

Defining a dichotomous indicator for population-level assessment of dietary diversity among pregnant adolescent girls and women: a secondary analysis of quantitative 24-h recalls from rural settings in Bangladesh, Burkina Faso, India and Nepal

Eric O Verger, Sabrina Eymard-Duvernay, Dang Bahya-Batinda, Giles T Hanley-Cook, Alemayehu Argaw, Elodie Becquey, Loty Diop, Aulo Gelli, Helen Harris-Fry, Shivani Kachwaha, et al.

▶ To cite this version:

Eric O Verger, Sabrina Eymard-Duvernay, Dang Bahya-Batinda, Giles T Hanley-Cook, Alemayehu Argaw, et al.. Defining a dichotomous indicator for population-level assessment of dietary diversity among pregnant adolescent girls and women: a secondary analysis of quantitative 24-h recalls from rural settings in Bangladesh, Burkina Faso, India and Nepal. Current Developments in Nutrition, 2024, 8 (1), pp.102053. 10.1016/j.cdnut.2023.102053 . hal-04322954

HAL Id: hal-04322954 https://hal.inrae.fr/hal-04322954

Submitted on 5 Dec 2023

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers. L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.



Distributed under a Creative Commons Attribution 4.0 International License

Defining a dichotomous indicator for population-level assessment of dietary diversity among pregnant adolescent girls and women: a secondary analysis of quantitative 24-h recalls from rural settings in Bangladesh, Burkina Faso, India and Nepal

Eric O. Verger, Sabrina Eymard-Duvernay, Dang Bahya-Batinda, Giles T. Hanley-Cook, Alemayehu Argaw, Elodie Becquey, Loty Diop, Aulo Gelli, Helen Harris-Fry, Shivani Kachwaha, Sunny S. Kim, Phuong Hong Nguyen, Naomi M. Saville, Lan Mai Tran, Rock R. Zagré, Edwige Landais, Mathilde Savy, Yves Martin-Prevel, Carl Lachat

PII: S2475-2991(23)26637-8

DOI: https://doi.org/10.1016/j.cdnut.2023.102053

Reference: CDNUT 102053

To appear in: Current Developments in Nutrition

Received Date: 19 September 2023

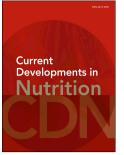
Revised Date: 21 November 2023

Accepted Date: 26 November 2023

Please cite this article as: E.O. Verger, S. Eymard-Duvernay, D. Bahya-Batinda, G.T. Hanley-Cook, A. Argaw, E. Becquey, L. Diop, A. Gelli, H. Harris-Fry, S. Kachwaha, S.S. Kim, P.H. Nguyen, N.M. Saville, L.M. Tran, R.R. Zagré, E. Landais, M. Savy, Y. Martin-Prevel, C. Lachat, Defining a dichotomous indicator for population-level assessment of dietary diversity among pregnant adolescent girls and women: a secondary analysis of quantitative 24-h recalls from rural settings in Bangladesh, Burkina Faso, India and Nepal, *Current Developments in Nutrition*, https://doi.org/10.1016/j.cdnut.2023.102053.

This is a PDF file of an article that has undergone enhancements after acceptance, such as the addition of a cover page and metadata, and formatting for readability, but it is not yet the definitive version of record. This version will undergo additional copyediting, typesetting and review before it is published in its final form, but we are providing this version to give early visibility of the article. Please note that, during the production process, errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

© 2023 Published by Elsevier Inc. on behalf of American Society for Nutrition.



Defining a dichotomous indicator for population-level assessment of dietary diversity among pregnant adolescent girls and women: a secondary analysis of quantitative 24-h recalls from rural settings in Bangladesh, Burkina Faso, India and Nepal

Author names and affiliations

Eric O. Verger^a, Sabrina Eymard-Duvernay^a, Dang Bahya-Batinda^a, Giles T. Hanley-Cook^b, Alemayehu Argaw^{b,c}, Elodie Becquey^d, Loty Diop^d, Aulo Gelli^e, , Helen Harris-Fry^f, Shivani Kachwaha^g, Sunny S. Kim^e, Phuong Hong Nguyen^e, Naomi M. Saville^h, Lan Mai Tranⁱ, Rock R. Zagré^d, Edwige Landais^a, Mathilde Savy^a, Yves Martin-Prevel^a and Carl Lachat^b.

^a MoISA, Univ Montpellier, CIRAD, CIHEAM-IAMM, INRAE, Institut Agro, IRD, Montpellier, France; ^b Department of Food Technology, Safety and Health, Faculty of Bioscience Engineering, Ghent University, Ghent, Belgium; ^c Department of Population and Family Health, Institute of Health, Jimma University, Jimma, Ethiopia; ^d International Food Policy Research Institute (IFPRI), Dakar, Senegal; ^e International Food Policy Research Institute (IFPRI), Washington, DC, USA; ^f Department Population Health, London School of Hygiene & Tropical Medicine, London, UK; ^g Johns Hopkins University, Baltimore, Maryland, USA; ^h UCL Institute for Global Health, London, United Kingdom; ⁱ Emory University, Atlanta, Georgia, USA

Corresponding author

Eric O. Verger, MoISA, IRD, 911 avenue d'Agropolis, 34000 Montpellier, France. Email: eric.verger@ird.fr

1 Abstract (300 words maximum)

Background: The Minimum Dietary Diversity for Women of Reproductive Age (MDD-W)
indicator was validated as a proxy of micronutrient adequacy among non-pregnant women in
low- and middle-income countries (LMICs). At that time, indeed, there was insufficient data to
validate the indicator among pregnant women, who face higher micronutrient requirements.

Objective: This study aimed to validate a minimum food group consumption threshold, out of
the 10 food groups used to construct MDD-W, to be used as a population-level indicator of
higher micronutrient adequacy among pregnant women aged 15-49 years in LMICs.

Methods: We used secondary quantitative 24-hour recall data from 6 surveys in 4 LMICs
(Bangladesh, Burkina Faso, India and Nepal, total n=4909). We computed the 10-food group
Women's Dietary Diversity Score (WDDS-10) and calculated the mean probability of adequacy
(MPA) of 11 micronutrients. Linear regression models were fitted to assess the associations
between WDDS-10 and MPA. Sensitivity, specificity and proportion of individuals correctly
classified were used to assess the performance of MDD-W in predicting an MPA >0.60.

Results: In the pooled sample, median values (interquartile range) of WDDS-10 and MPA were 3 (1) and 0.20 (0.34), respectively, while the proportion of pregnant women with an MPA >0.60 was 9.6%. The WDDS-10 was significantly positively associated with MPA in each survey. Although the acceptable food group consumption threshold varied between 4 and 6 food groups across surveys, the threshold of 5 showed the highest performance in the pooled sample with good sensitivity (62%), and very good specificity (81%) and percentage of correctly classified individuals (79%).

Conclusions: The WDDS-10 is a good predictor of dietary micronutrient adequacy among pregnant women aged 15-49 years in LMICs. Moreover, the threshold of 5 or more food group for the MDD-W indicator may be extended to all women of reproductive age, regardless of their physiological status.

26 Teaser Text

- 27 This study aimed to validate whether the threshold of 5 or more food group for the MDD-W
- indicator can to be used among pregnant women aged 15-49 years in low- and middle-income
- 29 countries.
- 30
- 31 Abbreviations:
- 32 BLUP: Best linear unbiased predictor
- 33 EAR: Estimated Average Requirement
- 34 LMIC: Low- and Middle-Income Country
- 35 MDD-W: Minimum Dietary Diversity for Women
- 36 MPA: Mean probability of adequacy
- 37 PA: Probability of adequacy
- 38 WDDS-10: 10-food group Women's Dietary Diversity Score
- 39
- 40 Keywords: dietary diversity; indicator; micronutrient adequacy; minimum dietary diversity for
- 41 women; pregnant; resource-poor settings.

