarios ($E = 32$).

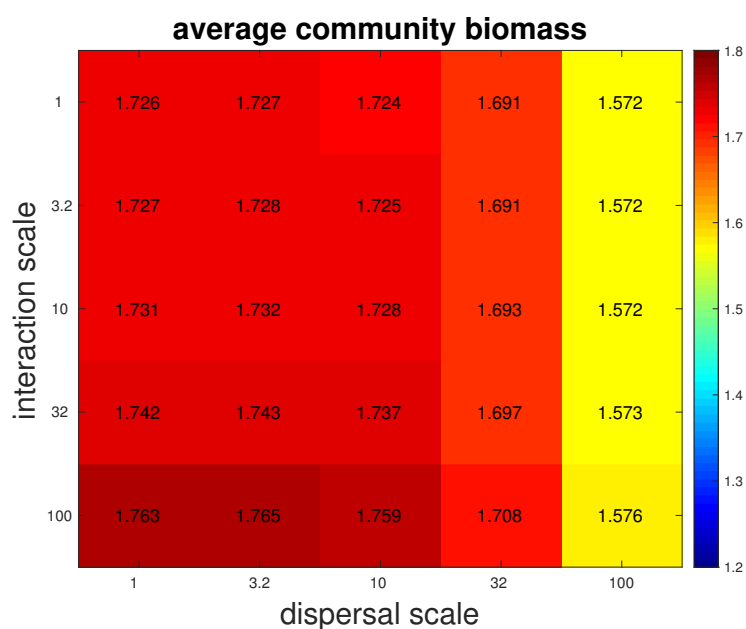


Figure S10: *Total community biomass*, averaged over domain, for the 5x5 scenarios ($E = 32$).

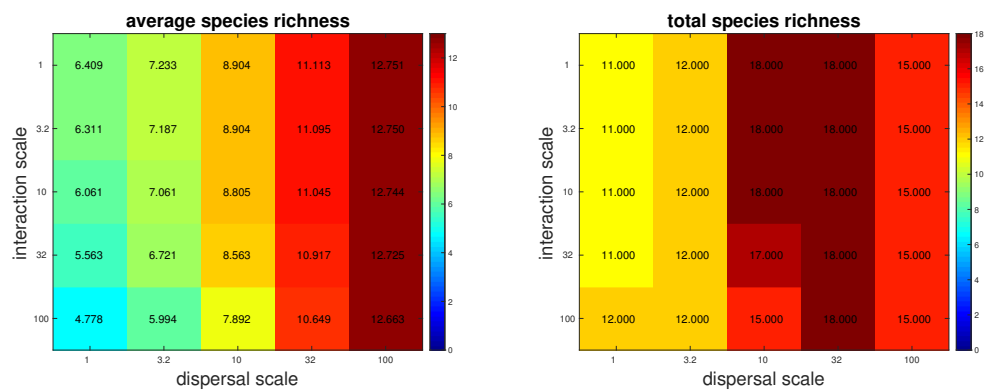


Figure S11: **Diversity plots.** Average local diversity of community (left) and total community diversity, (right) for the 5x5 scenarios ($E = 32$).

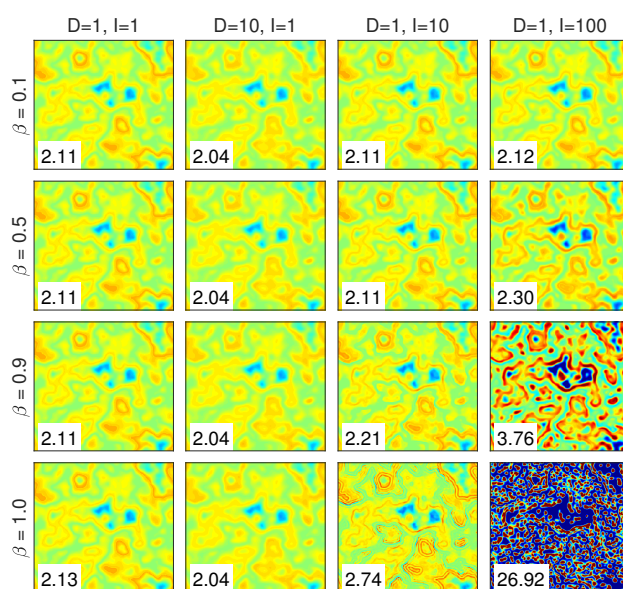


Figure S12: **Effect of changing the value of the parameter β ,** which determines the fraction of regional interactions. Each panel shows the spatial distribution of total biomass, with columns showing results for different values of I and D , while lower rows showing increasing values of β . The number in each panel shows the highest biomass density seen in the panel (where each panel's colors are scaled to that value to better show the spatial structure). For low values of β (top two rows) scale of interactions I has minimal effect (clearly seen by right column looking the similar to other columns). For values of β (bottom two rows) the effect of I becomes strong and clearly visible. However, for very high values of β (bottom row) the effect also includes very high densities of biomass, which is not very realistic. We therefore choose a high value of β but not so high as to lead to very high densities (leading us to the middle ground of $\beta = 0.9$).

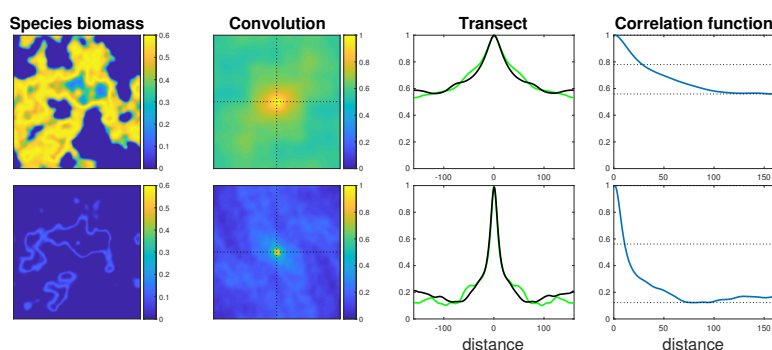


Figure S13: **Demonstration of calculation of correlation function.** The steps of calculating the correlation function are shown in the different columns, with the top (bottom) row corresponding to two different species in the same landscape used in Fig. 2. From left to right, the four columns correspond to: 1) The spatial distribution of biomass of a single species N_i . 2) Correlation map, which is the result of a convolution of this spatial distribution with itself. 3) Transects of the correlation map (horizontal and vertical, shown in green and black), also marked in previous column by dotted lines. 4) Averaging of transects resulting in the correlation function. Horizontal dotted lines show the highest and lowest values of the correlation function, along with the average of the two which is used as a threshold to determine the scale of correlation.