42 Introduction

Micronutrients are essential vitamins and minerals whose subclinical deficiencies contribute 43 44 to an increased risk of morbidity and mortality (1). A recent analysis suggested that two-thirds of non-pregnant women of reproductive age have one or more micronutrient deficiencies 45 worldwide, with higher prevalence in low- and middle-income countries (LMICs) (2). There 46 are important changes in dietary requirements driven by physiological processes during 47 pregnancy, including increased requirements for folate, iron, vitamin B12 and B6, and zinc 48 (3). These deficiencies are exacerbated during pregnancy due to an additional demand for 49 50 nutrients to support both fetal growth and development and maternal metabolism (4), and can result in adverse outcomes of pregnancy and birth (5), as well as maternal depression and 51 52 cognitive impairment (6).

Dietary diversification is a food-based strategy that has been widely promoted to address 53 micronutrient deficiencies (7). To help achieve healthy diets, eating a diversity of foods is 54 needed to help achieve healthy diets (8,9) as recommended by most dietary guidelines (10). 55 As a result, a large range of interventions and programmes to improve nutrition through 56 57 dietary diversification has been developed, and has subsequently triggered a demand for a set 58 of harmonized indicators to monitor progress. Subsequently, several simple indicators assessing dietary diversity were developed, primarily for use in global and national 59 monitoring, and in survey contexts where more detailed dietary methods that include 60 estimation of food quantities are infeasible. 61

In this context, the Women's Dietary Diversity Project developed and validated simple food
group indicators with consistent and relevant meaning across different contexts and over time.
The most recent example is the Minimum Dietary Diversity for Women of Reproductive Age
(MDD-W), a simple population-level dichotomous indicator expressed as the proportion of
non-pregnant women of reproductive age who consumed at least 5 out of 10 defined food

67	groups over the previous 24 hours (11). MDD-W was validated using nine datasets from 6
68	distinct LMICs as a proxy for a minimally acceptable level of intake adequacy of 11
69	micronutrients among non-pregnant women of reproductive age (12,13).
70	While the initial MDD-W validation study was able to assess the performance of the indicator
71	for non-pregnant non-lactating women and non-pregnant lactating women, this was not
72	possible for pregnant women due to the lack of data. Recent studies have used the threshold of
73	5 or more food groups to determine whether pregnant women had more adequate
74	micronutrient intakes but without further validation of this dichotomous indicator in this
75	population group (14–16). However, pregnant women generally have higher micronutrient
76	requirements than non-pregnant women (3), which may change the performance of food
77	group indicators in predicting adequate micronutrient adequacy in this specific population.
78	The only validation study among pregnant women we are aware of showed that an adapted 6
79	or more food group threshold markedly improved performance of the indicator in predicting
80	micronutrient adequacy among pregnant girls and pregnant women in Bangladesh (17). Using
81	secondary quantitative 24-hour recall data from 6 surveys in 4 LMICs, this study aimed to
82	validate a minimum food group consumption threshold, out of the 10 food groups used to
83	construct MDD-W, to be used as a population-level indicator of higher micronutrient
84	adequacy among pregnant women aged 15-49 years in LMICs. We followed the methods
85	used by previous studies on the development and validation of MDD-W to ensure
86	comparability of the analysis and facilitate the interpretation of findings (12,13).

87

88 Methods

89 Selection of surveys

90 This study was based on a pre-identified set of datasets that was completed by a systematic91 review of studies which collected dietary intakes from pregnant adolescent girls and women

92	in LMICs, using one or multiple 24-hour dietary recalls. Inclusion criteria were: (i) food
93	consumption data collected among pregnant women (15-49 y) in LMICs; (ii) quantitative
94	dietary intake data collected through one or multiple 24-hour dietary recalls; (iii) use of
95	relevant local food composition data with information on the 11 micronutrients included in
96	the initial development and validation of MDD-W (vitamin A expressed in retinol activity
97	equivalents (RAE), thiamin, riboflavin, niacin, vitamin B6, folate, vitamin B12, vitamin C,
98	calcium, iron and zinc); (iv) minimum sample size of 100 pregnant adolescent girls and
99	women and (v) repeated 24-hour dietary recalls from at least 10% of the study sample or
100	being able to be matched with a relevant dietary intake survey with two non-consecutive days
101	of recall to estimate external within-person variance.
102	
103	Study design and participants
104	Six datasets with quantitative 24-hour recall data collected from rural areas in Bangladesh in
105	2015 (17), Burkina Faso in 2017/2019/2020 (BF1, (18)), 2020 (BF2, (19)) and 2019/2021
106	(BF3, (20)), India in 2019 (21) and Nepal in 2015 (22) were selected for analysis. Each
107	dataset is described in more detail in Supplemental table 1, which includes their selection
108	process. Briefly, there were 5 pre-identified datasets (Bangladesh, BF1, BF2, BF3 and India)
109	and we undertook a literature research to identify others, leading to add the dataset from
110	Nepal. The included studies' primary objectives were to assess the feasibility and impact of
111	maternal nutrition packages or integrated agriculture-nutrition interventions (Bangladesh,
112	BF1, BF3 and India), to assess the efficacy of fortified balanced energy-protein
113	supplementation (BF2), or to characterize the status and determinants of intra-household food
114	and nutrient allocation, and test the effect of pregnancy interventions upon dietary intake
115	(Nepal). None of the study samples was nationally representative. Data quality control was
116	carried out by the data providers, including the exclusion of outliers. The representativeness

of each sample has been discussed in the original articles and primary study protocols for all
sites were approved by ethical review committees or institutional review boards (17–19,21–
25).

120

121 Dietary data collection

In all studies, dietary data were collected using one to three quantitative multiple-pass 24-hour 122 dietary recalls conducted by enumerators specially trained for this purpose (26). Participants 123 were asked to describe all foods and beverages consumed during the preceding 24 hours. 124 Recipes were usually collected from the household member who was responsible for cooking. 125 126 Portion sizes were estimated using methods best suited to local foods and contexts (e.g. previously distributed plates and bowls, common household measures, water volume, rice, 127 images, clay or wooden models, etc.). Only two datasets had repeated 24-hour dietary recalls 128 on non-consecutive days, with two recalls for 19% of the sample (BF1) and three recalls for 129 87% of the sample (Nepal). Dietary data were converted into nutrient intakes using country 130 specific food composition tables; the application of yields and nutrient retention factors was 131 done by data providers according to their own practice and information is available from 132 original studies (17-19,21-23). 133

134

135 MDD-W and WDDS-10

Among the various indicators with different food groupings developed and tested as part of the Women's Dietary Diversity Project I and II, the dichotomous MDD-W indicator has been shown to have a strong relationship to micronutrient adequacy and high consistency in terms of threshold which best discriminated higher versus lower micronutrient adequacy across various countries (12,27). The MDD-W was constructed considering 10 mutually exclusive food groups consisting of: 1) starchy staple foods, 2) pulses; 3) nuts and seeds; 4) dairy

products; 5) flesh foods; 6) eggs; 7) dark green leafy vegetables; 8) vitamin A-rich fruits and
vegetables; 9) other vegetables; and 10) other fruits. The 10 food groups are summed into a
score (WDDS-10) ranging from 0 to 10, starting with a score of 0 and adding 1 point per food
group consumed (if the total consumption of the foods in the food group was at least 15
g/day)¹. The WDDS-10 was computed using a single day recall (the first day in case of
repeated recalls). MDD-W was coded as 1 if WDDS-10 reached 5 food groups or more, and 0
if 4 or lower.

149

150 Micronutrient requirements, usual intakes and probability of adequacy

151 We used the Estimated Average Requirements (EAR) and coefficients of variations proposed

by Nguyen *et al.* (17), that are based on the information from the WHO/FAO (29), the

153 National Academy of Medicine (formerly the Institute of Medicine) (30,31) and the

154 International Zinc Nutrition Consultative Group (IZiNCG) (32). These requirements were

used regardless of the pregnancy trimester, age or country context of the participants

156 (Supplemental table 2). These requirements were chosen rather than those proposed by Allen

157 et al. (33) to enhance comparability and facilitate interpretation of findings with previous

studies on the development and validation of MDD-W (12,13,17).

159

Analogous to previous studies on the development and validation of MDD-W (12,13), we used the probability approach to estimate the micronutrient adequacies of each of the 11 micronutrients (28). This approach is based on information or assumption about both the distribution of nutrient requirements in the population and the day-to-day variations (within-

164 person) of nutrient intakes. We applied a Box-Cox transformation to the nutrient intake

¹ This is of course not easy to do in practice, when collecting data; therefore, what is recommended in the FAO MDD-W guidelines is to apply the 15g limit to each food. However, we decided here to stick to the methodology used for the validation of the MDD-W for the sake of comparability.

distribution of every micronutrient to obtain normal distributions. For each participant and 165 micronutrient in each separate dataset, we calculated the best linear unbiased predictor 166 (BLUP) of the individual's usual intake (34) which was then used to calculate the probability 167 of adequacy for every micronutrient (see Supplemental Methods). All usual nutrient intakes 168 have been calculated solely on the basis of food intakes, excluding intakes from food 169 supplements (e.g. fortified balanced energy-protein supplementation in BF2). When datasets 170 contained repeated 24-hour dietary recalls, the within person variance was defined as the 171 mean of squared intra-individual SDs. When datasets contained only one 24-hour dietary 172 recall, we used an external within-person variance estimate from a relevant dietary intake 173 174 survey with two non-consecutive days of recall (35,36). We used the external within-person 175 variance to between-person variance ratio multiplied by the between-person variance of our dataset as the within-person variance in the BLUP calculations. A relevant dietary intake 176 survey was defined as a survey conducted in the same geographical and seasonal context 177 among pregnant adolescent girls or women. For Bangladesh, we used the within-person 178 variance estimate from a subsample of the baseline study (~20%) that also participated in the 179 endline study conducted a year later (37). For BF2 and BF3, we used the within-person 180 variance estimate from BF1 due to the fact that these three surveys were conducted in the 181 182 same context (Boucle du Mouhoun, Centre-Ouest and Haut-Bassins for BF1, Haut-Bassins for BF2 and Boucle du Mouhoun for BF3) among pregnant adolescent girls and women. For 183 India, we used the within-person variance estimate from repeated 24-hour dietary recall used 184 185 to validate a Food Frequency Questionnaire among pregnant women living with or without HIV in Pune, India (38). 186

187

Probability of adequacy (PA) was calculated as the probability that a woman's usual intake
was at or above the EAR during pregnancy (28). For each individual, we averaged the mean

- 190 of the individual PAs for the 11 micronutrients to form the mean probability of adequacy
- 191 (MPA). Like individual PAs, the MPA has a possible range of 0–1.
- 192

193 Data analysis

Data were analyzed with Stata 17 (Statacorp, College Station, TX) and the Stata syntax that 194 was used for MDD-W validation in non-pregnant women (12,13), with a few minor revisions 195 to match the aims of our analyses. Descriptive statistics are reported as medians (interquartile 196 ranges) due to skewness of the distributions, except for age, height, weight and energy intake, 197 which are reported as means (SDs). Associations between the WDDS-10 and MPA (with or 198 without adjustment for total energy intake) were assessed by fitting simple linear regressions. 199 For the pooled sample, a mixed-effects regression model was used to examine the association 200 between WDDS-10 and MPA, with random effect at dataset level to take into account the 201 within-survey correlation. The MPA variable was previously transformed by BoX-Cox 202 203 transformation for all the regression models.

We used receiver operating characteristic (ROC) analysis and area under the curve (AUC) to 204 assess the diagnostic performance of WDDS-10 in predicting a MPA >0.60, with an AUC 205 206 >0.70 deemed acceptable for predictive capacity. We estimated sensitivity, specificity and percentage of correct classifications for MDD-W across datasets and in a pooled analysis. The 207 208 MPA level of 0.60, as well as the interpretation thresholds, were selected to ensure comparability with the previous analysis used to validate the MDD-W (12,13). Sensitivity 209 210 (i.e. ability to correctly detect a person with an MPA >0.60) is defined by the ratio between 211 the true positives and the sum of true positives and false negatives. Specificity (i.e. ability to correctly detect a person with an MPA ≤ 0.60) is defined by the ratio between the true 212 213 negatives and the sum of true negatives and false positives. A threshold was considered good 214 when both sensitivity and specificity were >0.60 and it was considered fair enough if only one

test characteristic was >0.60 and the other >0.50. Moreover, while we looked for the best 215 balance between sensitivity and specificity, we favored specificity over sensitivity when 216 trade-offs must be made, in order to be certain to identify the highest proportion of 217 participants with a MPA ≤ 0.60 . The percentage of correct classifications is defined by the 218 ratio between the sum of true positives and true negatives and the sum of true positives, false 219 positives, true negatives and false negatives. A threshold was considered as good when the 220 percentage of individuals correctly classified was >0.70 and it was considered fair enough if 221 >0.60. 222

In order to understand the implications of some methodological choices, we conducted 223 additional robustness analyses to estimate sensitivity, specificity and the percentage of correct 224 225 classifications for MDD-W across datasets and in a pooled analysis according to three distinct scenarios. In the first robustness analysis, we tested 3 scenarios (Sc1, Sc2 and Sc3) where 226 only 1 of the 3 Burkinabe datasets was included in the pooled analysis (BF1, BF2 and BF3, 227 respectively), in order to keep into account the potentially redundant nature of using three 228 surveys from Burkina Faso. In the second robustness analysis, we used the same 229 recommendations from WHO/FAO (29), the National Academy of Medicine (30,31) and the 230 IZiNCG (32) but took into account pregnancy trimester, age and level of bioavailability of 231 232 iron and zinc (see Supplemental table 3). In the third robustness analysis, we used the requirements proposed by Allen et al. (33) which take into account age and level of 233 bioavailability of iron and zinc but not pregnancy trimester (see **Supplemental table 4**). 234

235

236 **Results**

237 Characteristics of participants

Data were available for 4909 pregnant adolescent girls and women (Table 1), with sample
sizes of the datasets ranging from 452 (BF1) to 1912 (BF3). The mean (SD) age of

participants was 25.7 (6.2) years, with participants from Nepal being on average younger than 240 pregnant women from other countries. The inclusion of adolescent girls (15-18 years) across 241 studies varied from none (India) to up to 26% (Bangladesh), and was 7.1% in the pooled 242 sample. The pregnancy trimester distribution was highly variable across datasets, with a near-243 even distribution in BF1, whereas almost all participants were in their third trimester in Nepal. 244 Participants in their third trimester represented almost 60% of the pooled sample. Pregnant 245 women in the Burkinabé datasets were on average taller and heavier than participants from 246 other countries. 247

248

249 **Dietary diversity**

The median (interquartile range) WDDS-10 in the pooled sample was 3 (1), with higher 250 median scores in the Bangladeshi, Nepalese, and Indian datasets compared to the three 251 252 Burkinabe datasets (Table 1). Figure 1 shows the percentage of pregnant adolescent girls and women consuming each of the 10 food groups used to construct MDD-W across the six 253 datasets. Consistently across datasets, the diet of all participants was based on starchy staple 254 foods. Most participants consumed other vegetables, but with large variations ranging from 255 55% in BF1 and BF3 to 91% in Nepal. The prevalence of participants consuming pulses and 256 257 dairy products greatly differed across datasets: for pulses it was high in Nepal (over 80%), moderate in Bangladesh and India (59 and 46%, respectively), and low in the three Burkinabe 258 datasets (27% in BF1, 14% in BF2 and 15% in BF3). As for the prevalence of consumption of 259 260 dairy products, it was very high in India (over 80%), moderate in Nepal and Bangladesh (53 and 33%, respectively), and low in the three Burkinabe datasets (4% in BF1, 3% in BF2 and 261 11% in BF3). In contrast, the prevalence of participants consuming nuts and seeds, and dark 262 green leafy vegetables was higher in the three Burkinabe datasets. The prevalence of 263

264 participants consuming flesh foods, eggs, and other fruits was higher in the Bangladeshi265 datasets.

266

267 Energy and nutrient intakes and the probability of adequacy

268 The mean (SD) energy intake of the pregnant adolescent girls and women was 2068 (969)

kcal per day in the pooled sample (**Table 2**), ranging from 1816 (838) kcal in BF3 to 2473

270 (1482) kcal in BF2. For all micronutrients apart from zinc, median intakes in the pooled

sample were below the EAR (Supplemental table 5). However, there were differences

between datasets, with median intakes in the Nepalese and Bangladeshi datasets above the

EAR for 5 and 4 micronutrients, respectively. Accordingly, PAs varied widely across datasets

(Table 2). Across surveys, the PAs of vitamin A, riboflavin, folate, vitamin B12, calcium, and

iron were <0.50. The median (IQR) MPA of the participants was 0.20 (0.34) in the pooled

sample, ranging from 0.09 (0.21) in BF1 to 0.43 (0.32) in Nepal. The proportion of

participants with MPA above the threshold of 0.60 was low, at 9.6% in the pooled sample and

278 ranged from 2.4% (BF1) to 23.4% (Nepal).

279

280 Association between WDDS-10 and MPA

Figure 2 illustrates non-adjusted associations between WDDS-10 and MPA (see

282 Supplemental table 6 for details of the number of pregnant women consuming various

numbers of food groups by dataset). The WDDS-10 was significantly and positively

associated with the MPA in every dataset (all P <0.001) (**Table 3**). Unadjusted regression

coefficients ranged from 0.079 (95% CI: 0.070, 0.088) to 0.309 (95% CI: 0.250, 0.367) and

was 0.168 (95% CI: 0.157, 0.178) for the pooled sample. The unadjusted models explained

between 14% and 33% of the MPA variance, and 28% in the pooled sample. In models

including total energy intake (kcal/d) as covariate, associations were attenuated in all datasets

but remained highly significant. Energy adjusted regression coefficients ranged from 0.038

290 (95% CI: 0.028, 0.050) to 0.166 (95% CI: 0.114, 0.218) and was 0.079 (95% CI: 0.069,

291 0.088) in the pooled sample. The energy adjusted models explained between 29% and 66% of

- the MPA variance, and 41% in the pooled sample.
- 293

313

294 Food group indicator performance and identification of thresholds

The AUC value in the pooled sample was 0.78 (95% CI: 0.75, 0.80), which indicates an 295 acceptable predicting power, and ranged from 0.61 to 0.81 across datasets which indicates a 296 low to good performance in predicting a MPA >0.60, except for BF1 where the 95% CI (0.43, 297 298 0.78) included 0.50 which indicates no statistically significant predictive power (**Table 4**). In the sensitivity and specificity analyses in the pooled sample, the threshold of WDDS-10 \geq 5 299 food groups had the best performances in predicting an MPA >0.60 (i.e. both sensitivity and 300 301 specificity >0.60 and percentage of individuals correctly classified >0.70) with good sensitivity (62%) and very good specificity (81%) and percentage of individuals correctly 302 classified (79%). The threshold of \geq 4 food groups showed slightly lower performances with 303 very good sensitivity (84%), but fair enough specificity (55%) and a moderate percentage of 304 correctly classified participants (58%). The threshold of ≥ 6 food groups had lower 305 306 performances with low sensitivity (32%), but very good specificity (93%) and percentage of correctly classified participants (87%). The other thresholds had worse classification 307 properties. However, findings were heterogeneous across datasets. In summary, when 308 309 balancing sensitivity, specificity and percentage of correct classification, the most acceptable food group consumption threshold for predicting a MPA >0.60 was WDDS-10 \geq 4 in BF1, 310 BF2, and BF3, \geq 5 in India and Nepal, and \geq 6 in Bangladesh. 311 The three distinct scenarios from our robustness analyses returned similar findings, 312

confirming both the observed heterogeneity across countries and also that the threshold of

WDDS-10 ≥5 food groups had the best performance in predicting an MPA >0.60 in the
pooled sample (data not shown).

316

317 Discussion

Following the approach used for developing and validating MDD-W among non-pregnant 318 women (12,13), we analyzed six dietary datasets to determine the minimum number of food 319 groups consumed, out of the 10 food groups of the MDD-W, which best discriminates 320 between higher versus lower micronutrient adequacy among pregnant adolescent girls and 321 women in four LMICs. At least half of the women in each dataset had PAs of six 322 micronutrients at zero, highlighting the urgency of an emphasis on diet quality and nutrient 323 adequacy population group. Consequently, pregnant adolescent girls and women had low 324 325 nutrient adequacy, with median MPA values ranging from 0.09 to 0.43 across the datasets. These findings are consistent with those reported among lactating women, who also face 326 higher nutrient requirements, where the MPA ranged from 0.23 to 0.50 in nine datasets from 327 328 resource-poor settings (12,13). As with other population subgroups (12,13), the WDDS-10 329 was significantly and positively associated with MPA in each dataset. Similarly to the results found during the initial validation of MDD-W for non-pregnant women (12,13), our analyses 330 showed that across the pooled sample a threshold of 5 or more food groups had the best 331 performance in classifying pregnant adolescent girls and women as having a minimally 332 acceptable level of dietary micronutrient adequacy (i.e., MPA >0.60). 333 Nevertheless, we found evidence of heterogeneity across datasets, both in terms of dietary 334 patterns and in the optimal threshold of WDDS-10 to predict a minimally acceptable level of 335 micronutrient adequacy (which varied from 4 to 6). Pulses and dairy were more commonly 336 consumed in South Asian countries, whereas nuts, seeds, and green leafy vegetables were 337 more commonly consumed in Burkina Faso. This could be explained by geographical and 338

temporal differences, such as food availability, prices, budgets, and preferences. For example, 339 each dataset only captured certain months of the year while seasonality could affect food 340 availability and thus dietary diversity in these contexts (39,40). In terms of differences in 341 thresholds, it should be noted that even in the validation study that led to adopt the MDD-W 342 there were differences across datasets regarding the best threshold that predicted a MPA>0.60 343 - which varied from 4 to 6 as in the present study (12). Various food (sub)groups contribute 344 more or less to the MPA than others and/or can be consumed in larger or smaller quantities 345 according to the context. This heterogeneity is not specific to pregnant adolescent girls and 346 women. When recommending the threshold of 5 food groups, that work best in the pooled 347 348 sample in this study as well as across the 9 datasets of the MDD-W validation study (12), we are pretty confident that this threshold would most likely minimize the gap to the true, 349 context-specific and also probably season-specific optimal threshold that remains unknown in 350 351 many contexts but was found in the range of 4 to 6 in most if not all published studies (12, 17, 41).352

Measuring characteristics of diets and monitoring of their changes at global and national 353 levels are needed to support governments in establishing policies and programmes to promote 354 healthy diets, to assess the effectiveness of their actions and hold them accountable. This is 355 356 the spirit behind the development of the MDD-W (12,13). Although MDD-W is already widely collected in large multi-topic surveys, such as Demographic and Health Surveys and 357 Gallup World Poll, it only reflects dietary diversity which is one, albeit indispensable, 358 359 subconstruct of healthy diets (42,43). Other promising metrics were recently designed to assess in a synthetic manner several subconstructs of healthy diets. The Global Diet Quality 360 Score (GDQS), for example, is based on the consumption of 25 food groups that are globally 361 important contributors to nutrient intake, on the one hand, and/or to non-communicable 362 disease risk, on the other hand (44). Although it has been validated using several datasets 363

from various contexts, the validation was performed against several outcomes and by 364 comparisons with the performance of other metrics and not directly to nutrient adequacy. In 365 addition, the GDQS has not yet been widely used in large surveys, probably because some 366 appraisal of quantities or portions consumed is needed for its construction. The Global 367 Dietary Recommendations (GDR) score is another recently developed synthetic metric that 368 was designed to assess the adherence to a dietary pattern respecting 11 global dietary 369 recommendations from WHO, which include dietary factors protective against non-370 communicable diseases (45). Although the construction of the GDR score is based on a 371 standardized Diet Quality Questionnaire that was validated against 24h-recalls in three 372 373 different contexts, and has been used since in many other countries, as far as we know the GDR score itself was validated only with data from Brazil and the USA. Additional evidence 374 are needed to establish its validity in various contexts and its equivalence across contexts (43). 375 376 Thus, MDD-W arguably remains a statically robust and valid indicator, widely collected in large multi-topic surveys, to assess dietary diversity as a cornerstone of diet quality on a 377 global and national scale. This work contributes to ongoing efforts to validate MDD-W in 378 other populations such as adolescents and children (43). 379 The present analyses have some limitations. First, despite our efforts to obtain datasets from a 380

381 diversity of contexts, our study only includes data from rural contexts in four LMICs among two regions (sub-Saharan Africa and South Asia). Although our findings are not globally 382 representative, they are consistent with other analyses among non-pregnant women from more 383 384 settings (12,13). Furthermore, the rural locations included in our study are settings where valid scores are arguably much needed, as they typically have a high burden of undernutrition 385 and low dietary diversity (15,39,46,47). In the meantime, more datasets should be made 386 available in settings where a reasonable proportion of pregnant adolescent girls and women 387 reach an acceptable MPA, so that the best predictors of acceptable MPA can be further 388

studied. For example, in the BF1 sample of our study, only 11 (2.4%) pregnant women 389 reached an MPA \geq 0.60, which strongly limits the search for the best dichotomous indicator 390 predicting higher MPA. Another limitation concerns the use of an external within-person 391 variance estimate to calculate the MPA in four of the six datasets. This results in more reliable 392 prevalence estimates than when using a single day recall (36), but the use of within-person 393 variance estimates from repeated measures within the samples is preferable (35). Although we 394 tried to find and use an external estimate of within-person variance from a relevant food 395 intake survey, we were limited in our ability to find studies with the same geographical (e.g. 396 for India, the region of the external estimate study is 1500 km away from that of the dataset) 397 398 or temporal (different seasonality between BF1 and BF3) characteristics. Future analyses 399 from a wider variety of settings and with data containing repeated measures is recommended to confirm that a threshold of 5 or more groups is the best suited to indicate MPA >0.60. A 400 401 last limitation is the use of a set of nutrient requirements which did not take into account the pregnancy trimester, the age of the participants or the level of bioavailability of iron and zinc. 402 This simpler approach was preferred to take into account the fact that this information might 403 not be accurately collected in large surveys. Nevertheless, taking these characteristics into 404 account in three distinct robustness analyses did not affect our findings in terms of 405 406 determining the threshold of WDDS-10 with the best classification characteristics. 407 In conclusion, our study suggests that the WDDS-10 is a good predictor of dietary micronutrient adequacy among pregnant adolescent girls and women in LMICs, as it was 408 409 previously shown among non-pregnant and non-lactating women and lactating women (12,13). When a dichotomous indicator is preferred over a continuous measure, our results 410 suggest that the MDD-W may be used as a proxy indicator for higher micronutrient adequacy 411 in LMIC contexts in all women of reproductive age, regardless of physiological status. This 412 might be particularly useful for international comparisons and when the physiological status 413

of women is unknown, which is the case in many large surveys. However, our findings 414 suggest that context-specific thresholds might be more accurate and might therefore be 415 preferred for research purposes. Given the low micronutrient adequacy in the populations 416 studied, additional efforts are needed to enhance the diet of women of reproductive age. Even 417 though the threshold of 5 or more groups might not accurately predict micronutrient adequacy 418 in all contexts, the indicator allows tracking processes of such efforts over time and enables 419 benchmarking between populations. However, there is a need to provide complementary 420 assessment of other dimensions of diet quality, such as consumption of undesired foods, food 421 safety aspects, and within food group contribution of foods. In addition, in food environments 422 and diets with a considerable contribution of fortified foods, the validity of the 5 food group 423 thresholds might require careful reconsideration. 424

425

426 Acknowledgements

427 We are grateful to Kripa Rajagopalan (Cornell University) and Rupak Shivakoti (Columbia University) for sharing the within-person variance estimate from their FFQ validation study 428 among pregnant women living with or without HIV in Pune, India. The authors' 429 responsibilities were as follows: EOV, SED, DBB, GTH-C, EL, MS, YMP, and CL designed 430 the study; AA, EB, LD, AG, GTH-C, HH-F, SK, SSK, PHN, NMS, LMT and RRZ provided 431 the datasets; DBB harmonized the datasets; SBB analyzed data; EOV and SBB drafted the 432 figures, tables, and manuscript, and the other authors provided critical review; all authors read 433 and approved the final manuscript. 434

435

436 **Data Availability**

437 Data described in the manuscript, code book, and analytic code will be made available upon438 request pending application and approval by authors of the current study.

439

440 Funding

- 441 This publication has been produced with the financial support of the European Commission,
- under the project "Knowledge and research for nutrition" implemented by Agrinatura EEIG. Its
- 443 contents are the sole responsibility of the authors and do not necessarily reflect the views of the
- 444 European Union and of Agrinatura EEIG.

445

446 Author Disclosures

447 The authors report no conflicts of interest.

448

References

- 1. Bailey RL, West Jr. KP, Black RE. The Epidemiology of Global Micronutrient Deficiencies. Ann Nutr Metab 2015;66:22–33.
- 2. Stevens GA, Beal T, Mbuya MNN, Luo H, Neufeld LM, Addo OY, Adu-Afarwuah S, Alayón S, Bhutta Z, Brown KH, et al. Micronutrient deficiencies among preschool-aged children and women of reproductive age worldwide: a pooled analysis of individual-level data from population-representative surveys. Lancet Glob Health Elsevier; 2022;10:e1590–9.
- 3. Marshall NE, Abrams B, Barbour LA, Catalano P, Christian P, Friedman JE, Hay WW, Hernandez TL, Krebs NF, Oken E, et al. The importance of nutrition in pregnancy and lactation: lifelong consequences. Am J Obstet Gynecol 2022;226:607–32.
- 4. Abu-Saad K, Fraser D. Maternal Nutrition and Birth Outcomes. Epidemiol Rev 2010;32:5–25.
- 5. Keats EC, Haider BA, Tam E, Bhutta ZA. Multiple-micronutrient supplementation for women during pregnancy. Cochrane Database Syst Rev 2019;3:CD004905.
- 6. Bourassa MW, Osendarp SJM, Adu-Afarwuah S, Ahmed S, Ajello C, Bergeron G, Black R, Christian P, Cousens S, de Pee S, et al. Review of the evidence regarding the use of antenatal multiple micronutrient supplementation in low- and middle-income countries. Ann N Y Acad Sci 2019;1444:6–21.
- 7. Adrianopoli M, D'Acapito P, Ferrari M, Mistura L, Toti E, Maiani G, Thompson B, Amoroso L. Improving diets and nutrition: food-based approaches. Rome Italy CABI 2014;230–45.
- 8. Verger EO, Port AL, Borderon A, Bourbon G, Moursi M, Savy M, Mariotti F, Martin-Prevel Y. Dietary Diversity Indicators and Their Associations with Dietary Adequacy and Health

Outcomes: A Systematic Scoping Review. Adv Nutr [Internet] 2021 [cited 2021 Apr 22]; Available from: https://doi.org/10.1093/advances/nmab009

- 9. World Health Organization. WHO recommendations on antenatal care for a positive pregnancy experience [Internet]. Geneva: World Health Organization; 2016 [cited 2023 Jun 23]. 152 p. Available from: https://apps.who.int/iris/handle/10665/250796
- 10. Herforth A, Arimond M, Álvarez-Sánchez C, Coates J, Christianson K, Muehlhoff E. A Global Review of Food-Based Dietary Guidelines. Adv Nutr Bethesda Md 2019;10:590–605.
- 11. Food and Agriculture Organization. Minimum dietary diversity for women: An updated guide to measurement-from collection to action [Internet]. Rome: FAO; 2021. Available from: https://doi.org/10.4060/cb3434en
- 12. Martin-Prével Y, Allemand P, Wiesmann D, Arimond M, Ballard T, Deitchler M, Dop M-C, Kennedy G, Lee WT, Moursi M. Moving forward on choosing a standard operational indicator of women's dietary diversity. FAO; 2015.
- 13. Women's Dietary Diversity Project (WDDP) Study Group. Development of a Dichotomous Indicator for Population-Level Assessment of Dietary Diversity in Women of Reproductive Age. Curr Dev Nutr 2017;1:cdn.117.001701-cdn.117.001701.
- Gyimah LA, Annan RA, Apprey C, Edusei A, Aduku LNE, Asamoah-Boakye O, Azanu W, Lutterodt H. Dietary diversity and its correlates among pregnant adolescent girls in Ghana. PLOS ONE Public Library of Science; 2021;16:e0247979.
- 15. Shrestha V, Paudel R, Sunuwar DR, Lyman ALT, Manohar S, Amatya A. Factors associated with dietary diversity among pregnant women in the western hill region of Nepal: A community based cross-sectional study. PLOS ONE Public Library of Science; 2021;16:e0247085.
- 16. Yang J, Wang M, Tobias DK, Rich-Edwards JW, Darling A-M, Abioye AI, Noor RA, Madzorera I, Fawzi WW. Dietary diversity and diet quality with gestational weight gain and adverse birth outcomes, results from a prospective pregnancy cohort study in urban Tanzania. Matern Child Nutr 2022;18:e13300.
- 17. Nguyen PH, Huybregts L, Sanghvi TG, Tran LM, Frongillo EA, Menon P, Ruel MT. Dietary Diversity Predicts the Adequacy of Micronutrient Intake in Pregnant Adolescent Girls and Women in Bangladesh, but Use of the 5-Group Cutoff Poorly Identifies Individuals with Inadequate Intake. J Nutr 2018;148:790–7.
- Becquey E, Diop L, Awonon J, Diatta AD, Ganaba R, Pedehombga A, Gelli A. A Poultry Value Chain Intervention Promoting Diversified Diets Has Limited Impact on Maternal and Child Diet Adequacy during the Lean Season in a Cluster Randomized Controlled Trial. J Nutr 2022;152:1336–46.
- de Kok B, Argaw A, Hanley-Cook G, Toe LC, Ouédraogo M, Dailey-Chwalibóg T, Diop L, Becquey E, Kolsteren P, Lachat C, et al. Fortified Balanced Energy-Protein Supplements Increase Nutrient Adequacy without Displacing Food Intake in Pregnant Women in Rural Burkina Faso. J Nutr 2021;151:3831–40.
- 20. Kim SS, Zagré RR, Ouédraogo CT, Sununtnasuk C, Ganaba R, Zafimanjaka MG, Tharaney M, Sanghvi T, Menon P. Intensified Nutrition Interventions in Antenatal Care Services Increased Consumption of Iron and Folic Acid Supplements and Early Breastfeeding Practices in Burkina Faso: Results of a Cluster-Randomized Program Evaluation. J Nutr 2023;S0022-3166(23)72422-X.

- Nguyen PH, Kachwaha S, Tran LM, Avula R, Young MF, Ghosh S, Sharma PK, Escobar-Alegria J, Forissier T, Patil S, et al. Strengthening Nutrition Interventions in Antenatal Care Services Affects Dietary Intake, Micronutrient Intake, Gestational Weight Gain, and Breastfeeding in Uttar Pradesh, India: Results of a Cluster-Randomized Program Evaluation. J Nutr 2021;151:2282–95.
- 22. Harris-Fry HA, Paudel P, Harrisson T, Shrestha N, Jha S, Beard BJ, Copas A, Shrestha BP, Manandhar DS, Costello AM de L, et al. Participatory Women's Groups with Cash Transfers Can Increase Dietary Diversity and Micronutrient Adequacy during Pregnancy, whereas Women's Groups with Food Transfers Can Increase Equity in Intrahousehold Energy Allocation1. J Nutr 2018;148:1472–83.
- 23. Kim SS, Ouédraogo CT, Zagré RR, Ganaba R, Zafimanjaka MG, Tharaney M, Menon P. Multiple modifiable maternal, household and health service factors are associated with maternal nutrition and early breastfeeding practices in Burkina Faso. Matern Child Nutr 2023;19:e13457.
- 24. Harris-Fry H, Beard BJ, Harrisson T, Paudel P, Shrestha N, Jha S, Shrestha BP, Manandhar DS, Costello A, Saville NM. Smartphone tool to collect repeated 24 h dietary recall data in Nepal. Public Health Nutr England; 2018;21:260–72.
- 25. Harris-Fry H, Paudel P, Karn M, Mishra N, Thakur J, Paudel V, Harrisson T, Shrestha B, Manandhar DS, Costello A, et al. Development and validation of a photographic food atlas for portion size assessment in the southern plains of Nepal. Public Health Nutr Cambridge University Press; 2016;19:2495–507.
- 26. Gibson RS, Ferguson EL. An interactive 24-hour recall for assessing the adequacy of iron and zinc intakes in developing countries. [Internet]. Washington (DC) and Cali : International Food Policy Research Institute (IFPRI) and International Center for Tropical Agriculture (CIAT). 2008 [cited 2023 Jan 17]. 157 p. Available from: https://www.ifpri.org/publication/interactive-24-hour-recall-assessing-adequacy-iron-and-zinc-intakes-developing-countries
- 27. Arimond M, Wiesmann D, Becquey E, Carriquiry A, Daniels MC, Deitchler M, Fanou-Fogny N, Joseph ML, Kennedy G, Martin-Prevel Y, et al. Simple food group diversity indicators predict micronutrient adequacy of women's diets in 5 diverse, resource-poor settings. J Nutr 2010;140:2059S-69S.
- 28. National Research Council. Dietary Reference Intakes: Applications in Dietary Assessment [Internet]. Washington, DC: The National Academies Press; [cited 2014 Mar 7]. Available from: http://www.nap.edu/catalog.php?record_id=9956
- 29. World Health Organization. Vitamin and mineral requirements in human nutrition. World Health Organization; 2004.
- 30. Institue of Medicine. Dietary Reference Intakes: The Essential Guide to Nutrient Requirements [Internet]. The national academies press. Washington, D.C.; 2006 [cited 2023 Feb 3]. Available from: https://www.nap.edu/read/11537/chapter/1
- 31. Institute of Medicine (US) Committee to Review Dietary Reference Intakes for Vitamin D and Calcium. Dietary Reference Intakes for Calcium and Vitamin D [Internet]. Ross AC, Taylor CL, Yaktine AL, Del Valle HB, editors. Washington (DC): National Academies Press (US); 2011 [cited 2023 Feb 3]. Available from: http://www.ncbi.nlm.nih.gov/books/NBK56070/
- 32. International Zinc Nutrition Consultative Group (IZiNCG), Brown KH, Rivera JA, Bhutta Z, Gibson RS, King JC, Lönnerdal B, Ruel MT, Sandtröm B, Wasantwisut E, et al. International

Zinc Nutrition Consultative Group (IZiNCG) technical document #1. Assessment of the risk of zinc deficiency in populations and options for its control. Food Nutr Bull 2004;25:S99-203.

- 33. Allen LH, Carriquiry AL, Murphy SP. Perspective: Proposed Harmonized Nutrient Reference Values for Populations. Adv Nutr Bethesda Md 2020;11:469–83.
- Joseph ML, Carriquiry A. A measurement error approach to assess the association between dietary diversity, nutrient intake, and mean probability of adequacy. J Nutr 2010;140:2094S-101S.
- 35. French CD, Arsenault JE, Arnold CD, Haile D, Luo H, Dodd KW, Vosti SA, Slupsky CM, Engle-Stone R, The Variance Components of Nutrient Intakes Data Working Group. Within-Person Variation in Nutrient Intakes across Populations and Settings: Implications for the Use of External Estimates in Modeling Usual Nutrient Intake Distributions. Adv Nutr 2021;12:429–51.
- 36. Jahns L, Arab L, Carriquiry A, Popkin BM. The use of external within-person variance estimates to adjust nutrient intake distributions over time and across populations. Public Health Nutr Cambridge University Press; 2005;8:69–76.
- 37. International Food Policy Research Institute (IFPRI). A&T Bangladesh Maternal Nutrition Endline Survey 2016: Households - Pregnant Women [Internet]. Harvard Dataverse; 2021 [cited 2023 May 17]. Available from: https://dataverse.harvard.edu/dataset.xhtml?persistentId=doi:10.7910/DVN/PYLIY4
- 38. Rajagopalan K, Alexander M, Naik S, Patil N, Mehta S, Leu C-S, Bhosale R, Mathad JS, Caulfield LE, Gupta A, et al. Validation of New Interactive Nutrition Assistant - Diet in India Study of Health (NINA-DISH) FFQ with multiple 24-h dietary recalls among pregnant women in Pune, India. Br J Nutr 2021;126:1247–56.
- 39. Hanley-Cook GT, Argaw A, de Kok B, Toe LC, Dailey-Chwalibóg T, Ouédraogo M, Kolsteren P, Huybregts L, Lachat C. Seasonality and Day-to-Day Variability of Dietary Diversity: Longitudinal Study of Pregnant Women Enrolled in a Randomized Controlled Efficacy Trial in Rural Burkina Faso. J Nutr 2022;152:2145–54.
- 40. Saville NM, Cortina-Borja M, De Stavola BL, Pomeroy E, Marphatia A, Reid A, Manandhar DS, Wells JC. Comprehensive analysis of the association of seasonal variability with maternal and neonatal nutrition in lowland Nepal. Public Health Nutr 2021;25:1–16.
- 41. Diop L, Becquey E, Turowska Z, Huybregts L, Ruel MT, Gelli A. Standard Minimum Dietary Diversity Indicators for Women or Infants and Young Children Are Good Predictors of Adequate Micronutrient Intakes in 24-59-Month-Old Children and Their Nonpregnant Nonbreastfeeding Mothers in Rural Burkina Faso. J Nutr 2021;151:412–22.
- 42. Seligman HK, Levi R, Adebiyi VO, Coleman-Jensen A, Guthrie JF, Frongillo EA. Assessing and Monitoring Nutrition Security to Promote Healthy Dietary Intake and Outcomes in the United States. Annu Rev Nutr 2023;
- 43. Verger EO, Savy M, Martin-Préve Y, Coates J, Frongillo E, Neufeld L, Saha K, Hayashi C, Holmes B, Vogliano C. Healthy diet metrics: a suitability assessment of indicators for global and national monitoring purposes [Internet]. Geneva: World Health Organization; 2023. Available from: https://apps.who.int/iris/bitstream/handle/10665/371497/9789240072138-eng.pdf
- 44. Bromage S, Batis C, Bhupathiraju SN, Fawzi WW, Fung TT, Li Y, Deitchler M, Angulo E, Birk N, Castellanos-Gutiérrez A, et al. Development and Validation of a Novel Food-Based Global Diet Quality Score (GDQS). J Nutr 2021;151:75S-92S.

- 45. Herforth AW, Wiesmann D, Martínez-Steele E, Andrade G, Monteiro CA. Introducing a Suite of Low-Burden Diet Quality Indicators That Reflect Healthy Diet Patterns at Population Level. Curr Dev Nutr 2020;4:nzaa168.
- 46. Harris-Fry H, Azad K, Kuddus A, Shaha S, Nahar B, Hossen M, Younes L, Costello A, Fottrell E. Socio-economic determinants of household food security and women's dietary diversity in rural Bangladesh: a cross-sectional study. J Health Popul Nutr 2015;33:2–2.
- 47. Lander R, Hambidge K, Westcott J, Tejeda G, Diba T, Mastiholi S, Khan U, Garces A, Figueroa L, Tshefu A, et al. Pregnant Women in Four Low-Middle Income Countries Have a High Prevalence of Inadequate Dietary Intakes That Are Improved by Dietary Diversity. NUTRIENTS 2019;11.

ournal provide

Figures

Fig. 1. Percentage of participants having consumed the 10 food groups used to construct MDD-W in the previous 24-hours. BF1, rural Burkina Faso dataset (2017/2019/2020); BF2, rural Burkina Faso dataset (2020); BF3, rural Burkina Faso dataset (2019/2021).

Fig. 2. Average mean probability of adequacy by WDDS-10 score. BF1, rural Burkina Faso dataset (2017/2019/2020); BF2, rural Burkina Faso dataset (2020); BF3, rural Burkina Faso dataset (2019/2021). Error bars represent mean \pm standard error. Data points representing <10 participants are not shown. Details of the number of pregnant women by dataset are given in Supplemental Table 4.

Journal Pression

Tables

Table 1. Characteristics of pregnant women¹

Dataset	n	Repeated recall, $n (\%)^2$	Mean (SD) age, y	Adolescent, n (%)	First trimester, <i>n</i> (%)	Second trimester, n (%)	Third trimester, <i>n</i> (%)	Mean (SD) height, m	Mean (SD) weight, kg	Median (IQR) WDDS-10
Bangladesh	598	0 (0.0)	24.0 (5.6)	160 (26.0)	0 (0.0)	328 (54.8)	270 (45.2)	1.50 (0.06)	50.3 (8.1)	5 (2)
BF1	452	84 (18.6)	29.6 (5.3)	1 (0.2)	124 (27.4)	173 (38.3)	155 (34.3)	1.61 (0.07)	59.1 (8.0)	3 (2)
BF2	470	0 (0.0)	25.4 (6.4)	37 (7.9)	16 (3.4)	188 (40.0)	266 (56.6)	1.63 (0.06)	58.9 (8.7)	3 (2)
BF3 ³	1912	0 (0.0)	27.5 (6.6)	64 (3.4)	279 (14.7)	828 (43.8)	785 (41.5)	1.63 (0.01)	61.8 (2.5)	3 (2)
India	674	0 (0.0)	25.0 (4.0)	0 (0.0)	0 (0.0)	198 (29.4)	476 (70.6)	1.50 (0.06)	51.0 (8.5)	4 (2)
Nepal	803	745 (92.8)	21.5 (3.8)	88 (11.0)	0 (0.0)	1 (0.1)	802 (99.9)	1.51 (0.05)	52.1 (6.5)	4 (1)
Pooled ³	4909	N/A	25.7 (6.2)	350 (7.1)	419 (8.5)	1716 (34.9)	2754 (56.1)	1.58 (0.07)	56.7 (8.0)	3 (1)

¹ BF1, rural Burkina Faso dataset (2017/2019/2020); BF2, rural Burkina Faso dataset (2020); BF3, rural Burkina Faso dataset (2019/2021); SD, standard deviation; WDDS-10, 10-food group women dietary diversity score; IQR, Interquartile range. ² Women in the sample with more than one 24-hour dietary recall; ³ Information about the pregnancy trimester was missing for 20 participants.

Dataset	Energy intakes, kcal/d ²	Vitamin A ³	Thiamin ³	Riboflavin ³	Niacin ³	Vitamin B6 ³	Folate ³	Vitamin B12 ³	Vitamin C ³	Calcium ³	Iron ³	Zinc ³	MPA ³	MPA >0.60, <i>n</i> (%)
Bangladesh	2330 (822)	0.00 (0.60)	1.00 (0.22)	0.00 (0.25)	1.00 (0.00)	1.00 (0.00)	0.00 (0.00)	0.00 (0.07)	1.00 (0.00)	0.00 (0.00)	0.00 (0.00)	0.18 (0.70)	0.40 (0.19)	94 (15.7)
BF1	1950 (939)	0.00 (0.10)	0.00 (0.04)	0.00 (0.00)	0.00 (0.00)	0.00 (0.00)	0.00 (0.00)	0.00 (0.00)	0.00 (0.67)	0.00 (0.00)	0.00 (0.49)	0.40 (0.95)	0.09 (0.21)	11 (2.4)
BF2	2473 (1482)	0.00 (0.48)	0.00 (0.76)	0.00 (0.01)	0.00 (0.22)	0.00 (0.67)	0.00 (0.15)	0.00 (0.00)	0.01 (1.00)	0.00 (0.05)	0.00 (0.95)	0.97 (0.79)	0.16 (0.34)	69 (14.7)
BF3	1816 (838)	0.00 (0.00)	0.00 (0.01)	0.00 (0.00)	0.00 (0.03)	0.00 (0.04)	0.00 (0.00)	0.00 (0.00)	0.00 (0.05)	0.00 (0.01)	0.07 (1.00)	0.63 (0.95)	0.13 (0.21)	73 (3.8)
India	2122 (924)	0.00 (0.00)	0.80 (0.99)	0.00 (0.23)	0.20 (0.92)	0.00 (0.00)	0.00 (0.00)	0.00 (0.00)	0.04 (1.00)	0.00 (0.08)	0.00 (0.00)	0.62 (0.97)	0.20 (0.32)	35 (5.2)
Nepal	2254 (850)	0.05 (0.46)	0.96 (0.76)	0.06 (0.99)	0.64 (0.83)	1.00 (0.36)	0.00 (0.01)	0.00 (0.00)	1.00 (0.17)	0.00 (0.45)	0.00 (0.00)	0.99 (0.34)	0.43 (0.32)	188 (23.4)
Pooled	2068 (969)	0.00 (0.08)	0.03 (0.98)	0.00 (0.05)	0.03 (0.91)	0.00 (0.99)	0.00 (0.00)	0.00 (0.00)	0.02 (1.00)	0.00 (0.02)	0.00 (0.26)	0.69 (0.93)	0.20 (0.34)	470 (9.6)

Table 2. Energy intakes, probability of adequacy of individual micronutrients and mean probability of adequacy (MPA)¹

¹BF1, rural Burkina Faso dataset (2017/2019/2020); BF2, rural Burkina Faso dataset (2020); BF3, rural Burkina Faso dataset (2019/2021). ² Values are means (SD) calculated from a single 24-hour dietary recall (the first one in case of repetitions). ³ Values are medians (interquartile range).

		Unadjusted		Total energy (kcal/d) adjusted							
Dataset	WDDS-10	Constant	Adjusted R ²	WDDS-10	Energy intake, kcal/d	Constant	Adjusted R ²				
Bangladesh	0.079 (0.070, 0.088)	-1.06 (-1.11, -1.01)	0.333	0.055 (0.046, 0.063)	0.0001 (0.0001, 0.0001)	-1.23 (-1.27, -1.18)	0.529				
BF1	0.252 (0.195, 0.310)	-2.42 (-2.60, -2.24)	0.142	0.125 (0.067, 0.183)	0.0003 (0.0003, 0.0004)	-2.65 (-2.82, -2.48)	0.291				
BF2	0.309 (0.250, 0.367)	-2.20 (-2.38, -2.01)	0.185	0.166 (0.114, 0.218)	0.0002 (0.0002, 0.0003)	-2.37 (-2.52, -2.21)	0.431				
BF3	0.214 (0.194, 0.233)	-2.06 (-2.13, -2.00)	0.198	0.091 (0.074, 0.108)	0.0004 (0.0003, 0.0004)	-2.40 (-2.45, -2.34)	0.488				
India	0.162 (0.139, 0.186)	-1.73 (-1.83, -1.63)	0.214	0.049 (0.032, 0.067)	0.0003 (0.0003, 0.0004)	-2.00 (-2.07, -1.94)	0.662				
Nepal	0.082 (0.068, 0.095)	-0.93 (-0.99, -0.87)	0.149	0.038 (0.028, 0.050)	0.0002 (0.0001, 0.0002)	-1.11 (-1.16, -1.06)	0.465				
Pooled ³	0.168 (0.157, 0.178)	-1.74 (-1.96, -1.51)	0.286	0.079 (0.069, 0.088)	0.0003 (0.0002, 0.0003)	-2.03 (-2.27, -1.78)	0.411				

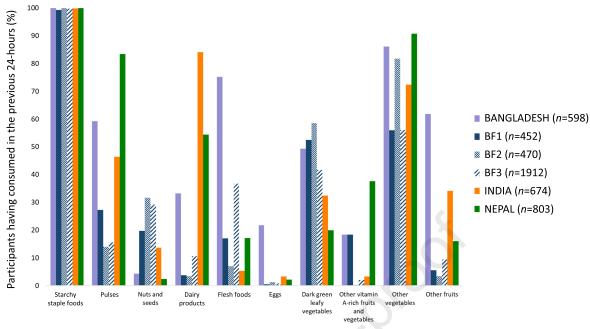
Table 3. Linear regression of V	DDS-10 with mean	probability of	of adequacy ^{1,2}

¹ Values are regression coefficients and (95% Confidence Intervals); WDDS-10, 10-food group women dietary diversity score; CI, confident interval; BF1, rural Burkina Faso dataset (2017/2019/2020); BF2, rural Burkina Faso dataset (2020); BF3, rural Burkina Faso dataset (2019/2021). ² The mean probability of adequacy after Box-Cox transformation was used as dependent variable in all the regression models. All *P*-values are <0.001. ³ A mixed-effects regression model, including a random intercept for survey, was fitted for the pooled sample.

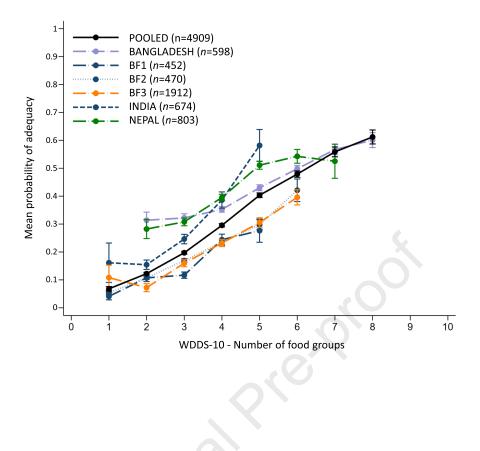
		WDDS-10 \geq 4			W	DDS-10 \geq 5		WDDS-10 ≥ 6		
Dataset	AUC	Sensitivity	Specificity	PCC	Sensitivity	Specificity	PCC	Sensitivity	Specificity	PCC
Bangladesh	0.81 (95% CI: 0.77, 0.85)	98.9	19.6	32.1	97.9	41.5	50.3	78.7	67.3	69.1
BF1	0.61 (95% CI: 0.43, 0.78)	54.6	69.4	69.0	9.10	93.4	91.4	0.00	98.9	96.5
BF2	0.71 (95% CI: 0.65, 0.78)	55.1	77.3	74.0	21.7	96.8	85.7	2.9	99.8	85.5
BF3	0.74 (95% CI: 0.69, 0.79)	63.0	70.3	70.0	31.5	91.3	89.0	17.8	98.4	95.3
India	0.79 (95% CI: 0.73, 0.86)	97.1	37.7	40.8	77.1	71.5	71.8	34.3	91.6	88.6
Nepal	0.74 (95% CI: 0.71, 0.78)	94.7	30.6	45.6	70.2	70.6	70.5	26.6	92.7	77.2
Pooled	0.78 (95% CI: 0.75, 0.80)	84.0	54.9	57.7	61.7	80.6	78.8	32.1	93.2	87.4

Table 4. Test characteristics of food group indicators for classifying mean probability of adequacy >0.60 for pregnant adolescents and women¹

¹ Values are percentages (except for the AUC values); AUC, area under the curve; CI, confident interval; BF1, rural Burkina Faso dataset (2017/2019/2020); BF2, rural Burkina Faso dataset (2020); BF3, rural Burkina Faso dataset (2019/2021). PCC, percentage correctly classified; WDDS-10, 10-food group Women Dietary Diversity Score.



vegetables



Declaration of interests

 The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

☑ The authors declare the following financial interests/personal relationships which may be considered as potential competing interests:

Eric Verger reports financial support was provided by European Commission.

<u>remis</u>